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Economic Crisis And the Fragility of World Wine Export

Imre Fertő

Institute of Economics, Centre for Economic and Regional Studies, Hungarian Academy of Sciences, and Kaposvár University, Hungary

Abstract

The paper investigates the impacts of economic crisis on the duration of world wine exports for 40 leading wine exporter countries between 2000 and 2012. We analyze three attributes of wine exports embodied in exports relationships: the initial value of exports, the growth of exports within a spell, and the hazard of exports ceasing. Our results indicate that wine exports are rather short lived even for leading wine exporter countries. Our estimations imply that economic crisis has not significant impacts on initial value of wine export starting before crisis, whilst it has negative effects on spells beginning after crisis. However, economic crisis does not influence the export growth, and decrease the probability of failure of wine exports.

Keywords

Duration of world wine exports, economic crises, leading wine exporters.

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Introduction

During the last two decades, the world's wine market has experienced a rapid process of globalization. Growth in the wine trade boomed during the nineties both in wine consumption in Northern Europe and in North America and in exports of the so-called New World wine producers. Recently, the wine international trade has been boosted by increasing demand in countries (Anderson et al., 2004; Anderson and Nelgen, 2011; Mariani et al., 2012). There is a growing literature on various aspects of recent developments on world wine markets including the impacts tariff and tariff frictions (Mariani et al. 2014; Dal Bianco et al., 2016), dynamics of world wine trade (Castillo et al., 2016), characteristics of intra-EU markets (Fertő et al., 2016; Lombardi et al., 2016).

However, one question is not yet addressed in empirical wine trade literature: when do countries trade and how long do their trade relationships last? Our analysis of this latter issue is, among other things, motivated by the finding of recent research that many countries do not trade in any given year and for any given product (Haveman and Hummels, 2004, Feenstra and Rose, 2000, Schott, 2004). As a consequence of it, a new literature focusing on the duration of international trade has emerged. Based on the surprising finding in Besedeš and Prusa (2006a) that US import flows have a remarkably

short duration, the question asked is: "which factors determine how long international trade relationships last?" From a policy-oriented point of view this is indeed an important question to ask. Trade will not grow very much if new products stop being exported after only a few years. Therefore, to better understand which factors may help countries increase their trade, and thereby potentially improve economic development, it is important to learn more about what determines the duration of trade flows. Recent studies provide evidence that trade relationships (e.g. Besedeš and Prusa, 2006b; Nitsch, 2009; Fertő and Soós, 2009; Brenton et al., 2010; Obashi, 2010; Cadot et al., 2013) are surprisingly short lived. Empirical studies usually confirm that exporter characteristics (such as GDP and language), product characteristics (such as unit values) and market characteristics (such as the import value, and market share) affect the duration of trade (Hess and Persson 2011, 2012). However all studies focus only manufacturing or all products except (Bojnec and Fertő, 2012).

In addition, being in or out of the market may be a particularly important issue when the market is affected by external shocks. The economic crisis in 2008 is good natural experiment. Our research focuses on the question how major wine exporter countries in the world market before and after the economic crisis. Specifically, we examine how economic crisis affect the value of wine export

at the start of a new wine export relationship, the length of wine export relationships, and how quickly wine export grows within a relationship.

Materials and methods

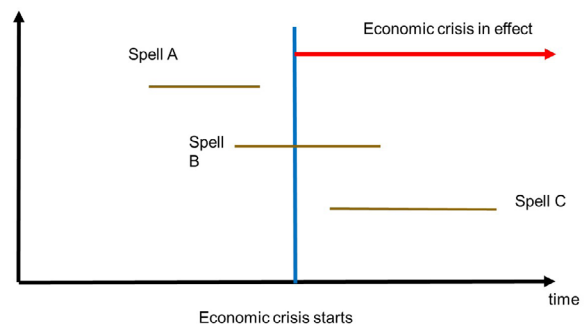
There are two empirical strands in the international trade literature on the duration of trade. The first analyses the duration of bilateral trade relations at the product (category) level and the second analyses the trade behaviour of firms, in particular switching of export products and destinations. This paper builds upon on the first strand of the literature focusing on country-product relations.

