

The World's Largest Open Access Agricultural & Applied Economics Digital Library

This document is discoverable and free to researchers across the globe due to the work of AgEcon Search.

Help ensure our sustainability.

Give to AgEcon Search

AgEcon Search
http://ageconsearch.umn.edu
aesearch@umn.edu

Papers downloaded from **AgEcon Search** may be used for non-commercial purposes and personal study only. No other use, including posting to another Internet site, is permitted without permission from the copyright owner (not AgEcon Search), or as allowed under the provisions of Fair Use, U.S. Copyright Act, Title 17 U.S.C.

Impacts of Chinese Tariff on World Soybean Markets

Ethan Sabala

Department of Agricultural and Applied Economics

Texas Tech University

Lubbock, Texas.

E-mail: ethan.sabala@ttu.edu

Stephen Devadoss

Department of Agricultural and Applied Economics

Texas Tech University

Lubbock, Texas.

E-mail: stephen.devadoss@ttu.edu

Selected Paper prepared for presentation at the Southern Agricultural Economics Association (SAEA) Annual Meeting, Birmingham, Alabama, February 2-5, 2019

Copyright 2019 by Ethan Sabala and Stephen Devadoss. All rights reserved. Readers may make verbatim copies of this document for non-commercial purposes by any means, provided that this copyright notice appears on all such copies.

Impacts of Chinese Tariff on World Soybean Markets

Introduction

The ongoing trade war between the United States and several other countries has been simmering since the U.S. steel tariff and intellectual property dispute between China and the United States. As a result of this trade war, several trading partners retaliated by imposing tariffs on U.S. goods, particularly agricultural commodities. On July 6, 2018, China implemented retaliatory tariffs on imports of many U.S. commodities worth \$50 billion; of which farm products, automobiles, and aquatic products account for \$34 billion (ChineseMinistry, 2018). The chief among China's targeted commodities is soybeans because it is the United States' leading agricultural commodity exported to China. Currently, the Chinese tariff on U.S. soybeans is 25%. According to FAS (2018c), the United States supplies 42% of Chinese soybean imports, which accounts for 62% of U.S. exports (FAS, 2018a). Soybean trade between the United States and China totaled \$14 billion last year alone (Good, 2018b), and consequently, a 25% soybean tariff would have dramatic effects on both countries' soybean markets. This tariff has already caused a sharp decline in U.S. soybean price, as well as an increase in Chinese soybean imports from Brazil (Good, 2018b,c), and these effects will only be exacerbated if the tariff is prolonged over a period of time.

Since the United States and China are major players in the world soybean market, they have significant market share (of exports by the United States and imports by China) in the world soybean market and can influence world price and trade volume, which impacts other exporting and importing countries' soybean markets. For example, China can divert its imports from the United States to U.S. competitors: Brazil and Argentina. This would cause the soybean price from these other exporters to rise while the U.S. soybean price falls. Consequently, other soybean importers, such as the European Union, may start purchasing soybeans from the United States. As a result, the United States would regain some of the lost export market in China. Furthermore, small exporters, such as Canada, that typically sell to these other importers would lose much of their export market, and therefore suffer the

consequences of a trade war that they had no part in. These are just a few of the reallocations that would occur as a result of the Chinese tariff. Additionally, trade changes occurring in each of these countries would likely impact prices, supply, and demand domestically, which will reverberate into the world soybean market. The spatial equilibrium model (SEM), pioneered by Samuelson (1952) and popularized by Takayama and Judge (1971), is highly suited to study price, bilateral trade flow, supply, and demand impacts resulting from policy changes.

The goal of this study is to analyze and quantify the effects of the Chinese soybean tariff on U.S., Chinese, and other major importers' and exporters' soybean markets. To achieve this goal, first, we develop a theoretical model and obtain analytical results of a tariff. Second, we construct an empirical SEM of the world soybean market. Third, we run a baseline using this empirical model to solve for prices, production, consumption, and bilateral trade flows. Finally, we run an alternate scenario with the Chinese 25% tariff and compare the values of the endogenous variables to those in the baseline to quantify the impacts. The next section briefly reviews the literature on the SEM and its applications to trade analysis, findings of recent studies that focused on the soybean tariff, and our contribution to the literature and policy debates. Section 3 develops a stylized spatial equilibrium theoretical model to analyze how policies impact the prices, trade flows, supply, and demand of each country. Section 4 presents the empirical SEM. Section 5 covers the data and sources and construction of supply and demand functions for the eleven countries included in the model. Section 6 presents the results of the baseline and alternate scenarios to quantify the effects of the Chinese tariff on the world soybean market. The final section discusses policy implications and concludes the paper.

Literature Review

The spatial equilibrium models of Samuelson (1952) and Takayama and Judge (1971) have been extensively utilized by economists to study the impacts of trade policies. The SEM is very useful to examine how trade flows are reallocated among trade partners due to domestic and/or trade policies. As discussed in the introduction, the trade flows of several countries would be affected by the implementation of the Chinese soybean tariff, and the SEM captures these ripple effects. However, with this capability comes a limitation in that the model assumes soybeans are a homogeneous commodity from all suppliers, and thus importers will buy soybeans based solely on the lowest purchase price. This may lead to trade reallocation results that do not strictly reflect real-world trade flows. For instance, the model may suggest that an entire export market will be lost as a result of the tariff, when in reality only a portion of the market is lost. This is because real-world trade decisions account for many non-economic factors such as time lag, trade loyalty, contractual agreements, and political incentives, that are not readily amenable for modeling. For this reason, we considered several other models and ascertained their suitability for analyzing the impacts of Chinese tariffs on the world soybean market, particularly the reallocation of trade flows. The various trade models we considered are the non-spatial equilibrium, Armington, and gravity trade models.

The non-spatial equilibrium model only allows for total exports or imports by a country, and does not determine bilateral trade flows between a pair of countries. Consequently, this model is highly unsuitable for quantifying the trade diversion from one country to another due to policy changes. The Armington model removes the homogeneity assumption and differentiates the commodity based on the country of origin. However, this model fails to recognize the trade reallocations that occur, the emergence of new markets, and the loss of old markets. Thus, the Armington model endures a similar pitfall as the non-spatial equilibrium model. The gravity model has become the workhorse in trade literature to model the trade flows among countries, particularly after the pioneering study by Eaton and Kortum (2002) which used productivity shocks to capture the comparative advantage. For example, Reimer and Li (2010) implemented this model to ascertain comparative advantage and trade cost effects on crop trade. The gravity model incorporates bilateral trade flows among trading partners, and econometrically estimates the impacts of various policies and exogenous variables. The nature of econometric estimation, unlike the optimization in the SEM, does

not allow the researchers to quantify the *trade flow reallocations*. Therefore, there is a trade off that we need to consider when deciding on which model to use. For this particular study, we decided that the benefits of determining trade reallocations outweigh the potential for over/under estimation of trade flow changes. Additionally, as described in the Data and Calibration section, in order to mitigate any exaggeration of trade reallocations, the model parameters are calibrated such that the base simulation results match real-world values.

Many studies have employed the SEM for policy analysis, it is not possible to do justice reviewing all of these studies. Consequently, we briefly review some key studies. Within one year of the initial work by Samuelson (1952), Fox (1953) utilized the SEM to analyze the livestock feed market among various regions of the United States. More recently, Devadoss et al. (2005) used the SEM to analyze the effects of disputes between the United States and Canada on the world softwood lumber market. Apart from direct application of the SEM, many economists have administered certain modifications in order to best fit the specific problem being studied. For example, Kawaguchi et al. (1997) applied the SEM to the Japanese dairy industry, but modified it in order to allow for several degrees of market competition. Von Oppen and Scott (1976) developed a SEM by integrating location and inter-regional trade aspects to simultaneously determine optimal inter-regional trade as well as plant location and size.

