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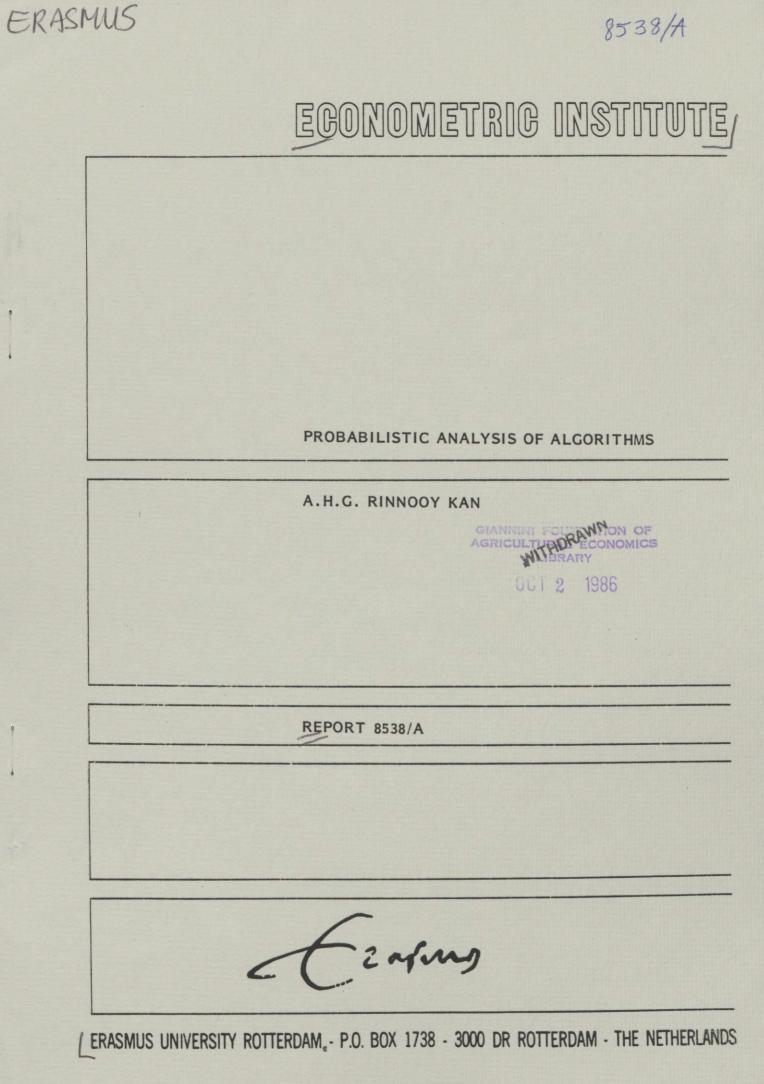
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PROBABILISTIC ANALYSIS OF ALGORITHMS by

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Abstract

An introductory and selective review is presented of results obtained through a probabilistic analysis of combinatorial algorithms. The emphasis is on asymptotic characteristics of optimal solution values, and on the relative and absolute error analysis for simple heuristics.

1. INTRODUCTION

Suppose that two thieves meet on a regular basis to divide the proceeds of their joint effort. Each stolen object has a specific dollar value and has to be assigned to one of the two. For obvious reasons, they are interested in a quick and fair partitioning scheme.

In spite of its apparent simplicity, the above combinatorial problem is not easy to solve if we insist on an optimal solution, i.e., one in which the difference between the values assigned to each thief is as small as possible. As is the case with most practical problems, this problem too is known to belong to the class of <u>NP-complete</u> problems. This implies that any optimization method for its solution could be expected to perform very poorly on some occasions: more formally, its <u>worst case running time</u> is likely to grow <u>exponentially</u> with problem size.

Hence, in choosing a solution scheme, the thieves will be forced into an unpleasant trade-off between two features of algorithmic quality: the <u>computational effort</u> (the smaller running time, the better) on one hand and the <u>computational result</u> (the smaller deviation from optimality, the better) on the other hand. Complexity theory indicates that we cannot insist on a simultaneous <u>absolute guarantee</u> for both, i.e., on a fast (polynomially bounded) running time as well as a zero deviation from the optimal solution value.

One possible way out of this dilemma is to change the perspective on the analysis by no longer demanding an <u>absolute</u> guarantee. For practical purposes an algorithm that, with respect to both effort and result, does well in the majority of cases or even on average might be perfectly acceptable.

<u>Probability theory</u> provides the natural setting for such an <u>analysis of</u> <u>algorithms</u>. This analysis starts from a specification of what an average problem instance would look like, in terms of a <u>probability distribution</u> over the <u>class of all instances</u>. The running time and the solution value of a particular algorithm are then considered as <u>random variables</u>, whose behaviour can be studied and evaluated. This approach can therefore be viewed as the analytical counterpart to the familiar <u>experimental</u> analysis in which an algorithm is tried out on a set of supposedly representative test problems and evaluated <u>statistically</u>. Here we obtain the rigor of mathematical analysis at the expense of a certain naivete, in that only relatively straightforward solution methods can be analyzed probabilistically in full detail.

Although the <u>probabilistic analysis of algorithms</u> has only recently become an active research area, it has already generated an impressive number of publications. A concise survey of this area would require the prior introduction of many techniques from probability theory and could hardly do justice to the diversity of ideas and approaches that one finds in the literature. Fortunately, a recent annotated bibliography [Karp et al. 1984] provides an up to date survey of the available articles and publications. In view of the existence of this source of detailed information, no attempt at completeness will be made in what follows below. Rather, the nature of the analysis and of the results will be illustrated by some typical examples.

In Section 2, we consider the problem of the two thieves in more detail. It is, of course, none other than the well known PARTITION problem in which one seeks to minimize the size of the largest share. In Section 3, we review some representative results that are known for problems defined in the <u>Euclidean</u> plane. In Section 4, we examine the fertile area of optimization problems defined on <u>graphs and networks</u>. In Section 5, we indicate some possibilities for future work in this very lively research area.