Besedeš and Prusa (2006a) distinguish homogeneous and differentiated goods using the Rauch (1999) classification. They find that homogeneous goods have higher hazard rates than differentiated goods and higher initial trade values increase survival. In addition their results indicate lower transportation costs, higher GDP, higher tariffs, and depreciation of the source country's currency all lead to longer durations. Nitsch (2009) applies also Cox proportional hazard models on the duration of German import relations between 1995 and 2005. He also concludes that GDP in the exporting country and a similar language lowers the hazard rate. This is also the case for the initial trade value and market share in the importing country. Brenton et al. (2009) analyse the duration of export flows at the 5-digit SITC level of about 80 exporting countries and 50 importing countries between 1985 and 2005. They also conclude that the initial trade value is important for survival. Hess and Persson (2011) focus on the imports of 15 EU-countries from 140 different exporting countries between 1962 and 2006 at the 4-digit SITC level. They conclude that the mean duration of import flows is only 1 year. Moreover they show that export diversification, which - both in terms of the number of products exported and the number of markets served with the given product - substantially lowers the hazard of trade flows dying. Notice, that these studies suffers from the lack of theoretical background. Existing theories based on heterogenous firms does not explain the short lived export relationships (Hess and Persson 2011). More recently Besedeš et al. (2016) provide a theory to explain some empirical regularities of short lived trade relationships.

Following their model we test the following hypotheses. First, the probability of an export relationship ceasing is decreasing in its size and age (or duration). Second, the growth rate of exports

is decreasing in size conditional on duration and decreasing in duration conditional on size. Third, larger initial volumes of exports are associated with longer lasting relationships and lower hazard rates. Beyond to theoretical considerations we focus on the impact of the economic crisis on the duration of wine exports. Uncertainties in demand side and cost conditions have important role in export decisions which cannot be resolved after entry of foreign markets. Nguyen (2012) argues that uncertainties may reconcile the high exit rates after first years. We can assume that economic crisis has increased the uncertainties especially on demand side, thus we can expect that crisis has negative impact on the duration of wine exports.

Figure 1 provides a schematic illustration of the types of trade spells a pair of countries can have as they relate to economic crisis, they enter into. The advent of crisis allows us to distinguish between three types of spells. There will be spells such as spell A, which begin and end before the crisis goes into effect. These spells are unaffected by the crisis. There are also spells such as spell B which start before the crisis, but do not end until after the crisis goes into effect. These spells will be directly affected by the crisis. Finally, there are also spells, such as spell C, which start after the crisis has been established.



Source: Own compilation

Figure 1: Effects of economic crisis on trade.

To properly identify the effects of economic crisis, we need to differentiate between spells active when the crisis begins and spells which begin after the crisis. In order to properly capture all effects of economic crisis we use two variables. One variable, labeled 'crisis in effect' (Spell B), identifies the years during which crisis is in force, thus identifying the differential effect of the crisis itself. Since model predicts that relationships or spells which start after the crisis are different from already active ones, we use a second dummy variable, 'Spell starts after crisis' (Spell C), which identifies all spells which started after the crisis is put in force. The 'crisis in effect' and 'Spell

starts after crisis variables in conjunction identify the effect on spells which begin after the crisis is in effect.

We are interested in the effect economic crisis have on three attributes of exports spells: the volume of exports in the first year, the growth of the volume of exports while the spell is active, and the conditional probability it will cease to be active or the hazard rate. We examine the effect on initial volumes and the growth of exports within an active spells by estimating two separate OLS regressions:

$$X(I)_{ikt} = \alpha_0 + \alpha_1 Crisis_t + \alpha_2 afterCrisis_{ikt} + \gamma_i + \mu_k + \eta_t + \varepsilon_{ikt} \quad (1)$$

$$Xgrowth_{ikt} = \alpha_0 + \alpha_1 Duration_{ikt} + \alpha_2 X_{ikt-1} + \alpha_3 Crisis_t + \alpha_4 afterCrisis_{ikt} + \gamma_i + \mu_k + \eta_t + \varepsilon_{ikt} \quad (2)$$

Where $X(I)_{ikt}$ denotes export values in initial period, $Xgrowth_{ikt}$ describes the export growth, X_{ikt-1} is lagged value of exports, $Crisis$ is a dummy takes 1 before 2009, otherwise zero, $afterCrisis$ is a variable reflecting how long the crisis has been in effect when a spell starts, $Duration$ the age of spell k in year t , γ_i is exporter fixed effects, μ_k is importer fixed effects and η_t , year fixed effects.

