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# Transition matrix for a bivariate normal distribution in Stata 

Marco Savegnago<br>Bank of Italy<br>Rome, Italy<br>marco.savegnago@bancaditalia.it


#### Abstract

. trabinor calculates the population transition matrix between two discretized variables when the original continuous variables follow a bivariate normal distribution. The user can specify the five parameters of bivariate normal and how to discretize the two variables by choosing either a given number of quantiles or a set of absolute boundaries.


Keywords: st0392, trabinor, transition matrix, bivariate normal, binormal()

## 1 Introduction

A transition matrix is a useful tool to describe a stochastic process that involves a finite number of states. Examples include transitions among employment statuses (unemployed, employed part-time, and full-time) for a population of workers or among rating classes for companies and governments. Often the variable of interest is continuous, such as income; for instance, one might be interested in the association between an individual's income and the income of his or her father. After a proper discretization in a finite number of classes, a transition matrix can help answer questions like the following: given that the father earns less than $x$, what is the probability that a child earns more than $y$ (that is, undergoes upward mobility) or less than $x$ itself (that is, falls into a "poverty trap")?
trabinor calculates the transition matrix between two variables (after being appropriately discretized) if they follow a bivariate normal distribution. It can be used mostly for the following two purposes:

- It can be used when sample size is too small to have a reliable estimate of the transition matrix, especially when there are many cells to estimate. If one is willing to assume that the two variables are jointly normal, then sample moments can be computed from data and passed to trabinor.
- It can be used in Monte Carlo simulations to study the properties of estimators related to a transition matrix. For example, in intergenerational mobility analysis, the trace of a matrix (that is, the sum of the elements of the main diagonal in a square matrix) is an important summary measure of persistence of socioeconomic status ${ }^{1}$ (for a recent survey, see Black and Devereux [2011]). The behavior of this type of estimator when variables are contaminated by measurement errors is studied in O'Neill, Sweetman, and Van de gaer (2007). trabinor can be used to compute the "true" (that is, population) value of such estimators under different parameter values.


## 2 Statistical background

Assume that

$$
\left[\begin{array}{l}
Y \\
X
\end{array}\right] \sim \mathcal{N}\left[\binom{\mu_{Y}}{\mu_{X}},\left(\begin{array}{cc}
\sigma_{Y}^{2} & \rho \sigma_{Y} \sigma_{X} \\
\rho \sigma_{Y} \sigma_{X} & \sigma_{X}^{2}
\end{array}\right)\right]
$$

where the parameters $\left(\mu_{Y}, \sigma_{Y}, \mu_{X}, \sigma_{X}, \rho\right)$ are known or can be estimated from data.
Let $F_{Y, X}(y, x)$ also denote the joint normal cumulative distribution function (CDF), where $F_{Y}(\cdot)$ and $F_{X}(\cdot)$ are the marginal CDF. Assume that $Y$ and $X$ are discretized according to the rules $\left(-\infty<Y \leq y_{1}, y_{1}<Y \leq y_{2}, \ldots, y_{K-1}<Y \leq y_{K}, y_{K}<Y<\right.$ $\infty)$ and $\left(-\infty<X \leq x_{1}, x_{1}<X \leq x_{2}, \ldots, x_{J-1}<X \leq x_{J}, x_{J}<X<\infty\right)$.

Let $\mathbf{M}$ be the transition matrix between the discretized versions of $Y$ and $X$,

$$
\mathbf{M}_{J+1, K+1}=\left(\begin{array}{cccc}
p_{1,1} & p_{1,2} & \cdots & p_{1, K+1} \\
p_{2,1} & p_{2,2} & \cdots & p_{2, K+1} \\
\vdots & \vdots & \ddots & \vdots \\
p_{J+1,1} & p_{J+1,2} & \cdots & p_{J+1, K+1}
\end{array}\right)
$$

where the generic element $p_{j k}$ gives the probability that $Y$ falls in the $k$ th class given that $X$ falls in the $j$ th class.

$$
\begin{aligned}
p_{j k} & =\operatorname{Pr}\left(y_{k-1}<Y<y_{k} \mid x_{j-1}<X<x_{j}\right) \\
& =\frac{\operatorname{Pr}\left(y_{k-1}<Y<y_{k}, x_{j-1}<X<x_{j}\right)}{\operatorname{Pr}\left(x_{j-1}<X<x_{j}\right)} \\
& =\frac{F_{Y, X}\left(y_{k}, x_{j}\right)+F_{Y, X}\left(y_{k-1}, x_{j-1}\right)-F_{Y, X}\left(y_{k}, x_{j-1}\right)-F_{Y, X}\left(y_{k-1}, x_{j}\right)}{F_{X}\left(x_{j}\right)-F_{X}\left(x_{j-1}\right)}
\end{aligned}
$$

The last equality exploits the Stata functions binormal ( $h, k, r$ ) (for the numerator) and normal ( $z$ ) (for the denominator).

