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The Vanishing Procyclicality of Labor Productivity

Jordi Gali and Thijs van Rens

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The Vanishing Procyclicality of Labor Productivity*

Jordi Galí

CREI, Universitat Pompeu Fabra and Barcelona GSE

jgali@crei.cat

Thijs van Rens

University of Warwick, Centre for Macroeconomics (LSE), IZA and CEPR

J.M.van-Rens@warwick.ac.uk

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Abstract

We document three changes in postwar US macroeconomic dynamics: *(i)* the procyclicality of labor productivity vanished, *(ii)* the relative volatility of employment rose, and *(iii)* the relative (and absolute) volatility of the real wage rose. We propose an explanation for all three changes that is based on a common source: the decline in labor market turnover, which reduced hiring frictions. We develop a simple model with hiring frictions, variable effort, and endogenous wage rigidities to illustrate the mechanisms underlying our explanation. We show that the decline in turnover may also have contributed to the observed decline in output volatility.

Keywords: labor hoarding, labor market turnover, hiring frictions, wage rigidities, effort choice

JEL classification: E24 E32

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1 Introduction

The nature of business cycle fluctuations changes over time. There is a host of evidence for changes in the dynamics of postwar US macroeconomic time series (Blanchard and Watson (1986), McConell and Pérez-Quirós (2000), Stock and Watson (2002), Hall (2007), Galí and Gambetti (2009)). The present paper documents and discusses three aspects of these changes.

- The correlation of labor productivity with output or labor input has declined, by some measures dramatically so.¹
- The volatility of labor input measures has increased (relative to that of output).²
- The volatility of real wage measures has increased, both in relative and absolute terms.³

All three of the above observations point towards a change in labor market dynamics. While each may be of independent interest and have potentially useful implications for our understanding of macro fluctuations, our goal in the present paper is to explore their possible connection. In particular, we seek to investigate the hypothesis that all three changes may be driven by the decline in labor market turnover in the US over this period, which reduced hiring frictions and allowed firms to adjust their labor force more easily in response to various kinds of shocks. In order to illustrate the mechanism behind this explanation, we develop a stylized model of fluctuations with labor market frictions and investigate how its predictions vary with the level of labor market turnover.

The decline in labor market turnover in the US is well documented, among others by Davis, Faberman, and Haltiwanger (2006), Davis (2008), Fallick and Fleischman (2004), Mukoyama and Şahin (2009), Faberman (2008), Davis, Faberman, Haltiwanger, Jarmin, and Miranda (2010), Davis, Faberman, and Haltiwanger (2012), Lazear and Spletzer (2012), Fujita (2011), Cairó and Cajner (2013), Cairó (2013) and Hyatt and Spletzer (2013), see Cairó and Cajner (2013) for an overview of this literature. In this paper, we explore the implications of this decline for business cycle dynamics. Our argument does not depend on the source of the decline in turnover and we leave this

¹As far as we know, Stroh (2009) was the first to provide evidence of a decline in the labor productivity-hours correlation. Gordon (2010), Barnichon (2010), Galí and Gambetti (2009), and Nucci and Riggi (2011), using different approaches, independently investigated the potential sources of that decline.

²To the best of our knowledge, Galí and Gambetti (2009) were the first to uncover that finding, but did not provide the kind of detailed statistical analysis found below. Independently, Hall (2007) offered some evidence on the size of the decline in employment in the most recent recessions that is consistent with our finding.

³As far as we know, this finding was not known previously, although it is reported in independent work by Gourio (2007) and Champagne and Kurmann (2013).

interesting question aside.⁴ We model the decrease in labor market turnover as a decline in the employment outflow probability or separation rate, which – as in much of the literature – is exogenous in our model. In response to the decline in turnover labor market frictions decrease *endogenously*. This effect arises because adjustment costs in employment in our model are convex, a relatively uncontroversial assumption. We show that the observed decline in turnover is sufficient to quantitatively generate the reduction in frictions needed to explain the changes in labor market dynamics.

The main intuition for our qualitative results is easy to describe. The idea goes back to a literature, starting with Oi (1962) and Solow (1964), which attributes the procyclicality of productivity to variations in effort, resulting in seemingly increasing returns to labor.⁵ Suppose that firms have two margins for adjusting their effective labor input: (observed) employment and (unobserved) effort, which we respectively denote (in logs) by n_t and e_t .⁶ Labor input (employment and effort) are transformed into output according to a standard production function,

$$y_t = (1 - \alpha)(n_t + e_t) + a_t$$

where a_t is log total factor productivity and α is a parameter measuring diminishing returns to labor.

Measured labor productivity, or output per person, is given by

$$y_t - n_t = -\alpha n_t + (1 - \alpha) e_t + a_t$$

Labor market frictions make it costly to adjust employment n_t . Since these adjustment costs are convex, frictions are higher when the average level of hiring is higher. Effort e_t provides an alternative margin of adjustment of labor input and is not subject to those frictions (or to a lesser degree). Thus, the larger the frictions, the less employment fluctuates and the more volatile fluctuations in effort. As a result, a decline in turnover reduces the average amount of hiring, reduces frictions, decreases the volatility of effort and therefore increases the relative volatility of employment with respect to output. The increased volatility of n_t also makes labor productivity less procyclical, and, in the presence of shocks other than shifts in technology, may even make productivity countercyclical, consistent with the evidence reported below.

⁴Recent papers have argued the decline is a result of decreased business volatility (Davis, Faberman, Haltiwanger, Jarmin, and Miranda (2010)), decreased job security (Fujita (2011)) or increased specificity of human capital (Cairó (2013)).

⁵Contributions include studies by Fair (1969), Fay and Medoff (1985), Hall (1988), Rotemberg and Summers (1990), Bernanke and Parkinson (1991), Shapiro (1993), Burnside, Eichenbaum, and Rebelo (1993), Bils and Cho (1994), Uhlig and Xu (1996), Basu (1996), Basu and Fernald (1997), Basu and Kimball (1997), Shea (1999), Gordon (2004), Wen (2004), Arias, Hansen, and Ohanian (2007), and Gordon (2010).

⁶To simplify the argument, we assume hours per worker are constant, consistent with the observation that in the US data most adjustments in total hours worked take place along the extensive margin.

In addition, as emphasized by Hall (2005), the presence of labor market frictions generates a non-degenerate bargaining set for the wage, i.e. a wedge between firms' and workers' reservation wages. Any wage within that bargaining set is consistent with labor market equilibrium. That feature makes room for wage rigidities. We model wages as rigid within the bargaining set, adjusting only when approaching its bounds. In our model, the reduction in labor market frictions resulting from the decline in labor market turnover narrows the bargaining set and therefore endogenously makes wages more sensitive to shocks, increasing the volatility of fluctuations in wages. If the rigidity is extended to the wages of newly hired workers, then the increased flexibility of wages may dampen the volatility of output and employment in response to shocks.⁷ That feature may help explain the observed decline in the volatility of those two variables in the recent US experience.⁸ Other authors have also argued that the Great Moderation may have been driven at least in part by increased wage flexibility (Gourio (2007), Champagne and Kurmann (2013), Nucci and Riggi (2011)). However, this paper is the first to show that such an increase in wage flexibility can arise endogenously from the decline in labor market turnover.⁹

The remainder of the paper is organized as follows. Section 2 documents the changes in the patterns of fluctuations in labor productivity, employment and wages. Section 3 develops the basic model. Section 4 describes the outcome of simulations of a calibrated version of the model, and discusses its consistency with the evidence. Section 5 concludes.

2 Changes in Labor Market Dynamics

We document three stylized facts regarding postwar changes in US economic fluctuations. The three changes that motivate our investigation pertain to the cyclical behavior of labor productivity, labor input and wages. None of the facts we report are new. However, and to the best of our knowledge, this paper is the first to provide a unified explanation for the three changes.

We use quarterly time series over the period 1948:1-2007:4 drawn from different sources (see below for a detailed description). To illustrate the changes in the different statistics considered, we split the sample period into two subperiods, pre-84 (1948:1-1983:4) and post-84 (1984:1-2007:4). The choice of the break date is fairly arbitrary, motivated by existing evidence on the timing of the Great Moderation (McConnell and

⁷This is clearly true for technology shocks. As argued in Blanchard and Galí (2007b), increased wage flexibility may also dampen the sensitivity of GDP and inflation to oil price shocks.

⁸A more flexible labor market does of course not make the economy immune to very large shocks like the recent financial crisis. Under our hypothesis, if the labor market were as rigid as it was in the early 80s, the current recession might have been even substantially more severe.

⁹Uren (2008) develops a model, in which a reduction in search frictions decreases output volatility. However, the reduction in frictions in that paper is exogenous and the mechanism, by which output volatility decreases (increased assortative matching), is completely different.

Pérez-Quirós (2000)), and we do not make any claims about the specific timing of the various changes in labor market dynamics.

Our evidence makes use of alternative measures of output and labor input. In all cases labor productivity is constructed as the ratio between the corresponding output and labor input measures. Most of the evidence uses output and hours in the private sector from the BLS Labor Productivity and Cost (LPC) program. We also use GDP as an economy-wide measure of output, with the corresponding labor input measures being total hours or employment. The time series for economy-wide hours is an unpublished series constructed by the BLS and used in Francis and Ramey (2009). The employment series is the usual one from the Current Employment Statistics (CES) establishment survey. In all cases we normalize the output and labor input measures by the size of the civilian noninstitutional population (16 years and older).