Next step we focus on the duration of world wine exports. Duration analysis of export (export > 0) is estimated by the survival function, $S(t)$, using the nonparametric Kaplan-Meier product limit estimator (Cleves et al., 2004). We assume that a sample contains n independent observations denoted $(t_i; c_i)$, where $i = 1, 2, \dots, n$, t_i is the survival time, and c_i is the censoring indicator variable C taking a value of 1 if failure occurred, and 0 otherwise of observation i . It is assumed that there are $m < n$ recorded times of failure. The rank-ordered survival times are denoted as $t_{(1)} < t_{(2)} < \dots < t_{(m)}$, while n_j denotes the number of subjects at risk of failing at $t_{(j)}$, and d_j denotes the number of observed failures. The Kaplan-Meier estimator of the survival function is then:

$$\hat{S}(t) = \prod_{t(i) \leq t} \frac{n_j - d_j}{n_j} \quad (3)$$

with the convention that $\hat{S}(t) = 1$ if $t < t_{(1)}$. Given that many observations are censored, it is then noted that the Kaplan-Meier estimator is robust to censoring and uses information from both censored and non-censored observations.

Beyond to descriptive analysis of duration of export, we are interested in the factors explaining the survival. Recent literature on the determinants of trade and comparative advantage duration uses

Cox proportional hazards models (e.g. Besedeš and Prusa, 2006; Bojnec and Fertő, 2012; Cadot et al., 2013). However, recent papers point out three relevant problems inherent in the Cox model that reduce the efficiency of estimators (Hess and Persson, 2011, 2012). First, continuous-time models (such as the Cox model) may result in biased coefficients when the database refers to discrete-time intervals (years in our case) and especially in samples with a high number of ties (numerous short spell lengths). Second, Cox models do not control for unobserved heterogeneity (or frailty). Thus, results might not only be biased, but also spurious. The third issue is based on the proportional hazards assumption that implies similar effects at different moments of the duration spell. Following Hess and Persson (2011), we estimate discrete-time models, namely probit specifications, where exporter country random effects are incorporated to control for unobservable heterogeneity.

More specifically, we estimate the hazard of exports ceasing at time t by estimating a discrete hazard using random effects probit specification controlling to origin's and destination's GDP, a vector of bilateral time-invariant gravity variables (distance, common border, and common language)

$$XD_{ikt} = \alpha_0 + \alpha_1 Duration_{ikt} + \alpha_2 X_{ikt-1} + \alpha_3 Crisis_t + \alpha_4 afterCrisis_{ikt} + \alpha_5 GDP_{it} + \alpha_6 GDP_{kt} + \alpha_7 Indistance_{ik} + \alpha_8 border_{ik} + \alpha_9 language_{ik} + \varepsilon_{ikt} \quad (3)$$

Our empirical analysis is based on a panel data set includes bilateral trade data of 40 major wine exporter countries and 216 trading partners between 2000-2012, giving 54587 observations. Wine export data comes from the World Bank World Integrated Trade Solution (WITS) database in HS-6 level, product code 2204, in thousand US dollars (World Bank, 2016a). Data for the other explanatory variables are obtained from the following data sources: GDP from the World Bank (2016b) database, trade costs variables including distance, common border and common language are from the CEPII database (Mayer and Zignano, 2011).

Results and discission

In our aim to explore the duration of wine trade in leading wine exporters countries, we start by performing a thorough descriptive analysis. Table 1 offers some initial summary statistics as to the length of wine exports flows. Notice first that there is a large number of observed spells

(over 6700). Table 1 shows that the mean duration of a spell in our benchmark data is less than 5 years for the leading wine exporters countries within a relatively short time span (13 years). The mean length of spells starting and ending before crisis (A sample) is rather short less than two years. The average duration of spells surviving crisis (B sample) is relatively long 7.3 years, whilst the length of spells entering into market after crisis is already short (1.8 years). The short lived duration is in line with findings of previous studies (Besedeš and Prusa, 2006b; Besedeš et al., 2016; Bojnec and Fertő, 2012; Nitsch, 2009; Fertő and Soós, 2009; Brenton et al., 2010; Obashi, 2010).

sample	Obs	Mean	Std. Dev.	Min	Max
total sample	6706	4.9	4.8	1	13
A sample	2966	1.9	1.6	1	9
B sample	3740	7.3	5.2	1	13
C sample	1625	1.8	1.0	1	4

Source: Own calculations based on World Bank WITS database (2017)