Until the Uruguay Round, much of the trade policy analysis using the SEM incorporated specific tariffs. However, once the Uruguay Round converted all tariffs and quotas into equivalent ad valorem tariffs, it became imperative that ad valorem tariffs be incorporated into the SEM. Toward this goal, the ad valorem tariff model of Takayama and Judge (1971) was made operational by Devadoss (2013). These studies have shown that when an ad valorem tariff is included in the model, the quasi-welfare function of the traditional SEM is no longer strictly concave. Consequently, instead of maximizing the quasi-welfare function, they propose maximizing the net revenue function. Furthermore, Devadoss (2013) demonstrated that optimization of the primal or dual approach can be also readily undertaken using the

mixed complementarity problem (MCP). The MCP solves the system of equations which are the first order conditions of either the primal or dual approach, as both approaches yield identical first order conditions. In this study we employ the MCP approach for the empirical analysis.

Within months of the initial tariff threats, a multitude of reports appeared in Choices magazine spearheaded by Marchant and Wang (2018), the popular press, and Farmdoc articles, which emphasizes the importance of this issue and the severity of the impacts of Chinese tariffs. The soybean tariff has been of particular interest because of the sheer volume of soybean trade between the United States and China. Several studies (Durisin and Dodge, 2018; Good, 2018b,a,c; Plume, 2018) reported qualitative impacts of the Chinese tariff on U.S. and Chinese soybean markets. In addition to these reports, several empirical studies have been conducted estimating the impacts of the Chinese soybean tariff. For example, Zheng et al. (2018) utilized the Global Simulation Model and Taheripour and Tyner (2018) used the Global Trade Analysis Project model to study the Chinese tariff. In the results section, we compare the results of our study and these studies to provide validity to our modeling framework and the findings.

The contributions of this study to the literature lie in the analytical results of the theoretical model and the ability of the SEM to capture bilateral trade flows and reallocation of trade arising from policy changes. The theoretical analysis clearly shows the adverse effects of the Chinese tariff on the U.S. soybean market, the advantages accrued to Brazil, and the mitigation of U.S. losses by trade reallocation. Applying the empirical SEM to this theoretical framework accurately quantifies the effects of the Chinese tariff on price, production, consumption, and bilateral trade flows. We also compute the welfare effects of this tariff using changes in producer surplus, consumer surplus, and tariff revenue. Therefore, the findings of our study, in conjunction with current literature, are valuable to soybean growers, agribusiness firms operating in the domestic and export markets, and policy makers.

Theoretical Analysis

The supply, demand, and trade flows of a many region (i, j = 1, ... n) model can be succinctly summarized by the following market clearing and spatial arbitrage conditions.

$$S_i(P_i^P) = D_i(P_i^C) + \sum_{j \neq i}^n X_{ij}, \ \forall \ i$$
 (1)

$$S_j(P_j^P) + \sum_{i \neq j}^n X_{ij} = D_j(P_j^C), \ \forall \ j$$
 (2)

$$P_i^C = P_i^P - s_i + t_{ij} + \tau_{ij}, \ \forall \ i, j$$

$$\tag{3}$$

where $S_i(\bullet)$ is the supply function in country i, $D_i(\bullet)$ is the demand function in i, P_i^P represents producer price in i, P_i^C denotes consumer price in i, X_{ij} indicates volume of trade from country i to country j, s_i measures subsidies provided by i, t_{ij} stands for transport costs from i to j, and τ_{ij} are tariffs levied by j on imports from i. Equation (1) states that for an exporting country i supply equals domestic demand plus exports to all other regions. Equation (2) indicates that for an importing country j, supply plus imports from all other countries equal domestic demand. In addition to these market clearing conditions, the model also incorporates the spatial arbitrage of prices between any pair of countries. Equation (3) captures this spatial arbitrage wherein consumer price in the importing country j is equal to producer price in the exporting country i minus production subsidy provided by i, plus transport costs incurred in moving the commodity from i to j, and tariffs imposed by j on imports coming from i.

This n country model contains a total number of equations $n + n^2$, i.e., n market clearing conditions for exporting and importing countries plus n^2 price linkage equations. With this large system of equations, obtaining analytical results is not plausible. Consequently, we simplify the model into four regions: two exporting—the United States (U) and Brazil (B)—and two importing—China (C) and the European Union (E). This stylized model allows us to examine how the Chinese tariff causes China to divert its imports from the United States to Brazil, and the United States to export more to the European Union to

mitigate the export loss to China. However, our empirical model encompasses a total of eleven regions—five exporting regions and six importing regions—to more accurately model the world soybean market. These regions are clearly delineated in the empirical analysis section.

In the appendix we present in detail how the four regions' market-clearing conditions and the corresponding spatial-arbitrage conditions are simplified into a four region model consisting of the following market-clearing conditions with spatial arbitrage embedded in them:

$$S_U(P_U^P) = D_U(P_U^P - s_U + t_{UU}) + X_{UC} + X_{UE}$$
(4)

$$S_B(P_U^P - s_U + t_{UC} + \tau_{UC} - t_{BC}) = D_B(P_U^P - s_U + t_{UC} + \tau_{UC} - t_{BC} + t_{BB}) + X_{BC} + X_{BE}$$
(5)

$$S_C(P_U^P - s_U + t_{UC} + \tau_{UC} + s_C - t_{CC}) + X_{UC} + X_{BC} = D_C(P_U^P - s_U + t_{UC} + \tau_{UC})$$
 (6)

$$S_E(P_U^P - s_U + t_{UE} - t_{EE}) + X_{UE} + X_{BE} = D_E(P_U^P - s_U + t_{UE})$$
(7)

Equation (4) states that U.S. soybean supply, which is a function of U.S. producer price P_U^P , is equal to U.S. domestic demand plus exports to China X_{UC} and exports to the European Union X_{UE} . The domestic demand is a function of consumer price P_U^C equal to producer price minus U.S. subsidy s_U plus internal transport cost t_{UU} . Equation (5) entails that Brazilian soybean supply, a function of Brazilian producer price P_B^P , is equal to Brazil's domestic demand, a function of Brazilian consumer price P_B^C , plus exports to China X_{BC} and the European Union X_{BE} . Brazilian producer price P_B^P is equal to U.S. producer price minus U.S. subsidy plus transport cost from the United States to China t_{UC} plus Chinese tariff τ_{UC} minus transport cost from Brazil to China t_{BC} . This price linkage equation is derived by combining the following two price linkages $P_C^C = P_U^P - s_U + t_{UC} + \tau_{UC}$ and $P_C^C = P_B^P + t_{BC}$. The first price linkage equation indicates that market price in China P_C^C is U.S. producer price minus U.S. subsidy plus transport cost from the United States to China and tariff imposed by China. The second price linkage denotes that market price in China P_C^C is equal

to producer price in Brazil plus transport cost from Brazil to China. We consider soybeans as a homogeneous product, so the market price in China is the same whether the soybeans are imported from the United States or Brazil. Equating these two price linkages and solving for P_B^P results in $P_B^P = P_U^P - s_U + t_{UC} + \tau_{UC} - t_{BC}$. Brazilian consumer price $P_B^C = P_B^P + t_{BB}$, where t_{BB} is the internal transport cost, and substitution of $P_B^P = P_U^P - s_U + t_{UC} + \tau_{UC} - t_{BC}$ results in the argument in the Brazilian demand.

