



**AgEcon** SEARCH  
RESEARCH IN AGRICULTURAL & APPLIED ECONOMICS

*The World's Largest Open Access Agricultural & Applied Economics Digital Library*

**This document is discoverable and free to researchers across the globe due to the work of AgEcon Search.**

**Help ensure our sustainability.**

Give to AgEcon Search

AgEcon Search  
<http://ageconsearch.umn.edu>  
[aesearch@umn.edu](mailto:aesearch@umn.edu)

*Papers downloaded from **AgEcon Search** may be used for non-commercial purposes and personal study only. No other use, including posting to another Internet site, is permitted without permission from the copyright owner (not AgEcon Search), or as allowed under the provisions of Fair Use, U.S. Copyright Act, Title 17 U.S.C.*

# TRANSGENIC TECHNOLOGY ADOPTION AND ON-FARM VARIETAL DIVERSITY

Vijesh KRISHNA\*

Department of Agricultural and Resource Economics  
University of California at Berkeley, 94720 CA

David ZILBERMAN

Department of Agricultural and Resource Economics  
University of California at Berkeley, 94720 CA

Matin QAIM

Department of Agricultural Economics and Rural Development  
Georg-August-University of Goettingen, 37073 Goettingen, Germany

*Contributed Paper prepared for presentation at the  
International Association of Agricultural Economists (IAAE) Conference  
Beijing, China, August 16-22, 2009*

*Acknowledgement: The financial support of the German Research Foundation (DFG) for this research is gratefully acknowledged.*

*Copyright 2009 by V. Krishna, D. Zilberman and M. Qaim. All rights reserved. Readers may make verbatim copies of this document for non-commercial purposes by any means, provided this copyright notice appears on all such copies.*

---

\* Corresponding author. Phone: + 1 510 643 0846; Fax: +1 510 643 8911; E-mail: [vijesh@berkeley.edu](mailto:vijesh@berkeley.edu)

# TRANSGENIC TECHNOLOGY ADOPTION AND ON-FARM VARIETAL DIVERSITY

**ABSTRACT:** Transgenic pest-resistant varieties are hypothesized to reduce farmers' demand for on-farm diversity through an act of substitution, as both serve as production risk reducing instruments. This adverse agro-biodiversity impact of technology adoption might be partially counteracted by an expanding seed sector, supplying a large number of transgenic varieties. The case of Bt cotton in India is taken for empirical illustration. The production function analyses show that both Bt technology and on-farm varietal diversity enhance yield, while reducing the production risk. With few Bt varieties available in the first years, technology adoption entailed a reduction in on-farm varietal diversity. This effect, however, was partially offset by more Bt varieties becoming available over time.

**KEY WORDS:** Agro-biodiversity; Bt cotton; India; Production risk.

## I. INTRODUCTION

From the very inception of transgenic technology in agriculture, concern about its potential environmental impacts has been taking the center stage in both public and scientific forums. In general, the potential biodiversity threats associated with introduction of genetically modified (GM) crops include (i) loss of original diversity through gene flow, (ii) damage infliction to non-target organisms, and (iii) erosion of crop varietal diversity (Dale et al., 2002; Ammann, 2005; Thomson, 2006). The present study addresses the third aspect that the adoption of few genetically uniform transgenic varieties erodes the on-farm diversity of crop varieties. Superficially, this argument bears significant resemblance to the conventional genetic erosion hypothesis (cf. Brush, 2000), which is arguably associated with the introduction of any high yielding crop variety. Nonetheless, this paper proposes that the mode in which GM technology impacts varietal diversity significantly differs from the conventional yield enhancing methods. We employ the case of transgenic pest resistance (PR) for illustration purposes, explicitly differentiating between demand and supply side aspects.

Crop genetic diversity is considered as a source of continuing advances in yield, pest resistance and quality improvement and thus acts as a contributing force towards sustainable agriculture (Birol et al., 2005; Jackson et al., 2007). There exists mixed evidence on the relationship between genetic

diversity and mean farm yields, but it is widely accepted that greater varietal and species diversity would enable the system to maintain productivity over a wide range of agro-ecological conditions (Chapin et al., 1997; Kinzig et al., 2002; Di Falco and Perrings, 2003). Further, Perrings and Gardgil (2002) indicated that the magnitude of agricultural risks typically varies inversely with crop genetic diversity. Transgenic variety adoption could be associated with a similar risk-reduction function. Researchers currently advise farmers to consider adoption of PR-GM as an insurance mechanism (Hyde et al. 1999). According to Crost and Shankar (2008), Bt technology has a risk-reducing effect on the cotton production sector of India. Thus, by adopting transgenic resistance, a farmer could reduce her dependence on varietal diversity to attain production risk reduction from biotic factors. Or, in other words, the farmer could switch over from diversity-based to technology-based crop insurance, taking into account the financial and transaction costs associated with each alternative. Hence, transgenic technology adoption might reduce farmers' demand for on-farm varietal diversity.

But how does the supply side look like, meaning the seed industry's provision of varietal diversity? If only one or a few transgenic varieties were available, the farmers might adopt and replace them for a large number of conventional crop varieties. However, if the conditions in a country are conducive, seed firms can backcross the transgenic trait into numerous varieties, thereby increasing the availability of transgenics. We argue here that, the farmers' motive behind cultivating multiple varieties on-farm plays a critical role in this regard. If diminishing returns to scale for land under single variety be the sole reason, then the increased corporate supply of varieties with Bt technology would restore on-farm diversity. The supply, however, very much depends on local breeding capacities, intellectual property rights regimes, biosafety systems, and other institutional conditions (Batie and Ervin, 2001; Qaim, 2005). Nevertheless, when the Bt technology is considered as an instrument to reduce production risk, the varietal diversity status could be expected to diminish with adoption, even when all the existing varieties are genetically modified.

