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# International Cross Section Estimates of Demand for Use in the GTAP Model by

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#### Abstract

The making of projections often requires an economy-wide perspective, and the estimation of consumer demands at the international level. In this paper, an implicit, directly additive demand system (AIDADS) is estimated using cross-country data on consumer expenditures from the International Comparison Program (ICP), and then from Global Trade Analysis Project (GTAP) data. The two data sets are found to produce results that are quite consistent despite their differing origins, and the fact that the former is based on consumer goods that embody wholesale/retail margins, while margin demands are treated separately in GTAP. Given the similarity of the results, the estimation based on GTAP data is favored because it is readily matched to input-output based production and trade data, and provides valuable new information concerning how aggregate margin expenditures are related to per capita income.

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#### 1. Introduction

There is widespread demand for projections of economic activity, by sector, for purposes of strategic planning as well as policy analysis. When such forecasts are related to national energy demand, employment, or environmental quality, researchers are usually forced to take an economy-wide approach to modeling, which in turn requires specification of a complete system of consumer demands. By linking these estimates of demand behavior to production data based on input-output accounts, a researcher can then estimate the direct and indirect effects of changes in final uses on specific industries and commodities (Lawson, 1997). Linking demand behavior to production and trade data is also necessary for testing certain economic theories (e.g., Reimer, 2003), as well as for understanding the process of economic development in general (Lluch, Powell, Williams, 1977). When such studies aim to determine global resource requirements, then an international demand system is required.

Unfortunately, specifying the demand parameters for a global model is a formidable challenge. One reason for this difficulty relates to the fact that there are few demand systems that are well behaved over the wide variation in per capita incomes inherent in a global model, let alone satisfactorily characterize demand behavior in such circumstances.<sup>1</sup> Another challenge stems from the fact that there are few publicly available data bases with information on per capita expenditures for multiple goods and countries. Perhaps the best known source of such data is the International Comparison Program (ICP), which was begun at the University of Pennsylvania and whose activities are currently coordinated by the World Bank. The ICP data have been used in numerous international demand studies, such as Theil and Clements (1987); Hunter and Markusen (1988); Theil, Chung, Seale (1989); Cranfield, Preckel, Eales, Hertel (2000); Regmi, Seale, Bernstein (2001); and Cranfield, Eales, Hertel, Preckel (2002).

Unfortunately, the ICP data pose a problem for researchers who need to match ICP-based demand estimates to production and trade information in data bases such as the Global Trade Analysis Project (GTAP), or the OECD Structural Analysis (STAN) data, which are based on input-output accounts (Dimaranan and McDougall, 2002; OECD, 2002). There are two primary reasons for this difficulty. The first is that production and trade data bases are generally evaluated in terms of producer prices, thereby excluding the wholesale/retail margins and transportation costs embodied in ICP consumer expenditures, which are evaluated at purchaser's prices. In theory, one should be able to reconcile consumption values with production values via use of a *margins matrix*, as in Lenzen (2000), Lawson (1997), Davis et al. (1997), and Ballard et al. (1985).<sup>2</sup> Such a matrix maps a portion of final goods expenditure to the wholesale/retail trade and transportation sectors. It thereby reflects the fact that to consume anything at the retail level requires consumption of the wholesaling, retailing, and transportation services which brought the product to the store. While margins matrices have been constructed and are available for certain, individual regions (e.g., ABS 1999a, for Australia; Lawson 1997, for the USA), they are quite difficult to construct (Piergiovanni and Pisani, 1998). The problems include issues such as collecting and compiling the large quantities of data that are

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<sup>&</sup>lt;sup>1</sup> Regarding the link between consumer demands and per capita income, an 1857 study by Prussian economist Ernst Engel may have been the first to explore the nature of this relationship. Using household data on 153 Belgian families, Engel made what is perhaps the first empirical generalization regarding consumer behavior: The proportion of total expenditure devoted to food declines as income rises. This hypothesis – now known as "Engel's Law" – has been verified in many subsequent studies, and suggests that consumer preferences tend to be non-homothetic. Classic studies in this area include Stigler (1954); Prais and Houthakker (1955); Lluch, Powell, and Williams (1977); Deaton and Muellbauer (1980); and Theil and Clements (1987).

<sup>&</sup>lt;sup>2</sup> Some authors call this a "distribution", "bridge", or "transformation" matrix.

necessary, reconciling firm level data with national accounts data, inconsistencies among data sources, issues with under-reporting by firms, and concerns about the transparency and reliability of the many assumptions that inevitably have to be made. These difficulties are particularly acute when country coverage is extensive. A single margins matrix is unlikely to satisfactorily represent a broad mix of countries, given the vast international differences in technology and economic structure that prevail. Furthermore, it would be extremely expensive in terms of time and funding to construct margins matrices for the many countries for which such information is currently unavailable.

The second reason for the difficulty in linking ICP consumer data to production data is that the classification schemes are typically very different. For example, the ICP category "Medical Care" includes expenditures on "Services of Nurses" as well as "Therapeutic appliances and equipment", for which International Standard Industrial Classification (ISIC) categories (upon which GTAP and OECD STAN data are based) have no direct equivalents.

These compatibility issues have been exacerbated in the most recent version of the ICP data, which corresponds to the year 1996. The problem relates to inconsistencies across regions in the way that the ICP data were collected and compiled (World Bank, 2002). For the data to be internationally comparable, it had to be aggregated from the original, base ICP categories (numbering in excess of 100), to a small number of categories (26 in the 1996 version). In general, the aggregation was carried out according to similarity in end use, meaning that expenditures on services were often combined with expenditures on physical goods. In contrast, production and trade data tend to be defined more nearly according to the factor used to produce them (e.g., labor-intensive activities tend to be distinguished from land- or capital-intensive activities). They also tend to have a small number of services categories, with a single category for wholesale/retail trade. Due to these issues, the most recent version of the ICP consumption data is even less compatible with production and trade data, than were earlier ICP versions.

One approach to dealing this problem begins with the observation that some input-output-based production and trade data sets also have information about per capita expenditures for a large number of countries. For example, the GTAP data contain information on consumption across 57 categories of goods and services, which for the purposes of demand system estimation, can be aggregated to a smaller number of categories that are representative of the consumption choices faced by households.<sup>4</sup> Over the past decade, GTAP data have gone through five public releases as well as extensive scrutiny by users, and their quality and usefulness are generally well regarded. If a global demand system could be satisfactorily estimated with these data, it would allow a researcher to bypass the severe mapping problems associated with the ICP data.

Based on these observations, the primary objective of this study is to estimate a global demand system with 1997 GTAP per capita consumption expenditure data, and compare these results to those arising from estimation of this demand system using the more conventional ICP data. Should the estimation based on GTAP data be found to perform as well as the ICP-based estimation, this new demand system would prove to be a boon to researchers needing to reconcile consumer demand estimates with production and trade data.

The study begins by first outlining the demand system to be estimated (AIDADS), and providing description of both the ICP and GTAP data bases. In the course of describing the data, the

<sup>&</sup>lt;sup>3</sup> Prior to the 1996 version, the most recent publicly released ICP data are for 1985. These have been used, for example, in the Cranfield et al. (2000, 2002) studies.

<sup>&</sup>lt;sup>4</sup> A producer category like "ferrous metals", which serves almost exclusively as an intermediate input, can be aggregated into a consumer category which primarily uses it, such as "manufactures/electronics".

reader is referred to appendices at the end of this study that describe the ICP terminology, and document several problems encountered with the 1996 ICP data, as well as how these problems were rectified. Subsequent sections briefly describe aggregation and estimation issues, before turning to the demand system results. An additional appendix provides a GAMS program that enables readers to calculate GTAP-based expenditure elasticities for countries outside of the GTAP sample, or to make projections of consumer demand and future expenditure elasticities for those countries within the sample.

This study also contains the results of an additional AIDADS estimation exercise that may be of interest to readers. In this section, the demand system is again estimated with GTAP data, but this time to study the composition of *regional final demand* with respect to its three primary components: Consumption (*C*), Investment (*I*), and Government (*G*). In particular, this section determines whether the shares of these final demand components vary systematically with income. This exercise has a number of interesting applications. For example, it provides evidence concerning hypotheses about final demand expenditures such as Wagner's Law, which states that government has a tendency to grow at a faster rate than the rest of the economy (Islam, 2001). In turn, this exercise may be quite useful for future applied general equilibrium modeling. Consider that the GTAP applied general equilibrium model, for example, employs the concept of a "regional household", which acts as a sort of regional clearinghouse regarding the allocation of regional expenditure. The exercise of this section will demonstrate that the behavior of such a regional household can be shown to systematically vary across the income spectrum.

The overall goal of this study, therefore, is to demonstrate that GTAP expenditure data (in conjunction with the AIDADS demand system, to be discussed below) are quite useful for the estimation of international demand behavior. It is hoped that the analyses presented below will stimulate further applications of the GTAP expenditure data by researchers, and strengthen current efforts to update and improve the GTAP data base.

#### 2. The AIDADS Demand System

It is assumed there is a single, representative private consumer in each country, whose preferences are identical to those of every other country's representative consumer. As such, issues related to aggregation over individual consumers are assumed away, which is necessary given the data available to us. Moreover, it is assumed that tastes are internationally identical – at least at a fairly aggregated level. This is an assertion for which Clements and Chen (1996) find strong empirical support.

With regard to choosing an appropriate demand system for estimation, first observe that there are several which could feasibly provide a basis for estimating consumer demands at the international level. We select among these by sifting through the various demand system characteristics that are desirable in this context. To begin, features such as ease of estimation, parsimonious parameterization, and the imposition of the economic restrictions of adding up, homogeneity, and symmetry are desired. Furthermore, the utility function underlying the demand system should be not homothetic, which eliminates the possibility that consumers adjust their purchasing behavior as their income changes. This requirement rules out forms such as the Homothetic Constant Elasticity of Substitution (HCES) and Cobb-Douglas. Additionally, the income consumption path should not be forced to be linear. In other words, the demand system should allow for more than just quasi-homotheticity, which severely limits the amount of demand response that is possible across the income spectrum (Deaton and

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<sup>&</sup>lt;sup>5</sup> The newly refined 1996 ICP data that are developed in this paper are available upon request.