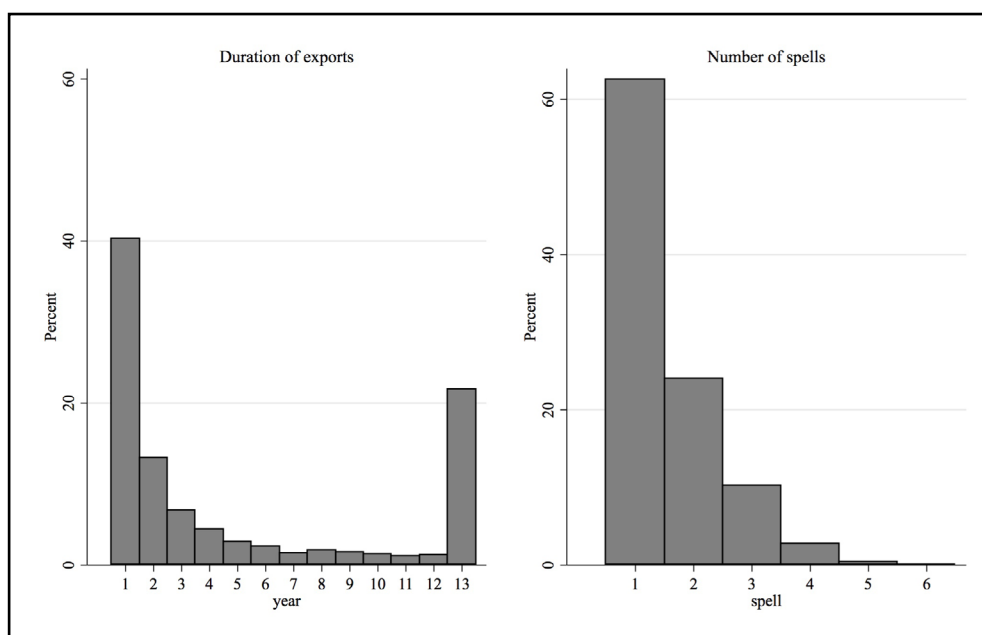
Table 1: Summary statistics of spells.

The left histogram in Figure 1 presents the distribution of the duration density of the number countries with exports > 0 over the thirteen years analyzed, which is slightly more concentrated on the left side, indicating fewer years continuously being at exports>0, than on the right side of the same histogram, indicating more years being continuously at exports > 0.

Around 22 percent of the wine products have a perfect continued survival rate in exports > 0 during the thirteen analyzed years.

The right histogram in Figure 2 presents the number of spells with exports > 0, focusing on the difference between single spells and multiple spells per a given wine product. First, the high share of a single spell with continuous exports > 0 indicates that most of the main wine exporter countries have a high percentage of wine products that survived a certain number of years in 2000-2012. During the analyzed 13-year period, the minimum length of a spell is one year, and the maximum length of a spell for a given wine exporter countries with continuous exports > 0 is 13 years. The average value of the length of spell is 4.9 years, while the median value of duration of the spell is only 2 years. Second, among the multiple spells with exports>0 per given wine product, two and three spells, and to a lesser extent four and five spells for a given wine product are identified.

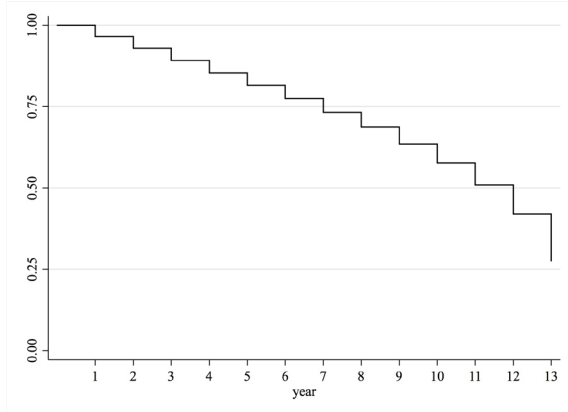
To be able to describe the trade flows with more information than a mere mean or standard deviation value will allow, we also plot a descriptive survivor function. Figure 3 depicts empirical survivor functions of wine exports spells. The x-axis plots the observed spell length, and the y-axis plots the fraction of observations whose observed spell of service exceeds a given length. The Kaplan-



Source: Own calculations based on World Bank WITS database (2017)

Figure 2: Histograms of the number of spells and duration of the exports.

Meier survival function indicate that in the first half of the period less than 25 per cent of spells have ceased, but this ratio has doubled in the second half of period. In other words more than 50% of all spells have ceased after economic crisis.



Source: Own calculations based on World Bank WITS database (2017)

Figure 3: Kaplan-Meier survival estimates.

Now we turn to determinants of duration of wine trade. First, we investigate the impact of economic crisis on the initial volume of wine export. Since we are examining a single value at the starting point of a spell, our ability to identify different effects of economic crisis is reduced. A spell either starts before or after the crisis. As a result, the impact of crisis taking effect only applies to spells starting after the crisis. We thus have two variables identifying the impacts of crisis: a dummy variable identifying the years when the agreement is in effect (crisis in effect) and a variable reflecting

how long the crisis has been in effect when a spell starts.