As a starting point for modeling the impacts of GM technology on crop varietal diversity, we present working hypotheses in the forthcoming section. Section III briefly outlines the development of PR-GM technology in the Indian cotton production sector, whereas section IV and V, respectively elaborates the procedure of data collection and econometric framework for analysis. The results of the empirical analysis are presented and discussed in section VI, while the last section concludes.

## II. RESEARCH HYPOTHESES

The broad objective of this study is to examine the impact of introduction of transgenic technology on (i) mean yield (ii) production risk and (iii) farmers' demand for on-farm varietal diversity. Based on this, following workable hypotheses are proposed:

- (i) Adoption of PR-GM technology enhances mean yield.
- (ii) Both on-farm varietal diversity and adoption of PR-GM technology reduce production risk.
- (iii) Adoption of transgenic technology reduces farmers' demand for varietal diversity.
- (iv) Supply of GM varieties by the seed industry is positively correlated with on-farm varietal diversity.

These hypotheses are tested using the data on Bt cotton in India collected at two points of time – one during the early stage of Bt adoption in India (year 2004), when only few GM cotton varieties were available, and the other after the number of available GM varieties had increased considerably (year 2006). The 15-fold increase in number of varieties available with Bt technology during this time makes the case of Bt cotton in India ideal for a study supply impacts of transgenic technology.

## III. CASE OF BT COTTON IN INDIA

Bollworm (mainly *Helicoverpa armigera*) infestation is reported to cause significant crop loss (worth Rs. 20 billion in the late 1990s as per BIRTHAL et al., 2000), despite the cotton cultivation sector being highly chemical intensive. Transgenic resistance development in cotton gained special relevance in this scenario. A gene (Cry1Ac) from the soil bacterium *Bacillus thuringiensis* (Bt) was transferred to the cotton genome, making it resistant to borer pests and reducing its dependency on chemical measures of pest control. The insect resistance in these hybrids was introgressed from a US Bt cotton variety (event MON531) developed by Monsanto (Qaim and Zilberman, 2000). In 2006, India became the world's fifth largest producer of GM crops, with a cultivated area of 9.4 million acres (3.8% of the world's total), and registered the highest growth rate of cultivated areas. From 2005 to 2006, the GM crop area in India increased by 192% (James, 2006).

Despite these significant economic benefits (cf. Qaim et al., 2006; Crost et al., 2007; Guere et al., 2008), the impact of Bt technology on on-farm varietal diversity is hardly examined. When this

technology was introduced in Indian cotton production sector, there were only three Bt hybrids approved. All of them were released by the joint venture of Monsanto with the Indian firm Maharashtra Hybrid Seed Company (MAHYCO), marketed under the trade name Bollgard<sup>TM</sup>. In 2004, an additional Bt variety, released by Rasi Seeds, became legally available for the farmers. The slow increase in number of transgenic varieties could be attributed to the fact that every single Bt variety has to be approved by the national Genetic Engineering Approval Committee (GEAC). However, within the next two years, the GEAC approved more than 60 Bt cotton hybrids developed by 13 seed companies. All of these, except four, have the MMB Bt technology (event MON531) that has been sublicensed to the respective seed companies (APCoAB, 2006). Also in 2006, MAHYCO has commercialized Bollgard II<sup>TM</sup>, which contains Bt genes Cry1Ac and Cry2Ab, imparting increased protection against a wider-spectrum of borer pests. A comparative evaluation of varietal diversity status in 2004 and 2006 would yield the supply-side impacts of technology provision, whereas a cross-sectional evaluation of cotton growing households would divulge the demand side impacts of technology adoption on varietal diversity.

#### IV. DATA

The south and central Indian states contribute to more than 60 per cent of the Indian cotton production (Qaim et al., 2006). When Bt cotton technology was legally introduced for the first time in India in 2002, it was limited to this region, and there was a rapid adoption of the technology here (Gandhi and Namboodiri, 2006). The data for the present study is based on the cotton farmers from the four states of south and central India, viz. Maharashtra, Karnataka, Andhra Pradesh and Tamil Nadu. Personal interview method was employed for data collection, with the help of a pre-tested questionnaire. Details on land allocation under different crops and cotton varieties, cost and returns from cotton cultivation, input use pattern, and farm household characteristics were collected. The purpose was to create a panel database to study the dynamics of cotton production after the introduction of transgenic technology.

Though the farm survey was conducted in three phases at an interval of two years starting from the first year of commercialization, this paper limits itself to the last two phases – of 2004 and 2006 – as detailed information on varietal use was made available for these phases only. All the surveys were conducted at the end of the cotton seasons of the respective years. In each of these phases,

data was collected by contacting the same farmers of the dataset gathered in year 2002. In addition, new households were included to rectify the missing ones in every round. Though all possible efforts were made to ensure the continuity by including all the samples of the first phase, there were some omissions due to reasons like demise of farmer, migration etc. Further, for the present analysis, we excluded the data on households that discontinued cotton cultivation in the respective years and hence reached the final sample size of 362 and 344 in the 2004 and 2006 surveys. However, input-output data was collected for both Bt and non-Bt varieties from farmers following partial adoption of the technology, making the sample size for production function analysis 464 and 369 during the respective years.

## V. ANALYTICAL FRAMEWORK

The analytical part consists of per-plot production function and per-farm Tobit model on factors contributing to varietal richness estimations. On-farm varietal diversity enters as one of the explanatory variables in production function, whereas it is the dependent variable in Tobit model. As simple count of varieties grown per-farm would not control for the scale of measurement (Nagarajan et al., 2007), we have employed Margalef varietal richness index (MVRI), which is a count of cotton varieties normalized by the scale of the cotton area. Following Magurran (1988),

$$MVRI_i = \frac{N_i - 1}{\ln(A_i)} \quad [1]$$

where  $N_i$  and  $A_i$  respectively are the number of cotton varieties cultivated and area under cotton on the  $i^{\text{th}}$  farm. MVRI as the diversity measure is widely used in literature (cf. Di Falco et al., 2007, and Natarajan et al., 2007). Unavailability of data on area under individual varieties was the main reason for not resorting to farm evenness indices for calculating the varietal diversity.<sup>1</sup> In order to maintain continuity in  $MVRI_i$  estimates, we have taken the unit of land area as 0.10 acres, smaller than the smallest farm size observed in the survey. About 36 (45) per cent of cotton farmers cultivated a single variety on their farms during 2004 (2006), which corresponds the zero varietal diversity ( $MVRI_i = 0.00$ ). The mean  $MVRI_i$  value was 0.33 in 2004, which has come down to 0.26 in 2006. Unsurprisingly, larger farms were cultivated with greater number of varieties.