Equation (6) shows that Chinese supply plus imports from the United States and Brazil equal Chinese demand, a function of Chinese consumer price P_C^C . Chinese producer price is $P_C^P = P_C^C + s_C - t_{CC}$ where s_C is the Chinese production subsidy and t_{CC} is the internal transport cost in China, and substitution for $P_C^C = P_U^P - s_U + t_{UC} + \tau_{UC}$ in this equation results in the argument in the Chinese supply. Finally, equation (7) asserts that EU supply plus imports from the United States and Brazil is equal to EU demand, a function of EU consumer price $P_E^C = P_U^P - s_U + t_{UE}$ where t_{UE} is transport cost from the United States to the European Union. EU supply is a function of EU producer price $P_E^P = P_E^C - t_{EE}$, and substitution for P_E^C results in the argument in the EU supply.¹

Totally differentiating the system of four equations (4) - (7) and arranging them in matrix form Ax = b gives the following.

$$\begin{bmatrix}
\left(\frac{\partial S_{U}}{\partial P_{U}^{P}} - \frac{\partial D_{U}}{\partial P_{U}^{C}}\right) & -1 & -1 & 0 \\
\left(\frac{\partial S_{B}}{\partial P_{B}^{P}} - \frac{\partial D_{B}}{\partial P_{E}^{C}}\right) & 0 & 0 & -1 \\
\left(\frac{\partial S_{C}}{\partial P_{C}^{P}} - \frac{\partial D_{C}}{\partial P_{C}^{C}}\right) & 1 & 0 & 1 \\
\left(\frac{\partial S_{E}}{\partial P_{E}^{P}} - \frac{\partial D_{E}}{\partial P_{E}^{C}}\right) & 0 & 1 & 0
\end{bmatrix}
\begin{bmatrix}
dP_{U}^{P} \\
dX_{UC} \\
dX_{UE} \\
dX_{BC}
\end{bmatrix} = \begin{bmatrix}
-\frac{\partial D_{U}}{\partial P_{E}^{D}} ds_{U} \\
\left(\frac{\partial S_{B}}{\partial P_{B}^{P}} - \frac{\partial D_{B}}{\partial P_{B}^{C}}\right) ds_{U} - \left(\frac{\partial S_{B}}{\partial P_{B}^{P}} - \frac{\partial D_{B}}{\partial P_{B}^{C}}\right) d\tau_{UC} \\
\left(\frac{\partial S_{C}}{\partial P_{C}^{P}} - \frac{\partial D_{C}}{\partial P_{C}^{C}}\right) ds_{U} - \left(\frac{\partial S_{C}}{\partial P_{C}^{P}} - \frac{\partial D_{C}}{\partial P_{C}^{C}}\right) d\tau_{UC} - \frac{\partial S_{C}}{\partial P_{C}^{P}} ds_{C} \\
\left(\frac{\partial S_{E}}{\partial P_{E}^{P}} - \frac{\partial D_{E}}{\partial P_{E}^{C}}\right) ds_{U}
\end{bmatrix},$$

$$(8)$$

The system of four equations in (4) - (7) contains five endogenous variables $(P_U^P, X_{UC}, X_{UE}, X_{BC}, X_{BE})$ and consequently, to solve the system we eliminate X_{BE} . This change allows us to analytically show that China reallocates soybean imports from the United States to Brazil, and the United States recoups some of the lost exports to China by redirecting its sales to the European Union.

where the determinant of the coefficient matrix A is negative:

$$|A| = -\left[\left(\frac{\partial S_U}{\partial P_U^P} - \frac{\partial D_U}{\partial P_U^C}\right) + \left(\frac{\partial S_B}{\partial P_B^P} - \frac{\partial D_B}{\partial P_B^C}\right) + \left(\frac{\partial S_C}{\partial P_C^P} - \frac{\partial D_C}{\partial P_C^C}\right) + \left(\frac{\partial S_E}{\partial P_E^P} - \frac{\partial D_E}{\partial P_E^C}\right)\right] < 0 \quad (9)$$

This determinant, consisting of the slopes of excess supply/demand, which in turn depends on each country's supply and demand conditions, indicates that the following comparative static results rely heavily on the magnitude of supply and demand elasticities in all countries.² Applying Cramer's Rule to the system of equations (8), we can solve for changes in the four endogenous variables in response to policy variables s_U , s_C , and τ_{UC} . However, since the focus of the analysis is on the effect of the Chinese tariff τ_{UC} , we present the comparative static analysis of only τ_{UC} on key endogenous variables. The comparative static results are presented below.

$$\frac{\partial X_{UC}}{\partial \tau_{UC}} = \frac{\left[\left(\frac{\partial S_U}{\partial P_U^P} - \frac{\partial D_U}{\partial P_U^C} \right) + \left(\frac{\partial S_E}{\partial P_E^P} - \frac{\partial D_E}{\partial P_E^C} \right) \right] \left[\left(\frac{\partial S_B}{\partial P_B^P} - \frac{\partial D_B}{\partial P_B^C} \right) + \left(\frac{\partial S_C}{\partial P_C^P} - \frac{\partial D_C}{\partial P_C^C} \right) \right]}{|A|} < 0 \qquad (10)$$

In equation (10), the numerator is positive and the denominator |A| is negative, indicating that the Chinese tariff reduces soybean imports from the United States. The rationale for this result is that the Chinese tariff will make U.S. soybeans more expensive, and consequently, Chinese imports of U.S. soybeans fall. The magnitude of change is determined by the excess supply/demand elasticities of the United States and the European Union multiplied by the excess supply/demand elasticities of Brazil and China, weighted by the value of the determinant |A| which captures the excess supply/demand elasticities of all four regions. Thus, changes in Chinese imports depend on the market conditions not only in China and the United States, but also in Brazil and the European Union. These inter-linkages are explicitly captured *only* by the SEM.

$$\frac{\partial P_U^P}{\partial \tau_{UC}} = \frac{\left[\left(\frac{\partial S_B}{\partial P_B^P} - \frac{\partial D_B}{\partial P_B^C} \right) + \left(\frac{\partial S_C}{\partial P_C^P} - \frac{\partial D_C}{\partial P_C^C} \right) \right]}{|A|} < 0 \tag{11}$$

²Note that the slopes in determinant |A| can be converted into elasticities.

Equation (11) shows that the Chinese tariff causes a decrease in U.S. producer price. This is because the Chinese tariff reduces Chinese imports from the United States, as demonstrated in equation (10), resulting in more availability to sell in the U.S. domestic market and thus lowering the U.S. price. The magnitude of this decrease depends on the value of the excess supply/demand elasticities of Brazil and China, weighted by the value of the determinant |A|.

$$\frac{\partial X_{UE}}{\partial \tau_{UC}} = \frac{-\left(\frac{\partial S_E}{\partial P_E^P} - \frac{\partial D_E}{\partial P_E^C}\right) \left[\left(\frac{\partial S_B}{\partial P_B^P} - \frac{\partial D_B}{\partial P_B^C}\right) + \left(\frac{\partial S_C}{\partial P_C^P} - \frac{\partial D_C}{\partial P_C^C}\right)\right]}{|A|} > 0$$
 (12)

Although the Chinese tariff is only intended to limit imports from the United States, a spillover effect of the Chinese tariff is revealed in equation (12) in that U.S. exports to the European Union increase. This is just one of the trade reallocations highlighted in the introduction, and is a result of the decreased U.S. producer price and additional U.S. soybeans available from lost exports to China. Here, the magnitude depends on the excess demand elasticity of the European Union multiplied by the excess supply/demand elasticities of China and Brazil, and weighted by |A|.

$$\frac{\partial X_{BC}}{\partial \tau_{UC}} = \frac{-\left(\frac{\partial S_B}{\partial P_B^P} - \frac{\partial D_B}{\partial P_B^C}\right) \left[\left(\frac{\partial S_U}{\partial P_U^P} - \frac{\partial D_U}{\partial P_U^C}\right) + \left(\frac{\partial S_E}{\partial P_E^P} - \frac{\partial D_E}{\partial P_E^C}\right)\right]}{|A|} > 0$$
(13)

Finally, the change in Brazil's exports to China in response to the Chinese tariff is positive. This is because the decrease in U.S. exports to China caused by the tariff leads to more opportunities for Brazil to export to China. Thus, Brazilian soybean producers benefit from the trade war between the United States and China. The magnitude of this change is dependent on the excess supply elasticity of Brazil multiplied by the excess supply/demand elasticities of the United States and the European Union, weighted by |A|.