We conclude this introduction by a short digression on modes of <u>stochastic</u> <u>convergence</u>, clearly an essential concept if we want to analyze the notion of a random variable such as the error of a heuristic going to 0 with increasing problem size.

<u>Almost sure convergence</u> of a sequence of random variables y_n to a constant c by definition means that $\Pr\{\lim_{n \to \infty} y_n = c\} = 1$; it is a strong form of stochastic convergence and implies the weaker <u>convergence in probability</u>,

which stands for $\lim_{n\to\infty} \Pr\{|\underline{y}_n - c| > \epsilon\} = 0$ for all $\epsilon > 0$. The reverse implication holds under the additional assumption that, for all $\epsilon > 0$,

$$\sum_{j=1}^{\infty} \Pr\{|\underline{y}_n - c| > \varepsilon\} < \infty .$$
 (1)

Similarly, convergence of y_n to c in expectation, i.e.

$$\lim_{n \to \infty} |E(\underline{y}_n) - c| = 0$$
⁽²⁾

also implies convergence in probability, with the reverse implication holding under additional boundedness assumptions on \underline{y}_n .

2. THE PARTITION PROBLEM

Perhaps the simplest possible way to solve the PARTITION problem of the two thieves is to allow each thief to choose a particular item in turn until they have all been assigned. If the j-th item has value a_j (j = 1, ..., n), then this amounts to ordering the items according to decreasing a_j values $a^{(n)} \geq a^{(n-1)} \geq \cdots \geq a^{(1)}$; one thief receives $a^{(n)} + a^{(n-2)} + \cdots$, the other $a^{(n-1)} + a^{(n-3)} + \cdots$

This is clearly a fast heuristic method that may, however, produce a very inequitable result: in the worst case, the first thief may receive up to 50 percent more than the optimal partition would grant him. (Take $a_1 = 2$, $a_{21} = a_{21+1} = 2^{-1}$ ($i \ge 1$.) How about its average case behaviour? To answer that question, we specify a probability distribution over all problem instances by assuming that robberies are so frequent and haphazard that the a_j can be thought of as independent draws from a uniform distribution on, say, [0,1]. (Actually, many of the results mentioned below hold under much more general assumptions, though independence is always required.)

Under the uniform assumption, the optimal solution value \underline{Z}_n^{OPT} of the partition problem (i.e., the smallest possible size of the larger share) turns out to be <u>almost surely</u> (a.s.) <u>asymptotic to</u> the lower bound $(\sum_{j=1}^{n} \underline{a}_j)/2$:

0.00

$$\frac{\underline{z_n^{n}}}{(\underline{z_{j=1}^n} \ \underline{a_j})/2} \neq 1 \qquad (a.s.)$$
(3)

This result provides a first example of <u>asymptotic probabilistic value</u> analysis: for n large enough, the optimal solution value Z_n^{OPT} can be guessed

with increasing (relative) accuracy. What about the heuristic solution value Z_n^H , i.e. the size of the larger share under the heuristic scheme proposed above? We know that

$$\underline{Z}_{n}^{H} = \frac{1}{2} \sum_{j=1}^{n} \underline{a}_{j} + \frac{1}{2} [(\underline{a}^{(1)} - \underline{a}^{(2)}) + (\underline{a}^{(3)} - \underline{a}^{(4)}) + \cdots], \qquad (4)$$

and it is not hard to show that the difference between \underline{Z}_n^H and $(\sum_{j=1}^n \underline{a}_j)/2$ a.s. converges to 1/4, so that also

$$\frac{\underline{z}_{n}^{H}}{(\sum_{j=1}^{n} \underline{a}_{j})/2} + 1 \quad (a.s.)$$
 (5)

and hence

$$\frac{\underline{z}_{n}^{H}}{\underline{z}_{n}^{OPT}} \rightarrow 1 \qquad (a.s.) \tag{6}$$

This implies that the heuristic is asymptotically optimal: its relative error (i.e., its percentage deviation from the optimum) $(\underline{Z}_n^H - \underline{Z}_n^{OPT})/\underline{Z}_n^{OPT}$ a.s. goes to 0. Hence, a probabilistic analysis leads to a much more optimistic conclusion than a worst case oriented one.

What about the <u>absolute error</u> $Z_n^H - Z_n^{OPT}$ or the <u>absolute difference</u> between the two shares? Neither of these two quantities goes to 0 for the above heuristic, so there is room for improvement. A slightly more sophisticated scheme would be to allow each thief in turn to select items until the value of his share exceeds the value of the current share of his colleague. Even from a worst case point of view, this is a much more reasonable approach: the larger share can never exceed its smallest possible size by more than 16 2/3 percent [Graham 1969]. (For a worst case example, take e.g., $a_1 = a_2 = 3$, $a_3 = a_4 = a_5 = 2$.) In a probabilistic sense, the difference is even more impressive. Of course, the relative error again goes to 0, but the absolute difference between the shares \underline{d}_n^H (and hence the absolute error) also satisfies

$$\frac{d^{H}}{n} \neq 0 \qquad (a.s.) \tag{7}$$

[Frenk and Rinnooy Kan 1983]. To prove this result, one observes that

$$\underline{d}_{n}^{H} \leq \max\{\underline{d}_{n-1}^{H} - \underline{a}^{(1)}, \underline{a}^{(1)}\}$$

$$(8)$$

which, after repeated application, yields that

$$\underline{d}_{n}^{H} \leq \max_{1 \leq k \leq n} \{ \underline{a}^{(k)} - \sum_{j=1}^{k-1} \underline{a}^{(j)} \}$$

$$\leq \underline{a}^{([\delta n])} + \max\{ \underline{a}^{(n)} - \sum_{j=1}^{[\delta n]} \underline{a}^{(j)}, 0 \}$$
(9)

for all $\delta > 0$. The first term converges a.s. to δ and can therefore be made arbitrarily small; for any fixed δ , the second term converges a.s. to 0 since $\sum_{j=1}^{\lfloor \delta n \rfloor} a^{(j)} = 0(n)$ and $a^{(n)}$ is o(n) for every distribution with finite first moment.