1. The higher the trace, the higher the elements on the first diagonal and the stronger the association between a father's and a son's socioeconomic statuses.

## 3 The trabinor command

### 3.1 Description

Besides the five parameters of the bivariate normal distribution $\left(\mu_{Y}, \sigma_{Y}, \mu_{X}, \sigma_{X}, \rho\right),{ }^{2}$ $Y$ and $X$ can be discretized by choosing either a given number of quantiles or a set of absolute boundaries. If the user sets a number $Q$ of quantiles, then the resulting transition matrix will obviously be of size $Q \times Q$. If the user sets $K$ absolute boundaries for $Y$ and $J$ absolute boundaries for $X$, then the resulting transition matrix will be rectangular of size $(J+1) \times(K+1) .{ }^{3}$ In this case, boundaries for $X$ must be chosen such that all the states of $X$ have a nonzero probability of occurrence; otherwise, we will be conditioning on an impossible event. ${ }^{4}$ A warning message appears if any of the marginal probabilities of (the discretized) $X$ are smaller than $5,000,000^{-1}$; a second message appears if this probability is so small that the corresponding row of the transition matrix cannot be computed.

If neither the number of quantiles nor the absolute boundaries are specified, the trabinor command computes a $5 \times 5$ quantile matrix by default.

Results are displayed in percentage points, so the row sums of the transition matrix sum to 100 .

### 3.2 Syntax

The syntax of trabinor is the following:

```
trabinor [, quant(#) y(numlist) x(numlist) muy(#) sy(#) mux(#) sx(#)
    rho(#) format(string)]
```


### 3.3 Options

quant(\#) specifies in how many quantiles $Y$ and $X$ should be divided. The default is quant (5). quant (\#) must be an integer greater than 1 and cannot be combined with the options $y()$ and $x()$.
y (numlist) sets the boundaries of the marginal distribution of $Y$. It cannot be combined with quant().
2. The default values are those of a standard bivariate normal with $\rho=0.5$.
3. To see this, imagine that we set the boundary 0 for both $Y$ and $X$; then, both variables will be divided in two levels, namely, $(-\infty, 0)$ and $(0, \infty)$.
4. From a theoretical point of view, the normal distribution assigns a nonzero probability on any interval defined on the real line, but these probabilities can be very small. For example, if $\Phi(\cdot)$ denotes the standard normal CDF, then $\Phi(-5) \approx 0.00000028665<3,000,000^{-1}$. Therefore, a wrong choice of the boundaries might induce some issue of numerical approximation, given that the conditional probabilities are computed as the ratios between the joint and marginal probabilities.
x (numlist) sets the boundaries of the marginal distribution of $X$. It cannot be combined with quant ().

If none of the options quant(), $y()$, or $x()$ are specified, trabinor computes a $5 \times 5$ quantile transition matrix (as if quant (5) were invoked). If only y() is specified, the same boundaries are applied to $X$; analogously, if only x() is specified, the same boundaries are applied to $Y$.
muy (\#) defines the mean of $Y$. The default is muy ( 0 ).
sy (\#) defines the standard deviation of $Y$. The default is sy(1).
$\operatorname{mux}(\#)$ defines the mean of $X$. The default is mux ( 0 ) .
$\mathrm{sx}(\#)$ defines the standard deviation of $X$. The default is $\mathrm{sx}(1)$.
rho(\#) defines the correlation coefficient between $Y$ and $X$. The default is rho(.5).
This must be between -1 and 1 .
format (string) controls how to display results. The default is format(\%9.3f).

### 3.4 Stored results

trabinor stores the following in $r()$ :

Scalars | $r($ muy $)$ | mean of $Y$ | $r(s y)$ |
| :--- | :--- | :--- |
| $r(\operatorname{mux})$ | mean of $X$ | $r(s x)$ |
| $r(r h o)$ | correlation coefficient between |  |
|  | $Y$ and $X$ | standard deviation of $Y$ |
|  | $\begin{array}{l}\text { standard deviation of } X\end{array}$ |  |
|  |  |  |

| Matrices <br> $r(M)$ | resulting transition matrix | $r(c h k 100)$ | consistency check on $r(M)$ (that <br> $r(P Y)$ |
| :---: | :---: | :--- | :---: |
|  | matrix of marginal row sums are equal to 100) |  |  |
| probabilities of $Y$ | $r(J)$ | matrix of joint probabilities |  |

## 4 Examples

Here I discuss three examples of trabinor. Let $\boldsymbol{\Theta}_{1}, \boldsymbol{\Theta}_{2}$, and $\boldsymbol{\Theta}_{3}$ represent the five parameters of the bivariate normal distribution in each of the three examples.