---

<sup>1</sup> A close association between farm evenness and richness, as observed globally by Jarvis et al. (2008), is indicative of the fact that the empirical findings would not change drastically if the evenness indices are used instead of richness.

For production function estimation, the generalized Cobb-Douglas model was employed, with the model specification of:

$$Y_i = \alpha + \sum_i \beta_i x_i^c + \sum_j \gamma_j x_j^p + \sum_k \delta_k D_k + \varepsilon \quad [2]$$

where  $Y_i$  is the per-acre yield of cotton and  $x_i^c$  is the cost incurred on conventional inputs (variables 2-4, and 6 in Table 2) all in logarithmic terms. The term  $x_j^p$  represents the damage control inputs (variables 5, 11-13).  $D_k$  is vector of household, geographic and time factors (variables 7-10, 15-19), and  $\varepsilon$  is random error term with mean zero. The descriptive statistics of cotton farming is provided for each year in Table 2. Due to multiple observations from a single farm, clustering is followed.

To mitigate the econometric problem with endogeneity of certain RHS variables, an instrumental variable approach is used for the variables viz. Bt adoption and pesticide cost. Instead of including these variables directly in the production function the predicted values are used, for which taluk (sub-district) average pesticide price and Bt/Non-Bt seed price ratio were used as instruments. In principle, similar problem might arise with the other inputs also, especially the on-farm varietal richness. Due to the unavailability of suitable instrumental variables, the effect of removing MVRI from production function was tested. Since the removal had only insignificant impact on remaining coefficients, it was concluded that there exist no serious correlation with the error term.

In spite of the popularity of Cobb-Douglas specification, there exists certain limitations in using this form in our present analysis due to the facts that (i) damage control inputs could differ fundamentally from other conventional inputs, and treating them alike could lead to biased estimates (ii) the focus of the study is also to estimate the factors resulting in yield variability. To address these problems, Damage Control function postulated by Lichtenberg and Zilberman (1986) is combined with the econometric framework proposed by Just and Pope (1978; 1979). Just-Pope approach provides a method for estimating the marginal risk effects of explanatory variables, and is used widely in associated literature (e.g., Traxler et al., 1995; Koundouri and Nauges, 2005). The production function form is:

$$Y = f(x_i^c, D_k)G(x_j^p) + h(x_i^c, D_k, x_j^p)v \quad [3]$$



Here  $f(.)G(.)$  is the mean function and  $h(.)$  is the risk function and we assume  $E(v) = 0$  and  $V(v) = 1$ . The mean function can be elaborated as:

$$f(x_i^c, D)G(x_j^p) = \alpha + \sum_i \beta_i x_i^c + \sum_k \delta_k D_k + g(x_j^p) + \varepsilon \quad [4]$$

Here  $g(.)$  is the natural logarithm of  $G(.)$ . Upon testing different specifications of  $g(.)$ , the logistic form was selected:

$$g(x_j^p) = \left[ 1 + \exp \left( \mu - \sum_j \theta_j x_j^p \right) \right]^{-1} \quad [5]$$

Equation [4] is estimated using non-linear methods, and predicted the vector of error terms,  $\varepsilon$ . Following Just and Pope (1979), Kumbhakar and Tveteras (2003) and Koundouri and Nauges, (2005), the variance is modeled as a Cobb-Douglas form of explanatory variables.

$$E(\varepsilon^2) = \sigma^2 = h(x_i^c, D, x_j^p) \text{ and}$$

$$\ln(\sigma^2) = \zeta + \sum_i \lambda_i x_i^c + \sum_j \xi_j x_j^p + \sum_k \tau_k D_k + \ln(v) \quad [6]$$

The variables are as defined before. The result permits asymptotically valid hypothesis tests about marginal risk effects. Following Traxler et al. (1995), predicted values from equation [9] are used as weights for generating generalized least squares (GLS) estimators for the mean output function.

The second part of the estimation deals with factors contributing to the on-farm varietal richness. Due to the large number of zero observations in the dependent variable, Tobit model (Tobin, 1958) was employed instead of ordinary least square model. The observed variable, MVRI, is defined to be equal to the latent variable whenever the latent variable is above zero and zero otherwise. Quadratic specification is assumed for the explanatory variables, viz. share of cotton area under Bt varieties and irrigation, and farm size owned by the household, and linear for the rest.

## VI. RESULTS AND DISCUSSION

There is a significant change in the pattern of Bt technology adoption between the years: percentage of non-adopters of technology has declined drastically from 54 in 2004 to 8 in 2006 amongst the sampled farmers, and correspondingly increased the full adopters' share from 16 to 74 per cent. About 30 per cent of farmers resorted to partial adoption in 2004, and their share has reduced to 18 per cent during the two year period. This change could be attributed to the reduction in seed price of Bt cotton through government regulations, and increased supply of GM varieties, suitable for various agro-climatic conditions. The next sub-section discusses the yield impacts of Bt adoption and on-farm varietal diversity, and the subsequent section analyzes the impact of Bt adoption on varietal richness.