The two results presented so far demonstrate the importance of the theory of <u>order statistics</u> for the analysis of heuristics that involve the sorting of numbers; <u>priority rules</u> generally fall into this class.

Can we do still better? One weakness of a result such as (7) is its <u>asymptotic</u> <u>nature</u>, i.e. its validity only for <u>sufficiently large</u> values of n. At the very least, one would like to know the rate at which $\frac{d^{H}}{n}$ converges to 0. It can be shown [Frenk and Rinnooy Kan 1984] that

$$\lim \sup_{n \to \infty} \frac{\underline{d}_n^{H}}{\log \log n/n} < \infty \qquad (a.s.), \qquad (10)$$

and a simple argument shows that the <u>rate of convergence</u> for this heuristic has to be at least 1/n in expectation [Karp, 1984]. This rate is good but not quite good enough: indeed, it is also known that $\frac{d^{OPT}}{n}$, the smallest possible absolute difference, satisfies

$$\lim \sup_{n \to \infty} \frac{\frac{d^{OPT}}{n^2 2^{-n}} < \infty \qquad (a.s.) \qquad (11)$$

[Karmarkar et al. 1984]. Hence, the exponential effort that may be required for the computation of the optimal partition is at least rewarded by an exponential decrease to 0 of the difference between the two shares. Can this also be achieved a.s. in polynomial time?

The answer to this question is unknown, but the previous heuristic can again be improved upon by essentially assigning two items at a time and compensating for their difference. More precisely, in the first iteration $a^{(n)}$ would be assigned to one thief, and $a^{(n-1)}$ to the other. The two terms would then be replaced by a single item of value $a^{(n)} - a^{(n-1)}$ and the process would be repeated on the new set of n-l items until only one item remains; its value represents the difference between the two shares. A simple backtracking procedure establishes the partition in terms of the original items.

In the worst case, this heuristic method is not better than the previous one. The probabilistic analysis of its performance is quite difficult: as on so many other occasions, each step in the algorithm conditions the probability distribution encountered in the succeeding steps in a complicated fashion. Since the difference of two independent uniformly distributed random variables follows a triangular distribution, there is no distributional invariance throughout the steps of this method and yet such invariance is an essential prerequisite for a successful analysis. One way to overcome this kind of obstacle is to change the algorithm so that (with high probability) the value produced will not be affected but its modified behaviour can be analyzed rigorously. In this particular case, the new versions of the algorithm works roughly as follows. In iteration m, it deals with numbers in an interval [0, α_m], which is first partitioned into subintervals of size α_{m+1} . The original differencing method is then applied to random pairs of numbers taken from the set S_i found in the i-th subinterval (i = 1, ..., α_m/α_{m+1}), until a set $S_i \subset [0, \alpha_{m+1}]$ is obtained and at most one of the numbers originally in S_i remains. The original differencing method is then used on the set of remaining numbers to reduce them to a single number in [0, α_{m+1}].

Now, before the method can be applied recursively to $[0, \alpha_{m+1}]$, the issue of distributional invariance has to be addressed. To do so, we assume (inductively) that $S = \cup S_i$ was divided into two subsets, a subset G of 'good' points that can be assumed to come from a uniform distribution over $[0, \alpha_m]$ and a subset B of 'bad' points. During the application of the differencing method to the sets S_i the numbers entering S'_i are labeled good only if they are obtained as the difference of two good numbers in S_i . As a result, we know that the good numbers in $S' = \cup S'_i$ follow a distribution related in the triangular one. In a final step, a subset of these numbers is relabeled 'bad' so that the remaining good ones are again uniformly distributed over $[0, \alpha_{m+1}]$. All that remains to be shown is that for appropriate choices of α_m , enough good numbers remain for the method to reach the number of iterations that is required for a good result.

Through this approach, it could be established in [Karwarkar & Karp 1982] that

$$\lim \sup_{n \to \infty} \frac{\underline{d}_n^H}{n^{-\log n}} < \infty \text{ (a.s.)}$$
(12)

Thus, in O(n log n) time this method guarantees a rate of convergence that is superpolynomial, yet subexponential. It is tempting to conjecture that this is best possible for a polynomial time heuristic!

We have dealt with this simple example in some detail, since it exhibits many of the ingredients typically encountered in a probabilistic analysis:

- a <u>combinatorial problem</u> that may be difficult to solve to optimality (the PARTITION problem is NP-complete);
- a probability distribution over all problem instances, that generates the problem data as realizations of independent and identically distributed (i.i.d.) random variables (the a_i are independent and uniform over [0,1]);
- a probabilistic value analysis that yields a description of the asymptotic optimal solution value as a simple function of the problem data $(\underline{Z}_n^{OPT}/(n/4) + 1 (a.s.));$
- simple, fast heuristics whose <u>relative error</u> or <u>absolute error</u> may decrease to 0 in some stochastic sense or may be otherwise well behaved;
- a rate of convergence analysis that allows further differentiation among the heuristics.

The state of the art for a particular problem class can conveniently be monitored by means of the above concepts. Consider, for example, the MULTIKNAPSACK problem $\max\{\sum_{j} c_{j}x_{j} \mid \sum_{j} a_{ij}x_{j} \leq b_{i} \ (i = 1, \dots, m), x_{j} \in \{0,1\} \ (j = 1, \dots, n)\}$, which is a generalization of PARTITION (take $m = 1, c_{j} = a_{1j} \ (j = 1, \dots, n), b_{1} = (\sum_{j=1}^{n} a_{1j})/2).$

Let us assume that the c_j and a_{ij} are i.i.d. uniform on [0,1] and that $b_i = n\beta_i$ is constant. As above, we are interested in the optimal solution value as a function of $\beta = (\beta_1, \dots, \beta_m)$ and in heuristics whose error vanishes asymptotically with high probability.