Using the above results, the comparative statics for several other variables such as supply and demand in each country can be obtained. For example, the effect of the Chinese tariff on U.S. demand is

$$\frac{\partial D_{U}}{\partial \tau_{UC}} = \frac{\partial D_{U}}{\partial P_{U}^{C}} \frac{\partial P_{U}^{C}}{\partial P_{U}^{P}} \frac{\partial P_{U}^{P}}{\partial \tau_{UC}} = \frac{\partial D_{U}}{\partial P_{U}^{C}} \left[\left(\frac{\partial S_{B}}{\partial P_{B}^{P}} - \frac{\partial D_{B}}{\partial P_{B}^{C}} \right) + \left(\frac{\partial S_{C}}{\partial P_{C}^{P}} - \frac{\partial D_{C}}{\partial P_{C}^{C}} \right) \right]}{|A|} > 0$$
(14)

Equation (14) reveals that demand in the United States increases as a result of the Chinese tariff. The rationale for this result is that the decrease in U.S. producer price, exhibited in equation (11), causes the consumer price to decrease and the domestic demand to rise. The effect of the Chinese tariff on U.S. supply is

$$\frac{\partial S_U}{\partial \tau_{UC}} = \frac{\partial S_U}{\partial P_U^P} \frac{\partial P_U^P}{\partial \tau_{UC}} = \frac{\partial S_U}{\partial P_U^P} \frac{\left[\left(\frac{\partial S_B}{\partial P_B^P} - \frac{\partial D_B}{\partial P_B^C} \right) + \left(\frac{\partial S_C}{\partial P_C^P} - \frac{\partial D_C}{\partial P_C^C} \right) \right]}{|A|} < 0 \tag{15}$$

The Chinese tariff reduces U.S. supply as it lowers U.S. producer price. This indicates that the decrease in U.S. exports to China shown in equation (10) dominates both the increase in exports to the European Union (Equation (12)) and the increase in domestic demand (Equation (14)). The effect of the Chinese tariff on Chinese soybean demand is³

$$\frac{\partial D_C}{\partial \tau_{UC}} = \frac{\partial D_C}{\partial P_C^C} \left(\frac{\left[\left(\frac{\partial S_U}{\partial P_U^P} - \frac{\partial D_U}{\partial P_U^C} \right) + \left(\frac{\partial S_E}{\partial P_E^P} - \frac{\partial D_E}{\partial P_E^C} \right) \right]}{\left[\left(\frac{\partial S_U}{\partial P_U^P} - \frac{\partial D_U}{\partial P_U^C} \right) + \left(\frac{\partial S_B}{\partial P_B^P} - \frac{\partial D_B}{\partial P_E^C} \right) + \left(\frac{\partial S_C}{\partial P_C^P} - \frac{\partial D_C}{\partial P_C^C} \right) + \left(\frac{\partial S_E}{\partial P_E^P} - \frac{\partial D_E}{\partial P_E^C} \right) \right]} \right) < 0 \tag{16}$$

The tariff increases Chinese consumer price, causing Chinese demand to fall. The effect of the Chinese tariff on supply of the U.S. competitor, i.e., Brazil is

$$\frac{\partial S_B}{\partial \tau_{UC}} = \frac{\partial S_B}{\partial P_B^P} \left(\frac{\left[\left(\frac{\partial S_U}{\partial P_U^P} - \frac{\partial D_U}{\partial P_U^C} \right) + \left(\frac{\partial S_E}{\partial P_E^P} - \frac{\partial D_E}{\partial P_E^C} \right) \right]}{\left[\left(\frac{\partial S_U}{\partial P_U^P} - \frac{\partial D_U}{\partial P_U^C} \right) + \left(\frac{\partial S_B}{\partial P_B^P} - \frac{\partial D_B}{\partial P_C^P} \right) + \left(\frac{\partial S_C}{\partial P_C^P} - \frac{\partial D_C}{\partial P_C^P} \right) + \left(\frac{\partial S_E}{\partial P_E^P} - \frac{\partial D_E}{\partial P_E^C} \right) \right]} \right) > 0 \tag{17}$$

Equation (17) shows that Brazilian supply expands because China diverts its imports from the United States to Brazil. Consequently, Brazilian exports rise, leading to a higher soybean price and supply in Brazil. The comparative statics for Chinese supply, Brazilian demand,

³The derivations of equations (16) and (17) are shown in the appendix.

and EU supply and demand can also be obtained, but are not presented here in the interest of space consideration.

Empirical Analysis

The theoretical model developed in the previous section is implemented using the SEM through either the primal, dual, or MCP approach. The primal approach maximizes the quasi-welfare function, subject to market clearing conditions, by optimizing over quantities. The dual approach minimizes cost, subject to (a) price linkages, (b) supply price-cost of production, and (c) demand price-willingness to pay, by optimizing over prices. Devadoss (2013) shows that the first order conditions of both the primal and dual problems lead to identical Kuhn-Tucker conditions and economic interpretations. The MCP solves the system of complementary-slackness equations associated with these Kuhn-Tucker conditions, and does not require an objective function. Although the MCP is scantly used in the literature, it is relatively simpler to implement empirically. For this reason, we employ the MCP and use GAMS software to solve the model. The MCP equations used in the empirical model along with their economic interpretations are given below.

$$MP_i^C \ge \alpha_i - \beta_i Q_i^D, \ \forall i$$
 (18)

where MP_i^C is the market demand price in i, α_i is the inverse demand intercept of i, β_i is the inverse demand slope of i, and Q_i^D is the quantity demanded in i. Equation (18) requires that the market demand price is on or above the demand curve. That is, market demand price is greater than (when demand quantity is zero) or equal to (when demand quantity is positive) the willingness to pay.

$$\gamma_i + \theta_i Q_i^S \ge M P_i^P, \ \forall i \tag{19}$$

where γ_i is the inverse supply intercept of i, θ_i is the inverse supply slope of i, Q_i^S is the quantity supplied in i, and MP_i^P is the market supply price in i. Equation (19) shows that

the market supply price must be on or below the supply curve. Therefore, the market supply price must be less than (when supply quantity is zero) or equal to (when supply quantity is positive) the marginal cost.

$$t_{ij} \ge MP_j^C * \frac{1}{1 + \tau_{ij}} - MP_i^P, \ \forall i, j$$
 (20)

where t_{ij} is the transport cost from country i to country j, and τ_{ij} is the ad valorem tariff levied by country j on imports from country i. Equation (20) is a price linkage equation that confines market demand price (including ad valorem tariff) in j minus the market supply price in i to be less than or equal to transport costs. This restricts exporters from charging a price that is less than the sum of their own producer price in i and the cost of transport from i to j. If the price charged in the importing country is more than the producer price in the exporting country plus transport cost, then profit opportunities exist which will entice other exporters to sell to this importing country until the profit opportunities are exhausted.

$$\sum_{i=1}^{n} X_{ij} \ge Q_i^D, \ \forall j \tag{21}$$

Equation (21) shows that demand must be met by domestic supply and foreign imports, so that there is no excess demand; otherwise, price will increase until the supply satisfies the demand.

$$Q_i^S \ge \sum_{j=1}^n X_{ij}, \ \forall i$$
 (22)

Equation (22) demonstrates that the quantity supplied should be at least as much as the quantity sold domestically and in foreign markets. As explained above, equations (18) - (22) hold with equality for interior solutions. It is worth noting that equations (18) - (22) are equivalent to the theoretical model presented in equations (4) - (7), as (18) and (19) are captured in the supply and demand functions in (4) - (7), the price linkage equation (20) is embedded in (4) - (7), equation (21) is the market clearing for importers given in (6) and (7),

and equation (22) is the market clearing for exporters shown in (4) and (5). More generally, equations (18) - (22) are directly comparable to equations (1) - (3) with linear supply and demand functions expressed in general functional form.

