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by<br>Richard J. Foote

In my opinion, much needless confusion exists in the minds of economists and statisticians when they think about least squares versus simultaneous equation techniques. Some analysts believe that the method of least squares now is completely outmoded; others feel that simultaneous equation methods ore so complex and computationally expensive that they should be avoided whenever possible. Each of these viewpoints is wrong. Simultaneous equation techniques are a useful addition to our kit of tools for use in probiems thet deal with regression analysis. When they are needed, they should be used, just as a hack saw is used to cut metal, whereas various woodsaws can be used to cut wocd. In systems of equations, one, several, or perhaps ail frequently can be fitted by least squares. Moreover, least squares equations are useful now, just as they always have been, in showing normal or average relationships that exist between sets of variables. Many problems that relate to regression analysis, such as choice of variables, location of data, choice of functional forms, and the testing and interpretation of results, are almost identical regardless of whether the equations are to be fitted by least squares or by simultaneous equation tecinniques.

Because of the confusion that exists, I propose to start with some extremely elementary concepts with which $I$ am sure you are all familiar. We shall then proceed step by step in such a way as to show precisely when and why simultaneous equation techniques are needed. In later sections, I shall discuss some computational aspects of regression analysis and some considerations relating to the degree of complexity that-may be desirable in formulating a system of economic relationships.

## Some Economic Considerations

In 1927, Elmer Working gave an excelleat discussion of what now is called the identification problem in his classic paper "What Do Statistical 'Demand Curves' Show?'. I/ He pointed out that when a research worker begins a demand study, he is confronted with a set of dots like that shown in section $A$ of the chart on p. 3. He knows that each can be thought of as the intersection of a demand and a supply curve, as in section $B$, but, without further information, neither curve can be determined from the data. Working then noted that if the demand curve has

[^0]shifted over time but the supply curve has remained relatively stable, as in section $C$, the dots trace out a supply curve; conversely, if the supply curve has shifted but the demand curve has remained stable, as in section $D$, the dots trace out a demand curve. If correlated shifts for each curve have taken place, as in section $E$, the dots trace out what may look like a structural demand or supply curve, but the slope will be too flat or too steep.

In many analyses of the demand for agricultural products, factors that cause the demand curve to shift over time are included as separate variables in a multiple regression equation. In effect, we are then able to derive from our estimating equation an average demand curve. This is indicated in a rough way in section $F$. In some analyses, we can assume that the quantity supplied is essentially unaffected by current price. When price is plotted on the vertical scale, the supply curve in such cases is a vertical line, and year-to-year shifts in the supply curve trace out a demand curve, just as they did in section D. Under these circumstances we may be able to obtain valid estimates of the elasticity of demand by use of a least squares multiple regression analysis for which price is the dependent variable and supply and some demand shifters are used as independent variables. This point was noted by Working in his 1927 paper $2 /$, emphasized by Ezekiel in a paper published in $19283 /$, and reconsidered in 1953 by Fox 4/ in the light of modern simultaneous equations theory. For many agricultural products, this set of circumstances permits us to estimate elasticities of demand with respect to price by use of single equation methods. Two points, however, should be kept in mind: (1) Price must be used as the dependent variable in order to obtain elasticity estimates that are statistically consistent since, to use the least squares technique, the supply curve must be a vertical line; and (2) an algebraic transformation must be made after the equation has been fitted to derive the appropriate coefficient of elasticity, since the definition is in terms of the percentage change in quantity associated with a given percentage change in price. Other circumstances under which least squares equations can be used to derive coefficients of elasticity are discussed in a later section.

What happens if we have a supply curve that is not a vertical line? If we consider any single point, as in section $G$, we have no way of know-: ing on which demand and supply curve of a whole family of curves it lies. The basic problem of indeterminateness is similar to that in which correlated shifts in the demand and supply curves take place. What is reeded... is some hypothesis, adequately tested and proven to be sound, as to the nature of the joint relationships between supply and demand. We should then be able to untangle the two and to obtain a reliable estimate of the slope of each curve. This is essentially what is involved in the simultaneous equations approach. This concept was set forth by Haavelmo in 1943. 5/ Staff meinoers of the Cowles Commission spent a considerable part of the next 10 years in showing how to implement it when working with actual data.