The analysis of this problem in [Meanti et al. 1984] is of interest in that exploits the close relationship (in probability) between certain difficult nonconvex combinatorial optimization problems such as MULTIKNAPSACK and their convex LP relaxations obtained by replacing the constraints $x_j \in \{0,1\}$ by $0 \leq x_j \leq 1$. It is easy to see that the absolute difference between the

solution values of these two problems is bounded by m, so that the relative error that we make by focusing on the LP relaxation goes to 0. But the LP relaxation (or, rather, its dual) is much easier to analyze: its value is given by

$$\min_{\lambda} \frac{L}{L} (\lambda)$$
(13)

with

$$\underline{\mathbf{L}}_{\mathbf{n}}(\lambda) = \Sigma_{\mathbf{i}=1}^{\mathbf{m}} \lambda_{\mathbf{i}} \mathbf{b}_{\mathbf{i}} + \max\{\Sigma_{\mathbf{j}}(\underline{\mathbf{c}}_{\mathbf{j}} - \Sigma_{\mathbf{i}=1}^{\mathbf{m}} \lambda_{\mathbf{i}} \underline{\mathbf{a}}_{\mathbf{i},\mathbf{j}}) \mathbf{x}_{\mathbf{j}} \mid 0 \leq \mathbf{x}_{\mathbf{j}} \leq 1\}$$
(14)

where the maximization problem in (14) is solved by setting

$$\underline{x}_{j}(\lambda) = \begin{cases} 1 & \text{if } \underline{c}_{j} - \sum_{i=1}^{m} \lambda_{i} \underline{a}_{ij} \geq 0 \\ 0 & \text{otherwise} \end{cases}$$
(15)

Results from convex analysis can then be used to establish that the optimal solution value $Z_n^{\mbox{OPT}}$ satisfies

$$\frac{Z^{\text{OPT}}}{n} \neq \min_{\lambda} L(\lambda) \quad (a.s.), \quad (16)$$

where

$$L(\lambda) = \Sigma_{i=1}^{m} \lambda_{i} \beta_{i} + \frac{1}{n} E_{\underline{c}}^{T} \underline{x}(\lambda) - \frac{1}{n} \Sigma_{i=1}^{m} \lambda_{i} E_{\underline{a}}^{T} \underline{x}(\lambda)$$
(17)

is a convex, twice differentiable function with a unique minimum that can actually be computed in closed form in some simple cases (e.g., for m=1).

Not surprisingly, successful heuristics for this problem also have a strong LP flavor (cf. [Frieze & Clarke 1981]). A natural one to consider is the <u>generalized greedy heuristic</u> in which x_j 's are set equal to 1 in order of non-increasing ratio's $c_j/\sum_{i=1}^m \lambda_i a_{ij}$. If the λ_i are chosen to be equal to the values minimizing the right hand side of (16), then the relative error of this greedy method goes to 0 a.s. A heuristic whose absolute error vanishes asymptotically is not known, however, and further analysis of the model reveals puzzling differences between the minimization and maximization version of MULTIKNAPSACK that still have to be resolved. None the less, the probabilistic analysis of the model yields surprisingly high returns.

What are the strong and the weak points of the approach? On one hand, the algorithmic insights nicely complement the more traditional worst case analysis, with an emphasis that is not so much on an exact <u>guarantee</u> that a certain running time or a certain error will not be exceeded, as on an <u>explanation</u> why the algorithm may perform so much better in practice than the worst case analysis would seem to suggest. On the other hand, the results are obtained under probabilistic assumptions that can always be questioned, and are usually only valid for 'sufficiently large' problem sizes, with little indication of how large these might have to be. Rate of convergence results somewhat compensate for this latter deficiency.

Apart from its contribution to error analysis, a probabilistic value analysis may, however, yield additional benifits. An estimate of the optimal solution value can, for instance, be used in a <u>branch and bound procedure</u> to replace weak upper or lower bounds, at the possible expense of sacrificing optimality but at the gain of a significant improvement in running time [Derigs 1984]. Or it may be used in a <u>two stage stochastic programming</u> problem, where the first stage decision determines the value of some parameters of the second stage problem. A probabilistic value analysis may then reveal how the second stage solution value depends on these parameters so that they can be given optimal values in the first stage [Stougie 1985].

As a fist step towards a classification of results in probabilistic analysis, let us note that the PARTITION problem and its generalization, the MULTIKNAPSACK problem, both have a problem input that consists of <u>numbers</u>. The probabilistic analysis of algorithms for these problems usually assumes that these are independently generated from a fixed distribution. To close this section, we briefly review some typical results that were obtained for other problems in this category.

The MULTIMACHINE SCHEDULING problem is an appropriate starting point, since it can be viewed as yet another generalization of PARTITION: it is the problem to distribute jobs with processing times a_1 , ..., a_n among m identical machines so as to minimize the maximal sum of processing times assigned to any machine (the <u>makespan</u>); the PARTITION problem corresponds to the case that m=2.

The second heuristic proposed for PARTITION can be viewed as a special case of the LPT (Largest Processing Time) heuristic for MULTIMACHINE SCHEDULING, in which jobs are assigned to the earliest available machine in order of nonincreasing a_j . If the a_j are i.i.d. uniform on [0,1], the proof technique based on order statistics can again be applied to show that the absolute error of the LPT heuristic goes to 0 a.s.; the optimal solution value is asymptotic to $\frac{1}{2}$ n/m. The order in which jobs are assigned to machines turns out to be really essential: an arbitrary list scheduling heuristic will have a relative error that goes to 0 a.s. but an absolute error that does not. The Karp-Karmarkar heuristic can be extended to the case of arbitrary m to improve the rate of convergence to optimality from 1/n for the LPT rule to $n^{-log n}$. We refer to [Coffman et al. 1983] for additional references.