This system of equations (18) - (22) is solved simultaneously using the parameters $(\alpha_i, \beta_i, \gamma_i, \text{ and } \theta_i)$ and exogenous variables $(t_{ij} \text{ and } \tau_{ij})$ to obtain the values of the endogenous variables $(MP_i^C, MP_i^P, Q_i^D, Q_i^S, \text{ and } X_{ij})$. This system is solved once with $\tau_{UC} = 0$ to find baseline values, and then again with $\tau_{UC} = 0.25$ to find the tariff scenario values. The empirical model consists of eleven regions: the United States, Brazil, Argentina, Paraguay, Canada, China, Mexico, the European Union, Japan, Taiwan, and ROW. Therefore, there are a total of 165 equations consisting of 44 demand price, supply price, demand quantity, and supply quantity conditions, plus 121 price linkage equations.

After solving for baseline and tariff scenarios for the values of endogenous variables, we compute changes in producer surplus, consumer surplus, tariff revenues, and net surplus. Change in producer surplus, ΔPS , for country i is calculated by integrating the supply function between producer prices in the baseline and tariff scenarios.⁴

$$\Delta PS = \int_{P_i^P}^{P_i^{P,\tau}} \Gamma_i + \Theta_i P dP, \tag{23}$$

where Γ_i and Θ_i are the supply intercept and slope in country i, and P_i^P and $P_i^{P,\tau}$ are country i's producer prices in the baseline and tariff scenarios, respectively. The change in producer surplus will be positive if $P_i^{P,\tau}$ is above P_i^P and negative if $P_i^{P,\tau}$ is below P_i^P . Similarly, the change in consumer surplus, ΔCS , is computed using

$$\Delta CS = \int_{P_i^{C,\tau}}^{P_i^C} A_i - B_i P dP, \tag{24}$$

⁴Note that the supply and demand functions in Q-P space in equations (23) and (24) correspond to the supply and demand functions in P-Q space in equations (18) and (19).

where A_i and B_i are the demand intercept and slope in country i, and P_i^C and $P_i^{C,\tau}$ are country i's consumer prices in the baseline and tariff scenarios, respectively. Change in consumer surplus is positive when P_i^C is above $P_i^{C,\tau}$ and negative when P_i^C is below $P_i^{C,\tau}$. Ad valorem tariff revenue is computed as quantity of imports times CIF price times tariff rate:

$$TR = \left(\sum_{i \neq j}^{n} X_{ij}\right) * \left(\left(MP_i^P + t_{ij}\right) * \tau_{ij}\right)$$
(25)

Tariff revenue is strictly non-negative. Net surplus is the sum of changes in producer surplus, consumer surplus, and tariff revenue, $NS = \Delta PS + \Delta CS + TR$, and can be positive or negative.

Data and Calibration

The data needed for empirically implementing the model are production and consumption quantities, domestic prices, transport costs, supply and demand elasticities, and realized trade flows. Production and consumption data for each country came from FAS (2018d). We collected these quantity data for the years 2015-2018, and used the average to smooth out unduly upward and downward swings in the data. Country level price data is not available from a single source and, consequently, was obtained from several sources. These include Table 29 of FAS (2018b) for the United States, Brazil, and Argentina, GAIN (2018) for Paraguay, and GFO (2018) for Canada. Domestic prices for China, Mexico, the European Union, Japan, and Taiwan were estimated using the average transport cost from countries they are importing from and world price. The average world price for 2015-2018 is from Macrotrends (2018) and is also used as the domestic price for ROW. Japan, the United States, and China provide production subsidies which were collected from Hudson (2018).

The transport costs were obtained using the data from WorldFreightRates (2018) for an average soybean price of \$401/metric ton with a 10,000 metric ton load. For countries with multiple ports, calculations were made based on the shortest port to port distance. For landlocked Paraguay, transport costs include costs from Asuncion, Paraguay to Buenos Aires, Argentina plus additional costs to the import destinations. Furthermore, we ensure

transport costs are such that there are no trans-shipments through a third country.

Supply and demand elasticities for most of the countries came from FAPRI (2018). U.S. and ROW elasticities were collected from Devadoss et al. (1989), and Mexican elasticities were obtained from Reimer et al. (2012). Additionally, Paraguay's elasticities are constructed using Argentinian elasticities because of their close proximity and similar cultivation practices.

To construct the supply equations, we utilize the supply elasticities, prices, and quantities. The slope Θ of a linear supply function can be obtained from the elasticity formula $\varepsilon = \frac{dQ_S}{dP} \frac{P}{Q_S} \text{ as}^5$

$$\frac{dQ_S}{dP} = \varepsilon \frac{Q_S}{P} = \Theta \tag{26}$$

Using Θ , the intercept, Γ , is computed using the supply function $Q_S = \Gamma + \Theta P$ as

$$\Gamma = Q_S - \Theta P. \tag{27}$$

Thus, the constructed supply function is $Q_S = \hat{\Gamma} + \hat{\Theta}P$. Following a similar approach, we construct the demand function as $Q_D = \hat{A} - \hat{B}P$. We then convert these to the inverse supply and demand functions $P^P = \hat{\gamma} + \hat{\theta}Q^S$ and $P^C = \hat{\alpha} - \hat{\beta}Q^D$, which are used in the empirical analysis. Running a baseline scenario with these supply and demand functions for equations (18) and (19), and (20) - (22), generates supply and demand quantities which are close to, but do not replicate, the actual quantities. This is because the elasticities obtained from the literature are based on econometric estimations involving a disturbance term, data inaccuracies, specification problems, etc. To overcome this problem, Paris et al. (2009) developed an approach to calibrate the parameters such that solved trade flows exactly match the realized trade flows. Therefore, for this calibration we need data on realized trade flows which are computed using the export quantity data from FAS (2018d) and 2016 bilateral export destination percentages from Simoes and Hidalgo (2018). The inverse sup-

⁵For China, Japan, and the United States, subsidies were taken into account in constructing their supply functions.

ply and demand functions constructed using these calibrated parameters are given in Table 1.

Simulation and Results

The inverse supply and demand equations and transport costs are used to run the baseline and tariff scenarios. The baseline simulation assumes free trade in the world soybean market and replicates the actual supply and demand quantities. The alternative tariff scenario results are compared with those of the baseline scenario to examine the impacts of this tariff on prices, production, consumption, and trade flows. We also undertake a welfare analysis of this tariff by computing change in producer surplus, consumer surplus, and overall welfare. Table 2 presents the effects of the Chinese tariff on prices, production, consumption, and welfare. Table 3 reports the baseline trade flows from exporters (rows) to importers (columns), and changes in trade flows caused by the Chinese tariff. The percent change is calculated using $\frac{(X_{ij}^r - X_{ij})}{X_{ij}}$ where X_{ij} is the baseline trade flow from i to j, and X_{ij}^{τ} is the tariff scenario trade flow from i to j. If the baseline trade flow X_{ij} is zero, i.e., country i does not export to country j, then percentage changes cannot be computed. In this case we report only the changes in trade flows in thousand metric tons.