[^1]
## SUPPLY - DEMAND RELATIONSHIPS



Suppose, however, that the analyst has no interest in the true demand and supply curves but only wants a methoci that will assist him in studying probsble future trends in prices. Working had some suggestions on this point, too. He said, "It does not follow from the forezoing analysis that, when conditions are such that shifts of the supply and ciemand curves are correlated, an attempt to construct a demand curve will give a result that will be useless. Even though shifts of the supply and demand curves are correlated, a curve which is fitted to the points of intersection will be usefui for purposes of price forecasting, provided no new factors ane introduced which did not affect the price during the period of study. Thus, so long as the shifts of the supoly and demand curves remain correlated in the same way, and so long as they shift through approximately the same range, the curve of regression of price upon quantity can be used as a means of estjmating price from quantity. "6/"

The problem here is that the shifts almost never "remain correlated ini the same way" over a sufficiently long period to generate enough data tc fit our equation. In some circuastances, changes in structure are so frequent that multiple regression equations almost always yield low correlations and frequently even "wrong" signs on the coefficients. This is perticularly apparent when we attempt, by the single equation approach, to study factors that affect volume of exports. In other cases, changes In structure are of minor importance and least squares equations may yiela completely satisfactory results in terms of showing relationships that have prevailed between simultaneously-determined economic variables over a considerable period of time. This is frequently true when we study relationships between prices at specifiec locations or at local market, whoiesale, or retail ievels. Even here, however, the analyst shouid examine his results closely, perhaps by plotting the data in scatter cilagrams, to determine whether changes in structure have affected the $e$ elationship. The in-between case is the one that can be dangerous. Here the coefficients may suggest that the analysis is satisfactory; it may, in fact, be of little value for the study of future trends.

Marschak 7/ gives an interesting example of the importance of changes in strucutre on the need for using a complete system of equations. He considers the old problem of taxation of a monopoly. He points out, "Knowledge is useful if it helps to make the best decisions". He considers, among other things, the kinds of knowledge that are useful to guide the firm in its choice of the most proritable output levei. If the tax rate has not changed in the past and is not expected to change, the firm can fit an empirical curve to observed data on output and profits and immediately derive the point of maximum revenue. If the tax rate has not changed in the past but is expected to change in the futire, the firm could, if it so desired, vary its output and profits under the new tax structure and derive a new empirical relation. But this takes time, and substantial losses might occur during the experimental period. If the firm had taken the trouble to derive the strucutral demand and net revenue

[^2]curves, it could immediately have deteritned its most profitable output under the new tax structure. If the tax rate had varied during the initial period, an empirical regression of net revenue on output and the tax rate could have been fitted and used to finc the most profitable output under the new tax structure. In many real. Iife situations, changes in structure are frequent. Hence Marschak concludes: "A theory may appear unnecessary for policy decisions (or forecisting) until a certain structural change is expected or intended. It becomes necessary then. Since it is difficult to specify in advance what structurei changes may be visualized later, it is almost certain that a broad anglysis of economic structure, later to be filled cut in detail according to neeas, is not a wasted effort".

This argument in no way invalidates the use of a single equation to estimate elasticities of demand in those cases in which the supply can be considered as unarfected by current price. In such cases, we may obtain estimates of the structural parameters that can be used in the same way as anvother statistically-valid estimates. Instead, Marschak is arguing that only rarely should the economic analyst be satisfied with a purely empirical fit if he can obtain structural relationships with some additional work.

## Some Statistical Consicerations

We now turn to some statistical considerations that have a bearing on the extent to which we can use least squares to estimate the coefficients in a given equation. In the relationshins diacussed in the preceding section, we have assumed, more or less implicitly, that the points lie exactly on the demana or supply curve... In actual statistical analyses, this is never true, since some variables that cause the curves to shift always are omitted and the precise shape of the curves to be fittited are not known. Thus we assume that we are dealing with stochastic rather than functional relations. A stochastic relation basiceliy is one that includes a set of unexplained residuals or error terms whose direction and magnitude are usually not known exactly for any particular set of calculations, but whose behavior on the average over repeated samples can be described or assumed.