The famous BIN PACKING problem is in a sense dual to MULTIMACHINE SCHEDULING: here, the makespan is fixed (say, equal to 1) and the objective now is to find the minimum number of bins (machines) into which the items (jobs) of size a1, ..., a_n can be packed. The probabilistic analysis of this problem is again usually carried out under the assumption that the jobs are i.i.d. uniform on [0,1]. It has yielded many beautiful results. To give one example, consider the heuristic that inspects the items in order of decreasing a i and matches each item a_k with the largest unassigned item a_k satisfying $a_k + a_k \leq 1$. To analyze the performance of this heuristic, we consider the set of all a on the interval [0,1] and replace each item a_k larger than 1/2 by 1 - a_k , marking it with a '+'; the items smaller than 1/2 get a '-' sign. The heuristic now amounts to pairing each '+' with the largest '-' to its left; the number of poorly filled bins (i.e., bins with only one item) is related to the excess of '+'s over '-'s. But the sequence of '+'s and '-'s can be viewed as generated by flips of a coin, and results from the theory of random walks can be invoked to show that the expected total number of bins used is $n/2 + O(\sqrt{n})$; since the expected optimal number of bins is known to be $n/2 + \Omega(\sqrt{n})$, the relative error of this heuristic goes to 0 in expectation [Frederickson 1980; Knodel 1981].

This analysis reveals an interesting connection between <u>matching problems</u> on the interval [0, 1/2] and bin packing methods. In a similar vein, a connection can be established between <u>on line</u> bin packing methods (i.e., methods in which items arrive in arbitrary order and have to be irrevocably assigned to a bin right away; the previous method is <u>off line</u>) and matching problems on a

square, where one of the dimensions corresponds to item size and the other dimension represents time, i.e., it indicates the order in which the items arrive for packing. In this case, feasibility dictates that each point can only be matched to a <u>right upward</u> neighbour; this matching problem was studied in a different context in [Karp et al. 1985], and the clever extension and refinement of these results to on-line bin packing can be found in [Shor 1984]. Further references can be found in the useful survey [Coffman et al. 1983] that was already quoted above.

A final number problem that deserves to be mentioned is LINEAR PROGRAMMING, not because it is a hard combinatorial problem (the Khachian method solves it in polynomial time) but because probabilistic analysis played such a vital role in understanding the excellent average performance of the <u>simplex method</u>. The history of the analysis illustrates the importance of an appropriate <u>probabilistic model</u>: ultimately, the concept of a random polytope as being generated by m fixed hyperplanes in \mathbb{R}^n and m coin flips to determine the direction of the corresponding inequalities reduced the computation of the average number of simplex pivots to a combinatorial counting argument. Within this model, various simplex variants admit of a quadratic upper bound on the expected number of iterations (including those in Phase I), which takes us very close to the behaviour observed in practice [Haimovich 1983; Adler et al. 1983; Adler & Megiddo 1983; Todd 1983].

In the next section, we turn to problems with a geometric flavor, whose probabilistic analysis involves random sets of points in the Euclidean plane.

3. EUCLIDEAN PROBLEMS

In this section we are concerned with problems whose input includes a set of n points in the Euclidean plane. The most famous problem of this type is surely the TRAVELLING SALESMAN problem, which is to find the shortest tour connecting the n points. It has a venerable history, of which the probabilistic analysis forms one of the most recent chapters (see [Lawler et al. 1975]).

To carry out such an analysis, one usually assumes the points to be uniformly distributed over a fixed region, e.g., the unit (1×1) square. Under such an assumption, it is not difficult to arrive at an intuitive probabilistic value

analysis for the TRAVELLING SALESMAN problem. For n large, the length of an optimal tour through a 2 × 2 square with 4n points will be approximately 4 times as large as the optimal tour in a unit square with n points. Scaling back the 2 × 2 square to a unit one, we conclude that the optimal tour length Z_n^{OPT} is likely to grow proportionally to \sqrt{n} . Indeed, a heuristic from [Few 1955] shows that its value is bounded deterministically from above by $\sqrt{(2n)}$, and it is not hard to show that there exists a positive constant c such that it is a.s. bounded by $c\sqrt{n}$ from below.

The actual convergence argument is much more difficult and was first provided in [Beardwood et al. 1959]: as expected,

$$\frac{Z_{n}^{OPT}}{\sqrt{n}} \neq \beta \quad (a.s.) \tag{18}$$

where β is a constant that has been estimated empirically to be about 0.765.

The proof of this result involves a technique useful in a broader context: rather than viewing the problem on a fixed set of n points, uniformly distributed on a square, they are assumed to be generated by a <u>Poisson process</u> of intensity 1. The advantage is that point sets in disjoint regions are now fully independent; the disadvantage that an expression for, say, the expected routelength F(t) in a square $[0,t] \times [0,t]$ has to be converted back into a result for the expected length EZ_n^{OPT} of a tour through n points, using the relation

$$F(t) = \sum_{n=1}^{\infty} t E \underline{Z}_{n}^{OPT} e^{-t^{2}} \frac{(t^{2})^{n}}{n!} .$$
 (19)

Certain Tauberian theorems allow one to do precisely that.

In the specific case of the TRAVELING SALESMAN problem, F(t) is computed through the use of a heuristic that embodies the intuitive insight mentioned above. The square $[0,t] \times [0,t]$ is divided into m^2 equal size subsquares; a tour in each subsquare has expected length F(t/m). These tours can be linked to form a feasible solution to the original problem by adding segments of length O(tm), and hence

$$F(t) \leq m^2 F(t/m) + O(tm)$$

(20)

It is not hard to see that this imples that $F(t)/t^2$ converges to a constant β and to conclude from (19) that $EZ_n^{OPT} = \beta/n$. The argument is completed by proving the variance of Z_n^{OPT} to be O(1) through the <u>Efron-Stein</u> inequality; Chebyshev's inequality and the Borel-Cantelli lemma then yield (18) [Karp & Steele 1985].

Given the result of the probabilistic value analysis, it now becomes attractive to search for heuristics whose absolute error is $o(\sqrt{n})$; their relative error then goes to 0 almost surely. As in the case of other Euclidean problems, <u>partitioning heuristics</u> do precisely that. Generally, in these heuristics the square is appropriately partitioned into subregions (e.g. rectangles), subproblems defined by the points in each subregion are analyzed separately, and a feasible solution to the problem as a whole is composed out of the separate results.