### VI.1. Yield impacts of Bt technology and on-farm varietal diversity

A cross-sectional comparison across the groups of farms with respect to number of varieties shows that there is a perceivable decline in coefficient of variation (CV) in yield with increase in MVRI (Table 1). At the same time, the impact of diversity on mean yield is ambiguous. Reduction in yield variability by on-farm varietal diversity are reported also by Smale et al. (1998) and Di Falco and Perrings (2003). Secondly, an inter-year comparison shows a significant increase in the cotton yield and a decline in CV in each variety class in 2006 as compared to 2004 data. Drastic increase in Bt cotton adoption could be one of the factors behind these changes. Though the positive yield impacts of Bt technology are widely discussed in the literature, its effect in reducing yield variability is yet to be subjected to in-depth analyses. The only existing empirical evidence on Indian production sector in our best knowledge (Croston and Shankar, 2008) supports the yield reduction function of Bt technology adoption.

With the objective of estimating the individual impacts of various production inputs on both mean and variance of cotton yield, GLS production functions were estimated (Table 3). The production inputs viz. chemical fertilizers, human labor, and irrigation are positively associated with yield, whereas farm yard manure and insecticide show negative yield effects. None of the household variables show significant impact on mean yield. We limit our discussion to Bt adoption and on-farm varietal richness. The Damage Control function provides comparatively higher marginal values for Bt adoption: the marginal impact of Bt adoption is computed as 4.60 Q in Cobb-Douglas

and 5.67 Q in Damage Control function. The corresponding values are 2.45 Q and 2.46 Q for varietal richness in the respective functions. It is noteworthy that unit increase in varietal richness indicates cultivation of additional 4 varieties at the mean cotton area per farm (5.80 acres). In sum, though both Bt adoption and varietal diversity are yield enhancing, when there are few Bt varieties available for cultivation, yield maximizing farmer would rather go for Bt adoption over maintenance of on-farm diversity. Nonetheless, production efficiency can be maximized by combining varietal richness with adoption of Bt technology, especially since the interaction term of Bt is positive (albeit statistically insignificant at 0.10 level). This is feasible only when the seed industry supplies multitude of locally-adapted Bt varieties.

In order to test the hypothesis on impact of on-farm varietal diversity and adoption of PR-GM technology on yield variability, variance function is estimated and presented in Table 3 (Col: iii). Chemical fertilizers and irrigation reduce whereas insecticide use enhances yield variability and hence the production risk. Household size, which acts as proxy for family labor availability, is also risk-reducing. Age of household head is found positively associated with variability. Varietal richness is found risk-reducing. This result validates the argument that one of the reasons for farmers to maintain on-farm varietal diversity is the insurance function it provides against potential yield loss from biotic and abiotic stresses. However, the impact of Bt adoption is much pronounced. In other words, technology based insurance is more efficient over diversity based, in the farmers' perspective. If the prime rationale behind conserving on-farm richness is risk-aversion for farmer, then adoption of Bt technology could result in drastic reduction in varietal diversity. However, since the magnitude of Bt-richness interaction is great (despite being statistically insignificant), the best strategy for risk minimization would be maintaining diversity among Bt varieties. Again, this possibility is dependent on a number of institutional factors.

## VI.2. Bt adoption and on-farm varietal richness

The previous sub-section dealt with the role of technology and diversity in production relations, and now the article proceeds by examining the inter-relationship between these factors. Quadratic trend lines are fitted to show the association between adoption of Bt technology and on-farm varietal count and richness in cotton farms of India (Figure 1 a & b). They suggest that, as a general rule, partial adopters of Bt technology maintained greater number of varieties as well as varietal richness on-farm, in comparison with both non- and full adopters. This effect is more prominent in year

2004, for which farmers' uncertainty about the yield impact of limited available Bt cotton varieties could be one of the major reasons. Nonetheless, as the share of partial adopters drastically reduces over time (30% in 2004 to 18% in 2006), *de facto* conservation efforts by this group cannot be given greater significance in policy formulations. More importantly, the impact of full adoption over non-adoption is significantly different across years. In 2004, the count and richness of varieties were low amongst the full adopters, whereas in 2006, the impact of full adoption on these parameters was opposite in direction. The first conclusion may be that, as suggested by Zilberman et al. (2007), increase in supply of GM varieties may have resulted in the positive association between adoption of GM varieties and varietal richness. On the other hand, the varietal count and richness have declined in farms with share Bt adoption less than 60 per cent over the two year period. The increased corporate focus on Bt seed production, and subsequent shortage in supply of non-Bt varieties could be one of the reasons for this reduction. Realizing the complex interaction of farm factors other than Bt adoption and household characteristics in determining the on-farm diversity, we proceed by employing more sophisticated econometric tools to the dataset.

Due to the large number of farmers following monoculture of varieties (MVRI = 0), the ordinary least square (OLS) estimation would lead to biased estimates (Greene, 2000). Hence, we resort to the Tobit model. The description of explanatory variables employed in the analysis is given in Table 2 (Col: ii). Along with farm variables viz. Bt adoption, farm size owned by the household, and share of cotton area under irrigation, the household characteristics are included in the model estimation. Since farm size is positively correlated with cotton area, a positive association with varietal richness would be unsurprising. As in the case of Bt adoption, a negative association of irrigation and varietal diversity is anticipated, as irrigation could substitute on-farm diversity as instrument for risk reduction.

The estimation results are presented in Table 4, which show that larger the farm area owned, greater would be varietal richness, and this positive relation holds for group of farms with acreage less than 25 acres, which includes 88 per cent of the sample. The decline of diversity in the largest 12 per cent farms could possibly be due to the risk-bearing ability of the wealthy households. Regarding irrigation facilities, the varietal richness is greatest in the farms with partial adoption of irrigation facilities, where cultivation of both drought-prone and tolerant varieties is possible. Low diversity status of rainfed farms could be due to the limited availability of drought-tolerant varieties. The statistical insignificance of other farm and household variables limits the insights into their patterns

of influence. The varietal diversity is shown as greatly determined by the geographic location, as indicated by the dummy variables for states.