Muellbauer, p. 145). This rules out the Linear Expenditure System (LES).<sup>6</sup> Another desirable feature is that the demand system be well-behaved across a substantial range of per-capita income. This consideration rules out the Almost Ideal Demand System (AIDS) and Working Preference Independence model, since the budget shares in these models can stray outside the unit interval under moderately large variations in per capita expenditure (Rimmer and Powell, p. 1614), and we will estimate using data from countries that span the world income spectrum.<sup>7</sup>

One demand system that possesses all of these desired properties is AIDADS (An Implicitly Directly Additive Demand System). Developed by Rimmer and Powell (1996), AIDADS is a generalization of the LES that allows for non-linear income consumption paths, while maintaining a parsimonious parameterization of preferences. AIDADS also has the advantage of being an effectively globally regular demand system. In a comparison test with four other demand systems (QUAIDS, QES, AIDS, and LES), AIDADS performed exceptionally well, particularly in situations where there are widely varying income levels (Cranfield et al., 2002). A possible disadvantage of AIDADS for some purposes is that it imposes implicit direct additivity, thereby limiting the range of substitution that is possible across goods. This feature is not critical to this study, however, since the goods represented correspond to relatively broad, aggregated categories, as opposed to narrowly defined ones. Instead of devoting more parameters to substitution responses, AIDADS captures a rich array of Engel effects by adding N - I income response parameters beyond that which is used by the LES, where N is the total number of goods. Therefore, the total number of parameters to be estimated is 3N + 1.

Rimmer and Powell (1996) derive the following system of budget share equations for AIDADS:

$$w_n^c = \frac{p_n^c \mathbf{g}_n}{M^c} + \left(\frac{\mathbf{a}_n + \mathbf{b}_n e^{u^c}}{1 + e^{u^c}}\right) \left(1 - \frac{\mathbf{p'?}}{M^c}\right),$$

where  $w_n^c$  is the share of good n in country c's private household consumption;  $\mathbf{g}_n$ ,  $\mathbf{a}_n$ , and  $\mathbf{b}_n$  are unknown parameters to be estimated;  $u^c$  is utility<sup>9</sup>;  $M^c$  is per-capita expenditure on private household consumption in country c;  $p_n^c$  is the price of good n in country c; and  $\mathbf{p}$  is the vector of goods prices. The term  $\mathbf{g}_n$  reflects the subsistence level of good i that all countries must consume. The term  $\mathbf{p}'$ ? represents the minimally sustainable per-capita expenditure in any country, and  $[M^c - \mathbf{p}']$  is discretionary income. The following parametric restrictions are used to ensure well-behaved demands:  $0 \le \mathbf{a}_n$ ,  $\mathbf{b}_n \le 1$  for all n, and  $\sum_n \mathbf{a}_n = \sum_n \mathbf{b}_n = 1$ . The regularity conditions of

<sup>&</sup>lt;sup>6</sup> Another drawback of the LES is that its income elasticities are highly counterintuitive, since for necessities they always increase with income, while for luxury goods they always decrease with income.

<sup>&</sup>lt;sup>7</sup> The AIDS demand system and variants of the Working model also have the disadvantage that they possess only local curvature properties, if they satisfy concavity at all.

<sup>&</sup>lt;sup>8</sup> This is because AIDADS is a rank 3 demand system, according to Lewbel's (1991) definition. Rank 1 demands are are independent of income, and therefore the most restrictive. Rank 2 demand systems do not force the Engel curve through the origin. Rank 3 (i.e., full rank) demand systems are the least restrictive, allowing for non-linear Engel responses.

<sup>&</sup>lt;sup>9</sup> It may be seem unusual to see utility in this function. In principle one could obtain uncompensated demands (and budget shares) that depend solely on prices and expenditure by substituting out utility, but this cannot be done analytically. The estimation process, documented in Cranfield et al., involves solving for utility numerically and making the necessary substitution.

consumer theory are satisfied in the price-income space for which discretionary income is strictly positive. Note that in the special case that  $\mathbf{a}_n = \mathbf{b}_n$  for all goods n, only the constant  $\mathbf{b}_n$  is left in the marginal budget share term, and AIDADS is equivalent to LES. Further description of AIDADS will be provided in later sections, when its estimation and the results are discussed.

#### 3. International Comparison Program Data

In this section, the background, characteristics, and present status of the International Comparison Program (ICP) data are described. Section 4 of the study will provide similar discussion of the Global Trade Analysis Project (GTAP) data. The origins of the ICP can be traced back to researchers working at the University of Pennsylvania in the late 1960s (Kravis, Heston, and Summers, 1975). Since that time the general objective of the ICP has been to develop internationally consistent price and quantity comparisons across countries regarding the components of GDP. Although these international comparisons could be made with respect to structure of production (i.e., GDP could be decomposed by industrial category), the ICP makes international comparisons via expenditure category breakdowns: household consumption (C), capital formation (f), government consumption (G), and net exports (United Nations, 1992). There is often a large amount of detail within these categories. Indeed, the number of household expenditure categories exceeds 100 in some versions of the ICP data. Another important characteristic is that the ICP data encompass many developing nations, in addition to those which are highly industrialized. Features such as these make the ICP data well suited to the study of international demand patterns.

A central feature of the ICP framework is its use of Purchasing Power Parity (PPP), a concept that can be traced back to sixteenth century work by Spanish scholars, and is widely used in the international finance literature (Daniels and VanHoose, 1999). A PPP rate is a certain type of exchange rate that compares the cost of a common basket of goods in two countries. In ICP terminology, a PPP rate indicates the amount of Local Currency Units (LCUs) needed to purchase a bundle of goods that is identical in quality and quantity to what can be bought with one unit of a base country's currency. Advocates of the PPP approach point to its several advantages over the alternative of conventional exchange rates for the study of international consumption patterns. First of all, PPPs do not fluctuate over time to the degree that conventional exchange rates do. PPP rates are also less likely to overstate poverty in developing countries, since PPPs allow for the fact that services may be cheaper there. Additionally, the ICP produces PPP rates that are specific to the individual components of a highly refined decomposition of GDP. A conventional exchange rate, in contrast, provides but a single conversion factor for all goods and services.<sup>10</sup>

Several recent international demand studies, such as Cranfield et al. (2000, 2002), have employed the 1985 version of the ICP data. These "Phase IV" data have information specific to 113 commodities and 64 countries, and are the most recent ICP release to have been widely distributed. While more recent versions of the ICP data exist, they have not been made widely available due to concerns about quality. Some of the problems may be traced back to a lack of funding, resulting in limited data collection in some parts of the developing world. Another issue has been the European

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<sup>&</sup>lt;sup>10</sup> For more information about the terminology of the ICP data, including some common misperceptions, please see Appendix C.

Union's desire to impose the "fixity principle", which has created structural problems in some versions of the ICP data (Seale, 1998). 11

In response to these problems, the World Bank, International Monetary Fund, and United Nations sponsored an independent review of the ICP project in the late 1990s. Consultant Jacob Ryten (1999) found the ICP's purpose to be important and timely, but acknowledged that there were significant financial and organizational obstacles facing the project, and made suggestions for improving the project's ability to generate credible and usable data. Fortunately for international demand analysts, a new round of the ICP has been launched with more resources and new organization. Unfortunately, this round will not yield new data for demand research until at least 2005 (World Bank, 2002).

Meanwhile, versions of the ICP data based on the efforts of the 1990s do exist. For example, the World Bank has made a data set corresponding to the year 1996 available to interested researchers. These data have been compiled in such a way to eliminate the concerns about quality in previous versions. Although some inconsistencies were encountered in preliminary examination of these data, it was possible to ascertain the source of these problems, and make the necessary adjustments (these are described in Appendix B). The resulting data now appear to be quite refined. For example, aggregate values are quite consistent with corresponding data from the World Bank's World Development Indicators (WDI), as well as GTAP data regarding household consumption. Additionally, it will later be seen that these data produce very credible estimates when used to estimate AIDADS.

The 1996 ICP data obtained for this study have complete information for 114 countries, the most ever made available. As indicated earlier, however, there is a major disadvantage to these data, since household consumption is decomposed into only 26 commodities, instead of the original decomposition of more than 100 commodities. This high level of aggregation was necessary because of difficulties in obtaining the appropriate disaggregated data from certain regions of the world. Consequently, unless one is willing to resort to the aging 1985 ICP data, the 26 commodity aggregation must be used. While this may be appropriate for some research objectives, if one eventually seeks to match demand estimates with supply characteristics, great difficulties will be encountered. This is because the ICP categories are incompatible with International Standard Industrial Classification (ISIC) categories, which is the basis for most production data sets.

The most significant problem is that both goods and services are classified together within each of the 26 ICP categories, whereas in ISIC categories, services are clearly distinguished from goods categories. Another issue is that the ICP categories of expenditure embody wholesale/retail trade margins as well as transportation costs, which are broken out in production and trade data. In theory, one could construct a transition matrix between producer and consumer categories to handle

<sup>&</sup>lt;sup>11</sup> The benefit of the "fixity principle" is that results obtained in a regional comparison (e.g., an EU comparison) remain unchanged when more countries are added to the sample, and a new comparison is made (e.g., an OECD comparison). The downside is that the imposition of fixity may prohibit the sum of subcategories from equaling the total for a category as a whole. While this may create problems for demand system estimation, it appears to have been mostly resolved in recent versions of the ICP data. See United Nations (1992) and Seale (1998) for more information.

<sup>&</sup>lt;sup>12</sup> These data were obtained from Mr. Yonas Biru of the World Bank, by way of Dr. Anita Regmi of the U.S. Department of Agriculture, Economic Research Service. At least two studies have made use of these data: Regmi, Seale, Bernstein (2001), and Regmi, Deepak, Seale, and Bernstein (2001).

<sup>&</sup>lt;sup>13</sup> While more disaggregated versions of the 1990s data exist and have been used by some researchers (e.g., Eaton and Kortum, 2002), they are generally only for one or two individual regions (in particular, OECD, African, and CIS countries), and are not necessarily compatible even with each other.

this issue (as in Ballard et al., 1985), but again this is unlikely to be a satisfying approach since much of the information that one wants is not available.

Despite these problems, the 1996 ICP data may be viewed as the "state of the art" regarding our understanding of international expenditure patterns. These data may have the most credible information regarding how prices vary across countries. They also contain the largest number of observations for the purpose of cross-country demand system estimation at one point in time.