For the TRAVELLING SALESMAN problem, one partitioning approach is to execute an alternating sequence of horizontal and vertical <u>cuts</u> through the point with current median vertical and horizontal coordinate respectively, until the resulting rectangles contain no more than $/(\log n)$ points. Each of these subproblems is solved to optimality by some enumerative technique (say, dynamic programming, which takes $O(n^{\varepsilon})$ time per rectangle, and hence $O(n^{1+\varepsilon})$ time overall, for all $\varepsilon > 0$). The resulting tours define a connected graph with even degree at each point; the <u>Euler walk</u> that visits each edge of this graph can be converted into a tour of no greater length by eliminating multiple visits to one point. The difference between the length of this tour and the optimal one can be shown to be of the same order as the total perimeter of the rectangles generated, which is easily seen to be o(/n) in this case. Thus, the relative error of the heuristic goes to 0 a.s. [Karp 1977, Karp and Steele 1985].

Not much is known about the rate of convergence to optimality of this heuristic, nor is any heuristic known whose absolute error goes to 0 asymptotically. The partitioning approach, however, has been generalized to a variety of other Euclidean problems. Consider, for example, the CAPACITATED ROUTING problem, where the points have to be equally distributed among salesmen, each of whom can visit no more than q customers before returning to

a common depot. If we assume that the i-th customer is at distance r_i from the depot ($\underline{r_i}$ i.i.d., with finite first moment \underline{Er}), then it is not hard to prove that the optimal solution value satisfies

$$Z_{n}^{OPT} \geq \max\{T_{n}^{OPT}, 2 \frac{\Sigma_{i=1}^{r} r_{i}}{q}\}$$
(21)

where T_n^{OPT} is the length of a single travelling salesman tour through the n customers. A tour partitioning heuristic, in which T_n is optimally divided into [n/q] consecutive segments containing q customers each, can be shown to yield a value no more than the sum of the two terms on the right hand side of (21), so that - in view of $(18) - \underline{Z}_n^{OPT}$ is a.s. asymptotic to 2nEr/q. As a byproduct of this result, one obtains that the tour partitioning heuristic (of the <u>'route first, cluster second</u>') type has relative error going to 0 a.s.. A similar result holds for certain region partitioning heuristics (<u>'cluster first, route second</u>'): as indicated above, their absolute error is dominated by the total perimeter of the subregions generated, which in this case is $o(\sqrt{(n/q)})$ and hence vanishes asymptotically relative to \underline{Z}_n^{OPT} .

All these results presuppose that q is a constant that does not grow with n; if it does, than the results hold as long as $q = o(\sqrt{n})$. Above this threshold, $\frac{Z^{OPT}}{n}$ behaves as in (18), since at that point the total cost of moving among the groups of customers, T_n , starts to dominate the total cost $2\Sigma_{i=1}^n r_i/q$ of reaching these groups from the depot (cf. (21)).

The ubiquitous presence of partitioning techniques in probabilistic Euclidean analysis points in the direction of a common generalization, which was indeed provided in [Steele 1981]. There, (18) is generalized to problems of arbitrary dimension, provided that the objective function is <u>subadditive</u> on sets of points (and that a few other technical conditions are satisfied). The TRAVELLING SALESMAN problem is one example, the MATCHING problem (find the segments linking each point to a neighbour, of minimum total length) is another one: see [Papadimitriou 1978] for a probabilistic value analysis and a heuristic that together establish the optimal solution value to be a.s. asymptotic to ε/n , with $\varepsilon \in [0.25, 0.40106]$.

We close this section by discussing an interesting class of problems that cannot quite be handled by Steele's techniques, i.e. LOCATION problems, in which k depots have to be located so as to minimize the average distance between each point and its closest depot (the k-MEDIAN problem) or the maximum of these distances (the k-CENTER problem). The probabilistic value analysis for both problems lead to surprisingly similar results, provided that k =0 (n/log n) (the case $k = \alpha n$ is partially open; see [Hochbaum & Steele 1981]). Both optimal solution values are asymptotically proportional to $1/\sqrt{k}$, albeit for different constants of proportionality.

The analysis leading to these results is of special interest, since it relies heavily on the similarity between the original (discrete) problem for large n and the <u>continuous</u> problem in which customer demand is not concentrated in a finite set of points but spread uniformly and continuously over the entire region. A simple partitioning heuristic in which a depot is located in each of k nonempty subsquares of size $1/\sqrt{k}$ by $1/\sqrt{k}$ each already provides an $O(1/\sqrt{k})$ upper bound on both optimal solution values. An asymptotically optimal heuristic, however, is only obtained by partitioning the region in regular <u>hexagons</u> (the <u>honeycomb</u> heuristic), with the constants of proportionality being determined by the optimal solution value of the continuous problem with k=l over one such hexagon. This heuristic actually solves the continuous problem to optimality, and a detailed error analysis shows that, for n sufficiently large, its relative error in the discrete case becomes vanishingly small [Haimovich 1984, Zemel 1984].

4. GRAPHS AND NETWORKS

We now turn to the rich area of combinatorial optimization problems defined on graphs and networks. One of the reasons for the wide variety of probabilistic results for this class of problems is the existence of a substantial theory dealing with <u>random graphs</u>. There are two popular definitions of this concept: $G_{n,p}$ is defined to be the (undirected) graph for n vertices for which each of the $\frac{1}{2}n(n-1)$ edges occurs independently with equal probability p; G_n^N is defined by assuming that each of the $\binom{n(n-1)/2}{N}$ undirected graphs on n vertices occurs with equal probability (see [Bollobas 1985] for a survey of the theory). Especially for <u>structural graph optimization</u> problems, in which we are interested in graph properties that depend only on the node-edge incidence structure, random graphs provide a natural probability distribution over the set of all problem intances of size n.

Continuing in the spirit of the previous two sections, we again refer to [Karp et al. 1985] for a list of recent references in this area and review only a few typical probabilistic analyses of heuristics for NP-complete problems, whose expected performance compares favorably with the limits set by worst case analysis. In doing so, we (reluctantly) exclude many beautiful results on problems of CONNECTIVITY and MATCHING that can be solved in worst case polynomial time.