However, the prime focus of the model is impact of Bt adoption on varietal richness. Due to the possible endogeneity associated with the technology adoption decision, instrumental variable approach is resorted to, as in the case of production function estimation. As shown in Figure 1 (b), the relationship is of quadratic form – the varietal richness increases at first, reaches the maximum when Bt adoption is partial (about 40% cotton area under Bt) and then declines with adoption. Overall, the marginal impact of Bt cotton adoption on varietal richness, which can be defined as the change in MVRI index corresponding to the full adoption of Bt cotton over non-adoption, is negative in year 2004 and positive in 2006. The possible trade-off between varietal diversity and transgenic technology in farmers' perspective could possibly be attributed for this effect. The coefficient of the interaction term between year dummy and share of area under Bt is indicative of a positive supply effect. A decrease in varietal richness in year 2006 among non-Bt farmers is observed, which could be the result of decline in supply of non-Bt varieties by the seed firms (by shifting seed market focus to more remunerative GM cultivars) in the corresponding year.

Since the supply of Bt varieties crucially determines the diversity, the limited supply of GM varieties would cost the agricultural sector in both direct and indirect manner. Wide adoption of few Bt varieties implies that at many locations agro-climatically incompatible varieties would be cultivated. The yield loss due to location-variety incompatibility is the direct impact of low diversity in Bt seed supply. This can be rectified through development of local GM varieties. The indirect impact is through the drastic decline in on-farm varietal richness with technology adoption. Richness is found as one of the production enhancing, risk reducing factors, and a reduction in it lead to economic welfare of farming community.

## VII. CONCLUSIONS

Biodiversity impact of transgenic technology is multi-dimensional. This paper has examined the particular aspect of on-farm crop varietal diversity conservation, taking the case of Bt cotton in Indian agriculture for the empirical evaluation. A combination of Damage Control and Just-Pope production function analyses on farm level data is carried out, in order to test the hypothesized adverse demand effect of PR-GM technology on diversity. Both Bt technology and varietal richness

are found to increase the mean yield and reduce yield variability, with the former being more efficient. Hence, the farmers' economic welfare can greatly be enhanced by ensuring diversity amongst the GM varieties. On the other hand, when farmers' motive for maintaining diversity is risk-aversion from pest infestation, and they realize the potential of Bt technology for the same purpose, then the on-farm diversity impact of technology adoption would invariably be negative. In other words, one of the factors behind the technology-diversity trade-off is that PR-GM technology could effectively provide a natural insurance against pests, which is also a function of agrobiodiversity, ultimately resulting in low yield variation.

The rate at which adoption of transgenic technology eliminates on-farm diversity depends on a multitude of factors. This paper particularly underlines the importance of the supply side, represented by the seed sector. Release of GM varieties by the seed industry is postulated as positively correlated with on-farm varietal diversity. Availability of two-year data before and after the expansion of GM seed industry enabled us to observe these supply-side impacts. The estimation results show that, in year 2004 when the supply of GM varieties was limited, the varietal diversity shrunk with full adoption of the technology. However, once the industry got expanded, the impact of Bt adoption turned marginally positive. At the same time, the varietal diversity has declined drastically among the non-adopters of technology, possibly due to decline in supply of non-Bt seeds by the firms. These results vividly support the study by Zilberman et al. (2007), which suggests the role of the existing institutional framework on ensuring diversity with technology development.

In sum, our findings complement those of Smale (1997) and Smale et al. (1998) on bread wheat cultivation, which indicated that, due to the very interplay of a multitude of factors, any causal relationship between technology adoption during Green Revolution and genetic erosion cannot be established. Here we extend her conclusions to Gene Revolution in cotton production sector of India. Since the agronomic impacts of GM crops could be potentially greater in magnitude, the concerned government policies regulating the introduction and dissemination of GM varieties should be more efficient. Development of transgenic technology prompts firms to release more varieties utilizing this technology. This increase in supply could partially ameliorate the adverse demand effect. Nevertheless, ensuring diversity among non-GM seed sector is also required in order to maintain diversity among all the farmer groups during the process of technology diffusion. Once the GM crop gets regulatory approval, release of more varieties at lower cost should be enhanced in order to reduce its negative impacts due to the declining farmer demand for

agrobiodiversity. Strong intellectual property rights regime in the seed sector and high biosafety regulation and/or transaction costs could hamper the firms from entering the GM seed market. In situations where supply of GM/non-GM varieties is restricted, possibly due to the informal structure of seed sector, government should take up direct agro-biodiversity conservation activities.