#### 4. Global Trade Analysis Project Data

In many respects, the progression of the GTAP data since the early 1990s has been in the opposite direction of the ICP data. The size of the GTAP consortium (19 national and international agencies as of 2003) and the ensuing stability in financial support has ensured rapid progress. During the past decade there was no official release of ICP data, but the GTAP data base went through five public releases. The GTAP data have generally received high marks in terms of credibility and usefulness by a large pool of researchers. The sources and procedures used to create the data are extensively documented (Dimaranan and McDougall, 2002), and new documentation is provided with each new release.

There are other fundamental differences between the ICP and GTAP data. While continually expanding, the country coverage of the GTAP data base is still less extensive than that of the ICP data. And while the ICP data document international differences with respect to consumption and prices, GTAP data contain information on many other economic characteristics, since they have been constructed to operationalize a global general-equilibrium model. For example, GTAP data incorporate country-specific information regarding production, trade, technology, endowments, transportation, and protection.

The two datasets are similar, however, in that they both have detailed information on household consumption for individual countries. In particular, GTAP data allow dis-aggregation of private household consumption in each country into 57 commodities. In conjunction with population data, per capita household expenditure can be calculated for each commodity. One can also calculate "comparative price levels" using GTAP data that are somewhat similar to those calculated using the ICP methodology. By way of GTAP's tariff information, a commodity's price in a particular country can be distinguished from the average world price for that commodity. This is accomplished by dividing a nation's expenditure on a commodity valued at domestic market prices, by its expenditure valued at c.i.f. prices.<sup>15</sup>

Whether or not one feels comfortable with this notion of comparative price levels, the issue of prices may not be of critical importance to global demand system estimation. This is because international incomes vary by a factor of several hundred, while prices tend to vary by a factor of one or two times at most. Additionally, the assumption of direct additivity in the AIDADS demand system limits the amount of substitution across goods that occurs because of price differences. As a result, the possibility that ICP comparative price levels are perhaps more credible than those of GTAP does not confer a significant advantage to the ICP data, for the purposes at hand.

<sup>&</sup>lt;sup>14</sup> We carried out the same sort of checks with the GTAP data that were done with the ICP data. For example, in comparing the sum of these per capita household expenditures to corresponding data from the World Bank WDI data, the values were very nearly the same across countries.

<sup>&</sup>lt;sup>15</sup> In GTAP notation, divide VIMS (imports valued at domestic market prices, summed over all sources) by VIWS (imports valued c.i.f., summed over all sources).

<sup>&</sup>lt;sup>16</sup> AIDADS was estimated both with and without price variation, and the results were nearly indistinguishable.

#### 5. Aggregation Issues

Now that background has been provided regarding the relevance of the ICP and GTAP data sets for estimating an international, cross section demand system, we turn to issues regarding the aggregation of countries and commodities. GTAP Version 5.0 data contain information specific to 52 countries, all of which are used to estimate AIDADS (Table A-3). The ICP data, on the other hand, enable AIDADS to be estimated with 114 national observations (Table A-4). The greater country coverage of the ICP data reflects differences in the history, objectives, and organizational structure of the respective data sets. 18

With regard to the commodities used to estimate the demand system, both data sets were aggregated up to 10 categories, so that N = 10, and 3N + 1 = 31 parameters need to be estimated. This number was chosen based on historical practice for the ICP data set, and also represents what may be the practical limit of AIDADS. The particular commodity aggregations used in GTAP and ICP are given in Tables A-1 and A-2, respectively. The ICP aggregation is very similar to that used in other ICP applications, such as Hunter and Markusen (1991), and Regmi, Seale, and Bernstein (2001). To the greatest extent possible, the GTAP data follow a similar aggregation. However, there are two reasons why this match is imperfect, which were alluded to earlier. First, the categories of the ICP data do not distinguish between goods and services, as is done in GTAP, which is based on International Standard Industrial Classification (ISIC) and Central Product Classification (CPC) definitions. In the ICP data, goods and services are grouped together according to similarity in final use. For example, one ICP category is "Medical goods and services". This category contains not only expenditure on pharmaceutical products, for example, but also the fees paid to doctors who prescribe those products. In the GTAP data, commodities are grouped more nearly according to the factor used to produce them (i.e., labor-intensive activities are distinguished from land- or capital-intensive activities). Therefore, service categories are clearly distinguished from goods categories. This characteristic is generally found in other internationally comparable production and trade data sets, such as OECD STAN (OECD, 2002).

A second reason why GTAP categories are different from ICP categories is that the latter incorporate the activities necessary to convert a raw, producer commodity into a final, consumer commodity. These activities are distinguished in GTAP data by the wholesale/retail trade and transportation categories that reflect the margins between producers and consumers. As a result, the ICP data can be thought of as "consumer" commodities, while the GTAP data represent "producer" commodities. Therefore, the reader should be mindful that in the following analysis, ICP and GTAP categories are distinct, even when they are similarly named.

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<sup>&</sup>lt;sup>17</sup> There are 66 regions in total in the GTAP data base, but many of these are regional composites and do not contain original data.

<sup>&</sup>lt;sup>18</sup> To facilitate comparison of the demand system estimation results we considered deleting the countries in the ICP data and the GTAP data that are not common to both data sets. However, the number of countries that are common to both is only 44. Deleting all other countries would have resulted in a loss of many degrees of freedom in each case. As a result, we elected to delete no country observations, using all 52 in the GTAP data set, and 114 of the ICP data set. In terms of demonstrating that the data sets provide similar qualitative results, the somewhat different mix of countries in the samples should only work against us.

#### 6. Demand System Estimation and Results

AIDADS is estimated using the procedure of Cranfield et al. (2000, p. 1909), and space constraints permit only a brief discussion of this approach. These authors formulate the problem as a mathematical programming model in GAMS, and minimize a concentrated log-likelihood function according to a constraint set involving non-linear equality and linear inequality constraints. Starting values for the non-linear estimation are chosen to be consistent with an LES demand system, since it is easily estimated, and is a special case of AIDADS. Upper and lower bounds on all choice variables are included to reduce the size of the feasible set. Cranfield et al. (2000) carry out 100 bootstrap replications and find their estimators have very little bias, as well as a high degree of efficiency.

We now turn to the estimation results. AIDADS parameter estimates based upon the ICP and GTAP data are presented in Tables 1 and 2, respectively. These tables also provide expenditure elasticities evaluated at the means of the data  $(e_n)$ , and correlations between the actual and fitted shares  $(r_n)^{21}$  The results for both sets of data are consistent with one's intuition regarding how the composition of consumption is likely to differ across income levels. Observe in Table 1, for example, that the estimated subsistence budget shares  $(\mathbf{g}_n)$  for "Meat, dairy, fish", "Home furnishings and appliances", and "Transport and communication" are zero (note that these also represent three of four active bounds in Table 1). In contrast, for "Grains, other crops" the subsistence share is 0.690. This implies that consumption of the former three categories is not necessary for survival at the lowest levels of income, while staple grains are essential. Likewise, in Table 2 it is seen that estimating AIDADS with GTAP data results in subsistence budget shares for "Meat, dairy, fish", "Utilities, other housing services", and "Transport, communication" that are also zero (these estimates similarly represent three of the four active bounds in Table 2). Similar to the ICP-based estimates, staple foods tend to have large subsistence budget shares, at 0.298 and 0.142 for "Grains, other crops" and "Processed food, beverages, tobacco", respectively. Thus, as expected, both data sets imply that staple food products are necessary for survival at the lowest levels of income, whereas other foods, transport, communications, and home furnishings are not.

We now move on to assess the estimates of the AIDADS parameters  $\boldsymbol{a}_n$  and  $\boldsymbol{b}_n$ , which represent the bounds of the marginal budget shares. Looking at Tables 1 and 2, the parameter estimates appear to be quite sensible for all commodities in both the ICP and GTAP cases. Consider, for example, the values corresponding to the "Grains, other crops" category in Table 1. Its  $\boldsymbol{a}_n$  estimate indicates that at low income levels, this category accounts for as much as 21.8 cents of each additional dollar of expenditure. However, its  $\boldsymbol{b}_n$  estimate of zero suggests that at higher income levels, "Grains, other crops" is no longer part of any increases in expenditure (this is the remaining active bound in Table 1). Likewise, the  $\boldsymbol{b}_n$  value corresponding to "Grains, other crops" for the GTAP data is zero (similarly, this is the remaining active bound in Table 2). If we compare the

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<sup>&</sup>lt;sup>19</sup> Thanks are given to Dr. John Cranfield for providing the GAMS code to estimate AIDADS.

<sup>&</sup>lt;sup>20</sup> Note that we attempted a bootstrapping exercise similar to Cranfield et al., but were unable to carry it out because of computational difficulties. Discussions with those authors suggested this is related to the fact that we are working with a larger number of goods (10 versus their 6). Despite our inability to generate standard errors, our estimation procedure is identical to that of Cranfield et al., and our results are very similar to the extent that they can be compared. Moreover, there are other means of evaluation at our disposal, and it will be shown that these lend a great deal of support to our results.

<sup>&</sup>lt;sup>21</sup> These correlation coefficients give an indication of goodness-of-fit, but do not capture the important non-linearities that dominate. In the estimation procedure, the bounds of only four of the parameters in each of Table 1 and 2 were active (these correspond to the zeros within the top three rows of each table).

estimates of  $a_n$  between the two tables, we see that it is higher for the ICP-based estimates. There appear to be two primary reasons for this. First, the ICP data include more countries than do the GTAP data, and many of these additional countries are at very low income levels where staple grains dominate the budget. Second,  $a_n$  is higher when AIDADS is estimated with ICP data because wholesale/retail trade margins are included within each category. Whereas "Grains, other crops" is a retail good in the ICP data, it reflects a producer good in the GTAP data. Since margins are a luxury (see  $\hat{e}_n$  in Table 2), when they are combined with grains, we expect the resulting good to have a higher marginal budget share.

Before moving on, note that in both Tables 1 and 2,  $a_n$  does not equal  $b_n$  for any category of expenditure. This suggests that there is a great richness of behavior that would have been missed had we assumed that  $a_n$  equal  $b_n$ , such that marginal budget shares are constant, as in the LES demand system estimated by Hunter and Markusen (1988), among others.