A typical example of a difficult structural problem is the CLIQUE problem of finding a <u>complete subgraph</u> of G that has maximal size $\omega(G)$. To carry out a probabilistic value analysis of $\omega(G_{n,p})$ for fixed p, we observe that the expected number of cliques of size k in such a graph is equal to $\binom{n}{k}p^{k(k-1)/2}$. We could expect the maximal clique of size k to occur when the expected number is approximately equal to 1, i.e., when (from Stirling's approximation of k!)

$$\frac{1}{\sqrt{2\pi k}} \left(\frac{nep^{(k-1)/2}}{k}\right)^k \approx 1.$$
 (14)

The left hand side of (14) decreases very rapidly as k increases and passes through the value 1 where

$$\frac{nep^{(k-1)/2}}{k} = 1$$
 (15)

i.e. when

$$k = 2 \log_{1/p} n + 2 \log_{1/p} e - 2 \log_{1/p} k + 1$$
 (16)

so that, approximately, $k \approx k(n,p)$ with

$$k(n,p) = 2 \log_{1/p} n - 2 \log_{1/p} \log_{1/p} n + 2 \log_{1/p} (e/2) + 1.$$
(17)

This estimate turns out to be very sharp indeed. In [Matula 1976], it is proved that, for all $\varepsilon > 0$,

$$\lim_{n\to\infty} \Pr\{[k(n,p) - \varepsilon] \leq \underline{Z}_n^{\text{OPT}} = \omega(G_{n,p}) \leq [k(n,p) + \varepsilon]\} = 1$$
(18)

so that, for large enough n, the size of the largest clique can be predicted to be one of two consecutive integers with high probability of success.

This precise probabilistic value analysis again encourages the search for a fast heuristic whose absolute error compares favorably to $2 \log_{1/p} n$. Consider, for instance, the <u>sequential greedy algorithm</u>, which considers the vertices of G in arbitrary order and adds a vertex to the current clique if it is adjacent to all its members. For an analysis of the performance of this method, one observes that the expected number of trials to increase the clique size from j to j+1 is $1/p^{j}$, so that we might guess the ultimate clique size Z^{H} to satisfy

$$\sum_{j=0}^{Z^{H}-1} 1/p^{j} = \frac{1-1/p^{Z^{H}}}{1-1/p} = n$$
(19)

i.e.,

$$Z^{H} \approx \log_{1/p}^{n}$$
 (20)

A more precise analysis shows that, indeed, this greedy approach a.s. yields a clique of size $(1/2 - \varepsilon)Z_n^{OPT}$ [Grimmett and McDiarmid 1975]. Thus, the relative error does not go to 0, but is almost surely close to 50 percent. (There is by the way, no known polynomial time heuristic with any constant worst case bound on the relative error.)

The above result has immediate implications for the problem to find the INDEPENDENT SET in G of maximal size; it coincides with the maximal size clique in the complement of G. Again, the sequential greedy approach, which picks up each successive vertex that is not adjacent to any member of the current independent set, produces an independent set whose size is a.s. close to 50 percent of the optimal value. The COLORING problem, which is to partition the vertices of G into the smallest possible number $\chi(G)$ of independent sets, is much harder to analyze: the asymptotic optimal solution value $\underline{Z}_{n}^{OPT} = \chi(G_{n,p})$ is known for p = 1/2 [Korzanov 1980], though that (Russian) announcement has not been verified. The heuristic method, which greedily finds an independent set as above, deletes it and repeats on the remaining graph does poorly [McDiarmid 1979] but good enough to get within a factor of $(2 + \varepsilon)$ a.s. [Grimmett and McDiarmid 1975].

The other class of structural graph problems for which probabilistic analysis has been successful is the HAMILTONIAN CIRCUIT problem of searching for a

simple cycle containing all vertices. The emphasis here is more on conditions under which such a cycle exists (e.g., a.s. in G_n^N when $N = \frac{1}{2} n \log n + \frac{1}{2} n \log \log n + cn$ [Komlos and Szemeredi 1983]) and less on the development of fast heuristics that, with high probability, would be successful in finding such a cycle if it existed. However, the heuristic principle of <u>extension</u> and <u>rotation</u> has been applied to this class of problems with considerable success [Posa 1976, Angluin & Valiant 1979]. The general idea is as follows. Given a path of vertices (v_0, \dots, v_k) , one of the neigbours of v_k , say w, sampled randomly and the edge $\{v_k, w\}$ is added to the path. If w is not in the path, it is adjoined and the method is applied to w. If $w = v_{\ell}$ ($0 \le \ell \le k-1$), then the edge $\{v_{\ell}, v_{\ell+1}\}$ is removed from the path and the method is applied to $v_{\ell+1}$. If N exceeds the threshold by a sufficient amount (e.g., N = cn log n, for large enough c) this method will be successful with high probability.

We now turn briefly to number problems on weighed graphs, i.e., graphs with weights on the edges, an area which mixes features addressed in Section 2 with the theory of random graphs. Here, most results refer to problems that admit of a worst case polynomially bounded algorithm. A typical example is provided by the LINEAR ASSIGNMENT problem of minimizing $\sum_{i} c_{i\pi(i)}$ over all permutations π . If the c_{ij} are i.i.d. with distribution F, then a probabilistic value analysis can be arrived at by viewing the problem as weighted matching problem on a directed bipartite graph, with weights \underline{b}_{ij} on edge (i,j) and \underline{d}_{ji} on edge (j,i) such that $\underline{c}_{ij} \stackrel{d}{=} \min{\{\underline{b}_{ij}, \underline{d}_{ji}\}}$. If we now remove all edges except the s outgoing ones of minimal weight at each vertex and disregard edge orientations in the resulting bipartite graph, a result by Walkup shows that a perfect matching will be present with probability 1 - O(1/n) if s=2 and $1 - O(2^{-n})$ if s > 2. Hence, Z_n^{OPT} is essentially determined by the small order statistics of F. In the case that the <u>c</u> are uniform on [0,1], this yields that $EZ_n^{OPT} \leq 3$ [Walkup 1979]; generally, EZ_{n}^{OPT} is asymptotic to $nF^{-1}(1/n)$ [Frenk & Rinnooy Kan 1985].