In order to have a concrete example about which to talk, let us consider the following equation:

$$
\begin{equation*}
Y=a+b_{1} Z_{2}+b_{2} Z_{2}+u \tag{I}
\end{equation*}
$$

Here $\underline{Y}$ is the variable for which as estimate is desired, the $\underline{Z}^{\prime}$ s are 2 variables which are known to affect $\underline{Y}$, and $\underline{u}$ is an error term. We assume that for a number of periods we know the value of $\underline{Y}$ and the $\underline{Z}$ 's and we wish to estimate $2, \underline{b}_{1}$, and $\underline{b}_{2}$. We do not know the value of $\underline{u}$ but can estimate it in a rough way for any given period as the difference between the value of $\underline{y}$ computed from the equation and its actual value.

We know that estimates of the regression coefficients will differ for different sets of observations. However, we would like to estimate them in such a way that the average value for a large number of periods or samples equals the value that would be obtained from a similar calculation based on the combined evidence of all possible samples. Estimates of this sort are known statistically as unbiased estimates. We also would like the variation of the estimates about their average or true value to be as small as possible, since under this circunstance we would have more confidence in any single estimate than if we had a large amount of variation. Estimating procedures that give the smallest possible variance are known as best estimates. Despite their name, such estimates possess no more desirable properties than many alternative estimates. So the choice of terminology is unfortunate, but it has become firmly established. In certain circumstances, we may be unabie to obtain best unbiased estimates but may be able to obtain estimates that are consistent and efficient. A consistent estimate is one that is unoiased when we work with all the possible data; it may or may not be biased in small samples. In acturl practice, of course, we never have all possible data, but estimation procedures that give consistent estimates presumably are better even with small samples than are those that are known to be biased even with an infinitely large sample. Efficient estimates are similar to "best" estimates, except that they are known to give the smallest possible variance only when we work with all possible data.

If we use the method of least squares to estimate the coefficients in equation ( 1 ), we obtain best unbiased estimates if the u's and $\underline{Z}$ 's meet certain rather rigid specifications. Some of the specifications that relate to the u's are difficult to state precisely in nonmathematical terms, but essentially they require that the u's follow some (not necessarily a normal) probability distribution, that their average or expected value be zero, that their variance be finite and independent of the $Z$ 's (the latter is the property of homoscedasticity, for those of you who remember your elementary statistics), and, finally, they must be serially independent. When working with economic data, we usually assume that these assumptions hold; but we may test at least the one regarding serial independence of the residuals after we have run the analysis. In some cases, we transform the data to (1) logarithms, which frequently helps to render the varience of $\underline{u}$ less dependent upon the $Z$ 's, or (2) first differences, which when working with economic data frequently tends to reduce the serial correlation in the u's. As an alternative we may, of course, use first differences of logarithms.

In addition there is a specification regarding the $\underline{Z}$ 's that is easily stated but frequently disregarded by economists. To be certain that the least squares approach will give best unbiased estimates, eack Z must be a set of known numbers, in contrast to a random variable. When attempting to obtain elasticities of demand, this is true only in rare instances. The only case of which I can think is the one for which prices are arbitrarily set at certain levels, as in a retail store experiment, and the quantity bought by consumers at these prices are recorded. I know of only one experiment that has been conducted in this way--that for oranges ky Godwin. 8 / The least squares method was developed for use in connection with experiments

8/Godwin, Marshall R. Customer Response to Varying Prices for Florida Oranges. Fla. Agr. Expt. Sta. Bul. 508. 1952.
in the physical sciences, where the independent variables frequentiy are sets of known numbers, or to study relationships between variables, such as heights of fathers and beights of sons, where no. "structural" coefficients are involved.

Econometricians have shown that the least squares approach will give estimates of the regression coefficients that are statistically consistent and efficient, provided the u's meet approximately the same requirements as for the previous case, if the $\underline{Z}$ 's are predetermined variables. A simplified proof of this is given by Klein 9/, although a fairly advanced knowledge of calculus and of the principles of probability given in chapter 2 of his book are required to follow his development. Just as the term "best" as used by statisticians is an unfortunate one, the term "predetermined" as used by econometricians is unfortunate. But as it now is commonly used in the iiterature, it seems advisable to learn what it means so that we can continue to use it. Predetermined variables include those that the analyst takes as given, while endogenous variables are those that are determined simultaneously by the same set of economic forces and, therefore, are to be estimated from the model. The predetermined variables cormonly are divided into exogenous variables and lagged values of endogenous variables. Lagged values of the endogenous variables, naturally, are values of these variables for a previous time period. Exogenous variables include all other variables that might enter into an analysis. Weather is a commonly cited exogenous variable. For many variables, the exact economic structure of the segment of the economy being studied must be carefully considered to determine whether they should be classified as endogenous or exogenous, and frequently different analysts will classify them in different ways.