Tables 1 and 2 also report expenditure elasticities evaluated at the means of the data  $(\hat{e}_n)$ . Again, the ICP and GTAP results are similar. For example, "Grains, other crops" has the lowest expenditure elasticity in both cases (it is 0.286 for ICP, 0.403 for GTAP). In contrast, expenditure categories relating to housing, health, and education services tend to have expenditure elasticities well above 1.00. For example, the ICP "Rent and housing utilities" and "Medical products and services" categories have expenditure elasticities of 1.200 and 1.322, respectively. Likewise, the GTAP category "Housing, education, health, public services" has an expenditure elasticity of 1.275. Overall, expenditure elasticities generated with GTAP versus ICP data are qualitatively similar.

We now move on to consider Figures 1a and 1b, which plot the ICP-based fitted budget shares evaluated at mean prices against the log of per capita household expenditure in 1996 U.S. dollars (the results have been split into two figures for clarity). First of all, note that in contrast to what would happen with demand systems of lower rank (e.g., AIDS, or Working's model), the fitted budget shares never become negative, even at extreme income levels. We see three basic shapes to the fitted shares: monotonically increasing, monotonically decreasing, and a *non*-monotonic pattern in which a good's share of the budget rises at low income levels, then falls at higher income levels. The most dramatic change over the observed income range occurs in the "Grains, other crops" category in Figure 1a, which goes from being the most important component of the consumption bundle at low income levels (43%), to the least important at high income levels (2%). In contrast to most of the food- and apparel-related commodities of Figure 1a, the commodities in Figure 1b have to do with services. Thus we see that budget shares increase in Figure 1b, and in the case of "Other goods and services" there is an interesting S-shaped curve. Clearly, AIDADS is a very flexible functional form when it comes to Engel responses.

Figures 2a and 2b plot the fitted budget shares when AIDADS is estimated with the GTAP data. In many respects the results are quite similar to those when ICP data is used, although one must be careful in making comparisons since none of the categories are identical, even those having the same name. Recall that whereas the ICP data include wholesale/retail trade in each of the consumption goods, the GTAP data distinguish those services as an individual category. This particular category, in fact, is quite an interesting feature and a unique contribution of the GTAP-based results. We see in Figure 2b that the proportion of the budget allocated to "Wholesale/retail trade" varies substantially over the income spectrum. It is estimated to be only 12% of the budget in the poorest countries, but takes up as much as 20% of the budget in the richest countries. This reflects the costs of overhead that might arise in the retailing of food, for example. Whereas a shopper in a developing country might purchase from a curbside vendor, a shopper in a rich country likely visits a

vast, air-conditioned supermarket with computerized checkout facilities. Indeed, to our knowledge, this is the first explicit evidence as to how wholesale/retail margins vary internationally over the per capita income spectrum.

We now turn to Figures 3a and 3b. As with Figures 1 and 2, these plot the fitted AIDADS budget shares evaluated at mean prices against the log of per capita household expenditure. These figures are different, however, in that they have been designed to facilitate a more direct comparison of the ICP- and GTAP-based results. The graphs have been limited to three commodities that correspond quite closely to each other. In some cases the fitted budget shares of one commodity have been added to another commodity to overcome problems with definitions.

The first "commodity" to be compared is an aggregate of "Grains, other crops" and "Processed food, beverages, tobacco". We see in Figures 3a and 3b that "Grains, other crops; Processed food, beverages, tobacco" has an exceptionally high share of the budget in the poorest countries (54% and 42% for ICP and GTAP, respectively), and a very low share of the budget at the highest income levels (6% and 8%). The decline in this category's share of the budget appears to be roughly linear in each case. The results would be even more similar if we had information on the wholesale/retail margins embedded within the ICP data, and incorporated these into the GTAP shares.

Figures 3a and 3b further reveal that "Meat, dairy, fish" has an interesting non-monotonic path. In the ICP data, this category's share of the budget is only 11% in the poorest countries, but then rises across the income spectrum, peaking at 15% near the income level of Indonesia. It then falls in importance until it takes up only 2% of the budget at the highest income level. Impressively, this same non-monotonic path is captured using the GTAP data. There the respective percentages for "Meat, dairy, fish" are 7%, 10%, and 6%. The fact that this subtle pattern is picked up by both of the datasets speaks well of the performance of each one. It is also remarkable, given the very different provenance of these two data sets.

The third commodity that can be compared is an aggregate relating to housing, health, education and other expenditures (Figures 3a and 3b). The ICP-based estimates suggest that this aggregate is about 15% of the budget in poor country households, but nearly half of the budget in the rich country households. The GTAP-based estimates suggest that the corresponding figures are about 14% and 36%. Again, the lower GTAP values reflect the fact that margins have a separate category, as well as any remaining differences in the definition of the category.

In summary, for many applications there is no indication that one would get substantively different results by using GTAP data in place of ICP data, or vice versa. The results of these two estimations are remarkably similar considering that goods definitions are not precisely the same, numerous countries are represented in one but not the other, and the raw data are collected from independent sources and compiled using distinct methodologies.<sup>23</sup> In this regard, the data sets strongly "corroborate" each other.

Given this finding we may want to ask: Which set of data is appropriate for endowing the consumption side of a general equilibrium model?<sup>24</sup> The ICP data have the advantage of more

<sup>&</sup>lt;sup>22</sup> Although each name is common to the data sets, there are still differences in definition, and combining them in this particular way limits the problems that otherwise result.

<sup>&</sup>lt;sup>23</sup> Additionally, our two sets of results compare favorably to the results of Cranfield et al. (2000, 2002), to the extent that comparisons are possible. Those studies were based upon 1985 ICP data, were aggregated into six goods having somewhat different definitions from ours, and were estimated with data from a somewhat different set of countries.

<sup>&</sup>lt;sup>24</sup> In order to actually use either of these demand systems in a general equilibrium model one further step will be necessary. One must calibrate the model to observed consumption shares. This step necessarily results in a country-specific set of preference parameters.

observations, and likely has a better representation of price differences across countries. However, in estimating AIDADS with and without any price variation, we find that there are almost no qualitative or quantitative differences in the Engel responses. Moreover, the GTAP data have the great advantage of already being in ISIC producer categories, making them compatible with a large number of production and trade data bases, such as GTAP and OECD STAN. Therefore, if a researcher needs to estimate a demand system for analyses involving more than just the study of consumption, then the GTAP data are quite appealing as a source of per capita consumption information. Goods and services are not combined within GTAP categories, wholesale/retail margins are already broken out, and no transition matrix between consumer and producer commodities has to be constructed.<sup>25</sup>

Furthermore, applying the ICP-based estimates directly to producer goods categories is likely to result in under-estimation of the aggregate demand for wholesale/retail trade services as income rises, since the GTAP-based results suggest that wholesale/retail margins become a larger part of expenditure as countries develop. This also means that use of a single transformation matrix between consumer and producer goods would be inappropriate.

#### 7. Using GTAP Data to Estimate the Components of Final Demand

In this section we move on to consider another way by which the AIDADS demand system can be estimated using GTAP expenditure data. This exercise further develops two of the themes of this study: (i) there are important, systematic differences in the way that countries allocate their expenditures, and (ii) GTAP expenditure data in conjunction with the AIDADS demand system provide an excellent means of understanding these patterns. Whereas previous sections of this study focused on the composition of household consumption, this section focuses on the composition of overall, regional final demand. Specifically, the patterns by which three components of final demand vary over the income spectrum are examined: household consumption (C), investment (I), and government (I).

As suggested in the Introduction, an exercise of this sort has a number of useful applications. For example, consider the theory of the "regional household" employed in the GTAP global AGE model (Hertel, 1997). In this model, each nation is assumed to have an aggregate "household" that acts as a sort of clearinghouse, whose behavior is governed by an aggregate utility function, specified over the three components of final demand. To operationalize this concept, GTAP modelers have limited the behavior of the regional household by way of restrictive assumptions. In this section it will be shown, however, that the regional household's behavior can be estimated, using the AIDADS demand system in conjunction with C, I, G information from the GTAP data base. It will be shown that there are substantial differences in final demand allocation related to per capita income.

This exercise can also shed light on a number of hypotheses proposed by economists concerning the composition of final demand. One such hypothesis is Wagner's Law, which states that government spending tends to grow at a higher rate than does GDP (Islam, 2001). In estimating AIDADS, we would have empirical support for Wagner's Law if government expenditure is found to have an increasing share of final demand among successively higher income countries.

<sup>&</sup>lt;sup>25</sup> This is not to suggest that a margins matrix (if it were available) would not provide useful information about the structure of demand. Ideally one would have a well-specified margins sector that handles these activities.

<sup>&</sup>lt;sup>26</sup> Here, "government expenditure" refers to direct spending on goods and services; it excludes transfers of income among citizens.

<sup>&</sup>lt;sup>27</sup> Actually, "Savings" is the second component of GTAP regional household expenditure, not "Investment". We choose to work in terms of investment, however, since savings is calculated as a residual in GTAP, and for some countries savings is negative.

We can also formulate some additional hypotheses about how the relative sizes of C, I, and G might vary across per capita income levels. Consider that the shapes of their fitted expenditure shares may be related to the stage of a country's economic development. Additionally, the shapes could be related to demography, since poor countries tend to have very young populations, middle-income countries have a high percentage of working people, and rich countries tend to have aging populations.<sup>28</sup> For purely economic reasons, there is likely little investment in very poor countries because of their low stage of development. Indeed, population growth may be outstripping economic growth, with only a small share of the population at working age. Additionally, the scale of government activity may be low, since the effectiveness of taxation is limited, organized interest groups are few, and the provision of public goods is not yet a priority. In contrast, the relative size of investment in middle income countries may be quite large, since these economies are likely to be growing at a rapid rate, and may have a higher proportion of the population in the workforce than countries at other stages of development.<sup>29</sup> In turn, a very different situation might prevail in high income countries. There, economic growth is likely to have slowed, and so the importance of investment in final demand may also be declining. Furthermore, Wagner's Law suggests that government's share of overall expenditures will be quite high, perhaps even overtaking investment as the second largest component of C+I+G.

Now that we have motivated why we should be interested in estimating these components of final demand, and have described what we might expect, we turn to the results. As before, they were generated using the maximum-likelihood technique of Cranfield et al. (2000) in conjunction with C, I, and G data for the 52 GTAP countries identified in Table A-3. Table 3 provides estimates of the AIDADS parameters  $\boldsymbol{a}_n$ ,  $\boldsymbol{b}_n$ ,  $\boldsymbol{g}_n$ , as well as expenditure elasticities evaluated at the means of the data  $(\boldsymbol{e}_n)$ .