In the first example, trabinor computes the quantile transition matrix of size 4 for a bivariate standard normal distribution with $\rho=0.5$. Moreover, it uses the output program to obtain the trace of the transition matrix. Interpreting the results for the quantile transition matrix is straightforward: the probability that $Y$ is smaller than its first quartile, given that $X$ is smaller than its first quartile, is $48.1 \%$. Analogously, using a formal notation, $\operatorname{Pr}\left\{F_{Y}^{-1}(0.50)<Y \leq F_{Y}^{-1}(0.75) \mid X \leq F_{X}^{-1}(0.25) ; \boldsymbol{\Theta}_{1}\right\}=16.8 \%$, and so on.


In the second example, let's specify different parameters for the bivariate distribution. The variable $X$ is discretized in 4 classes $[(-\infty,-3],(-3,0],(0,3],(3, \infty)]$ and $Y$ in 6 classes $[(-\infty, 2],(2,3],(3,5],(5,7.5],(7.5,10],(10, \infty)]$. From the output, we observe that $\operatorname{Pr}\left(Y<2 \mid X<-3 ; \boldsymbol{\Theta}_{2}\right)=0.4 \%$, or that $\operatorname{Pr}\left(7.5 \leq Y<10 \mid-3 \leq X<0 ; \boldsymbol{\Theta}_{2}\right)=23.0 \%$, and so on. We then look at the marginal probabilities for both variables, presented as row vectors: $\operatorname{Pr}\left(7.5 \leq Y<10 ; \boldsymbol{\Theta}_{2}\right)=15.5 \%$ and $\operatorname{Pr}\left(-3 \leq X<0 ; \boldsymbol{\Theta}_{2}\right)=43.3 \%$.


```
. matrix list px
px[1,4]
\begin{tabular}{rrrrr} 
& \(r 1\) & \(r 2\) & \(r 3\) & \(r 4\) \\
\(c 1\) & 6.6807201 & 43.31928 & 43.31928 & 6.6807201
\end{tabular}
```

In the last example, we pass the empirical sample moments to trabinor using a dataset containing measures of blood pressure in 2 time periods (bp_before and bp_after) for a sample of 120 patients characterized by high blood pressure. Then, we look at the conditional probability of being in hypertension (bp_after $\geq 140$ ) given the level of bp_before. We note that this conditional probability is $80 \%$ given hypertension in the past and $69 \%$ given no hypertension in the past.

```
. sysuse bpwide, clear
(fictional blood-pressure data)
. quietly summarize bp_before
. scalar mu_0 = r(mean)
. scalar sd_0 = r(sd)
. quietly summarize bp_after
. scalar mu_1 = r(mean)
. scalar sd_1 = r(sd)
. quietly correlate bp_before bp_after
. scalar corr = r(rho)
. trabinor, muy(`=mu_1') sy(`=sd_1`) mux(`=mu_0`) sx(`=sd_0`) rho(`=corr')
> x(140)
    Mean Std. Dev. Corr(Y,X)
\begin{tabular}{llll}
Y & \(151.36 \quad 14.178 \quad .15912\)
\end{tabular}
\begin{tabular}{lll}
X & 156.45 & 11.39
\end{tabular}
    Discretization of Y and X
Y 140
    140
    Population Transition Matrix of Y given X
```

| $M[2,2]$ |  |  |
| :--- | ---: | ---: |
|  | $y 1$ | $y 2$ |
| $x 1$ | 30.654 | 69.346 |
| $x 2$ | 20.389 | 79.611 |

## 5 Acknowledgments

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## 6 References

Black, S. E., and P. J. Devereux. 2011. Recent developments in intergenerational mobility. In Handbook of Labor Economics, ed. O. Ashenfelter and D. Card, vol. 4B, 1487-1541. San Diego: Elsevier.

O'Neill, D., O. Sweetman, and D. Van de gaer. 2007. The effects of measurement error and omitted variables when using transition matrices to measure intergenerational mobility. Journal of Economic Inequality 5: 159-178.

## About the author

Marco Savegnago is a junior economist at the Bank of Italy.