We discuss the results of Tables 2 and 3 in tandem, because of the symbiotic relationship among price, supply, demand, and trade flows. The simulation results in both tables confirm that the quantitative impacts of the Chinese tariff are in line with the directional impacts derived in the theoretical analysis for prices, supply, demand, and trade flows. The producer and consumer prices are the same for all countries, except for China, Japan, and the United States, because demand and supply functions are measured at the same market level. The reason for the producer and consumer price differences in China, Japan, and the United States are the production subsidies provided by these countries which cause a wedge between these two prices.⁶

⁶Note the baseline trade flows are not exactly representative of realized trade flows and are instead the optimal trade flows solved by the model. This is because realized trade flows account for non-economic factors such as political motives and prearranged contractual agreements.

The tariff increases China's consumer price by 6.77%, and in response to this change in consumer price (through the domestic price linkage equation), the producer price increases by 4.70%. The higher consumer price decreases consumption by 1.36%, while the higher producer price expands production by 1.57%. China's price and quantity changes decrease consumer surplus by \$3.03 billion but increase producer surplus by \$370 million, resulting in a net welfare loss of \$2.66 billion. Because of the 25% tariff, China stops importing from the United States and diverts all of its imports of U.S. soybeans to other regions: Brazil, Argentina, Paraguay, Canada, and ROW (See Table 3). To put the surplus changes into perspective, China's \$3.03 billion decrease in consumer surplus equates to 6.73% of Chinese soybean consumption value, in spite of the increased imports from Brazil, Argentina, Paraguay, Canada, and ROW. China's producer surplus increase is 4.68% of total value of production. However, it is important to note that China's total value of consumption is far greater than the total value of production, and these percentage changes, though similar, are misleading. That is, 6.73% of China's consumption value is much larger than 4.68% of its production value.

Since the United States is exporting less to China, its producer and consumer prices decline by 11.92% and 12.56%, respectively, because more is available for domestic sales. The drop in U.S. prices cause a 3.96% decrease in production and a 3.07% increase in consumption. Table 3 shows that the United States mitigates some of its export losses to China by increasing exports to Mexico, the European Union, Japan, Taiwan, and ROW. Regardless of these trade reallocations, however, welfare in the United States suffers more than any other country as a result of the Chinese tariff, with producer surplus falling by \$5.52 billion, consumer surplus rising by \$2.8 billion, and net surplus declining by \$2.72 billion. The \$5.52 billion decrease in U.S. producer surplus is 11.86% of the total value of U.S. soybean production, whereas the \$2.80 billion increase in U.S. consumer surplus is

⁷Note that the ad valorem tariff, τ_{UC} , is zero in the baseline simulation, and U.S. exports to China, X_{UC} , are zero in the tariff scenario simulation, causing tariff revenues to be zero in both cases. For this reason, tariff revenue does not affect net surplus and is not included in Table 2.

12.77% of the total value of U.S. soybean use. Taheripour and Tyner (2018) used the GTAP model to evaluate the impacts of Chinese tariffs on many commodities. Their results for the soybean market show that U.S. and Chinese net surplus declines by -\$2.55 and -\$2.55 billion dollars, respectively. These results corroborate our findings of U.S. and Chinese net surplus changes of -\$2.72 and -\$2.66 billion dollars, respectively.

Brazil is the second largest soybean producer in the world,⁹ and, as a result of the tariff, China reallocates much of its imports from the United States to Brazil. Specifically, Brazil's exports to China rise by 62.94\%, which causes price in Brazil to go up by 7.70\%. As a consequence of this price increase, Brazil experiences a 2.74% increase in production and a \$3.23 billion increase in producer surplus which is 7.81% of the value of production. To meet the Chinese demand, Brazil diverts 100% of its exports from the European Union, Japan, and ROW to China. The price increase causes Brazilian demand to decrease by 1.34%, resulting in \$1.27 billion loss in consumer surplus which is 7.66% of the value of consumption. Overall, Brazil amasses \$1.96 billion in welfare, and therefore benefits the most by virtue of this trade war. The aggregate ROW region has the second largest gain from the Chinese tariff with a \$520 million increase in net surplus, but these benefits are spread out among many countries. This increase in welfare is caused by the 5.46% decrease in consumer and producer prices, which increase consumption by 1.97% and consumer surplus by \$940 million. ROW production falls by 1.16%, and producer surplus accrues a \$420 million loss. Table 3 shows that ROW decreases domestic sales by 23.53% and exports 4,049,000 MT to China due to the tariff.¹⁰

Because of the Chinese tariff, considerable reallocation of world trade occurs in the soybean market as the United States diverts its exports to the remaining importing countries: Mexico, the European Union, Japan, and Taiwan. The producer and consumer prices

⁸Zheng et al. (2018) found changes in U.S. net welfare of -\$1.8 billion.

 $^{^9}$ In 2016 and 2017, Brazil produced 114,100,000 MT and 117,000,000 MT, respectively, just behind U.S. production of 116,920,000 MT and 119,518,000 MT in the same years (FAS, 2018b).

¹⁰ROW is comprised of some exporting and importing countries, so as the world price changes, this region could switch between exporting and importing.

in Mexico, the European Union, and Taiwan decrease by 11.07%, 5.81%, and 11.93%, respectively. In Japan, producer price decreases by 2.78% and consumer price falls by 11.89% (Recall, producer and consumer prices differ in Japan due to the production subsidy). As a consequence of these price changes, Mexican, EU, Japanese, and Taiwanese production decrease by 2.60%, 1.45%, 0.80%, and 1.78% while consumption increases by 2.50%, 1.29%, 4.74%, and 3.55%, respectively. Each country's gain in consumer surplus exceeds its loss in producer surplus, and the net surplus gain in Mexico, the European Union, Japan, and Taiwan are \$220, \$350, \$160, and \$130 million, respectively.

Argentina is the third largest producer of soybeans, but is far behind the United States and Brazil, and thus plays only a modest role in the world soybean market. Paraguay is an exporter, but a relatively small player in the soybean market. Both countries reallocate 100% of exports from ROW to China, causing Argentinian and Paraguayan prices to increase by 7.37% and 7.24%, respectively. As a result, Argentina sustains a 2.42% increase in production and 1.93% decrease in consumption, while Paraguay experiences a 2.40% increase in production and a 1.94% decrease in consumption. Because of the changes in prices, supply, and demand, Argentina gains \$1.42 billion in producer surplus and loses \$1.26 billion in consumer surplus. Paraguay gains only \$250 million in producer surplus and loses \$100 million in consumer surplus. Although producer and consumer surplus changes are much higher in Argentina than they are in Paraguay, the overall gain in welfare is nearly identical for the two exporters with Argentina and Paraguay collecting \$160 and \$150 million, respectively.

Canada, unlike other exporters, loses as a result of the Chinese tariff. The rationale for this result is that Canada traditionally exports to the European Union; but with the Chinese tariff on U.S. soybeans, the United States diverts exports to the European Union at a price that Canadian exporters cannot compete with. This causes Canada to lose its EU export market, thus increasing its availability to sell domestically and decreasing Canadian prices by 1.98%. These reduced prices cause Canadian production to decrease by 0.67% and its consumption to increase by 0.55%. Although this result is different from those of other

exporters, it follows intuitively with the trade reallocation scenario given in the introduction. Estimation of this type of reallocation are possible only with the spatial equilibrium model. Ultimately, Canada endures a net surplus loss of \$40 million, making it the only country besides the United States and China to suffer a net surplus loss as a consequence of the tariff.