Overall, the results are consistent with one's intuition about the relative size of C, I, and G at different per capita income levels. First, consider the "subsistence" expenditure shares  $\mathbf{g}_n$  in Table 3. Observe that the estimate is 0.000 and 0.050 for Investment and Government, respectively, but is 1.078 for private household consumption. This implies that Investment and Government expenditure are much less important at the lowest levels of income, while (unsurprisingly) private household expenditure is important. Now move on to consider the estimates of  $\mathbf{a}_n$  and  $\mathbf{b}_n$  in Table 3. Recall that these represent the bounds of the marginal expenditure shares in the AIDADS share equation. We first inspect the results for Government. Its  $\mathbf{a}_n$  estimate indicates that at low income levels, Government accounts for as little as 12.3 cents of each additional dollar of expenditure. However, its  $\mathbf{b}_n$  estimate suggests that the public sector accounts for as much 21.9 cents of each additional dollar of expenditure at higher income levels. Thus, our results are quite consistent with Wagner's Law, which states that Government is an increasingly larger share of GDP as countries develop. The high expenditure elasticity associated with Government (1.149) is also consistent with Wagner's Law.

<sup>&</sup>lt;sup>28</sup> Bloom and Williamson (1998) provide an interesting account of how demography relates to economic growth. To our knowledge the proposed relationship between demographic structure and investment has not been suggested elsewhere.

<sup>&</sup>lt;sup>29</sup> As an example, according to the U.N. World Population Prospects (http://esa.un.org/unpp), the number of children (aged 0-14) and retired people (aged 60+) per 100 people in the workforce is just 39 for South Korea (a fast-growing, middle income country), while the corresponding number is 52 for the USA (a slower growing, high income country) and 55 for Indonesia (a lower income country).

<sup>&</sup>lt;sup>30</sup> The only active bound during estimation was the subsistence expenditure share parameter for Investment (Table 3).

For Consumption, the marginal expenditure share results tend to move in the same direction as for Government (Table 3), but here  $\mathbf{a}_n$  and  $\mathbf{b}_n$  are quite close in value to one another (0.598 and 0.645). With regard to Investment's marginal expenditure shares, we see that at low income levels this category accounts for as much as 28.0% of each additional dollar of expenditure, but the corresponding value at higher expenditure by less is only 13.6%. This suggests that Investment is a smaller component of the expenditures of higher income countries.

We now move on and consider Figure 4, which plots the fitted shares of C, I, and G against the log of per capita C+I+G in 1997 U.S. dollars. Observe that there is a considerable amount of variation in these three components across the income spectrum. Household consumption's share of final demand is as high as 76.1% in the very poorest countries, but this share generally declines as countries become more developed, dropping as low as 61.5% in upper-middle income countries. Government spending, in contrast, claims a successively larger share of GDP, as suggested by Wagner's Law. Its extreme values occur in the poorest and richest countries of the sample (8.7% and 18.3%, respectively).

Finally, Figure 4 shows that Investment follows an interesting, non-monotonic path, and has a hump-backed shape. In the poorest countries, Investment's share is as low as 15.1%, and is the least important component of C+I+G. Its importance increases until it reaches a peak of 25.2% at the median income level, which corresponds to Brazil in this sample. Investment's importance then declines as countries reach higher stages of development, falling to 18.8% in the richest country. In short, there is strong, empirical confirmation of our earlier conjectures about how the fitted expenditure shares of final demand may vary across income levels. Long run economic projections that abstract from the relationships between final demand and per capita income found here, may miss an important aspect of the international economy.

#### 8. Conclusions

A variety of applications in economic research call for estimation of consumer demands at the international level, and for this behavior to be related back to production or trade data based on input-output accounts. In this study, a non-homothetic AIDADS demand system is estimated using cross-country, 1996 ICP data on per capita expenditures for each of 10 goods. This demand system is then re-estimated using per capita expenditure data from the GTAP version 5 data base, which corresponds to the year 1997. Both estimations make clear that the common assumption of homotheticity in demand is at odds with the data, as fitted budget shares vary to a great extent across the income spectrum. For example, the GTAP-based fitted budget shares for all food and beverages decline from 49% in the poorest countries, to only 13% of private household spending in the richest countries, which is a strong confirmation of Engel's Law. Per capita income is also found to explain much of the variation in international consumption for other goods and services. Thus we lend strong empirical support to Heston and Summers' contention that "the cliché that no single number tells more about households than their income also holds true for whole economies" (1996).

Another key finding is that the AIDADS results based on the GTAP data are remarkably consistent with those based on the ICP data, even though their origins and treatment of margins are quite different. In fact, for those categories that can be compared more or less directly, the GTAP data capture the same shapes in the fitted budget shares that are produced by the ICP data, even in those cases where the budget shares are non-monotonic in income. Given the similarity of results produced by the two sets of data, the estimation based on GTAP data is favored because its commodity definitions can be much more easily matched to production and trade data derived from input-output accounts. In addition, we find that the application of ICP-based demand estimates to such input-

output based modeling would likely be misleading when it comes to the consumption of margins and transportation services. This is due to the fact that GTAP-based results show wholesale/retail margins occupying an increasing share of consumer expenditure as per capita income rises. This also implies that use of a single transformation matrix between producer and consumer expenditures would not be suitable for international studies.

In an additional exercise, it is found that estimation of AIDADS with the GTAP data is an effective means for studying how the shares of consumption, investment, and government spending in final demand vary across countries. Empirical support for Wagner's Law is found, which suggests that the scale of government activity expands as countries grow richer. Additionally, an interesting, non-monotonic shape is uncovered regarding the shares of investment in final demand. The pattern that emerges appears to be related to a country's economic development, and to its demographic composition.

In summary, this technical paper offers a comprehensive analysis of alternative approaches to the estimation of consumer demand behavior across the income spectrum. The preferred demand system for use in the GTAP model is that based directly on GTAP consumption data. In the future, researchers may wish to incorporate the estimated AIDADS demand system directly into the GTAP model. In the mean time, a more practical approach involves using the fitted income elasticities of demand to calibrate the parameters for the existing GTAP model. Appendix D provides the GAMS code to generate these new elasticity targets. An electronic version of this file is also available for downloading.

Table 1. ICP-Based AIDADS Estimates: Household Consumption Expenditures

	Grains, other crops	Meat, dairy, fish	Processed food, beverages, tobacco	Apparel, footwear	Rent and housing utilities	Home furnishings and appliances	Medical products and services	Transport and communication	Recreation and education	Other goods and services
$\hat{m{a}}_n$	0.218	0.205	0.176	0.089	0.083	0.060	0.016	0.076	0.055	0.020
$\hat{m{b}}_n$	0.000	0.012	0.028	0.044	0.207	0.069	0.132	0.144	0.177	0.186
$\hat{m{g}}_n$	0.690	0.000	0.022	0.077	0.052	0.000	0.081	0.000	0.024	0.042
$\hat{m{e}}_n$	0.286	0.415	0.557	0.810	1.200	1.044	1.322	1.155	1.239	1.336
$\hat{m{r}}_n$	0.839	0.644	0.622	0.296	0.552	0.282	0.731	0.547	0.633	0.703

*Notes*:  $\hat{\boldsymbol{e}}_n$  are expenditure elasticities evaluated at the means of the data.  $\hat{\boldsymbol{r}}_n$  are correlation coefficients between the actual and fitted budget shares. Other estimates correspond to parameters in the AIDADS share equation.

Table 2. GTAP-Based AIDADS Estimates: Household Consumption Expenditures

		_						1	1	
	Grains, other crops	Meat, dairy, fish	Processed food, beverages, tobacco	Textiles, apparel, footwear	Utilities, other housing services	Whole- sale/retail trade	Manu- factures, elect- ronics	Transport, commun- ication	Financial and business services	Housing, education, health, public services
$\hat{m{a}}_{\scriptscriptstyle n}$	0.084	0.122	0.138	0.068	0.035	0.132	0.169	0.115	0.030	0.108
$\hat{m{b}}_n$	0.000	0.026	0.032	0.030	0.047	0.238	0.099	0.097	0.118	0.313
$\hat{m{g}}_n$	0.298	0.000	0.142	0.030	0.000	0.078	0.002	0.000	0.014	0.086
$\hat{\boldsymbol{e}}_n$	0.403	0.649	0.645	0.784	1.092	1.164	0.867	0.964	1.337	1.275
$\hat{m{r}}_n$	0.852	0.452	0.632	0.379	0.618	0.497	0.378	0.524	0.449	0.542

*Notes:* As discussed in the text, the ten GTAP-based goods of this table do not correspond directly to the ten ICP-based goods of Table 1.

Table 3. GTAP-Based AIDADS Estimates: Overall Final Demand Expenditures

	Consumption	Investment	Government
$\hat{\boldsymbol{a}}_{\scriptscriptstyle n}$	0.598	0.280	0.123
$\hat{m{b}}_n$	0.645	0.136	0.219
$\hat{m{g}}_n$	1.078	0.000	0.050
$\hat{m{e}}_n$	1.012	0.864	1.149

Figure 1a. ICP-Based Fitted AIDADS Budget Shares (Goods 1-5)

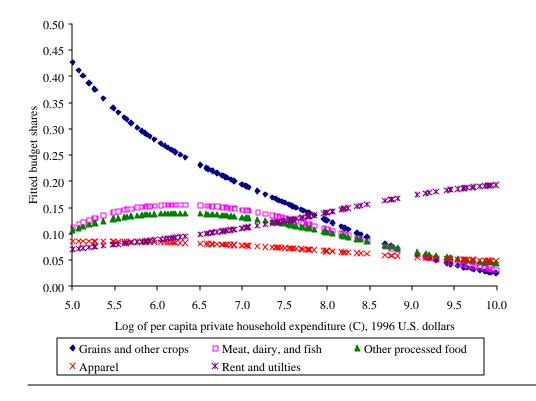


Figure 1b. ICP-Based Fitted AIDADS Budget Shares (Goods 6-10)