We now can reconsider an example cited earlier. In section $F$ we showed a diagrammatic representation of a situation for which the least squares method could be used to estimate the slope of the demand curve. We now know that this estimate will be statistically consistent and efficient only if the quantity consumed and the demand shifters each can be classified as an exogenous or lagged endogenous variabie. Fox io/ has argued that this is approximately true for a considerable number of agricultural products, including meat, poultry and eggs, feed grains, and a number of fresh fruits and vegetables. Under the assumptions of Fox, market price is used as the dependent variable and the independent variables are supply (which is assumed to be highly correlated with consumption) and some relevant demand shifters, including usually disposable income. Another situation under which we can use the method of least squares to estimate elasticities of demand is that for which data are available on purchases or consumption of individual consumers, as prices that confront consumers are determined chiefly by factors other than those that affect their purchases. In this case, consumption is taken as the dependent variable and retail prices of the various items, family income, and perhaps other household characteristics; are taken as independent variables.

[^3]One further problem needs to be considered before leaving the statistical aspects of this subject. In all cases given previously, we have assumed implicitly that the variables are known without error. Any analyst who has been connected with the compilation of data from original sources knows that errors of one sort or another always creep in. These can result from nemory bias on the part of respondents, inability to find all of the people in a complete census, errors of sampling, and a host of other reasons beyond the control of the most careful investigator. Whenever we work with economic series, nonnegiigible errors in the data are known to exist.

Fox has suggested an easy way to correct for this, provided we are willing to make an iniormed guess about the average magnitude of the errors. 11/ Such guesses or estimates can be made by carefully reading how the series was compiled, or preferable, by talking to the person in charge of the compilation. For most series, we have a fairly good idea as to whether the error is in terms of, say, a few percent, 10 to 20 percent, or something higher. If we wish, we can experiment with various assumptions about the level of error, compare the effect on our coefficients, and then arrive at a rather good idea as to the effect which Iikely errors may have on their magnitude.

Because of the importance of a correction of this sort, I would like to show you exactly what is involved in an extremely simple case. All of you know, I am sure, that when we express variables in terms of deviations from their respective means, the least squares regression coefficient between two variables is given by:

$$
b_{y x}=\frac{\Sigma y x}{\Sigma x^{2}}
$$

Suppose that $\underline{y}$ and $\underline{x}$ are each subject to error and that the magnitude of the error is given $\bar{b} y \underline{d}$ and e, respectively. The sample regression coefficient then equals:

$$
b_{y x} \frac{1}{y}=\frac{\Sigma(y+d)(x+e)}{\Sigma(x+e)^{2}}
$$

Let us now expand the numerator and denominator. We obtain:

$$
\begin{aligned}
\Sigma(y+d)(x+e) & =\Sigma y x+\Sigma y e+\Sigma d x+\Sigma d e \\
\Sigma(x+e)^{2} & =\Sigma x^{2}+2 \Sigma x e+\Sigma e^{2}
\end{aligned}
$$

[^4]If $\underline{d}$ and $\underline{e}$ are randorn and independent, and we have a lange sampie, any summation term that involves ore or both of them as a cross-product equais approximetely zero. Sumation ferms that involve their squares, however, do not equal zero. Applying this princirie, when the variables are subject to error, we can write the value of the regression coefiicient in the following way:

$$
{ }_{\mathrm{byx}}^{*}=\frac{\Sigma \mathrm{x} x}{\Sigma x^{2}+\Sigma e^{2}}
$$

This gives a biased estimate of byx because the denominato is too large. But if we reduce the sum of squares for the independent veriable by an amount proportional to the assumed percentage error, then we obtain an estimate of the regression coefficient that is approximately unbiased provided the only source of bias is that due to errors in the data. Such a correction can be made easily. In working with multiple regression coefficients, some formulas involve sums of squares for the dependent variable also, so all sums of squares that enter into the aralysis should be corrected in a similar way. I see no reason why the same sort of correction could not be applied to the sums of squares that ere involved in equations to be fitted by the simuitaneous equations approach.