## REFERENCES

- Ammann, K., Effects of biotechnology on biodiversity: herbicide-tolerant and insect-resistant GM crops. *TREDS in Biotechnology*, 23(8):388-394.
- APCoAB, 2006, Bt cotton in India: a status report. Asia-Pacific Consortium on Agricultural Biotechnology, New Delhi, India.
- Batie, S.S. and Ervin, D.E., 2001, Transgenic crops and the environment: missing markets and public roles. *Environment and Development Economics*, 6: 435-457.
- Birol, E., Smale, M. and Gyovoi, A., 2005, Farmer management of agricultural biodiversity in Hungary's transition economy. In Smale, M. (ed), *Valuing crop biodiversity: on-farm genetic resources and economic change*. Wallingford, UK: CAB International Publishing.
- Birthal, P.S., Sharma, O.P., Kumar, S. and Dhandapani, A., 2000, Pesticide use in rainfed cotton: frequency, intensity and determinants. *Agricultural Economics Research Review*, 13(2): 107-122.
- Brush, S.B. 2000. The issues of *in situ* conservation of crop genetic resources. In Brush, S.B., (ed.) *Genes in the Field: On-Farm Conservation of Crop Diversity*. IDRC/IPGRI/Lewis Publishers.
- Chapin, F.S., Walker, B.H., Hobbs, R.J., Hooper, D.U., Lawton, J.H., Sala, O.E., and Tilman, D., 1997, Biotic control over the functioning of the ecosystem. *Science*, 277(5325): 500-504.
- Crost, B., Shankar, B., Bennett, R. and Morse, S., 2007, Bias from farmer self-selection in genetically modified crop productivity estimates: evidence from Indian data. *Journal of Agricultural Economics*, 58(1): 24-36.
- Crost, B., and Shankar, B., 2008, Bt cotton and production risk: panel data evidence. *International Journal of Biotechnology*, 10 (2/3): 122-131.
- Dale, P.J., Clarke, B., Fontes, E.M.G., 2002, Potential for the environmental impacts of transgenic crops. *Nature Biotechnology* 20, 567 – 574.
- Di Falco, S. and Perrings, C., 2003, Crop genetic diversity, productivity and stability of agroecosystems. A theoretical and empirical investigation. *Scottish Journal of Political Economy*, 50(2): 207-216.
- Di Falco, S., Chavas, J-P., Smale, M., 2007, Farmer management of production risk on degraded lands: the role of wheat diversity in the Tigray region, Ethiopia. *Agricultural Economics*, 36: 147-156.
- Gandhi, V.P. and Namboodiri, N.V., 2006, The adoption and economics of Bt cotton in India: preliminary results from a study. Working paper number 2006-09-04, Indian Institute of Management, Ahmedabad, India.
- Greene, W.H., 2000, *Econometric analysis*, 4<sup>th</sup> edition. Prentice-Hall Inc., New Jersey.
- Guere, G.P., Mehta-Bhatt, P., Sengupta, D., 2008, Bt cotton and farmer suicides in India: reviewing the evidences. Environment and Production Technology Division, IFPRI Discussion Paper 00808, International Food Policy Research Institute, Washington DC.
- Hyde, J., Martin, M.A., Preckel, P.V. and Edwards, C.R., 1999, The economics of Bt corn: valuing protection from the European corn borer. *Review of Agricultural Economics*, 21:442-454.
- Jackson, L.E., Pascual, U., and Hodgkin, T., 2007, Utilizing and conserving agrobiodiversity in agricultural landscapes. *Agriculture, Ecosystems and Environment*, 121: 196-210.
- James, C., 2006, Global status of commercialized biotech/GM crops: 2006. ISAAA Briefs, No. 35. International Service for the Acquisition of Agri-biotech Applications, Ithaca, NY.
- Jarvis, D.I., Brown, A.H.D., Cuong, P.H., Collado-Panduro, L., Latournerie-Moreno, L., Gyawali, S., Tanto, T., Sawadogo, M., Mar, I., Sadiki, M., Hue, N, T., Arias-Reyes, L., Balma, D., Bajracharya, J., Castillo,

- F, Rijal, D., Belqadi, L., Rana, R., Saidi, S., Ouedraogo, J., Zangre, R., Rhrib, K., Chavez, J.L., Schoen, D., Sthapit, B., De Santis, P., Fadda, C. and Hodgkin, T., 2008, A global perspective of the richness and evenness of traditional crop-variety diversity maintained by farming communities. *Proceedings of the National Academy of Sciences (PNAS)*, 105(14): 5326-5331.
- Just, R.E and Pope, R.D., 1978, Stochastic representation of production functions and econometric implications. *Journal of Econometrics*, 7(1): 67-86.
- Just, R.E and Pope, R.D., 1979, Production function estimation and related risk considerations. *American Journal of Agricultural Economics*, 61(2): 276-284.
- Kinzig, A., Pacala, S. and Tilman, D., 2002, *The functional consequences of biodiversity: empirical progress and theoretical extensions*. Princeton, NJ: Princeton University Press.
- Koundouri, P. and Nauges, C., 2005, On production function estimation with selectivity and risk considerations. *Journal of Agricultural and Resource Economics*, 30(3): 597-608.
- Kumbhakar, S.C. and Treterås, R., 2003, Risk preferences, production risk and firm heterogeneity. *Scandinavian Journal of Economics*, 105(2): 275-293.
- Lichtenberg, E. and Zilberman, D., 1986, The economics of damage control: why specification matters. *American Journal of Agricultural Economics*, 68: 262-273.
- Magurran, A., 1988, *Ecological density and its measurement*. Princeton University Press, Princeton, NJ.
- Natarajan, L., Smale, M. and Glewwe, P., 2007, Determinants of millet diversity at the household-farm and village community levels in the drylands of India: the role of local seed systems. *Agricultural Economics*, 36: 157-167.
- Perrings, C. and Gardgil, M., 2002, Conserving biodiversity: reconciling local and global public benefits. In Kaul (ed.), *Providing global public goods: making globalization work for all*. Oxford University Press, UK.
- Qaim, M., 2005, Agricultural Biotechnology Adoption in Developing Countries. *American Journal of Agricultural Economics*, 87(5): 1317-1324.
- Qaim, M. and Zilberman, D., 2003, Yield effects of genetically modified crops in developing countries. *Science* 299: 900-902.
- Qaim, M., Subramanian, A., Naik, G. and Zilberman, D., 2006, Adoption of Bt cotton and impact variability: insights from India. *Review of Agricultural Economics*, 28 (1), 48–58.
- Smale, M., 1997, The Green Revolution and wheat genetic diversity: some unfounded assumptions. *World Development*, 25(8): 1257-1269.
- Smale, M., Hartell, J., Heisey, P.W. and Senauer, B., 1998, The contribution of genetic resources and diversity to wheat production in the Punjab of Pakistan. *American Journal of Agricultural Economics*, 80: 482-493.
- Thirtle, C., Beyers, L., Ismail, Y., and Piesse, J., 2003, Can GM technologies help the poor? The impact of Bt cotton in Makhathini Flats, KwaZulu-Natal. *World Development*, 31(4): 717-732.
- Thomson, J.A., 2006, *GM crops: the impact and the potential*. Collingwood, Australia: CSIRO Publishing. pp:65-77.
- Tobin, J., 1958, Estimation for relationships with limited dependent variables. *Econometrica*, 26 (1): 24–36
- Traxler, G., Falck-Zepeda, J.I, Ortiz-Monasterio, R., and Sayre, K., 1995, Production risk and the evolution of varietal technology. *American Journal of Agricultural Economics*, 77: 1-7.
- Zilberman, D., Ameden, H. and Qaim, M., 2007, The impact of agricultural biotechnology on yields, risks, and biodiversity in low-income countries. *Journal of Development Studies*, 43(1): 63-78.