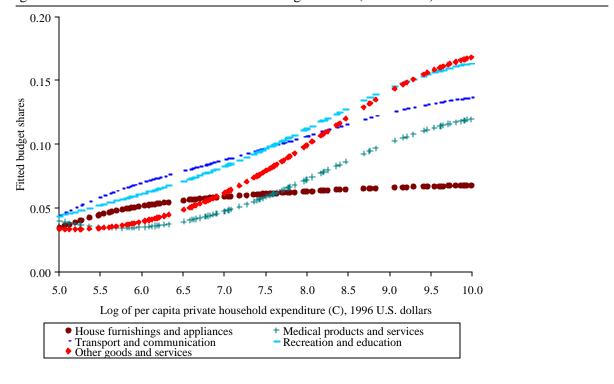


Figure 2a. GTAP-Based Fitted AIDADS Budget Shares (Goods 1-5)

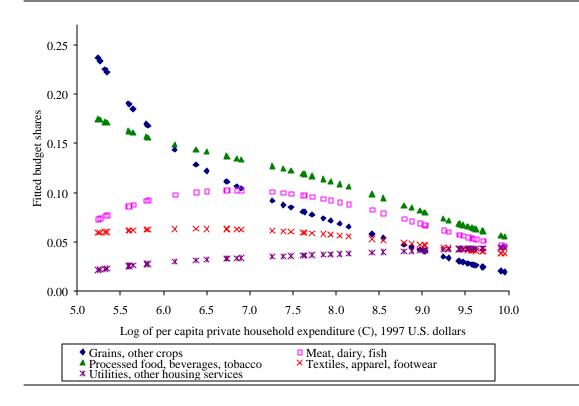


Figure 2b. GTAP-Based Fitted AIDADS Budget Shares (Goods 6-10)

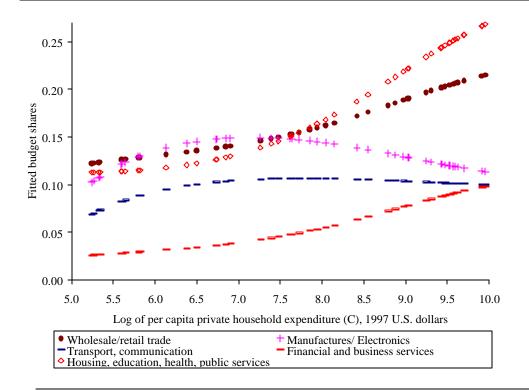


Figure 3a. Selected ICP-Based Fitted Budget Shares

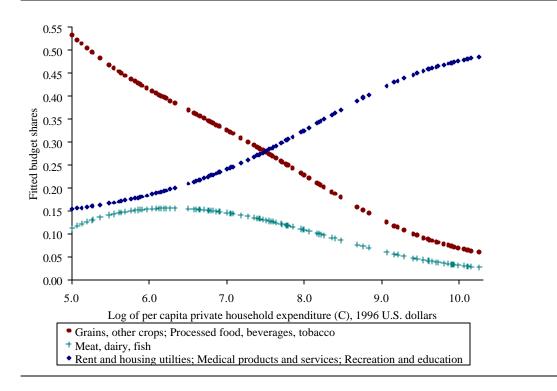


Figure 3b. Selected GTAP-Based Fitted Budget Shares

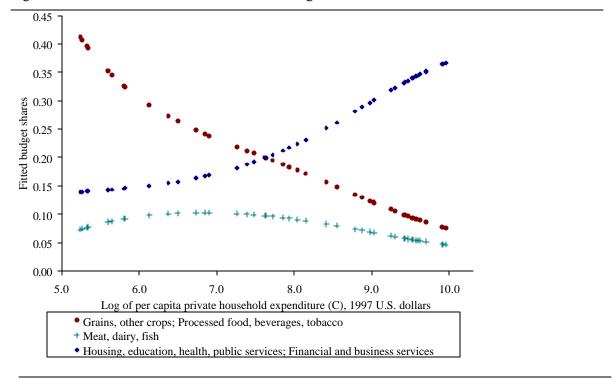
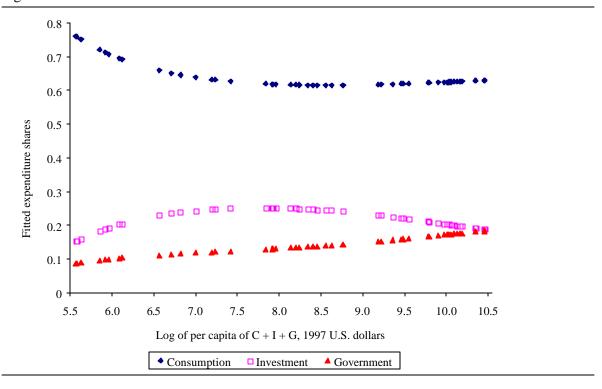


Figure 4. GTAP-Based Fitted Shares of Final Demand



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### Appendix A: Aggregation of Goods and Regions In this appendix several tables are presented that identify the countries and commodities used

In this appendix several tables are presented that identify the countries and commodities used in estimating the AIDADS demand system. Further description of these data is provided in the main text.

GTAP Codes	ISIC Revision 3 Codes	Central Product Classification Codes
1. Grains, other crops (G		
1 pdr	• •	011*
2 wht		011*
3 gro		011*
4 v_f		012, 013
5 osd		014
6 c_b		018
8 ocr		015, 016, 017, 019*
23 pcr		231
2. Meat, dairy, fish (Mea	ntDairy)	
9 ctl	(Duny)	021*, 029*
10 oap		021*, 029*
11 rmk		029*
14 fsh	015, 05	02)
19 cmt	010, 00	211*, 216*
20 omt		211*, 216*
22 mil		22
3. Processed food, bever	ages, tobacco (OthFoodBev)	
21 vol		216*, 217, 218
24 sgr		235
25 ofd		21*, 23*
26 b_t		24, 25
4. Textiles, apparel, footy	wear (TextAppar)	
7 pfb		019*
12 wol		029*
27 tex	17, 243	
28 wap	18	
29 lea	19	
5. Utilities, other housing	g services (HousUtils)	
15 col	101, 102	
17 gas	111*, 112*	
43 ely	401	
44 gdt	402, 403	
45 wtr	41	
46 cns	45	
53 isr	66	
6. Wholesale/retail trade		
47 trd	50, 51, 52*, 55	

Table A-1.	<b>GTAP Commodity</b>	Aggregation	Continued

GTAP Codes	ISIC Revision 3 Codes	Central Product Classification Codes				
7. Manufactures, electro	7. Manufactures, electronics (Mnfcs)					
18 omn	12, 13, 14					
30 lum	20					
31 ppp	21, 221*, 222, 223					
33 crp	241, 242, 25					
34 nmm	26					
35 i_s	271, 2731					
36 nfm	272, 2732					
37 fmp	28					
38 mvh	34					
40 ele	30, 32					
41 ome	29, 31, 33					
42 omf	36, 37					
8. Transport, communic						
16 oil	111*, 112*, 103					
32 p_c	231, 232, 233					
39 otn	35					
48 otp	60, 63					
49 wtp	61					
50 atp	62					
51 cmn	64					
9. Financial and busines	s services (FinService)					
52 ofi	65, 67					
54 obs	70, 71*, 72, 73, 74					
10. Housing, education,	health, public services (HousOthServ					
55 ros	92, 93, 95					
56 osg	75, 80, 85, 90, 91, 99					

<sup>56</sup> osg 75, 80, 85, 90, 91, 99
57 dwe
\* Denotes that only a sub-section of this category is included.

Table A-2.	ICP Commodity Aggregation
ICP Aggregate	
1. Grains, other	
	Bread and cereals
	Fruit, vegetables and potatoes
2. Meat, dairy, a	nd figh
2. Meat, dairy, a	Meat
	Fish
	Milk, cheese and eggs
	Mirk, cheese and eggs
3. Processed for	od, beverages, tobacco
	Oils and fats
	Other food
	Non-alcoholic beverages
	Alcoholic beverages
	Tobacco
4. Apparel, foot	wear
	Clothing, including repairs
	Footwear, including repairs
5. Rent and hous	
	Gross rent and water charges, including maintenance
	Fuel and power
6. Home furnish	ings and appliances
	Furniture and floor coverings, including repairs
	Other household goods and services
	Household appliances and repairs
7. 16. 17. 1	
/. Medical prod	ucts and services
	Medical care, including both services and goods
8. Transport and	communication
	Personal transportation equipment, repairs
	Operation of transportation equipment
	Purchased transport services
	Communication
9. Recreation an	ad education
	Recreation and cultural goods and services
	Education services
10. Other goods	and services
	Restaurants, cafes and hotels
	Other goods and services, nec

Table A-3. GTAP Regions for Demand System Estimation

No.	Region name	Code	Per Capita Household Expenditure (1997 US\$)	Per Capita GDP (1997 US\$)
1	Tanzania	TZA	189	216
2	Mozambique	MOZ	193	219
3	Vietnam	VNM	205	274
4	Malawi	MWI	209	290
5	India	IND	269	336
6	Uganda	UGA	271	356
7	Bangladesh	BGD	283	416
8	Zambia	ZMB	331	446
9	China	CHN	337	695
10	Zimbabwe	ZWE	460	720
11	Sri Lanka	LKA	588	841
12	Indonesia	IDN	664	1,042
13	Philippines	PHL	836	1,097
14	Morocco	MAR	843	1,280
15	Botswana	BWA	941	1,566
16	Rest of world	XRW	994	2,361
17	Thailand	THA	1,423	2,658
18	Colombia	COL	1,619	2,664
19	Peru	PER	1,767	3,082
20	Malaysia	MYS	2,033	3,116
21	Turkey	TUR	2,070	3,222
22	Poland	POL	2,253	3,676
23	Venezuela	VEN	2,568	4,028
24	Mexico	MEX	2,800	4,140
25	Hungary	HUN	2,802	4,825
26	Brazil	BRA	3,108	4,896
27	Chile	CHL	3,453	5,208
28	Uruguay	URY	4,517	5,837
29	South Korea	KOR	5,160	9,138
30	Portugal	PRT	6,553	9,138 9,687
31	Argentina	ARG	7,148	9,901
32	Taiwan	TWN	7,146	10,961
33	Greece	GRC	8,293	13,248
33 34			8,385	13,747
3 <del>4</del> 35	Spain Ireland	ESP	8,383 10,346	
		IRL		17,303
36	New Zealand	NZL	10,991	18,149
37 38	Finland	FIN	12,339	19,608
	Canada	CAN	12,352	21,047
39	Italy	ITA	12,490	21,198
40	Australia	AUS	12,974	21,818
41	Sweden	SWE	13,668	22,703
42	Netherlands	NLD	13,781	22,719
43	France	FRA	14,330	23,007
44	Austria	AUT	14,352	23,485
45	United	GBR	14,363	23,600
46	Germany	DEU	14,693	24,426
47	Belgium	BEL	15,105	25,117
48	Luxembourg	LUX	16,261	25,729
49	Denmark	DNK	16,341	29,260
50	Japan	JPN	20,086	31,149
51	United States	USA	20,237	33,750
52	Switzerland	CHE	21,059	35,190