We apply three alternative transformations on the logarithms of all variables in order to render the original time series stationary. Our preferred transformation uses the bandpass (BP) filter to remove fluctuations with periodicities below 6 and above 32 quarters, as in Stock and Watson (1999). We also apply the fourth-difference (4D) operator, which is the transformation favored by Stock and Watson (2002) in their analysis of changes in output volatility, as well as the more common HP filter with smoothing parameter 1600.¹⁰

2.1 The Vanishing Procyclicality of Labor Productivity

Figure 1 shows the fluctuations at business cycle frequencies in labor productivity in the US over the postwar period. It is clear from the graph that in the earlier part of the sample, productivity was significantly below trend in each recession. However, in the later years this is no longer the case. When we calculate the correlation of productivity with output or employment, as in Figure 2, it is clear that there is a sharp drop in the cyclicity of productivity. The correlation of productivity with output, which used to be strongly positive, fell to a level close to zero, while the correlation of productivity with employment, which was zero or slightly positive in the earlier period of the sample, became negative.

These findings are formalized in Table 1, which reports the contemporaneous correlation between labor productivity and output and employment, for alternative transformations and time periods. In each case, we report the estimated correlation for the pre and post-84 subsamples, as well as the difference between those estimates. The standard errors, reported in brackets, are computed using the delta method.¹¹ We now turn to a

¹⁰We also showed the results are robust to using shorter sample periods, centered around the break date 1984. These estimates, which are suppressed for brevity, are available in the online appendix to this paper.

¹¹We use least squares (GMM) to estimate the second moments (variances and covariances) of each pair of variables, as well as the (asymptotic) variance-covariance matrix of this estimator. Then,

short discussion of the results in this Table.

2.1.1 Correlation with Output

Independently of the detrending procedure, the correlation of output per hour with output in the pre-84 period is high and significantly positive, with a point estimate around 0.60. In other words, in the early part of the sample labor productivity was clearly procyclical.

In the post-84 period, however, that pattern changed considerably. The estimates of the productivity-output correlation dropped to a value close to (and not significantly different from) zero. The difference with the corresponding pre-84 estimates is highly significant. Thus, on the basis of those estimates labor productivity has become an acyclical variable (with respect to output) over the past two decades.

When we use an employment-based measure of labor productivity, output per worker, the estimated correlations also drop substantially but remain significantly greater than zero in the post-84 period. This should not be surprising given that hours per worker are highly procyclical in both subperiods and that their volatility relative to employment-based labor productivity has increased considerably.¹²

2.1.2 Correlation with Labor Input

The right-hand side panels in Table 1 display several estimates of the correlation between labor productivity and labor input. The estimates for the pre-84 period are low, and in the case of output per hour, insignificantly different from zero. Thus, labor productivity was largely acyclical with respect to labor input in that subperiod. This near-zero correlation is consistent with the evidence reported in the early RBC literature, using data up to the mid 80s.¹³

As was the case when using output as the cyclical indicator, the estimated correlations between labor productivity and employment decline dramatically in the post-84 period. In fact these correlations become significantly negative for output per hour, with a point estimate ranging from -0.40 to -0.54 , depending on the filter. The change with

we calculate the standard errors for the standard deviations, the relative standard deviations and the correlation coefficient using the delta method.

¹²Letting n and h denote employment and total hours respectively, a straightforward algebraic manipulation yields the identity:

$$\rho(y - n, y) = \frac{\sigma_{y-h}}{\sigma_{y-n}} \rho(y - h, y) + \frac{\sigma_{h-n}}{\sigma_{y-n}} \rho(h - n, y)$$

Thus, even in the case of acyclical hours-based labor productivity, i.e. $\rho(y - h, y) \simeq 0$, we would expect $\rho(y - n, y)$ to remain positive if hours per worker are procyclical, i.e. $\rho(h - n, y) > 0$.

¹³Christiano and Eichenbaum (1992) used data up to 1983:4 (which coincides with the cut-off date for our first subperiod), but starting in 1955:4. Their estimates of the correlation between labor productivity and hours were -0.20 when using household data and 0.16 using establishment data.

respect to the pre-84 period is highly significant. In other words, labor productivity in the past two decades has become strongly countercyclical with respect to labor input.

2.2 The Rising Relative Volatility of Labor Input

The left-hand panel of Table 2 displays the standard deviation of several measures of labor input in the pre and post-84 periods, as well as the ratio between the two. The variables considered include employment in the private sector, hours in the private sector (employment times hours per worker) and economy-wide hours. The decline in the volatility of hours, like that of other major macro variables, is seen to be large and highly significant, with the standard deviation falling between 35% and 49% and always significantly so.

A more interesting piece of evidence is the change in the *relative* volatility of labor input, measured as the ratio of the standard deviation of labor input to the standard deviation of output. These estimates are presented in the right-hand panel of Table 2. Without exception, all labor input measures have experienced an increase in their relative volatility in the post versus pre-84 period. In other words, the decline in the variability of labor input has been less pronounced than that of output. The increase in the relative volatility of hours worked ranges from 30% to 48% in the private sector and from 7% to 30% in the total economy. The corresponding increase for employment is slightly smaller, ranging from 23% to 43% in the private sector, but is still statistically significant.

The previous evidence points to a rise in the elasticity of labor input with respect to output. Put differently, firms appear to have relied increasingly on labor input adjustments in order to meet their changes in output.

2.3 The Rising Volatility of Wages

Next we turn our attention to the volatility of (real) wages, both in absolute and relative terms. We consider four different wage measures. The first three are constructed as real compensation per hour. The first difference is in the measure of compensation, which is measured either from the national income and product accounts (NIPA) or as earnings from the CES establishment survey. The second difference is in the measure of hours, which refers to the private sector or to the total economy.¹⁴ The fourth measure is usual hourly earnings (or usual weekly earnings divided by usual weekly hours) from the Current Population Statistics (CPS). For all measures, compensation or earnings are deflated using the compensation deflator from the LPC, but the results are robust to deflating with the consumer price index (CPI-U) as we show below.

¹⁴Our baseline measure uses compensation from the NIPA and hours in the private sector and corresponds to compensation per hour from the BLS labor productivity and cost program (LPC).

2.3.1 Average Wages

The left-hand panel of Table 3 displays the standard deviation for each wage measure for different detrending procedures. Our statistics uncover a surprising finding: despite the general decline in macro volatility associated with the Great Moderation, the volatility of several wage measures has *increased* in absolute terms. The estimated increase the standard deviation is fairly large, between 10 and 42% and mostly significant for the bandpass filtered NIPA-based wage measures. Using fourth differences, the increase is much smaller and no longer significant and by some periods there seems to have been a (small) decrease in wage volatility.

Using earnings per hour from the CES, however, there seems to be a large and highly significant reduction in wage volatility.¹⁵ One difference between the two measures is that the NIPA compensation measure includes non-wage payments and, in particular, employee stock options. Mehran and Tracy (2001) have argued that since these options are recorded when they are realized rather than when they are handed out to employees, the NIPA measure gives a misleading picture of the evolution and volatility of compensation in the 90s. However, using the CPS measure of usual hourly earnings, presented in Table 4, which includes non-wage compensation but not stock options, we again observe a fairly large increase in the volatility of wages.¹⁶ Given the short time series available for these data, it is remarkable that the increase in volatility is (borderline) significant for the bandpass and HP filtered series. Using fourth differences, wage volatility seems roughly constant.

Our finding that wage volatility increased or at least did not decrease around the time of the Great Moderation, although with a caveat, is consistent with the results in Champagne and Kurmann (2013), who also use the CPS to show that the increase in wage volatility is not driven by compositional changes in the labor force. To the best of our knowledge, this result was not previously known.¹⁷

An immediate implication of the previous finding, and the one that we want to emphasize here, is the possibly very large increase in the relative volatility of wages

¹⁵This finding is not driven by the fact that the CES earnings measure is only available after 1965. If we restrict the sample for the NIPA based wage measures to the 1965-2004, the volatility statistics for these measures look very similar those in the table.

¹⁶We use data from the CPS outgoing rotation groups. Since these data are available only from 1979 onwards, we compare the volatility over the 1980-1984 period (allowing for fourth differences) with that of the 1985-2005 period. For comparison, the first panel of Table 4 presents the volatility of our baseline measure for compensation per hour for this period. The second panel presents comparable statistics from the CPS series. Because the CPS wage series is based on a fairly small cross-section of workers, there is substantial measurement error in these series. Therefore, the standard deviations of the fourth differenced data are biased upward, see Haefke, Sonntag, and van Rens (2013) for details. There is no reason however, why the ratio of the standard deviations before and after 1984 would be biased. In addition, the bandpass filtered data, which do not include the high frequencies induced by the measurement error, are not subject to this bias.