TABLE 1. On-farm varietal richness and cotton yield

Year	No. of varieties/farm	% of farms	MVRI	Share of Bt adoption	Cotton yield [Q/acre]*	CV in yield
2004 [N =362]	1	36.19	0.00 (0.00)	0.42 (0.49)	5.45 (3.80)	0.70
	2	31.49	0.28 (0.05)	0.21 (0.31)	5.79 (3.75)	0.65
	3	14.09	0.51 (0.07)	0.21 (0.25)	5.81 (2.66)	0.46
	4	7.18	0.69 (0.07)	0.19 (0.23)	6.12 (2.49)	0.41
	5	3.87	0.92 (0.10)	0.19 (0.21)	5.93 (2.45)	0.41
	> 5	7.18	1.18 (0.25)	0.13 (0.16)	4.14 (1.66)	0.40
	Overall	100.00	0.33 (0.36)	0.28 (0.38)	5.58 (3.40)	0.61
2006 [N =343]	1	44.90	0.00 (0.00)	0.83 (0.37)	7.27 (3.85)	0.53
	2	26.53	0.28 (0.04)	0.88 (0.24)	8.31 (3.47)	0.42
	3	14.87	0.50 (0.06)	0.85 (0.23)	8.96 (3.29)	0.37
	4	8.45	0.72 (0.09)	0.82 (0.23)	8.59 (2.13)	0.25
	5	4.37	0.84 (0.08)	0.89 (0.17)	8.45 (2.87)	0.34
	> 5	0.87	1.29 (0.21)	0.76 (0.21)	8.71 (2.06)	0.24
	Overall	100.00	0.26 (0.29)	0.85 (0.30)	7.97 (3.54)	0.44

Figures in parentheses show standard deviation (SD)

N stands for the sample size. 1 Quintal (Q) = 100kg

TABLE 2. Summary statistics

Variables	Description (units)	i. Per plot		ii. Per farm	
		2004 [N = 464]	2006 [N = 369]	2004 [N = 362]	2006 [N = 343]
1. Yield	Cotton yield (Q/acre)	6.19 (3.18)	8.13 (3.56)	5.58 (3.40)	7.97 (3.54)
2. Manure	Cost of farm yard manure (Rs/acre)	277.56 (803.77)	868.62 (1877.50)		
3. Fertilizer	Cost of chemical fertilizers (Rs/acre)	1745.59 (1251.05)	1883.86 (2801.37)		
4. H.Labor	Cost of human labor employed (Rs/acre)	1808.10 (1047.06)	2755.23 (2062.26)		
5. Insecticide	Cost of insecticides used (Rs/acre)	2115.65 (1812.13)	1166.80 (1675.77)		
6. Other	Cost of other production inputs (Rs/acre)	104.13 (1206.34)	196.57 (846.01)		
7. Farm size	Farm size owned by the household (acres)	14.88 (17.41)	12.31 (13.55)	13.57 (16.21)	12.04 (13.43)
8. Household	Number of members in the household	6.79 (4.02)	6.26 (4.05)	6.63 (3.77)	6.18 (4.07)
9. Age	Age of head of the household (years)	44.96 (12.75)	44.68 (13.11)	44.87 (12.77)	44.64 (13.03)
10. Education	Years of schooling obtained by head of the household	7.42 (4.90)	6.82 (5.46)	7.27 (4.89)	6.85 (5.46)
11. Irrigation	Share of cotton area under irrigation (0-1)	0.39 (0.46)	0.51 (0.48)	0.39 (0.46)	0.52 (0.48)
12. Richness	Margalef Varietal Richness Index	0.39 (0.36)	0.27 (0.29)	0.33 (0.36)	0.26 (0.29)
13. Bt	Dummy for Bt cotton	0.36	0.85		
14. Bt share	Share of cotton area under Bt cotton per farm (0-1)			0.28 (0.38)	0.85 (0.30)
15. Year	Dummy for year of 2006		0.44		0.49
16. Maharashtra	Dummy [ <i>reference</i> ] if farmer is from Maharashtra	0.33	0.32	0.33	0.32
17. Karnataka	Dummy if farmer is from Karnataka	0.29	0.29	0.31	0.29
18. AP	Dummy if farmer is from Andhra Pradesh	0.31	0.35	0.30	0.35
19. TN	Dummy if farmer is from Tamil Nadu	0.07	0.04	0.08	0.04

Figures in parentheses show standard deviation of sample mean. N stands for the sample size.