Many other problems in this category have also been succesfully analyzed. For instance, in [Frieze 1982], the MINIMUM SPANNING TREE problem is studied under the assumption that the graph is complete and the edge weights are i.i.d. uniform on [0,1]; the expected optimal solution value is equal

to $\sum_{k=1}^{\infty} 1/k^3 \approx 1.2202$. We also refer to [Perl 1977] for an analysis of the SHORTEST PATH problem; further references can be found in [Karp et al. 1985].

An NP-complete problem that belongs in this category is the ASYMMETRIC TRAVELLING SALESMAN problem, defined on a complete directed graph. We refer to [Steele 1985] for a probabilistic value analysis for a Euclidean variant of this problem n. The optimal solution value is asymptotic to ζ/n in expectation; see also [Karp 1979] for a heuristic that patches the subcycles appearing in the linear assignment relaxation together to achieve a relative error going to 0 in expectation. Perhaps the most peculiar result has been obtained for its generalization, the QUADRATIC ASSIGNMENT problem

 $\max \{ \sum_{i=1}^{n} \sum_{j=1}^{n} \sum_{k=1}^{n} \sum_{\ell=1}^{n} c_{ij} d_{k\ell} x_{ik} x_{j\ell} \mid \sum_{i=1}^{n} x_{ij} = 1, (j = 1, ..., n); \\ \sum_{j=1}^{n} x_{ij} = 1 (i = 1, ..., n), x_{ij} \in \{0,1\} \}.$

In [Burkhard and Fincke 1982; Frenk et al. 1982], it is shown that for this problem with \underline{c}_{ij} and $\underline{d}_{k\ell}$ i.i.d., the ratio of the best and the worst possible solution value tends to 1 in probability. It shows an unexpected side benefit of probabilistic analysis, in that it clearly indicates how not to generate test problems for an empirical analysis of heuristic solution methods!

5. CONCLUDING REMARKS

In this final section, we first discuss a limitation of the preceding survey in that the probabilistic behaviour discussed there was only caused by factors extraneous to the algorithm itself. The algorithm itself could also contribute to the randomness of its outcome by containing <u>random steps</u>, i.e., steps whose result depends partially on a random mechanism.

A very early example of an <u>algorithm that flips coins</u> is the quicksort method for SORTING. In each step, the algorithm picks a number a_j from the set $\{a_1, \ldots, a_n\}$ to be sorted; it divides the set into the subset smaller than a_j and the subset larger than a_j , and repeats recursively on each subset of cardinality greater than two. The worst case and best case running time of the method are easily seen to be $O(n^2)$ and $O(n \log n)$ respectively; the average running time is $O(n \log n)$ under a variety of distributional assumptions but also under the assumption that for a <u>fixed input</u> the splitting number a_j is chosen randomly from the set of candidates. Notice that the behaviour of the algorithm is now a random variable independent of any distributional assumption, avoiding the always controversial issue of what a random instance of a particular problem looks like.

The formal study of randomized algorithms is far from complete, and in particular the real power of randomization remains a mysterious issue; for instance, it is not clear to what extent (if any) the class of problems that can be solved in randomized polynomial time (i.e. fast with high reliability) strictly includes the class of problems that can be solved in worst case polynomial time. A recent annotated bibliography [Maffioli et al. 1985] provides a useful survey of the area. We restrict ourselves once again to a discussion of a few examples that highlight the variety of results produced so far.

Historically, PRIMALITY TESTING was the first successful application of randomization. In the algorithm in [Rabin 1980], a number is submitted to k tests and declared to be prime if it passes all of them, with the probability of it being composite after all equal to 2^{-k} . Such an algorithm is called a <u>Monte Carlo method</u>, in contrast to a <u>Las Vegas method</u> in which the algorithm never produces an incorrect answer, but may - with small probability - produce no answer at all. The method for GRAPH ISOMORPHISM in [Babai 1979] is of this nature. The two examples above are of special interest in that they concern two problems whose computational complexity (polynomially solvable or NP-complete) is still unknown.

Generally, randomization may produce a speed-up of a polynomial computation at the expense of a little uncertainty, as for instance for a fast BIPARTITE MATCHING algorithm in [Ibarra & Moran 1981] with a small probability of error, or an $O(|V|^2)$ expected time algorithm for CIRCUIT COVERING ([Itai & Rodeh 1978]) that with small probability will not terminate. For NP-hard problems, there are other potential benefits. <u>Statistical inference</u> has been suggested as a way to estimate the optimal solution value from a sample of (supposedly independent) local minima, i.e., feasible solutions whose value cannot be further improved by a <u>local improvement method</u> ([Dannenbring 1977]). And the fashionable simulated annealing approach can be viewed as a randomized version

of such a local improvement heuristic, in which neighbouring solutions that decrease the quality of the current feasible solution are also accepted, albeit with a small probability that is appropriately decreasing over time ([Kirkpatrick et al. 1983]).

It should be clear by now that the area of probabilistic analysis still harbors many interesting research challenges. The purpose of the preceding sections has, again, not been to provide an exhaustive review, but to provide some typical examples that convey the flavour of this area. They have ranged from the very complete insight we have into various solution methods for the PARTITION problem to the less satisfactory state of the art for the CLIQUE and the COLORING problem. Clearly, a lot of problems and a lot of algorithms await investigation. It is not hard to formulate open questions for probabilistic analysis; so far, however, it has turned out to be quite hard to come up with satisfactory answers for any but the simplest heuristics.

A particularly fascinating possibility is the development of a complexity theory that would lead to a class of problems for which solution to optimality in polynomial expected time is as unlikely as the equality of P and NP. A first step in that direction can be found in [Levin 1984], where a TILING problem is introduced, together with a probability distribution over its problem instances, such that any other problem with a (mildly restricted type of) probability distribution is reducible to the TILING problem.

To establish completeness for other problems in this class is a major challenge of considerable interest. After all, the reasonable average behaviour of enumerative methods (and the remarkable success of a nonenumerative method based on computations in an integer lattice [Lagarias & Odlyzko 1983]) to solve some NP-complete problems and the apparent impossibility to find such algorithms for other NP-complete problems, still defy theoretical explanation!

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