¹⁷Stock and Watson (2002) uncover breaks in the volatility of a long list of macro variables, but they do not provide evidence for any wage measure.

with respect to output or labor input, as shown in the right-hand panels of Tables 3 and 4. The relative volatility of wages with respect to output more than doubled for the NIPA-based measures and for the CPS wage. We interpret this evidence as being consistent with a decline in the significance of real wage rigidities.¹⁸

2.3.2 Wages of Newly Hired Workers

In a frictional labor market, the average wage is not allocative, since the frictions drive a wedge between the reservation wages of firms and workers. Therefore, in order to assess the implications of increased wage flexibility for other labor market variables, we also consider the volatility of the wage of newly hired workers as suggested by Haefke, Sonntag, and van Rens (2013) and Pissarides (2009).

Table 4 presents volatility statistics for the wage of new hires, constructed from the CPS as in Haefke, Sonntag, and van Rens (2013).¹⁹ The first thing to notice is that the wage of newly hired workers is much more volatile than the average wage in the entire labor force. This is consistent with the results in Haefke, Sonntag, and van Rens (2013), who argue that, in the post-84 period, wages of newly hired workers are perfectly flexible, in the sense that they respond one-to-one to changes in labor productivity. Here, we focus on the change in the volatility of wages over time.

The absolute volatility of the wage of newly hired workers, unlike the average wage, decreased substantially and significantly between the pre and post-84 periods. As a result, the increase in the relative volatility with respect to output is much smaller for new hires, ranging between 3% and 69%, depending on whether we use the mean or median wage and on the filter used. Although the increase in the relative volatility of the wage of newly hired workers is much less pronounced, there is some evidence that wages fluctuated more between recessions and booms also for this group of workers. This finding is consistent with the evidence presented in Haefke, Sonntag, and van Rens (2013, section 3.4) and points towards a decrease in wage rigidity that may be important for employment fluctuations.

2.3.3 Correlation of Wages with Productivity and the Labor Wedge

In macroeconomic models with a perfectly competitive labor market and a standard production function, wages are proportional to labor productivity.²⁰ Our evidence makes

¹⁸Blanchard and Galí (2007b) argue that a reduction in the rigidity of real wages is needed in order to account for the simultaneous decline in inflation and output volatility, in the face of oil price shocks of a similar magnitude.

¹⁹But unlike in that paper, we do not correct fluctuations in the CPS wage series for changes in the composition of the labor force by demographic characteristics, education level and experience for comparability with the other wage measures. Doing so however, makes very little difference for the conclusions presented here.

²⁰With competitive labor markets, the wage equals the marginal product of labor, and with a Cobb-Douglas production function, the marginal and average product of labor are proportional to each other.

clear that the extent to which this is true in the data depends very much on the period one looks at. From Table 3, we see that the relative standard deviation of wages with respect to labor productivity was about 0.3 – 0.7 in the pre-84 period. In the post-84 period, this relative standard deviation more than doubled to about 1.2 – 1.3. This is consistent with the evidence in Haefke, Sonntag, and van Rens (2013, section 3.4), who show that the elasticity of wages with respect to productivity in 1984 increases from 0.2 to 0.4 for all workers, and from 0.3 to 0.8 for newly hired workers.

The increased responsiveness of wages to movements in productivity has implications for the volatility of the labor wedge. However, the direction of the effect is ambiguous and depends on the relative importance of technology and non-technology shocks. If business cycles are driven by technology shocks, then more flexible wages dampen the volatility of employment fluctuations and make the labor wedge less countercyclical. If business cycles are primarily due to non-technology shocks, then more flexible wages amplify the volatility of employment and the labor wedge. In the data, the labor wedge became more volatile over time, see Shimer (2010, Figure 1.2), suggesting that the latter effects dominates the former.

2.4 Evidence for Europe

Although in this paper we focus on the US, it is worth exploring whether the basic patterns hold for other countries as well. In particular, given the importance of labor market frictions for our story, one might expect similar or even stronger changes in labor market dynamics in European countries, where labor market frictions are generally believed to be more important. For many countries, data are not available for our sample period. However, Ohanian and Raffo (2012) collected data on output, employment and hours worked from the OECD Economic Outlook database and national statistics offices, for some countries starting in 1960. Table 5 reports the cyclicalities of labor productivity and the relative volatility of labor input for the four major European economies using these data.²¹ For comparison, we also report the statistics for the US over the same period. The changes in labor market dynamics look very similar for these five countries. Unfortunately, lack of data does not allow us to investigate whether the European countries, like in the US, experienced a decline in labor market turnover that may have driven these changes.²²

This well-known argument, which does not rely on business cycles being driven by productivity shocks, was used recently by Rogerson and Shimer (2010, section 1.3.2).

²¹To the best of our knowledge, quarterly data for employee compensation are not available for these countries over this period. Therefore, we are not able to explore whether the volatility of wages rose in these countries.

²²International time series data for worker flows for these countries were calculated by Elsby, Hobijn, and Şahin (2013), but for most countries (including Germany and Italy) these data start in 1983. For France and the UK the data start in 1975 and 1969 respectively and do not show a secular decline in the employment exit probability over this period.

2.5 Conclusions

Summarizing, we showed that labor productivity in the US became less procyclical or acyclical with respect to output, and countercyclical with respect to employment. In addition, the relative volatility of both employment and wages increased. For completeness, we also report that the relative volatility of labor productivity increased, and the correlation between employment and output decreased slightly.²³ These changes in business cycle dynamics, which seem to have happened in European countries as well, coincided with the reduction in volatility in output and most other macroeconomic aggregates, the so called Great Moderation.

From the vantage point of the early 80s – the period when the seminal contributions to RBC theory were written – the procyclicality of labor productivity was a well established empirical fact. This observation lent support to business cycle theories that assigned a central role to technology shocks as a source of fluctuations. The relative volatility of labor input in these models was lower than in the data, which posed one of the main challenges for these models, see King and Rebelo (1999) or Hall (1997). From today’s perspective, things look distinctly worse for real business cycle theory. The relative volatility of labor input increased even further and productivity is barely procyclical anymore. A possible reason is that over time non-technology shocks became more important drivers of business cycles, as Barnichon (2010) argues. However, this does not explain why similar changes in dynamics are observed also when conditioning on particular shocks, as in Galí and Gambetti (2009).

In the remainder of this paper, we explore whether the observed changes in business cycle dynamics may be explained by a structural change in the labor market. We show that the vanishing procyclicality of labor productivity and the increasing relative volatility of employment, can be explained by the decline in labor market turnover, which resulted in a reduction in hiring costs. If wages are set competitively, then a reduction in hiring costs, which makes employment faster to adjust, should make wages relatively less volatile. Although the relative standard deviation of wages decreased in some datasets, we argued above that the majority of the evidence seems to point towards a large *increase* in the relative volatility of wages. Therefore, we explore how the model can be extended with endogenous wage rigidities in order to generate this result.

²³These observations, which are reported in the online appendix to this paper, are completely determined by the statistics already reported and do not contain independent information. We emphasized the statistics that we consider easiest to interpret.

3 A Model of Fluctuations with Labor Market Frictions and Endogenous Effort

Having documented in some detail the changing patterns of labor productivity, labor input, and wages, we turn to possible explanations. More specifically, and as anticipated in the introduction, we explore the hypothesis that all three observed changes documented above may have, at least partly, been caused by the same secular change: a decline in labor market turnover.

To formalize this explanation, we develop a model of fluctuations with labor market frictions, modelled as adjustment costs in employment (hiring costs). The crucial element in this model is an endogenous effort choice, which provides an intensive margin for labor adjustment that is not subject to the adjustment costs. Since the purpose of the model is to illustrate the main mechanisms at work, we keep the model as simple as possible in dimensions that are likely to be orthogonal to the factors emphasized by our analysis. Thus, we abstract from endogenous capital accumulation, trade in goods and assets with the rest of the world, and imperfections in the goods and financial markets. We also ignore any kind of monetary frictions, even though we recognize that these, in conjunction with changes in the conduct of monetary policy in the Volcker-Greenspan years, may have played an important role in accounting for the decline in macro volatility.²⁴

3.1 Households

Households are infinitely-lived and consist of a continuum of identical members represented by the unit interval. The household is the relevant decision unit for choices about consumption and labor supply. Each household member's utility function is additively separable in consumption and leisure, and the household assigns equal consumption C_t to all members in order to share consumption risk within the household. Thus, the household's objective function is given by,²⁵

$$E_0 \sum_{t=0}^{\infty} \beta^t \left[\frac{Z_t C_t^{1-\eta}}{1-\eta} - \gamma L_t \right] \quad (1)$$

where $\beta \in (0, 1)$ is the discount factor, $\eta \in [0, 1]$ is the inverse of the intertemporal elasticity of substitution, $\gamma > 0$ can be interpreted as a fixed cost of working and Z_t is a preference shock. The second term in the period utility function is disutility from effective labor supply L_t , which depends on the fraction N_t of household members that

²⁴See, e.g. Clarida, Galí, and Gertler (2000) for a discussion of the possible role of monetary policy in the Great Moderation.