## Some Econometric Consicerations

Discussion in the precejing section suggests, and econometricians have shown, that if we use the least squares approach to estimate the regression coefficients in an equation that contains current values of 2 or more endogenous variables, we obtain estimates that are statistically biased. The mathematicai nature of the bias has ceen shown by a number of authors in a supposealy popular way, i2/ but i have yet to find an explanation that is completely satisfactory for a nonmatnematician. We do, however, have some experimental evideuce of the kind of bias that results when the method of ieast squares is applied in such cases. Methods which have been deveioped to handle equations that contain 2 or more endogenous variaioles are known to give estimates that are statistically consistent, but methods are not now available that are known to be statistically unbiased. Thus, we cannot say for sure what happens when we work. with the small samples that usually are involved in economic research.

An experiment was designed to measure the kind or bies that arises when we apply these methods to small samples 13/ and, as a byproduct of this experiment, we have some concrete evidence of the kind of bias that may arise when we use the method of least squares instead.

[^5]In this experiment, a simple 3 -equation model was formuiated with known coefficients. Variables generated by the mode? were obtained and random error terms added to them. Two thousand observations were obtained in this way, and they were then divided into 100 samples of 20 observations each. The first equation contains 2 endogenous variables. Regression coefficients were obtained for this equation for each sample by the limited information approach, which I will discuss later, and by the method of least squares. Since there were 100 samples; 100 separate estimates of the single regression coefficient involved was obtained by each approach. Frequency distributions of these estimates are shown in section $H$ of the chart, together with the true value of the coefficient. Each method gives estimates that are biased, but the 3 highest frequencies for the limited information approach are grouped about the true value, whereas the 3 highest frequencies for the least squares approach each are to the right of the true value. The average bias of the least squares estimates was almost 3 times as large as that for the limited information approach. In fairness, I should point out that for a second version of this model the 2 frequency distributions are more nearly similar, but the average bias in the least squares estimates still was aimost double that for the limited information estimates. This study was carried out at Stanford University by making use of a largescale electronic computor. It is to be hoped that further experiments of similar character, with different sorts of models, will be undertaken.

In discussing methods that are used to handle equations that involve more than 1 endogenous variable, it is convenient to introduce a mathematical concept that deals with the degree of identification. We saw earlier that it is sometimes impossible to estimate the coefficients in certain structural equations with the kind of statistical data available to do the job. Such equations are said to lack identifiability or to be underidentified. Identifiable equations, however, may be just identified or overidentified. In the discussion that foilows, we deal only with identifiable equations. The degree of identification relates to individual equations in a system, not to the entire system.

By algebraic manipulation, we always can write down the equations in a complete system so that the number of equations equals the number of endogenous variables. We then can think of these as $n$ equations in $n$ unknown endogenous variables, and we can always solve the equations so that each endogenous variable is expressed as a function of all of the predetermined variables in the system. These are called reduced form equations. Since each reduced form equation contains only a single endogenous variable, we can obtain estimates of the regression coefficients in these equations that are statistically consistent and efficient by use of the method of least squares.

If each equation in the system is just identified, there is always a unique transformation by which we can go from the coefficients in the reduced form equations to the coefficients in the structural equations. I am sure that many of you have seen this transformation for simple systems of equations, but perhaps it is worthwhile to run through an example to make sure that its nature is clear to everyone. Suppose that we have the following structural demand and supply equations, where
$\underline{q}$ and $p$ have their usual meaning, $y$ is consumer income, $w$ represents important weather factors that affect suppiy, and each väriable is expressed in terms of deviations from its respective mean:

$$
\begin{align*}
& p=b_{11} q+b_{1.2} y  \tag{2}\\
& q=b_{12} p+b_{22} W \tag{3}
\end{align*}
$$

We have 4 variables in the system and 3 variables in each equation; since 4-3 equals the number of endogenous variables in the system minus one, we can assume that each equation is just identified. Several rules of thumb like this are available to determine the degree of identifi" cation; more exact rules depend on the rank of certain matrices.