Table A-4. ICP Regions for Demand System Estimation

No.	Region Name	Per Capita Real Household Expenditure (1996 US\$)	Per Capita Nominal Household Expenditure (1996 US\$)
1	Tanzania	424	160
2	Nigeria	600	180
3	Tajikistan	700	149
4	Zambia	720	281
5	Yemen	720	348
6	Malawi	792	194
7	Madagascar	794	255
8	Mali	826	214
9	Mongolia	889	297
10	Benin	1,017	319
11			239
	Kenya	1,129	
12	Nepal	1,197	168
13	Sierra Leone	1,209	192
14	Turkmenistan	1,258	148
15	Congo	1,356	529
16	Senegal	1,445	419
17	Bangladesh	1,462	273
18	Vietnam	1,496	241
19	Pakistan	1,694	390
20	Azerbaijan	1,776	369
21	Cote D'Ivoire	1,817	500
22	Kyrgyzstan	1,957	357
23	Uzbekistan	1,995	472
24	Cameroon	2,008	499
25	Moldova	2,009	336
26	Bolivia	2,103	764
27	Ecuador	2,138	1,100
28			
	Sri Lanka	2,223	561
29	Armenia	2,238	441
30	Jordan	2,362	1,183
31	Albania	2,567	737
32	Indonesia	2,587	719
33	Jamaica	2,631	1,094
34	Zimbabwe	2,648	486
35	Guinea	2,713	456
36	Syria	2,871	816
37	Georgia	2,899	670
38	Ukraine	3,182	875
39	Philippines	3,455	863
40	Peru	3,490	1,816
41	Botswana	3,501	956
42	Panama	3,528	1,688
43	Thailand	3,567	1,659
<del>4</del> 3	Morocco	3,656	919
44	Venezuela	3,687	1,900
46	Macedonia	3,856	1,160
47	Belize	3,863	2,003
48	Egypt	3,869	899
49	St. Vincent & Gren.	3,894	1,943
50	Swaziland	4,093	774
51	Lebanon	4,181	2,942
52	Belarus	4,237	1,439

Table A-4. ICP Regions for Demand System Estimation Continued

Capita Nominal cold Expenditure (1996 US\$)  2,360 994 1,561 3,168 997 1,745 2,858
2,360 994 1,561 3,168 997 1,745
994 1,561 3,168 997 1,745
1,561 3,168 997 1,745
3,168 997 1,745
997 1,745
1,745
2,412
1,979
2,243
1,626
1,146
1,000
2,563
3,524
3,183
4,417
2,396
2,973
2,414
2,069
1,312
3,678
4,244
2,202
2,966
5,903
3,828
6,844
9,932
3,608
9,518
6,398
8,614
10,760
12,200
2,423
6,335
12,748
12,772
16,785
9,528
19,960
16,411
14,913
19,153
17,133
17,993
17,993 21,761
17,993 21,761 14,990
17,993 21,761 14,990 19,836
17,993 21,761 14,990

Table A-4. ICP Regions for Demand System Estimation Continued

No.	Region Name	Per Capita Real Household Expenditure (1996 US\$)	Per Capita Nominal Household Expenditure (1996 US\$)
106	Canada	15,861	13,884
107	Bermuda	16,287	24,303
108	Hong Kong	16,444	14,794
109	Switzerland	16,527	28,450
110	Barbados	16,566	4,770
111	Iceland	16,683	19,832
112	Denmark	16,836	24,646
113	Luxembourg	20,235	25,672
114	USA	20,592	20,592

#### Appendix B: Modifications Needed for the 1996 ICP Data

This appendix documents several problems concerning the 1996 ICP data set contained within the "ToolPak" aggregation program obtained from the World Bank. These 1996 ICP data (described earlier in the main text) were not as refined as the 1985 Phase IV ICP data used by Cranfield et al. (2000, 2002) and others. Below we document our approach to dealing with these problems. In the end we believe that we have a very high quality data set, which is available to readers on request.

We first discovered errors concerning the "real expenditure" data that are calculated using the 1996 ICP ToolPak aggregation software. In particular, the PPP rates for several Asian nations were not expressed relative to the U.S. dollar (as in all other countries), but according to the Hong Kong dollar.<sup>31</sup> Following the approach of Regmi, Seale, and Bernstein (2001), we used 2001 World Bank "World Development Indicators" (WDI) data to get an estimate of Hong Kong's 1996 per capita expenditure relative to U.S. per capita expenditure in 1996. We calculated this ratio to be 79.86%. Accordingly, the real expenditures of all the mischaracterized Asian countries were multiplied by 24.78. <sup>32</sup> The approach and accuracy of the resulting real expenditure figures were verified as correct through graphical comparison with WDI and GTAP data, as well as in correspondence with researchers at the U.S. Dept. of Agriculture, Economic Research Service.

The next two corrections concern the "nominal expenditure" data calculated with the 1996 ICP ToolPak software. The problem is that the nominal expenditure data were still in Local Currency Units (LCUs). For all but six nations we were able to convert the figures to the correct units using exchange rates out of the 2001 World Bank WDI data for 1996. For the remaining six countries we obtained other sources of exchange rate information. Average 1996 exchange rates for four countries (Bahamas, Bermuda, Qatar, and Kazakhstan) were found in Oanda (2002). Average 1996 exchange rates for the remaining two countries (Oman and Turkmenistan) were obtained from the Central Intelligence Agency World Fact Book (2002).

Even after adjusting nominal expenditure with exchange rate data, four nations (Vietnam, Belarus, Bulgaria, Russia) still had total nominal per capita expenditures that were large outliers in plots of the data. Comparison of these nations' ICP-based expenditures to 1996 WDI household per capita final consumption expenditure in current (1996) LCUs as well as in current U.S. dollars revealed the nature of the disparity. The values for these four nations deviated by a factor of 1000, meaning that the problem was likely related to units (i.e., measurement in thousands versus millions). To solve this problem we multiplied the LCU nominal expenditures for Vietnam by 0.001, and the LCU nominal expenditures for Bulgaria, Belarus, and Russia by 1000.

The final problem that we encountered concerned the expenditure category "Restaurants, Cafes, and Hotels" for Egypt and Morocco. In contrast to all other values within the 1996 ICP data, these values were negative (-2,756 and -968, respectively). This does not make economic sense, and would also hinder the estimation of a demand system. In correspondence with the World Bank (2002), it was suggested that this was in fact an error, and suggested that these two values be "dropped" from the data. Before doing this, however, we inspected the 1985 ICP data set used by Cranfield et al. (2000, 2002). We found that in disaggregated 113-commodity form, this data set also contained a

<sup>32</sup> This is because the unmodified ICP per capita expenditure for Hong Kong is 664. It needs to be scaled up to 16,444 in order to be exactly 79.86% of the U.S. ICP per cap expenditure. 664 multiplied by 24.78 is 16,444. Note that one should multiply only real expenditure (not nominal expenditure or PPP rates) by 24.78.

<sup>&</sup>lt;sup>31</sup> The set of nations for which this is a problem includes Bangladesh, Fiji, Hong Kong, Indonesia, Iran, Korea, Nepal, Pakistan, Philippines, Singapore, Sri Lanka, Thailand, and Vietnam.

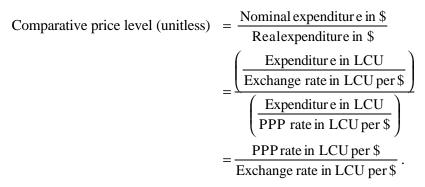
number of negative values. These authors did not mention this issue, perhaps because in aggregating the data up to six categories, the negative values were cancelled out and no longer apparent. Since we also needed to aggregate the 1996 data, we decided to not drop these two values. In aggregating the ICP data from 26 to 10 categories, the two negative values were combined with larger positive values, and the resulting commodity aggregates were no longer negative.

Upon making all of the above corrections we again checked whether several aggregated values, such as per capita household expenditure and GDP, were comparable to those of the World Bank WDI and GTAP data. In scatterplots and in tabular comparisons we found no problems of significance. We also used our refined data to estimate a global AIDADS demand system, and obtained results that are very reasonable, and similar in many ways to results based on previous ICP data sets, as well as the GTAP data. As a result, we believe that our refined 1996 ICP data set is an excellent resource for many types of international consumption analysis.

### Appendix C: Notes on ICP Terminology and the Calculation of Budget Shares

The terminology of the International Comparison Program (ICP) data can be somewhat confusing because its concept of prices and its use of the terms "real" and "nominal" are somewhat different from that of other areas of economics. To contribute to a better understanding of ICP concepts, this appendix briefly describes some of the terminology used, and relates concepts such as "comparative price levels" and "purchasing power parity" (PPP) to each other. We also discuss the proper calculation of budget shares using ICP data. Most of this section draws from the *Handbook of the International Comparison Program* (United Nations, 1992). This document is an authoritative source, and provides more detail than this appendix, should that be desired.

First of all, in contrast to its meaning in other contexts, "nominal expenditure" in ICP terminology refers to expenditure in a base country's currency units (such as in U.S. or "international" dollars). In particular, "nominal expenditure" is found by dividing expenditure in local currency units (LCUs) by the appropriate international exchange rate. The ICP concept of "real expenditure" is also expressed in terms of a base country's currency units, but it is found by dividing LCU expenditures by ICP PPP rates. So, both "nominal" and "real" expenditure are expressed in the same currency units (and in contrast to other areas of economics, have nothing to do with adjustments for inflation, for example). Note, in turn, that "real expenditure" is sometimes referred to as "real product", "real output", or even "quantity". Dividing "nominal expenditure" by "real expenditure" results in what is called the "comparative price level". Thus, it is incorrect to refer to the "comparative price level" as the "PPP rate". In fact, the "comparative price level" is the PPP rate divided by the exchange rate. These relationships can be summarized by the following equations:



In passing, note this implies that the product of ICP "real expenditure" and the "PPP rate" should equal the product of ICP "nominal expenditure" and the "exchange rate".