²⁵We assume utility is linear in effective labor for simplicity. The implication that the Frisch elasticity of labor supply is infinity is of course counterfactual, but our results are very similar if we assume a Frisch elasticity of 0.25, as advocated by Chetty, Guren, Manoli, and Weber (2012).

are employed, as well as on the amount of effort \mathcal{E}_{it} exerted by each employed household member i . Formally,

$$L_t = \int_0^{N_t} \frac{1 + \zeta \mathcal{E}_{it}^{1+\phi}}{1 + \zeta} di = \frac{1 + \zeta \mathcal{E}_t^{1+\phi}}{1 + \zeta} N_t \quad (2)$$

where the second equality imposes the equilibrium condition that all working household members exert the same level of effort, $\mathcal{E}_{it} = \mathcal{E}_t$ for all i . The parameter $\zeta \geq 0$ measures the importance of effort for the disutility of working, and the elasticity parameter $\phi \geq 0$ determines the degree of increasing marginal disutility from exerting effort. For simplicity we assume a constant workweek, thus restricting the intensive margin of labor input adjustment to changes in effort.

The household maximizes its objective function above subject to the sequence of budget constraints,

$$C_t = \int_0^{N_t} W_{it} di + \Pi_t \quad (3)$$

where Π_t represents firms' profits, which are paid out to households in the form of lump-sum dividends, and W_{it} are wages accruing to employed household member i . The household takes into account the effect of its decisions on the level of effort exerted by its members.

3.2 Firms

Firms produce a homogenous consumption good using a production technology that uses labor and effort as inputs,

$$Y_t = A_t \left(\int_0^{N_t} \mathcal{E}_{it} di \right)^{1-\alpha} = A_t \left(\mathcal{E}_t N_t \right)^{1-\alpha} \quad (4)$$

where Y_t is output, \mathcal{E}_{it} is effort exerted by worker i , $\alpha \in (0, 1)$ is a parameter that measures diminishing returns to total labor input in production, $\in [0, 1]$ measures additional diminishing returns to effort, and A_t is a technology shock common to all firms. Since all firms are identical, we normalize the number of firms to the unit interval, so that Y_t and N_t denote output and employment of each firm as well as aggregate output and employment in the economy. The second equality imposes the equilibrium condition that all workers in a firm exert the same level of effort, $\mathcal{E}_{it} = \mathcal{E}_t$ for all i .

Firms choose how many workers to hire H_t in order to maximize the expected discounted value of profits,

$$E_0 \sum_{t=0}^{\infty} Q_{0,t} [Y_t - W_t N_t - g(H_t)] \quad (5)$$

where the function $g(\cdot)$, with $g' > 0$ and $g'' > 0$, represents the costs (in terms of output)

of hiring new workers, subject to a law of motion for employment implied by the labor market frictions,

$$N_t = (1 - \delta) N_{t-1} + H_t \quad (6)$$

where δ is the gross separation rate (employment exit probability) and $Q_{0,t}$ is the stochastic discount factor for future profits. The stochastic discount factor is defined recursively as $Q_{0,t} \equiv Q_{0,1} Q_{1,2} \dots Q_{t-1,t}$, where

$$Q_{t,t+1} \equiv \beta \frac{Z_{t+1}}{Z_t} \left(\frac{C_t}{C_{t+1}} \right)^\eta \quad (7)$$

measures the marginal rate of substitution between two subsequent periods. Like the household, the firm takes into account the effect of its decisions on the level of effort exerted by its workers.

3.3 Effort Choice and Job Creation

The household and the firm jointly decide the wage and the level of effort that the worker will put into the job. In equilibrium, the effort level of all workers is set efficiently, maximizing the total surplus generated by each match.²⁶ This efficient effort level, in each period and for each worker, equates the cost of exerting more effort, higher disutility to the household, to the benefit, higher production and therefore profits for the firm.

Consider a worker i , who is a member of household h and is employed in firm j . The marginal disutility to the household from that worker exerting more effort, expressed in terms of consumption, is obtained from equation (2) for total effective labor supply and equals:

$$\frac{\gamma C_{ht}^\eta}{Z_t} \frac{\partial L_{ht}}{\partial \mathcal{E}_{it}} = \frac{(1 + \phi) \zeta}{1 + \zeta} \frac{\gamma C_{ht}^\eta \mathcal{E}_{it}^\phi}{Z_t} di \quad (8)$$

The marginal product of that additional effort to the firm is found from production function (4):

$$\frac{\partial Y_{jt}}{\partial \mathcal{E}_{it}} = (1 - \alpha) A_t \left(\int_0^{N_{jt}} \mathcal{E}_{vt} dv \right)^{-\alpha} \mathcal{E}_{it}^{-(1-\alpha)} di \quad (9)$$

In equilibrium, the marginal disutility from effort must equal its marginal product for all workers i . Also, because all firms and all households are identical, it must be that $C_{ht} = C_t$ and $N_{jt} = N_t$ in equilibrium. Therefore, it follows that all workers exert the same level of effort in equilibrium, $\mathcal{E}_{it} = \mathcal{E}_t$ for all i . Imposing this property, we obtain the following equilibrium condition for effort,

$$\mathcal{E}_t = \left[\frac{(1 - \alpha) (1 + \zeta)}{(1 + \phi) \zeta} \frac{Z_t}{\gamma C_t^\eta} A_t N_t^{-\alpha} \right]^{\frac{1}{1 + \phi - (1 - \alpha)}} \quad (10)$$

²⁶ Suppose not. Then, household and firm could agree on a different effort level that increases total match surplus, and a modified surplus sharing rule (wage) that would make both parties better off.

or, using production function (4) to simplify:

$$\mathcal{E}_t^{1+\phi} = \frac{1+\zeta}{1+\phi} \frac{Z_t}{\zeta} \frac{(1-\alpha) Y_t}{\gamma C_t^\eta N_t} \quad (11)$$

When considering whether to hire a worker, firms take into account the impact of the resulting increase in employment on the effort level exerted by their workers. Thus, the marginal product of a new hire is given by,²⁷

$$\frac{dY_{jt}}{dN_{jt}} = \frac{\partial Y_{jt}}{\partial N_{jt}} + \frac{\partial Y_{jt}}{\partial \mathcal{E}_{jt}} \frac{\partial \mathcal{E}_{jt}}{\partial N_{jt}} = (1 - \Psi_F) \frac{(1-\alpha) Y_t}{N_t} \quad (12)$$

where $\Psi_F = \frac{\alpha\psi}{1+\phi-(1-\alpha)}$ measures the additional (negative) effect from a new hire on output that comes from the endogenous response of the effort level in the firm.

Maximizing the expected net present value of profits (5), where output is given by production function (4) and the stochastic discount factor by (7), subject to the law of motion for employment implied by the matching technology (6) and the equilibrium condition for effort (11), gives rise to the following first order condition,

$$g'(H_t) = S_t^F \quad (13)$$

where S_t^F is the marginal value to the firm of having an additional worker in period t , which is given by,

$$S_t^F = (1 - \Psi_F) \frac{(1-\alpha) Y_t}{N_t} - W_t + (1 - \delta) E_t [Q_{t,t+1} S_{t+1}^F] \quad (14)$$

$$= E_t \sum_{s=0}^{\infty} (1 - \delta)^s Q_{t,t+s} \left[(1 - \Psi_F) \frac{(1-\alpha) Y_{t+s}}{N_{t+s}} - W_{t+s} \right] \quad (15)$$

where the second equality follows from iterating forward (and defining $Q_{t,t} = 1$). This is a job creation equation, which states that the marginal costs of hiring a new worker $g'(H_t)$, must equal the expected net present value of marginal profits (additional output minus the wage) of the filled job, S_t^F .

3.4 The Bargaining Set

Employment relationships generate a strictly positive surplus. This property of our model comes from the assumption that wages and effort levels are determined after

²⁷With a slight abuse of notation, \mathcal{E}_{jt} denotes the effort level exerted by all workers (from different households) in a particular firm j . Firm j considers employing N_{jt} workers, given that all other firms employ the equilibrium number of workers N_t . Because there are infinitely many firms, firm j 's decision to employ $N_{jt} \neq N_t$ workers does not affect the fraction of household h 's members that are employed, so that by the assumption of perfect risk-sharing within the household, the consumption of workers in firm j , $C_{ht} = C_t$, is not affected. Therefore, the relation between effort and employment that the firm faces if all other firms (and all households) play equilibrium strategies, is given by equation (10), keeping C_t fixed. See appendix A for details on the derivation of equation (12).

employment adjustment costs are sunk: if firm and worker cannot agree to continue their relationship, then the firm has to pay the hiring costs again in order to find another worker to match with. We make this timing assumption in order to generate wage setting under bilateral monopoly, as in a search and matching model, which we believe to be a realistic feature of the labor market.²⁸ Firms and households bargain over the wage as a way to share the match surplus. These negotiations are limited only by the outside option of each party. The lower bound of the bargaining set is given by the reservation wage of the household, the wage offer at which the household is indifferent between accepting the offer and looking for another job. Similarly, the upper bound of the bargaining set is the reservation wage of the firm, the wage offer that makes the firm indifferent between accepting the offer and hiring a different worker. Within these bounds, any wage is consistent with equilibrium, see Hall (2005). Clearly, the bounds of the bargaining set are endogenous variables, which we now derive before introducing an equilibrium selection rule for the wage within the bargaining set.