If we substitute the right-hand side of equation (3) for $q$ in equation (2) and the right-hand side of equation (2) for $p$ in equation (3) and simplify terms, we obtain the following:

$$
\begin{align*}
& \mathrm{p}=\frac{\mathrm{b}_{11} \mathrm{~b}_{22}}{1-\mathrm{b}_{11} \mathrm{~b}_{21}} w+\frac{\mathrm{b}_{12}}{1-b_{11} b_{21}}  \tag{4}\\
& q=\frac{\mathrm{b}_{22}}{1-b_{11} b_{21}} w+\frac{\mathrm{b}_{122^{b}} 1-b_{11} b_{21}}{y} \tag{5}
\end{align*}
$$

Since the denominator of each term on the right of the equality sign is identical, we can ignore these denominators for the moment. If we divide the coefficient of $W$ in equation (4) by the coefficient of $W$ in equation (5), we obtain an estimate of bll. If we divide the coefficient of $y$ in equation (5) by the coefficient of $y$ in equation (4), we obtain an estimate of $b_{21}$. Given an estimate of $b_{1} \bar{y}^{2}$ and $b_{21}$, we can estimate $\underline{b}_{12}$ from the coefficient of $y$ in equation (4) and $\overline{\mathrm{b}}_{22}$ from the coefficient of $\underline{w}$ in equation (5). This gives the 4 regression coefficients needed for our structural equations. Estimates that are uniquely equivalent are obtained by any alternative algebraic manipulation. Since the b's are known to be statistically consistent estimates, the estimates of the structural coefficients obtained in this way are statistically consistent.

Computationally, we may wish to estimate the coefficients in another way, but the answers obtained are identical to those that would have been gotten by an algebraic manipulation of the regression coefficients from the reduced form equations.

Let us now consider an equation that is overidentified. Suppose that our supply equation contains a second predetermined variable, $\underline{z}$, that represents lagged values of prices. We now have 5 variables in the system. The supply equation still is just identified, since 5-4 equals the number of endogenous variables in the system minus one. However, the demand equation is overidentified, since 5-3 is greater than 2-I.

The new reduced form equations can be obtrined by the same general approach as used previously, but the result now looks like this:

$$
\begin{align*}
& p=\frac{b_{11} b_{22}}{1-b_{11} b_{21}} w+\frac{b_{11} b_{23}}{1-b_{11} b_{21}} z+\frac{b_{12}}{1-b_{11} b_{21}} y  \tag{6}\\
& q=\frac{b_{22}}{1-b_{11} b_{21}} w+\frac{b_{23}}{1-b_{11} b_{21}} z+\frac{b_{12} b_{21}}{1-b_{11} b_{21}} y \tag{7}
\end{align*}
$$

With this set of equations, by could be estimated either by dividing the coefficient of $w$ in equation (b) by the coefficient of $w$ in equation (7) or by dividing the coefficient of $z$ in equation (6) by the coefficient of $\bar{z}$ in equation (7). Different answers are obtained from the 2 estimates. It is in this way that overidentified equations differ from just identified ones; for overidentified equations, we have an oversufficiency of information and no direct way to decide which answer to use. In fact, neither answer obtained by the use of reduced form equations is statistically consistent.

It would be possible to solve the 2 structural equations directly for the several coefficients involved by making use of a maximum likelihood approach. Maximum likelihood estimates are known to be statistically consistent and efficient. They are used widely in statistical work because the necessary equations always can be derived by performing certain mathematical operations that involve the maximization of the socalled likelihood function. The general approach is the same as for any maximization process by use of calculus and it is not difficult. For complex systems of equations, however, the mathematics involved in solving the resulting equations is generally complex. That part of Klein referred to in footnote 9 involved the derivation of maximum likelihood estimates. Methods for obtaining maximum likelihood estimates based on a simultaneous solution for all of the structural equations are discussed by Klein and in Cowles Commission Monograph 14 and are called full-information maximum likelihood estimates but, to quote Klein, the computations involved in general are "formidable." Hence this method is seldom used.