We now move on from discussion of terminology and consider the issue of how budget shares should be calculated when ICP data are used. This issue arises, for example, when one needs estimates of household budget shares to serve as a dependent variable in demand system estimation. We first assume that the "true" budget shares are those that would be calculated using a country's expenditure in local currency units (LCUs). However, an ICP data set does not always contain the original expenditures in a country's own currency. The question then arises: Should budget shares based on ICP "nominal" or "real" expenditures be used? We propose that "nominal expenditures" be used for calculating budget shares in international demand analysis. The reason is that budget shares calculated using "nominal" expenditure should be identical to the budget shares calculated using a country's own currency. This is because the exchange rate used to convert LCU expenditure into

"nominal" expenditure does not vary by good. In contrast, "real" expenditure is converted from LCU expenditure using PPP rates, which differ across the disaggregated components of national GDP, including the goods comprising household consumption.

#### Appendix D: Household Consumption Expenditure Elasticity Estimation

In this appendix a GAMS program is presented that has been designed to compute expenditure elasticities for any country for which the per capita expenditures of the 10 categories in Table A-1 are known. This can be carried out for a GTAP country outside of the sample, or for the projected expenditure for one of the countries in the sample. The GAMS program uses AIDADS parameter estimates generated with the GTAP data as described in the main text, but these estimates can be easily changed as needed. The GAMS program requires an input file called "data.prn", which is essentially a table that has the VIMS, VIWS, and VPA data described below. These data come from the GTAP Baseview.har file that results from aggregating data in GTAPAgg as shown in Table A-1. Once the GAMS program runs successfully, one only needs to look at the output at the very bottom for the parameter "Elas" to get the expenditure elasticities specific to the 10 goods. The program is set up to work for one country at a time (this could be changed if desired).

```
* GAMS program to calculate expenditure elasticities for one country at a time
* Written by Jeff Reimer, November 2002
option decimals=4 ;
sets i /c1*c10/
    k /VIMS, VIWS, VPA/;
* To run this program, this is the information required from the GTAP database:
  (1) VIMS (imports valued at domestic market prices, summed over all sources)
  (2) VIWS (imports valued cif, summed over all sources)
  (3) VPA (household expenditure valued at agent's prices, in Baseview.har
      this is Cost Structure of Consumption, Sum DIR, Sum PURCHVALUE)
* (4) Population (the unit is a single person, NOT in millions or otherwise)
* Each time you change the country, you must change the values within the
* "data.prn" file. This can be modified using Microsoft Excel. Make sure that
* the path is correct.
$include c:\data\reimer\GTAPtech\data.prn
parameter alpha(i)
 / c1 0.084, c2 0.122, c3 0.138, c4 0.068, c5 0.035, c6 0.132, c7 0.169
  c8 0.115, c9 0.030, c10 0.108 / ;
parameter beta(i)
 / c1 0.000, c2 0.026, c3 0.032, c4 0.030, c5 0.047, c6 0.238, c7 0.099
  c8 0.097, c9 0.118, c10 0.313 / ;
parameter gamma(i)
 / c1 0.298, c2 0.000, c3 0.142, c4 0.030, c5 0.000, c6 0.078, c7 0.002
  c8 0.000, c9 0.014, c10 0.086 / ;
scalar kappa
/ 1.977 / ;
alias(i,j);
Parameters u, price(i), nomexp(i), realexp(i), x(i), m, phi(i), what(i), xhat(i)
          mbs(i), elas(i);
price(i) = data(i, "VIMS")/data(i, "VIWS");
nomexp(i) = 1000000*data(i, "VPA") ;
realexp(i) = nomexp(i)/price(i) ;
x(i) = realexp(i)/population/100;
```

The table below gives the expenditure elasticities for countries within the GTAP sample. For a given commodity, the elasticity will vary across countries according to a country's per capita income, and its price for that commodity. Country names and per capita incomes are provided in Table A-3.

Table D-1. Expenditure elasticities evaluated at observed, country-specific price levels

Region code	Grains, other crops	Meat, dairy, fish	Processed food, beverages, tobacco	Textiles, apparel, footwear	Utilities, other housing services	Whole- sale/ retail trade	Manu- factures, electronics	Tran- sport, commun- ication	Financial and business services	Housing, education, health, public services
TZA	0.342	1.633	0.765	1.128	1.662	1.106	1.625	1.649	1.246	1.011
MOZ	0.361	1.591	0.761	1.071	1.620	1.097	1.584	1.608	1.233	1.007
VNM	0.376	1.566	0.726	1.083	1.596	1.106	1.561	1.582	1.241	1.022
MWI	0.384	1.538	0.737	1.060	1.569	1.100	1.535	1.556	1.232	1.019
IND	0.433	1.359	0.778	1.052	1.399	1.088	1.364	1.382	1.214	1.040
UGA	0.388	1.377	0.836	1.079	1.418	1.102	1.383	1.400	1.232	1.055
BGD	0.440	1.327	0.798	1.046	1.369	1.085	1.334	1.351	1.211	1.044
ZMB	0.484	1.243	0.844	1.040	1.291	1.045	1.254	1.271	1.173	1.047
CHN	0.415	1.308	0.811	1.076	1.356	1.117	1.318	1.335	1.245	1.089
<b>ZWE</b>	0.520	1.164	0.839	1.017	1.227	1.088	1.182	1.200	1.206	1.077
LKA	0.558	1.092	0.865	0.995	1.173	1.084	1.118	1.139	1.228	1.113
IDN	0.613	1.054	0.849	0.985	1.147	1.080	1.085	1.108	1.232	1.121
PHL	0.587	1.022	0.872	0.976	1.135	1.096	1.061	1.089	1.261	1.155
MAR	0.614	1.026	0.853	0.973	1.133	1.093	1.063	1.089	1.252	1.148
BWA	0.651	0.983	0.865	0.957	1.112	1.089	1.029	1.059	1.262	1.158
XRW	0.623	0.994	0.853	0.963	1.121	1.099	1.039	1.069	1.271	1.169
THA	0.643	0.936	0.837	0.930	1.104	1.115	0.998	1.038	1.306	1.208
COL	0.648	0.904	0.832	0.915	1.097	1.120	0.977	1.022	1.319	1.221
PER	0.651	0.896	0.833	0.911	1.095	1.119	0.972	1.018	1.319	1.221
MYS	0.611	0.869	0.818	0.898	1.097	1.135	0.959	1.011	1.342	1.248
TUR	0.620	0.871	0.807	0.901	1.098	1.136	0.960	1.012	1.343	1.249
POL	0.621	0.865	0.791	0.897	1.099	1.141	0.957	1.011	1.350	1.256
VEN	0.615	0.824	0.783	0.871	1.093	1.144	0.934	0.995	1.353	1.263
MEX	0.594	0.810	0.773	0.866	1.094	1.149	0.927	0.992	1.360	1.271
HUN	0.596	0.815	0.770	0.869	1.096	1.150	0.930	0.995	1.362	1.272
BRA	0.590	0.790	0.759	0.852	1.093	1.152	0.918	0.987	1.361	1.275
CHL	0.567	0.768	0.743	0.841	1.093	1.156	0.908	0.982	1.361	1.279
URY	0.511	0.718	0.702	0.815	1.094	1.163	0.888	0.973	1.360	1.285

Table D-1. Expenditure elasticities evaluated at observed, country-specific price levels

			D 1		*******	****	<u> </u>		E' '1	Housing,
	Crains	Moot	Processed	Toutiles	Utilities,	Whole-	Manu	Tran-	Financial	education,
Region	Grains, other	Meat, dairy,	food, beverages,	Textiles, apparel,	other housing	sale/ retail	Manu- factures,	sport, commun-	and business	health, public
code	crops	fish	tobacco	footwear	services	trade	electronics	ication	services	services
KOR	0.458	0.698	0.681	0.807	1.098	1.171	0.883	0.973	1.366	1.293
PRT	0.402	0.649	0.647	0.785	1.090	1.161	0.868	0.964	1.330	1.269
ARG	0.385	0.638	0.637	0.780	1.090	1.162	0.865	0.963	1.328	1.269
TWN	0.352	0.624	0.623	0.775	1.089	1.160	0.862	0.962	1.320	1.264
GRC	0.335	0.617	0.622	0.775	1.085	1.153	0.862	0.962	1.302	1.251
ESP	0.332	0.616	0.620	0.774	1.086	1.154	0.862	0.962	1.303	1.251
IRL	0.271	0.598	0.607	0.773	1.079	1.142	0.862	0.961	1.271	1.228
NZL	0.257	0.595	0.605	0.773	1.077	1.138	0.863	0.962	1.263	1.221
FIN	0.224	0.589	0.600	0.775	1.074	1.132	0.865	0.963	1.248	1.209
CAN	0.225	0.584	0.594	0.770	1.076	1.136	0.862	0.961	1.257	1.217
ITA	0.221	0.590	0.602	0.776	1.073	1.130	0.866	0.963	1.244	1.206
AUS	0.214	0.590	0.602	0.777	1.072	1.128	0.867	0.963	1.239	1.202
SWE	0.199	0.590	0.604	0.780	1.070	1.124	0.869	0.964	1.229	1.195
NLD	0.196	0.590	0.603	0.780	1.069	1.123	0.869	0.964	1.229	1.194
FRA	0.187	0.591	0.605	0.783	1.068	1.120	0.871	0.964	1.222	1.189
AUT	0.187	0.591	0.606	0.783	1.067	1.120	0.871	0.964	1.221	1.188
GBR	0.186	0.590	0.604	0.782	1.068	1.121	0.870	0.964	1.223	1.189
DEU	0.182	0.591	0.606	0.784	1.067	1.119	0.872	0.965	1.218	1.186
BEL	0.175	0.592	0.607	0.786	1.066	1.117	0.873	0.965	1.214	1.182
LUX	0.158	0.596	0.612	0.791	1.063	1.111	0.877	0.966	1.202	1.173
DNK	0.159	0.596	0.612	0.791	1.062	1.110	0.877	0.966	1.201	1.172
JPN	0.100	0.599	0.616	0.800	1.059	1.103	0.884	0.969	1.185	1.159
USA	0.116	0.612	0.630	0.808	1.054	1.096	0.888	0.970	1.171	1.147
CHE	0.090	0.591	0.602	0.796	1.060	1.105	0.881	0.968	1.188	1.162