The part of the match surplus that accrues to the firm S_t^F , as a function of the wage, is given by equation (14). In order to derive a similar expression for the household's part of the surplus S_t^H , we must first calculate the marginal disutility to the household of having one additional employed member, taking into account the endogenous response of effort. This marginal disutility of employment, expressed in terms of consumption, is given by,²⁹

$$\frac{\gamma C_t^\eta}{Z_t} \frac{dL_{ht}}{dN_{ht}} = \frac{1}{1+\zeta} \frac{\gamma C_t^\eta}{Z_t} \left(1 + \zeta \frac{(1+\phi) \Psi_H}{\mathcal{E}_t^{1+\phi}} \right) = \frac{1}{1+\zeta} \frac{\gamma C_t^\eta}{Z_t} + \Psi_H \frac{(1-\alpha) Y_t}{N_t} \quad (16)$$

where the second equality follows from substituting equation (11), and where $\Psi_H = \frac{(1-\eta)(1+\phi)-}{1+\phi}$ captures the effect on utility of one more employed member in the household through the endogenous response of effort. Using this expression, we can take a derivative of the household's objective function (1) with respect to N_t and divide by the marginal utility of consumption, to obtain the following expression for S_t^H .

$$S_t^H = W_t - \frac{1}{1+\zeta} \frac{\gamma C_t^\eta}{Z_t} - \Psi_H \frac{(1-\alpha) Y_t}{N_t} + (1-\delta) E_t [Q_{t,t+1} S_{t+1}^H] \quad (17)$$

The value to the household of having one more employed worker, equals the wage minus

²⁸Specifically, the within-period timing we assume is the following. First, aggregate shocks realize and a randomly chosen fraction δ of employed workers is separated from their jobs. Second, firms that want to hire pay employment adjustment costs $g(H_t)$ and are randomly matched with H_t non-employed workers. Third, firm and worker bilaterally and with full commitment decide on the effort the worker will put into the job and the wage she will be paid for doing it. If a firm and a worker cannot agree, the worker is placed back into the unemployment pool and the firm pays $g'(H_t)$ in order to get another random draw from that pool. Since all unemployed workers are identical, this never happens in equilibrium. When a firm and worker do reach an agreement, the worker is hired and added to the pool of employed workers. Finally, production, consumption and utility are realized.

²⁹The derivation of this expression is similar to that of equation (12), see appendix A for details.

the disutility expressed in terms of consumption, plus the expected value of still having that worker next period, which is discounted by the probability that the worker is still employed next period.

The upper bound of the bargaining set W_t^{UB} is the highest wage such that $S_t^F \geq 0$, whereas the lower bound W_t^{LB} is the lowest wage such that $S_t^H \geq 0$. Using equations (14) and (17), we get $S_t^F = W_t^{UB} - W_t$ and $S_t^H = W_t - W_t^{LB}$. Substituting back into equations (13), (14) and (17), we can explicitly write the equilibrium of the model in terms of the wage and the bounds of the bargaining set.

$$g'(H_t) = W_t^{UB} - W_t \quad (18)$$

$$W_t^{UB} = (1 - \Psi_F) \frac{(1 - \alpha) Y_t}{N_t} + (1 - \delta) E_t [Q_{t,t+1} (W_{t+1}^{UB} - W_{t+1})] \quad (19)$$

$$W_t^{LB} = \frac{1}{1 + \zeta} \frac{\gamma C_t^\eta}{Z_t} + \Psi_H \frac{(1 - \alpha) Y_t}{N_t} + (1 - \delta) E_t [Q_{t,t+1} (W_{t+1}^{LB} - W_{t+1})] \quad (20)$$

Everything else equal, the more rigid is the wage in response to technology or preference shocks that shift the bounds of the bargaining set, the more volatile is hiring H_t and therefore employment N_t in response to those shocks, see equation (18). We now turn to various possibilities for how wages are determined within the bargaining set.

3.5 Wage Determination

One possible criterion for wage determination that we can interpret as flexible wages in a model with a frictional labor market, is period-by-period Nash bargaining. Nash bargaining assumes that the wage is set such that the total surplus from the match is split in equal proportions between household and firm.³⁰ It is straightforward to see that in our framework, this assumption implies that the wage is the average of the lower and upper bounds of the bargaining set, $S_t^H = \frac{1}{2} (S_t^H + S_t^F) = \frac{1}{2} (W_t^{UB} - W_t^{LB})$. Then,

$$W_t^* = \frac{1}{2} (W_t^{UB} + W_t^{LB}) \quad (21)$$

where W_t^* denotes the Nash bargained wage.

Shimer (2005) and Hall (2005), among others, have argued that period-by-period Nash bargaining generates too volatile a wage in equilibrium, relative to what is observed in the data. As discussed below, in our model period-by-period Nash bargaining leads to fluctuations in the (log) wage of the same amplitude as labor productivity, and perfectly correlated with the latter. This is at odds with the data, where wages are about half as volatile as labor productivity in the pre-84 period, with the correlation between the two variables much smaller than one. Both the relative volatility of wages and their

³⁰The symmetry assumption is not crucial, but simplifies the solution of the model substantially. The online appendix presents shows that our results are virtually unchanged for bargaining power well below and above 0.5.

correlation with labor productivity increases significantly in the post-84 period. This motivates the introduction of a wage setting mechanism that departs from period-by-period Nash bargaining.³¹ We use the following wage determination process,

$$W_t = R_{t-1}W_{t-1} + (1 - R_{t-1})W_t^* \quad (22)$$

where R_t measures the degree of wage rigidity, which is endogenous.

Only wages within the bargaining set are renegotiation-proof and therefore consistent with equilibrium, see Barro (1977). The maximum degree of wage stickiness that respects this condition, is for the wage to remain fixed inside the bargaining set, but to be adjusted when it hits either bound. This is the wage determination mechanism in Thomas and Worrall (1988), MacLeod and Malcomson (1993) and Hall (2003). We use a convex version of this wage rule, in order to be able to solve the model using perturbation methods, and assume the following reduced-form equation for wage rigidity,

$$R_t = \bar{R} \left(1 - \frac{W_t - W_t^*}{\frac{1}{2}(W_t^{UB} - W_t^{LB})} \right)^{2\rho} \quad (23)$$

where $\rho \in \mathbb{N}_0^+$. This wage rule captures the idea that the wage is more likely to adjust when it is closer to the bounds of the bargaining set. The parameter ρ captures the degree of non-linearity in this relation. For $\rho = 0$, $R_t = 0$ and $W_t = W_t^*$, i.e. wages are flexible. For $\rho \in \mathbb{N}^+$, the degree of wage rigidity is endogenous, with wages being perfectly flexible at the upper or lower bound of the bargaining set and most rigid at the Nash-bargained wage W_t^* . As ρ becomes larger, wages are rigid in a larger part of the bargaining set. The limiting case for $\rho \rightarrow \infty$ and $\bar{R} = 1$ captures the case where the wage is fixed within the bargaining set but adjusts when it has to in order to avoid inefficient match destruction as in Hall (2003). We consider a flexible wage regime with $\rho = 0$ and regime with endogenous wage rigidity, $\rho \in \mathbb{N}^+$ and $\bar{R} = 1$.

The crucial insight for our purposes is that with this type of wage rule, the degree of wage rigidity depends endogenously on the size of the frictions. If frictions decrease, the width of the bargaining set decreases as well, so that there is less room for wage rigidity. Notice also that this type of wage rigidity can never lead to inefficient match destruction.

3.6 Equilibrium

We conclude the description of the model by listing the conditions that characterize the equilibrium. Vacancy posting decisions by firms are summarized by the job creation

³¹This is consistent with the recent trend in the literature, see e.g. Hall (2005), Blanchard and Gali (2007a), Hall and Milgrom (2008), Christoffel and Kuester (2008), Gertler and Trigari (2009), and Shimer (2010).

equation (18).

$$g'(H_t) = W_t^{UB} - W_t \quad (24)$$

The equilibrium level of effort is determined by efficiency condition (11),

$$\mathcal{E}_t^{1+\phi} = \frac{1+\zeta}{1+\phi} \frac{Z_t}{\zeta} \frac{(1-\alpha)Y_t}{\gamma C_t^\eta N_t} \quad (25)$$

and wage negotiations are described by equations (22) and (23) for the equilibrium selection rule, and stochastic difference equations for the upper and lower bounds of the bargaining set (19) and (20).

$$W_t = R_{t-1}W_{t-1} + (1 - R_{t-1}) \frac{1}{2} (W_t^{UB} + W_t^{LB}) \quad (26)$$

$$R_t = \bar{R} \left(1 - \frac{W_t - \frac{1}{2} (W_t^{UB} + W_t^{LB})}{\frac{1}{2} (W_t^{UB} - W_t^{LB})} \right)^{2\rho} \quad (27)$$

$$W_t^{UB} = (1 - \Psi_F) \frac{(1-\alpha)Y_t}{N_t} + (1-\delta) E_t [Q_{t,t+1} (W_{t+1}^{UB} - W_{t+1})] \quad (28)$$

$$W_t^{LB} = \frac{1}{1+\zeta} \frac{\gamma C_t^\eta}{Z_t} + \Psi_H \frac{(1-\alpha)Y_t}{N_t} + (1-\delta) E_t [Q_{t,t+1} (W_{t+1}^{LB} - W_{t+1})] \quad (29)$$

Employment evolves according to its law of motion (6).

$$N_t = (1 - \delta) N_{t-1} + H_t \quad (30)$$

Finally, goods market clearing requires that consumption equals output minus hiring costs.

$$C_t = Y_t - g(H_t) \quad (31)$$

Output is defined as in production function (4), and the stochastic discount factor as the marginal rate of intertemporal substitution (7).

$$Y_t = A_t (\mathcal{E}_t N_t)^{1-\alpha} \quad (32)$$

$$Q_{t,t+1} = \beta \frac{Z_{t+1}}{Z_t} \left(\frac{C_t}{C_{t+1}} \right)^\eta \quad (33)$$

and the parameters $\Psi_F = \frac{\alpha\psi}{1+\phi-(1-\alpha)}$ and $\Psi_H = \frac{(1-\eta)(1+\phi)-}{1+\phi-(1-\alpha)}$ are functions of the structural parameters. In total, we have 8 equations in the endogenous variables H_t , \mathcal{E}_t , W_t , R_t , W_t^{UB} , W_t^{LB} , N_t and C_t , or 10 equations including the definitions for Y_t and $Q_{t,t+1}$.