1 US\$ = Rs. 48 (in June 2009)

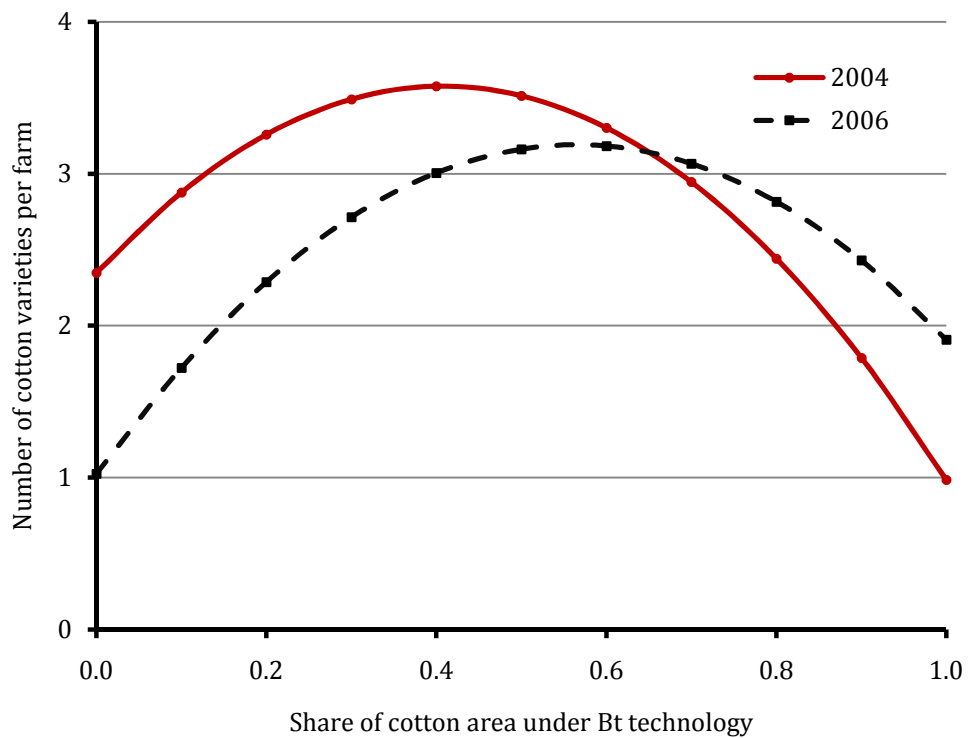
TABLE 3. Production function analysis

Variables	i. Cobb-Douglas			ii. Damage Control			iii. Yield variance		
	Coefficient	Std. Err.	p-value	Coefficient	Std. Err.	p-value	Coefficient	Std. Err.	p-value
<i>ln</i> (Manure)	-0.027	0.007	0.00	-0.026	0.007	0.00	0.019	0.028	0.49
<i>ln</i> (Fertilizer)	0.023	0.020	0.27	0.028	0.017	0.09	-0.287	0.101	0.01
<i>ln</i> (H.Labour)	0.036	0.016	0.03	0.033	0.016	0.04	-0.107	0.075	0.16
<i>ln</i> (Other)	0.010	0.011	0.40	0.009	0.011	0.39	0.043	0.042	0.31
<i>ln</i> (Insecticide)	-0.063	0.052	0.23				0.630	0.217	0.00
Irrigation	0.352	0.063	0.00				-0.493	0.239	0.04
Bt	0.567	0.183	0.00				-2.230	0.813	0.01
Richness	0.255	0.094	0.01				-0.930	0.382	0.02
Bt*Richness	0.227	0.217	0.30				-0.784	0.600	0.19
<i>ln</i> (Farm size)	-0.026	0.021	0.22	-0.029	0.020	0.16	0.063	0.090	0.49
<i>ln</i> (Household size)	-0.013	0.056	0.81	8.E-08	0.052	1.00	-0.579	0.227	0.01
<i>ln</i> (Age)	-0.026	0.079	0.75	-0.032	0.079	0.69	0.481	0.287	0.09
<i>ln</i> (Education)	0.031	0.025	0.21	0.028	0.024	0.24	-0.043	0.097	0.66
Karnataka	0.142	0.062	0.02	0.146	0.061	0.02	0.157	0.268	0.56
AP	0.356	0.109	0.00	0.409	0.094	0.00	-1.522	0.415	0.00
TN	-0.475	0.161	0.00	-0.494	0.119	0.00	1.040	0.682	0.13
Year	-0.308	0.145	0.03	-0.476	0.167	0.01	1.593	0.692	0.02
Damage control $\mu$				0.041	0.565	0.94			
<i>ln</i> (Insecticide)				0.155	0.066	0.02			
Irrigation				-0.698	0.156	0.00			
Bt				-1.713	0.563	0.00			
Richness				-0.453	0.164	0.01			
Bt*Richness				-1.078	1.255	0.39			
Intercept	1.449	0.383	0.00	2.279	0.433	0.00	-3.580	1.473	0.02
Adjusted R <sup>2</sup>	0.162			0.167			0.058		

TABLE 4. Determinants of on-farm varietal richness

		Tobit model			Test statistic		
		Coefficient	Std. Err	p-value	H <sub>0</sub>	F(2, 691)	p-value
Bt share	[ $\beta_1$ ]	0.617	0.403	0.13	$\beta_1 = \beta_2 = 0$	4.85	0.01
Square of Bt share	[ $\beta_2$ ]	-1.297	0.529	0.02			
Farm	[ $\beta_3$ ]	0.016	0.003	0.00	$\beta_3 = \beta_4 = 0$	15.43	0.00
Square of Farm	[ $\beta_4$ ]	-1.E-04	4.E-05	0.00			
Irrigation	[ $\beta_5$ ]	0.751	0.217	0.00	$\beta_5 = \beta_6 = 0$	6.95	0.01
Square of Irrigation	[ $\beta_6$ ]	-0.784	0.214	0.00			
Household		-0.003	0.004	0.48			
Age		-0.002	0.001	0.18			
Education		3.E-04	0.004	0.94			
Karnataka		-0.470	0.050	0.00			
AP		-0.056	0.044	0.20			
TN		-0.572	0.090	0.00			
Year		-1.639	0.481	0.00			
Year*Bt share		2.312	0.694	0.00			
Intercept		0.320	0.107	0.00			
$\sigma$		0.386	0.015	0.00			
Log-pseudo-likelihood ratio		-385.46					
F(14, 691)		18.85					
Prob. > F		0.000					
Pseudo R <sup>2</sup>		0.280					

FIGURE 1. Impact of Bt cotton adoption on:  
a. Number of varieties on-farm



b. Varietal richness index