We also aggregated the production, consumption, and surplus measures for the entire world, which are shown in the last row of Table 2. Total world production and consumption remains almost unchanged, increasing by only 0.006%. This is because the increased production of China and several of the exporting countries nearly proportionately matches the decrease in production of the other countries. Similarly, increased consumption in the United States, Canada, and several of the importing countries is roughly analogous to the decreased consumption of the remaining countries. Table 2 shows that both producer and consumer surplus decrease after the implementation of the Chinese tariff. Consequently, the world loses \$1.75 billion in total welfare. This result is similar to the findings of Taheripour and Tyner (2018) who showed that the total world welfare loss as a result of the Chinese tariff is \$1.49 billion.

We also undertook analyses of 10%, 15%, and 20% tariff scenarios, which leads to world welfare losses of \$0.94, \$1.04, and \$1.42 billion, respectively. These results clearly indicate the world's welfare inefficiencies escalate significantly as the tariffs are progressively increased and underscores the importance of moving toward freer trade rather than pursuing protectionist policies. This finding is congruent with economic trade theory in that larger tariffs create greater inefficiencies. The trade reallocations also vary considerably across different tariff scenarios. For example, under the 25% tariff rate scenario, the United States fully diverts its exports from China to other countries, and furthermore new trade between ROW and China emerges. However, under the 20% tariff rate, the United States diverts only 89.78% of its exports to China to other importers, and there is no emergence of trade between ROW and China. Additionally, under the 15% tariff rate, there is no emergence of

trade between Canada and China, and Canada loses only 2.10% (rather than 100%) of its exports to the European Union. Finally, under the 10% tariff rate, U.S. exports to China decline by only 73.80%, but the emergence of trade between Paraguay and China (which were experienced in the 25%, 20%, and 15% scenarios) no longer occurs.

Though the current study covers only the raw soybean market, further extensions of this study would benefit from including additional sectors, specifically the oil and meal sectors, which would help to capture the interlinkages among the primary commodity soybeans and final products soymeal and soyoil. Including these sectors would allow the model to quantify changes such as China reducing imports of U.S. soybeans and augmenting the imports of soymeal and soyoil from other countries. Furthermore, it would allow the model to quantify price, production, and consumption in these additional sectors, and these results will be useful to the producers and consumers of these allied sectors.

Conclusion

The trade war between the world's economic superpowers, the United States and China, will have drastic effects on the world economy. Though these effects span across hundreds of traded goods and impart spillover effects on many countries, we focus on soybeans because the United States and China have a particularly strong presence in the world soybean market. Consequently, it is worth studying the effects of the Chinese tariff on the world soybean market. Analyzing this tariff can be performed using several trade models such as non-spatial and gravity models; however, the spatial equilibrium model is most ideally suited because it can quantify the tariff impacts on prices, production, consumption, and bilateral trade flows simultaneously. In doing so, this model captures the trade reallocations occurring in the world market.

We develop a theoretical model using general functional forms for supply and demand functions and demonstrate qualitative impacts of the tariff. The empirical analysis implements the theoretical model by applying the SEM to the world soybean market using the MCP approach. The baseline and tariff scenarios solve for the impacts of the Chinese tariff on the endogenous variables: prices, supply, demand, and trade flows. These quantitative solutions confirm the qualitative results of the theoretical analysis. Using the solved values of the endogenous variables, we compute welfare measures.

The results exhibit the widespread repercussions of the Chinese tariff on the world soybean market, including countries that are not directly involved in this trade war. The Chinese tariff on U.S. soybeans inflicts net losses on three countries: China, the United States, and Canada. China and the United States endure billions of dollars in losses, thus manifesting the self-destructive economic consequences of protectionist policies. China loses because the higher prices resulting from the tariff harms consumers more than it helps producers. In contrast, the United States losses are due to lower prices which hurt producers more than it benefits the consumers. Canada, on the other hand, incurs losses because the United States displaces some of the Canadian exports to the European Union.

Brazil, which is the leading competitor to the United States in the world soybean market, is the largest beneficiary of this trade war as it captures much of the United States' lost market in China. Brazil has been increasing production and exports at an alarming rate in recent years, and this trade litigation could propel Brazil to surpass the United States as the largest soybean producer in the world. Though Brazil and a few other countries may benefit from the Chinese tariff, the results show that the world as a whole incurs welfare loss, thus creating economic inefficiency.

References

- ChineseMinistry (2018). Announcement of the customs tariff commission of the state council on adding tariffs to imports of us \$50 billion. Technical report, Ministry of Finance of the People's Republic of China, State Council Tariff Commission. http://gss.mof.gov.cn/zhengwuxinxi/zhengcefabu/201806/t20180616_2930325.html, accessed on July 12, 2018.
- Devadoss, S. (2013). Ad valorem tariff and spatial equilibrium models. *Applied Economics* 45(23), 3378–3386.
- Devadoss, S., A. H. Aguiar, S. R. Shook, and J. Araji (2005). A spatial equilibrium analysis of US-Canadian disputes on the world softwood lumber market. *Canadian Journal of Agricultural Economics/Revue canadienne d'agroeconomie 53* (2-3), 177–192.
- Devadoss, S., P. C. Westhoff, M. D. Helmar, E. Grundmeier, K. D. Skold, W. H. Meyers, and S. R. Johnson (1989). The FAPRI modeling system at CARD: a documentation summary.
- Durisin, M. and S. Dodge (2018). Why soybeans are at the heart of the U.S.-China trade war. *Bloomberg*. https://www.bloomberg.com/graphics/2018-soybean-tariff/, accessed on July 10, 2018.
- Eaton, J. and S. Kortum (2002). Technology, geography, and trade. *Econometrica* 70(5), 1741–1779.
- FAPRI (2018, June). Elasticity database. http://www.fapri.iastate.edu/tools/elasticity.aspx, accessed on June 13, 2018.
- FAS (2018a). Global agricultural trade system. Technical report, United States

 Department of Agriculture, Foreign Agricultural Sevice.

 https://apps.fas.usda.gov/gats/ExpressQuery1.aspx, accessed on June 17, 2018.

- FAS (2018b). Oilseeds: World markets and trade. Technical report, United States

 Department of Agriculture, Foreign Agricultural Service.

 http://usda.mannlib.cornell.edu/usda/fas/oilseed-trade//2010s/2018/oilseed-trade-0510-2018.pdf, accessed on June 14,
 2018.
- FAS (2018c). Production, supply, and distribution. Technical report, United States

 Department of Agriculture, Foreign Agricultural Service.

 https://apps.fas.usda.gov/psdonline/app/index.html/app/home, accessed on June 17, 2018.
- FAS (2018d). PS&D data sets-oilseeds. Technical report, United States Department of Agriculture, Foreign Agricultural Service.

 https://apps.fas.usda.gov/psdonline/app/index.html/app/downloads, accessed on July 15, 2018.
- Fox, K. A. (1953). A spatial equilibrium model of the livestock-feed economy in the United States. *Econometrica: Journal of the Econometric Society*, 547–566.
- GAIN (2018). Paraguay: Oilseeds and products annual. Global agricultural information network, United States Department of Agriculture, Foreign Agricultural Service. https://www.fas.usda.gov/data/paraguay-oilseeds-and-products-annual-2, accessed on May 24, 2018.
- GFO (2018, May). Daily commodity report.

 http://gfo.ca/marketing/daily-commodity-report/, accessed on May 30, 2018.
- Good, K. (2018a). Corn, soybean prices flounder; Sec. Perdue reiterates promise to stand with farmers. Farm Policy News. https://farmpolicynews.illinois.edu/2018/06/cornsoybean-prices-flounder-sec-perdue-reiterates-promise-to-stand-with-farmers/, accessed on June 30, 2018.