Without an endogenous effort choice ($\phi = 0$ so that effort is not useful in production, $\Psi_F = \Psi_H = 0$, and $\mathcal{E}_t = 0$ for all t in equilibrium), and with flexible wages ($\bar{R} = 0$ so

that $R_t = 0$ for all t), the model reduces to a standard RBC model with labor market frictions. However, unlike in the standard model, fluctuations in our model are driven by technology shocks as well as non-technology shocks or preference shocks. The two driving forces of fluctuations, log total factor productivity $a_t \equiv \log A_t$ and log preferences over consumption $z_t \equiv \log Z_t$ follow stationary $AR(1)$ processes,

$$a_t = \rho_a a_{t-1} + \varepsilon_t^a \quad (34)$$

$$z_t = \rho_z z_{t-1} + \varepsilon_t^z \quad (35)$$

where ε_t^a and ε_t^z are independent white noise processes with variances given by σ_a^2 and σ_z^2 respectively.

4 Implications of the Decline in Labor Market Turnover

We now proceed to use this model to analyze the possible role of the decline in labor market turnover in generating the observed changes in the cyclical patterns of output, labor input, productivity, and wages. We start by looking at a version of the model with zero separation rate. In this case, small employment adjustments are costless so that the labor market effectively becomes frictionless. This model provides a useful benchmark that we can solve for in closed form. Then, we rely on numerical methods to simulate the model for different values of the separation rate.

4.1 Zero Labor Market Turnover

Consider the limiting case of an economy with a separation rate of zero. In this case, hiring costs are zero in steady state. As a result, the width of the bargaining set collapses to zero, and the job creation equation (24) and the wage block of the model, equations (26), (27), (28) and (29), imply

$$W_t = W_t^{UB} = W_t^{LB} = (1 - \Psi_F) \frac{(1 - \alpha) Y_t}{N_t} = \frac{1}{1 + \zeta} \frac{\gamma C_t^\eta}{Z_t} + \Psi_H \frac{(1 - \alpha) Y_t}{N_t} \quad (36)$$

for all t . Employment becomes a choice variable, so that its law of motion (30) is dropped from the system and employment is instead determined by the static condition (36).

$$N_t = (1 - \alpha) (1 - \Psi_F - \Psi_H) \frac{(1 + \zeta) Z_t Y_t}{\gamma C_t^\eta} \quad (37)$$

Substituting into the equilibrium condition for effort (25), we obtain

$$\mathcal{E}_t^{1+\phi} = \frac{1}{1 + \phi} \frac{1}{\zeta} \frac{1}{1 - \Psi_F - \Psi_H} \quad (38)$$

implying an effort level that is invariant to fluctuations in the model's driving forces. Since effort has stronger diminishing returns in production and stronger increasing marginal disutility than employment, this intensive margin of adjustment is never used if the extensive margin is not subject to frictions.

Without hiring costs, the aggregate resource constraint (31) reduces to $C_t = Y_t$. Combining the resource constraint and equations (37) and (38) with the production function (32), we can derive closed-form expressions for equilibrium employment, output, wages and labor productivity. Using lower-case letters to denote the natural logarithms of the original variables, ignoring constant terms and normalizing the variance of the shocks,³² we get:

$$n_t = (1 - \eta) a_t + z_t \quad (39)$$

$$y_t = a_t + (1 - \alpha) z_t \quad (40)$$

$$w_t = y_t - n_t = \eta a_t - \alpha z_t \quad (41)$$

A useful benchmark is the model with logarithmic utility over consumption ($\eta = 1$). In this case, employment fluctuates in proportion to the preference shifter z_t but does not respond to technology shocks.³³

From the previous equations, it is straightforward to calculate the model's implications for the second moments of interest. In particular we have

$$\text{cov}(y_t - n_t, y_t) = \eta \text{var}(a_t) - \alpha(1 - \alpha) \text{var}(z_t) \quad (42)$$

$$\text{cov}(y_t - n_t, n_t) = \eta(1 - \eta) \text{var}(a_t) - \alpha \text{var}(z_t) \quad (43)$$

In the absence of labor market frictions, labor productivity is unambiguously countercyclical in response to preference shocks. The intuition for this result is that output responds to preference shocks only through employment, and this response is less than proportional because of diminishing returns in labor input ($\alpha > 0$). Since productivity is unambiguously procyclical in response to technology shocks, the unconditional correlations depend on the relative variances of the shocks and the model parameters. For a wide range of parameter values, e.g. with logarithmic utility over consumption ($\eta = 1$), productivity is procyclical with respect to output but countercyclical with respect to employment.

The relative volatility of employment and wages with respect to output are given by

³²If the original shocks are \tilde{a}_t and \tilde{z}_t , then we define $a_t = \Omega \tilde{a}_t$ and $z_t = \Omega \tilde{z}_t$, where $\Omega = 1/[1 - (1 - \alpha)(1 - \eta)]$.

³³This result is an implication of the logarithmic or 'balanced growth' preferences over consumption in combination with the absence of capital or any other intertemporal smoothing technology, and is similar to the 'neutrality result' in Shimer (2010).

the following expressions:

$$\frac{\text{var}(n_t)}{\text{var}(y_t)} = \frac{(1 - \eta)^2 \text{var}(a_t) + \text{var}(z_t)}{\text{var}(a_t) + (1 - \alpha)^2 \text{var}(z_t)} \quad (44)$$

$$\frac{\text{var}(w_t)}{\text{var}(y_t)} = \frac{\eta^2 \text{var}(a_t) + \alpha^2 \text{var}(z_t)}{\text{var}(a_t) + (1 - \alpha)^2 \text{var}(z_t)} \quad (45)$$

The size of the relative volatility measures above depends again on the relative importance of the shocks, as well as on the size of α , the parameter determining the degree of diminishing returns to labor.

4.2 Infinite Labor Market Turnover

We can contrast the predictions of the frictionless model above, with the opposite extreme case of an infinitely large separation rate, resulting in employment being fixed at zero, see equation (30). In this case, no workers will be hired, so that by the aggregate resource constraint (31) $C_t = Y_t$, as in the case of zero turnover. Combining the production function (32) with the equilibrium condition for effort (25), and taking logarithms, ignoring constant terms and normalizing the variance of the shocks,³⁴ we get:

$$e_t = (1 - \eta) a_t + z_t \quad (46)$$

$$y_t = y_t - n_t = (1 + \phi) a_t + (1 - \alpha) z_t \quad (47)$$

Since employment is fixed, effort is now procyclical in response to both types of shocks, as all of the adjustment of labor input occurs on the intensive margin. With an infinitely large separation rate, labor productivity is perfectly (positively) correlated with output. The correlation between productivity and employment, as well as the relative volatility of employment with respect to output equal zero. Finally, since the bargaining set is now infinitely wide, wages may be arbitrarily rigid, depending on the model parameters, so that the relative volatility of wages is also arbitrarily close to zero.

4.3 Preview of the Results

Comparing the predictions of the model with zero separations to the model with an infinitely large separation rate, it is clear that for a sufficiently large decline in labor market turnover:

1. Labor productivity becomes less procyclical with respect to output.
2. Labor productivity goes from acyclical to countercyclical with respect to employment, depending on parameter values (a sufficient condition is logarithmic utility

³⁴In this case, the normalization factor is $1/[1 + \phi - (1 - \alpha)(1 - \eta)]$.

over consumption).

3. The relative volatility of employment increases.
4. The relative volatility of wages increases.

All four of these predictions are consistent with the data, as we documented in section 2. Four elements of our model are crucial for these results: convex employment adjustment costs, multiple shocks, endogenous effort and endogenous wage rigidity.

We are not arguing, of course, that labor market turnover fell from infinity to zero. Rather, the argument so far is meant to illustrate that if the decline in labor market turnover was large enough, it can qualitatively explain the patterns we observe in the data. To answer the question whether we can also quantitatively match those patterns for reasonable parameter values, we now turn to a numerical analysis of the full model.

4.4 Calibration

We simulate data at quarterly frequency and calibrate accordingly. The calibration is summarized in Table 6. Many of the model's parameters can be easily calibrated to values that are standard in the literature. In this vein, we set the discount factor β equal to 0.99, assume logarithmic utility over consumption ($\eta = 1$), and assume $\alpha = 1/3$ for the curvature of the production function to match the capital share in GDP. In the model there is no difference between unemployment and non-participation. Therefore, we set the marginal utility from leisure γ to match the employment-population ratio. Since the amount of labor market frictions affects this ratio as well, we calibrate to an employment-population ratio of 0.7 in the frictionless model.