Another method, developed by staff members of the Cowles Commission, is called the single-equation limited-information maximumlikelihood method. In this approach, equations are fitted one at a time and information regarding variables that appear in each of the other equations of the system is disregarded. Use is made, however, of all endogenous and predetermined variables that appear in the equation and of all other predetermined variables that appear in the system. This method is known to give estimates of the regression coefficients that are statistically consistent and as efficient as any other method that utilizes the same amount of information. A.
compromise between this and the full-information method is one calied the limited-information subsystem method, in which. selected groups of equations are fitted as a unit. Most of the systems of simultaneous equations that have been fitted and taat involve overidentified equations have been based on the single.equation limited information approach. This is the approach that generally is meant when reference is made to the use of the limited information method.

For some appiications:of the limited information approach, I would like to call your attention to Cowles Commission Monograph 15 by Hildreth and Jarrett 14/ dealing with the livestock economy; Iowa Research Bulletin 410 by Nordin, Judge, and Wahby $15 /$ dealing with livestock products; and USDA Technical Bulletin $11 \frac{1}{36}$ by Meinken 16/ on wheat, to be issued by us this fall, This approach has been used in several other studies, but each of these devotes considerable attention to broad methodological aspects.

## Some Computational Considerations

If certain equations in a system of equations involve only a single endogenous variable, these can be fitted by the least aquares approach. If other equations are just identified, they can be fitted by the method of reduced forms. If still other equations are overidentified, they can be fitted by: the single-equation imited information approach. If the full information approach were to be used, all equations would be fitted simultaneously. But, even if, in essence, each equation is to be fitted separately, similar computations based on the same data frequently are involved for each analysis. : Hence considerable computational time may be saved by handing the computations as a unit. Even if several equations can be fitted directly by least squares and no other equations are involved, time frequently is saved by carrying out the computations as a simultaneous operation for the group. Whichever approach is used, the answer obtained is the same; here we are concerned only with computational efficiency.

If time permitted, I would give some concrete examples to illustrate this point. We have in process a computational handbook in which several such examples are discussed in detail. 17/ I would like to tell you just what is included in this handbook. First we describe a new method for handiling ordinary multiple-regression analyses. To quote the preface, "This method is believed to be more efficient than

[^6]those now commonly in use and is easier for beginners to understand than those based essentially on the Doolittle approach, since no back solution is required." Then we consider methods for handing equations that involve 2 or more endogenous variables, using essentially the same procedure for equations that are either just identified or overidentified. We discuss in detail the exact steps required for equations containing specified numbers of endogenous and predetermined variables. Then we consider the previcusly mentioned examples of complex systems of equations. In each case, we show how to obtain the coefficients and their respective standard errors. Finally, we take up tests that should be made after the analyses are completed, correlation concepts that apply to systems of equations, and the use of systems of equations for analytical purposes or forecasting. To quote the preface again, "This handbook is designed to provide a complete description of the steps involved in the more common types of problems and to illustrate them in a way that will be clear to research and clerical workers who have an acquaintance only with standard methods for handing singleequation multiple-regression analyses." The handbook should be available late this fall. Carbon copies of it are already in use in our central computing unit.

## Formulation of Models

In my opinion, there should be no question in the minds of research analysts as to whether they should use single-equation or simultaneous-equation methods for particular equations or groups of equations, but decisions must be reached regarding the complexity of the model to be used and the degree of aggregation, Klein has a useful suggestion here in connection with his discussion of sector models. $18 /$ He suggests that a master model for the entire national economy be formulated. As I am sure most of you know, he and staff workers at the University of Michigan have made some progress in formulating models of this type. Models for individual industries, individual commodities, or individual regions can then be formulated and fitted as separate entities in such a way that they can be "grafted" onto the master model. Grafting of this sort is being carried out at the University of Michigan and elsewhere. In our work for example, we have tended to treat variables that relate to the national economy, such as disposable income, as though they were predetermined. With Klein's approach, (1) the computed value of disposable income for each observation, based on the master model, could be used as a predetermined variable, or (2) disposable income could be treated as an endogenous variable and the preuetermined variables on which it depends brought in as predetermined variables in the system ior the sector. Either method yields consistent estimates, whereas the treatment of disposable income, as such, as a predetermined variable results in some bias.