- Good, K. (2018b). In retaliation for U.S. tariffs, China targets soybeans, corn. Farm Policy News. https://farmpolicynews.illinois.edu/2018/06/in-retaliation-for-u-s-tariffs-china-targets-soybeans-corn/, accessed on June 30, 2018.
- Good, K. (2018c). Looming U.S.- China trade battle: Soybean trade flows and substitutes. Farm Policy News. https://farmpolicynews.illinois.edu/2018/07/looming-u-s-china-trade-battle-soybean-trade-flows-and-substitutes/, accessed on July 12, 2018.
- Hudson, D. (2018). Update to foreign crop subsidy database. Technical report, Texas Tech University. https://www.depts.ttu.edu/ceri/assets/pdf/database.pdf, accessed on July 20, 2018.
- Kawaguchi, T., N. Suzuki, and H. M. Kaiser (1997). A spatial equilibrium model for imperfectly competitive milk markets. American Journal of Agricultural Economics 79(3), 851–859.
- Macrotrends (2018). Soybean prices 45 year historical chart.

 https://www.macrotrends.net/2531/soybean-prices-historical-chart-data, accessed on June 5, 2018.
- Marchant, M. A. and H. H. Wang (2018). Theme overview: U.S.-China trade dispute and potential impacts on agriculture. *Choices* 33(2).
- Paris, Q., S. Drogué, G. Anania, et al. (2009). Calibrating mathematical programming spatial models. Ag-Food Trade Working Paper 10.
- Plume, K. (2018). As U.S. and China trade tariff barbs, others scoop up U.S. soybeans. Reuters. https://www.reuters.com/article/us-usa-trade-china-soybeans/as-u-s-and-china-trade-tariff-barbs-others-scoop-up-u-s-soybeans-idUSKBN1HF0FQ, accessed on May 30, 2018.

- Reimer, J. J. and M. Li (2010). Trade costs and the gains from trade in crop agriculture.

 American Journal of Agricultural Economics 92(4), 1024–1039.
- Reimer, J. J., X. Zheng, and M. J. Gehlhar (2012). Export demand elasticity estimation for major US crops. *Journal of Agricultural and Applied Economics* 44(4), 501–515.
- Samuelson, P. A. (1952). Spatial price equilibrium and linear programming. *The American economic review* 42(3), 283–303.
- Simoes, A. J. G. and C. A. Hidalgo (2018, June). The economic complexity observatory:

 An analytical tool for understanding the dynamics of economic development.

 https://atlas.media.mit.edu/en/, accessed on June 14, 2018.
- Taheripour, F. and W. E. Tyner (2018). Impacts of possible Chinese 25% tariff on U.S. soybeans and other agricultural commodities. *Choices* 33(2).
- Takayama, T. and G. G. Judge (1971). Spatial and Temporal Price Allocation Models.

 Amsterdam: North Holland Publishing Company.
- Von Oppen, M. and J. T. Scott (1976). A spatial equilibrium model for plant location and interregional trade. *American Journal of Agricultural Economics* 58(3), 437–445.
- WorldFreightRates (2018, June). Freight calculator. http://worldfreightrates.com/freight, accessed on June 16, 2018.
- Zheng, Y., D. Wood, H. H. Wang, and J. P. Jones (2018). Predicting potential impacts of China's retaliatory tariffs on the U.S. farm sector. *Choices* 33(2).

Table 1: Inverse Supply and Demand Functions

	Supply Function	Demand Function
USA	$P^P = -857.112 + 0.011Q_S$	$P^C = 2007.455 - 0.029Q_D$
Brazil	$P^P = -658.059 + 0.009Q_S$	$P^C = 2457.750 - 0.046Q_D$
Argentina	$P^P = -741.625 + 0.021Q_S$	$P^C = 1745 - 0.029Q_D$
Paraguay	$P^P = -711.875 + 0.109Q_S$	$P^C = 1675 - 0.341Q_D$
Canada	$P^P = -773.5 + 0.162Q_S$	$P^C = 1820 - 0.545Q_D$
China	$P^P = -1373.444 + 0.135Q_S$	$P^C = 2478 - 0.019Q_D$
Mexico	$P^P = -1455.128 + 4.268Q_S$	$P^C = 2424.876 - 0.415Q_D$
EU	$P^P = -1302 + 0.685Q_S$	$P^C = 2384.562 - 0.119Q_D$
Japan	$P^P = -5739.980 + 24.799Q_S$	$P^C = 1452.5 - 0.3Q_D$
Taiwan	$P^P = -2368.689 + 484.511Q_S$	$P^C = 1811.333 - 0.555Q_D$
ROW	$P^P = -1568.842 + 0.110Q_S$	$P^C = 1594.667 - 0.029Q_D$

		Table 2: I	Table 2: Baseline Values,		Tariff Scenario Percent Changes, and Surplus Changes	rcent Cha	nges, and Sur	plus Cha	nges		
	Producer Price	er Price	Consumer Price	er Price	Production	tion	Consumption	otion		Surplus	
	Baseline (\$/MT)	Change (%)	Baseline (\$/MT)	Change (%)	Baseline (1000 MT)	Change (%)	Baseline (1000 MT)	Change (%)	Producer (\$ billion)	Consumer (\$ billion)	Net (\$ billion)
USA	415.46	-11.92	394.16	-12.56	113,752	-3.96	55,631	3.07	-5.52	2.80	-2.72
Brazil	364.24	7.70	364.24	7.70	113,589	2.74	45,511	-1.34	3.23	-1.27	1.96
Argentina	362.28	7.37	362.28	7.37	52,567	2.42	47,680	-1.93	1.42	-1.26	0.16
Paraguay	353.73	7.24	353.73	7.24	9,776	2.40	3,875	-1.94	0.25	-0.10	0.15
Canada	394.75	-1.98	394.75	-1.98	7,211	29.0-	2,615	0.55	-0.06	0.03	-0.04
China	596.53	4.70	414.53	6.77	13,244	1.57	108,604	-1.36	0.37	-3.03	-2.66
Mexico	447.28	-11.07	447.28	-11.07	446	-2.60	4,765	2.50	-0.02	0.24	0.22
EU	434.15	-5.81	434.15	-5.81	2,535	-1.45	16,390	1.29	-0.06	0.42	0.35
Japan	1773.43	-2.78	414.05	-11.89	248	-0.80	3,461	4.74	-0.01	0.17	0.16
Taiwan	414.91	-11.93	414.91	-11.93	9	-1.78	2,516	3.55	-0.0002	0.13	0.13
ROW	422.27	-5.46	422.27	-5.46	18,101	-1.16	40,427	1.97	-0.42	0.94	0.52
World	NA	NA	NA	NA	331,475	0.006	331,475	0.006	-0.81	-0.94	-1.75

	VSA	Brazil	Argentina	Paraguay	Canada	China	Mexico	EU	Japan	Taiwan	ROW
USA	55,631					51,292	4,320	0	0	2,510	0
(Change) (3	(3.069%)					(-100%)	(3.030%)	(+14,104)	(+3,379)	(3.557%)	(+27,380)
Brazil		45,511				44,068		9,259	3,213		11,538
(Change)		(-1.340%)				(62.939%)		(-100%)	(-100%)		(-100%)
Argentina			47,680			0					4,887
(Change)			(-1.930%)			(+7,078)					(-100%)
Paraguay				3,875		0					5,905
(Change)				(-1.939%)		(+6,212)					(-100%)
Canada					2,615	0		4,596			
(Change)					(0.549%)	(+4,534)		(-100%)			
China						13,244					
(Change)						(1.568%)					
Mexico							446				
(Change)							(-2.602%)				
EU								2,535			
(Change)								(-1.452%)			
Japan									248		
(Change)									(-0.800%)		
Taiwan										9	
(Change)										(-1.775%)	
ROW						0					18,101
(Change)						(+4.049)					(-23.527%)