The calibration of the labor market frictions is crucial for the simulation exercise. Estimates of the convexity of employment adjustment costs vary, with the exponent $1+\mu$ of the cost function $g(H) = \frac{\kappa}{1+\mu}H^{1+\mu}$ ranging from 1.6 to 3.4. The lower end of this range corresponds to a specification, in which we interpret the adjustment costs as search frictions, vacancy posting costs are linear and the matching function has an elasticity with respect to unemployment of 0.6, as in Mortensen and Nagypal (2007). The upper end of the range is the point estimate of the convexity of employment adjustment costs in Merz and Yashiv (2007). In our benchmark specification, we use the midpoint of this range and assume an exponent of $1+\mu = 2.5$, but we explore the implications for our results if adjustment costs are less or more convex than that.³⁵ We calibrate κ such that

³⁵Despite the estimates in Merz and Yashiv (2007), in later work Yashiv (2012) and Faccini, Millard, and Yashiv (2012) favor a less convex specification closer to our benchmark specification. Two important points should be noted here. First, a crucial insight from Merz and Yashiv (2007) and Yashiv (2012) is that it is important to allow for an interaction in adjustment costs in capital and labor in order to get correct estimates for these adjustment costs. We ignore this interaction because our model does not have capital. This should not cause any problems because we are not estimating the model. Second, by assuming adjustment costs are convex, we are ruling out a large class of adjustment cost functions.

hiring costs are 3% of output in calibration for the pre-84 period, consistent with the estimates in Silva and Toledo (2009), see also Hagedorn and Manovskii (2008, p.1699).

The employment outflow rate declined by about 50%, from 4% per month in the early 1980s to 2% per month in the mid-1990s (Davis, Faberman, Haltiwanger, Jarmin, and Miranda (2010), Fujita (2011), Cairó and Cajner (2013)).³⁶ Using these estimates, we calibrate the gross separation rate δ in our model to 35% per quarter for the pre-84 subsample and to 20% per quarter for the post-84 period.³⁷ In equilibrium, the decline in the separation rate implies a decline in job creation, because the amount of replacement hiring that is necessary to maintain a certain level of employment decreases. This effect is dampened, however, by the lower cost of hiring, which raises equilibrium employment by about 14%.

For the model's driving forces, we assume high persistence in both shocks, setting $\rho_a = 0.97$ to match the first-order autocorrelation in Solow residuals, and $\rho_z = 0.97$ to make sure that none of the results are driven by differences in persistence. Given those values, we calibrate σ_a^2 and σ_z^2 so that the frictionless version of the calibrated model matches the relative volatility of employment and predicts a standard deviation of log output of 1%. The first target is justified by the observation that in this very simple model, preference shocks are a stand-in for all sources of misspecification that result in the unemployment volatility puzzle. The second target is arbitrarily chosen to emphasize that we consider this model mostly illustrative and not able to generate realistic predictions for the overall level of volatility in the economy.

For the parameters related to effort, we have very little guidance from previous literature. We normalize $\phi = 0$ and ζ such that effort is expressed in utility units and equals 1 in the frictionless steady state. We treat the curvature of the production function in effort as a free parameter. Since we are mostly interested to illustrate

This assumption is justified by various studies that find that, while non-convexities are important at the plant level, convex adjustment costs provide a good approximation for the aggregate dynamics for capital (Cooper and Haltiwanger (2006), Khan and Thomas (2008)) and employment (Cooper and Willis (2004)).

³⁶The estimates in Fujita (2011) differ from those in Davis, Faberman, Haltiwanger, Jarmin, and Miranda (2010) and Cairó and Cajner (2013) because Fujita calculates worker flows from matching the labor force status of workers in the monthly CPS files, whereas the other two studies use data on unemployment duration following Shimer (2012). The size of the proportional decline in the separation rate is very similar in both approaches, but the level of the separation rate is different. Starting with Shimer (2005), it is common in the literature to calibrate models to the level of the separation rate as calculated from the unemployment duration data, resulting in a post-war sample average of about 3% per month.

³⁷The quarterly separation probability is the probability that a worker who is employed at the beginning of the quarter is no longer employed at the end of the quarter. Using a monthly job finding probability of $f_m = 0.45$, see Shimer (2012), and a monthly separation probability of $s_m = 0.04$, we get a quarterly separation probability of $s = s_m (1 - f_m)^2 + (1 - s_m) s_m (1 - f_m) + (1 - s_m)^2 s_m + s_m^2 f_m = 0.07$ and a quarterly job finding probability of $f = f_m (1 - s_m)^2 + (1 - f_m) f_m (1 - s_m) + (1 - f_m)^2 f_m + f_m^2 s_m = 0.80$. The gross separation rate is the average number of times that a worker who is employed at the beginning of the quarter loses her job over the quarter. Since workers that are separated in a given quarter may find another job within that quarter, the quarterly gross separation rate is given by $\delta = s / (1 - f) = 0.35$.

the qualitative changes in the business cycle moments that the model can generate, we set this parameter fairly arbitrarily to $\alpha = 0.3$, so that the model roughly replicates the second moments in the data. The testable prediction here is not whether the model can quantitatively match some or most of the second moments, but whether it can qualitatively generate *all* observed changes, changing *only* the separation rate.

4.5 Simulation Results

We now simulate the calibrated model in order to calculate the second moments of interest. We start with the model with flexible wages and show that a decline in labor market turnover of the same size as observed in the US, roughly matches the change in the cyclicalities of labor productivity and the relative volatility of labor input in the data. Then we consider endogenous wage rigidity and show that the model can also generate an increase in the relative volatility of wages and a reduction in output volatility, although not of the same magnitude as observed in the data.

The model with flexible wages is close to log-linear and a first-order approximation captures well the dynamics generated by the model. However, the model with endogenous wage rigidity is non-linear, because the wage depends on the past wage W_{t-1} and the bounds of the bargaining set W_t^{UB} and W_t^{LB} , multiplied by the degree of rigidity R_t , which itself depends again on W_t and W_t^{UB} and W_t^{LB} , see (26) and (27). If we were to log-linearize this wage rule, it would reduce to a partial adjustment rule with a constant degree of wage rigidity. Therefore, we use a second-order approximation of the policy functions. As an accuracy check, Figure 3 shows that a second-order approximation captures well the non-linear wage rule for $\rho = 1$, but for $\rho = 2$ or larger, a higher-order approximation is needed. We simulate the second-order approximation of the model 201,000 periods, discarding the first 1,000 observations to eliminate the effect of the initial conditions. The results of this exercise are reported in Table 7.

4.5.1 Flexible Wages

Labor productivity is strongly procyclical in terms of its correlation with output in the model and its procyclicality falls substantially as we reduce labor market turnover. The correlation of productivity with employment also falls, from around zero in the labor market with high turnover to a negative value in the calibration with low turnover. Both observations are qualitatively as well as quantitatively consistent with the evidence. The reason for the decline in the procyclicality of productivity, is the increase in the relative volatility of employment, a result that is consistent with the data as well.

Three elements in the model are crucial for these results. First, the convexity of the employment adjustment costs implies that hiring costs fall from 3% to around 1% of output with the decline in labor market turnover. Second, the effort choice provides an intensive margin of adjustment for labor input. As frictions fall, it becomes optimal

to adjust labor more through employment and less through effort. Thus, the volatility of employment increases more than that of output. Note that the model also predicts that the volatility of effort falls, but this prediction is not testable without observable measures of effort.³⁸

The third element in the model that is important for the results is that fluctuations in the model are driven by two types of shocks: technology shocks and preference shocks or labor supply shocks. In a one-shock model, the correlations between all variables would be close to either 1 or -1 .³⁹ In addition, if fluctuations were driven only by technology shocks then productivity could never be countercyclical, since employment would only fluctuate because of changes in labor demand, and the direct effect of technology on productivity would always prevail over the indirect effect of employment. It is important to stress, however, that our results are *not* driven by changes in the relative importance of both shocks, which we keep constant, but by the reduction in frictions, which changes the response of the economy conditional on each shock.

The model also predicts a small decrease in the relative volatility of wages and a small increase in the volatility of output. As argued in section 2.3, the first prediction is arguably not consistent with the data. The second one clearly is in contradiction with the well-documented reduction in output volatility, the so called Great Moderation. The decrease in the relative volatility of wages is driven by the fact that the wage is approximately proportional to the marginal product of labor.⁴⁰ Since the marginal product of labor is proportional to output, but inversely proportional to employment, an increase in the relative volatility of employment must necessarily also decrease the relative volatility of wages. The increase in the volatility of output simply stems from the fact that reducing adjustment costs amplifies fluctuations in employment and therefore in output as well. In the next subsection we show that endogenous wage rigidities can reverse the predictions of the model for wages and possibly also bring the prediction for output volatility closer to the evidence.