The same general approach could be used in narrower fields :ror example, we plan to fit some national aggregate simultaneous equation models for the feedilivestock economy in which all feeds and ell livestock and livestock products will be aggregated. The model will be similar to
that used by Hildreth and Jarrett but will contain what we believe to be some important modifications. At the same time we will be fitting models that relate to individual feeds and to individual types of livestock, either for the country as a whole or for specified regions. Later, less aggregative models may be fitted for separate regions, with separate equations for each major type of livestock, and these regional models then could be grafted onto the master feed-livestock economy model, which in turn could be grafted onto a master model for the entire economy. Thus, over time, a series of studies could be built up, each of which would be manageable as a research unit, but all of which would tie together to make a united whole.

## Concluding Remarks

I regret that time has not permitted the presentation of concrete examples of the formulation and use of simultaneous systems of equations, but the references which I cited earlier contain some excellent examples of these. A complete description of any one of them would have required at least the full time allotted to me today. I hope that my remarks have cleared up some of the concepts and terminology used in this area by modern econometricians. In formulating and fitting complex models, the nonmathematician will do well to seek the advice of a competent expert in this area. Fortunately, our graduate schools are turning out an increasing number of such experts. However, it is my belief that anyone who is willing to give the matter careful thought can understand the basic concepts involved and can conduct applied research in such a way that he can use these concepts to advantage in his work with only the occasional help of a mathematician.


[^0]:    * Prepared for presentation before the American Farm Economic Association, August 1, 1955. The author wishes to thank Glen Burrows of the Agricultural Marketing Service for a number of helpful suggestions on the statistical aspects of this paper.

    1/ Quart. Jour. Econ. 41:212-235, 1927. A similar line of reasoning is followed by Koopmans, Tjaijing C. Identification Froblems in Economic Model Construction, pp. 27-35. This is given as Chap. II of Studies in Econometric Method. Cowles Commission for Research in Economics Monogr. 14, 1953.

[^1]:    2/ Op. Cit., p. 223. 3/Ezekiel, Mordecai. Statistical Analyses and the "Laws" of Price. Quart. Jour. Econ. 42:199-225. 4/Fox, Karl A. The Analysis of Demand for Farm Products. U. S. Dept. Agr. Tech. Bul. 1081. 5/ Haavelmo, Trygve. The Statistical Implications of a System of Simultain. neous Equations. Econometrica. 11:1-12, 1943. A more complete discussion by the same author is given in The Probability Approach to Econometrics. Econometrica, Vol. 12, Supplement, 1944.

[^2]:    6/ Op. Cit., p. 227.

    1) Marschak, Jacob. Economic Measurements for Policy and Prediction, Chap. I in Studies in Econometric Method, Cowles Commission for Research in Economics Monogr. 14. 1953.
[^3]:    9/Klein, Lawrence.R. A Textbook of Econometrics...Row, Peterson and Co., 1953, pp. 80-85.
    10/ Op. Cit.

[^4]:    11/ This approach and some exampies based on it are described in some detail by Foote, Richard J. and Fox, Karl A. Analytical Tools for Measuring Demand. U.S. Dept. Agr. Agr. Handb. 64, 1954, pp. 29-35.

[^5]:    12 See, for example, Bronfenbrenner, Jean. Sources and Size of LeastSquares Bias in a Two-Equation Model, Chap. IX in Studies in Econometric Method. Cowles Commission for Research in Economics Monogr. 14, 1953; Bennion, E.G. The Cowles Commission's "Simultaneous-Equation Approach": A Simplified Explanation. Review Econ. and Statis., 34:49-56, 1952; and Meyer, John R. and Miller, Henry Lawrence Jr. Some Comments on the "Simultaneous-Equations" Approach. Review Econ. and Statis, 36:88-92.1954.

    13/ Wagner, Harvey M. A Monte Carlo Study of Estimates of Simultaneous Linear Structural Equations. Tech. Rpt. 12, Dept. of Econ., Stanford Univ. 1954.

[^6]:    $14 / \mathrm{Hildreth}, \overline{C l i f f} \overline{\mathrm{Id}, \text { and Jarrett, Frank. A Statistical Study of }}$ Iivestock Production and Marketing. Cowles Commission for Research in Economics Monog. 15. 1955.
    $15 /$ Nordin, J. A., Judge, George G., and Wahby, Omar. Application of Econometric Procedures to the Demands for Agricultural Products. Iowa Agr. Expt. Sta. Research: Bul. 410, July 1954.

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