The first and second columns of Table 2 collect the results from estimating equation (1). Using only the dummy variable identifying when the crisis is in effect we have not find significant impact. We then add the variable measuring how long the crisis was in effect when the spell started. The augmented model show that in a fixed (with respect to time) effect of the crisis decreasing initial volumes by 3,300 thousand dollars, as well as a time-dependent effect which decreases initial volumes by 894 thousand dollars for every year of the crisis being in force. In other words. The impacts of economic crisis was higher by 3.5 for pre-crisis initial values than for post-crisis initial values.

We now turn to examining the impact of economic crisis on the growth of wine exports embodied in active spells. In particular, we investigate the growth of wine exports conditional on spell survival. Our results imply that the rate of growth of wine exports within a spell decreases the longer the duration of the spell, just as Besedeš et al. (2016) model predicts (third and fourth columns). Larger spells grow faster. The impacts of economic crisis are insignificant for both crisis and post-crisis dummies.

We estimate the hazard of wine exports ceasing by estimating equation (3) using random effects probit, which allows us to take into account unobserved heterogeneity. To estimate the hazard we include the standard gravity variables, GDP

	Initial value	Initial value	growth	growth	duration	duration
Crisis _t	458.509	-3.3e+03***		-8.9e+04		-0.121***
afterCrisis _{ikt}		-893.583***		4.0e+04		-0.086***
Duration _{ikt}			-1.8e+05***	-1.9e+05***	-0.110***	-0.073***
X _{ikt-1}			2.8e+05***	2.7e+05***	-0.617***	-0.609***
lnGDP _{it}					-0.067***	-0.064***
lnGDP _{kt}					-0.080***	-0.076***
Indistance _{ik}					0.130***	0.125***
border _{ik}					-0.167**	-0.166**
language _{ik}					-0.079**	-0.072*
constant	-1.7e+06***	-485.560	-2.7e+05***	2.7e+05***	3.153***	3.083***
N	54587	50388	50388	50388	46865	46865
R ²	0.1436	0.0972	0.2350	0.1658		
Chi ²					4183.966	4426.496
Rho					0.154	0.149

Note: * p<0.1; ** p<0.05; *** p<0.01
Source: Authors' own calculations

Table 2: Estimation results.

of both the importer and the exporter, distance between the two, as well as a dummy indicating the existence of a common border and a common language that the two countries share. Our results indicate that the hazard are decreasing in duration, indicating that longer lived spells are less likely to cease and also grow less (fifth and sixth columns) confirming findings by Besedeš et al. (2016). The size of GDP for both exporter and importer sides decrease the probability of wine export ceasing. In addition, exports are less sensitive to the GDP of the exporting country than to the income of the importing countries. Similarly to earlier studies (Brenton et al., 2009; Hess and Persson, 2011, 2012; Besedeš et al., 2016) estimations suggest that the common border and common language decreases, whilst the distance increases the likelihood of failure in the wine export relationships in both specifications.

Conclusion

The paper analyses the impact of economic crisis on the duration of word wine exports for 40 leading wine exporter countries between 2000 and 2012. More specifically, we investigate three attributes of wine exports embodied in exports relationships: the initial value of exports, the growth of exports within a spell, and the hazard of exports ceasing. Trade relationships between the 40 leading wine exporters and their partners and are often dynamic, with numerous entry and exits and spells of service. During the 2000 to 2012 period, there were 6706 different exporter trade relationships for leading wine exporters. While two-third of these trade

relationships had a single spell of service, one-third of all trade relationships have multiple spells of service. Approximately 40 per cent of all spells of service last for just a single year, and approximately 55 per cent of all spells of service last for three years or less. In line with the literature on the trade duration, we find that wine exports are rather short lived even for leading wine exporter countries. Compared with other duration studies, we find relatively longer spells of service for wine exports.

Our estimations imply that economic crisis has not significant impacts on initial value of wine export starting before crisis, whilst it has negative effects on spells beginning after crisis. Surprisingly, economic crisis does not influence the export growth, and decrease the probability of failure of wine exports. This puzzling results can be explained by quick recovery of world wine exports after economic crisis. The effects are usually smaller for spells starting after crisis. This suggests that we find two types of spells according to starting before and after crisis. Estimations indicate that longer lived spells are less likely to cease and also grow less. We find the standard “gravity” determinants of trade including market size and trade costs do affect on the export duration at the world wine market.

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Corresponding author:

Professor Imre Fertő

Institute of Economics, Centre for Economic and Regional Studies, Hungarian Academy of Sciences, H-1097 Budapest, Tóth K. u. 4. and Kaposvár University, H-7400 Kaposvár Guba S. 40, Hungary

E-mail: fertő.imre@krtk.mta.hu

ORCID: orcid.org/0000-0002-3742-5964

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