4.5.2 Endogenous Wage Rigidity

The third panel in Table 7 presents the simulated second moments for the model with endogenous wage rigidity ($\rho = 1$ and $\bar{R} = 0.95$). Comparing these moments to those for

³⁸Some authors have used hours per worker as an observable proxy for effort, see e.g. Basu, Fernald, and Kimball (2006). However, the crucial characteristic of the effort margin in our model is that it is not subject to adjustment costs, or at least that adjusting this margin is less costly than adjusting along the extensive margin. It is not clear that adjusting work hours is costless. In fact, van Rens (2012) argues that adjustment costs in hours per worker may be larger rather than smaller than adjustment costs in employment in OECD countries.

³⁹This is exactly true in a static, linear model. Our model is close to (log)linear and the version without capital and with flexible wages has only one state variable (employment), which has very fast transition dynamics.

⁴⁰In a frictional labor market, the wage is not equal to marginal product of labor, but as long as γ is not too large, they are still proportional since workers and firms share the match surplus in equal proportions.

the flexible wage model, we see that the previously described predictions of the model for the cyclicalities of labor productivity and the relative volatility of employment remain largely unchanged. The reason is that the fact that wages adjust when they get close to the bounds of the bargaining set mitigates the allocative effect of wage rigidity.

The prediction of the model for the relative volatility of wages, however, is reversed. The reduction in labor market frictions now increases the volatility of wages, although not by as much as in the data. To understand the mechanism behind this result, it is useful to consider the extreme case of endogenous wage rigidity ($\rho \rightarrow \infty$), in which wages are completely fixed within the bargaining set, but adjust when they hit the bounds of the bargaining set. If frictions are high enough, so that the bargaining set is very wide, wages never adjust and their volatility is zero. On the other extreme, in a frictionless labor market, the bargaining set reduces to a point and wages behave as if they were flexible. Of course this effect is counteracted by the fact that the bounds of the bargaining set themselves are less volatile when frictions are lower. However, for our calibration of the parameters, the first effect dominates, as illustrated in Figure 4.

The small increase in output volatility in response to the reduction in labor market frictions is reversed to a (very small) decrease for the model with endogenous wage rigidity. The reason is that increased wage flexibility dampens fluctuations in output in response to technology shocks.⁴¹ This result is consistent with a possible role of a decline in labor market frictions as a source of the Great Moderation. However, the effect is again much smaller than in the data. Thus, whereas we find this last result intuitively compelling, it is not clear whether it is quantitatively important.

5 Conclusions

In this paper, we documented three changes in labor market dynamics over the post-war period in the US: the strong procyclicality of labor productivity has vanished, the volatility of employment has increased with respect to output, and the volatility of wages has increased relative to output and possibly even in absolute terms. We presented a model to argue that the decline in labor market turnover, modelled as a reduction in the employment exit probability, can explain all three facts. In addition, we showed that – in principle – the decline in turnover may also have contributed to the reduction in output volatility, which happened around the same time.

The intuition for why a decline in labor market turnover increases the relative volatility of employment and reduces the procyclicality of labor productivity is straightforward and compelling. If employment adjustment costs are convex, then lower turnover implies lower hiring costs. If there is another input into production that can be used at least partly as a substitute for labor, then a reduction in hiring frictions will make that

⁴¹In response to preference shocks, which affect labor supply rather than labor demand, the reverse is true. However, the effect of technology shocks dominates that of preference shocks in our calibration.

input less volatile, so that employment becomes more volatile with respect to output. In this paper, we refer to this other factor input as effort, but a very similar argument can be made for capacity utilization of capital. Given that capital does not fluctuate much at business cycle frequencies, the fact that the comovement of labor and output – and therefore labor productivity – has changed almost unavoidably leads to the conclusion that there must be another input into the production process.

Our argument that the decline in labor market turnover is also responsible for the increase in the relative volatility of wages is more contentious. Our simulations of the model with endogenous wage rigidity show that this effect is qualitatively present, and can be made to dominate the direct effect on wages for reasonable parameter values. However, the increase in wage volatility in our simulations is much smaller than in the data.⁴² It is possible that this effect could be amplified by using a more non-linear wage rule, but we do not pursue this possibility. While there is a compelling argument that wage rigidity is, at least to some degree, endogenous, it remains an open question quantitatively how important this mechanism is.

What have we learned about the possible causes of the Great Moderation? We showed that a decline in labor market turnover, through a decrease in hiring frictions and an endogenous decrease in wage rigidity, can lead to a decrease in output volatility. However, this effect is small in our simulations, partly because the decrease in wage rigidity is relatively small (at least much smaller than in the data), but also because there is a direct effect of the reduction in hiring frictions that counteracts the effect of decreased wage rigidity: reduced frictions make labor more volatile, reducing the volatility of its marginal product. A further caveat against the conclusion that the Great Moderation is driven by the decline in labor market turnover is that this story only works if technology shocks are relatively important as a source of business cycle fluctuations. The reason is that preference shocks act as labor supply shocks, so that wage rigidity dampens rather than amplifies fluctuations in employment in response to these shocks. Finally, reduced wage rigidity can only lead to reduced output volatility if the wage is allocative, i.e. if the reduction in wage rigidity applies to the wage of newly hired workers. We documented an increase in the relative volatility of wages of new hires, see section 2.3.2, but that increase is substantially smaller than for the average wage in the economy, casting further doubt on the importance of this mechanism for the Great Moderation.

⁴² At least than the data from the Labor Productivity and Cost program and the Current Population Survey. The data from the Current Employment Statistics show a decrease in the relative volatility of wages, see section 2.3, which our model has no trouble generating.

A Marginal Product and Disutility of Effort

This appendix derives the marginal product of employment to the firm, equation (12), and the marginal disutility from employment, expressed in consumption terms, to the household, equation (16), if effort adjusts endogenously. From equations (4) and (2), it is straightforward differentiation to decompose the total effect of employment on output and total effective labor supply into a direct effect and an effect through the endogenous response of effort.

$$\frac{dY_{jt}}{dN_{jt}} = \frac{\partial Y_{jt}}{\partial N_{jt}} + \frac{\partial Y_{jt}}{\partial \mathcal{E}_{jt}} \frac{\partial \mathcal{E}_{jt}}{\partial N_{jt}} = \frac{(1-\alpha)Y_{jt}}{N_{jt}} \left(1 + \frac{N_{jt}}{\mathcal{E}_{jt}} \frac{\partial \mathcal{E}_{jt}}{\partial N_{jt}} \right) \quad (48)$$

$$\frac{dL_{ht}}{dN_{ht}} = \frac{\partial L_{ht}}{\partial N_{ht}} + \frac{\partial L_{ht}}{\partial \mathcal{E}_{ht}} \frac{\partial \mathcal{E}_{ht}}{\partial N_{ht}} = \frac{1}{1+\zeta} \left[1 + \zeta \mathcal{E}_{ht}^{1+\phi} \left(1 + (1+\phi) \frac{N_{ht}}{\mathcal{E}_{ht}} \frac{\partial \mathcal{E}_{ht}}{\partial N_{ht}} \right) \right] \quad (49)$$

Here, \mathcal{E}_{jt} denotes the effort of all workers i that are employed in firm j and \mathcal{E}_{ht} the effort of all workers that are members of household h .

To find the response of effort to changes in employment that firm and household face, we use the condition that the marginal disutility from effort of a given worker i (expressed in consumption terms) from equation (8), in equilibrium must equal the marginal productivity of that worker to the firm from equation (9).

$$\mathcal{E}_{it}^{1+\phi-} = \frac{(1+\zeta)}{(1+\phi)\zeta} \frac{Z_t}{\gamma C_{ht}^\eta} (1-\alpha) A_t \left(\int_0^{N_{jt}} \mathcal{E}_{vt} dv \right)^{-\alpha} \quad (50)$$

First, suppose firm j considers employing N_{jt} workers, given that all other firms employ the equilibrium number of workers N_t . Because there are infinitely many firms, firm j 's decision to employ $N_{jt} \neq N_t$ workers does not affect the fraction of household h 's members that are employed, so that by the assumption of perfect risk-sharing within the household, the consumption of workers in firm j is not affected, $C_{ht} = C_t$. Substituting this, as well as the condition that all workers in firm j exert the same amount of effort, $\mathcal{E}_{it} = \mathcal{E}_{jt}$ for all $i \in [0, N_{jt}]$, the effort condition becomes,

$$\mathcal{E}_{jt}^{1+\phi-} = \frac{(1+\zeta)}{(1+\phi)\zeta} \frac{Z_t}{\gamma C_t^\eta} (1-\alpha) A_t \left(\mathcal{E}_{jt} N_{jt} \right)^{-\alpha} \quad (51)$$

so that the elasticity of effort in a given firm j with respect to employment in that firm, is given by

$$\frac{N_{jt}}{\mathcal{E}_{jt}} \frac{\partial \mathcal{E}_{jt}}{\partial N_{jt}} = - \frac{\alpha}{1+\phi-(1-\alpha)} \quad (52)$$

Substituting this elasticity into equation (48) above, gives expression (12) in the text.

Next, suppose household h considers having N_{ht} employed workers, given that all other households have N_t employed workers. Because there are infinitely many households, household's h 's decision to have a fraction of $N_{ht} \neq N_t$ of its members employed,

does not affect the level of employment in any firm $N_{jt} = N_t$. Furthermore, although the effort level of worker i may change because of household h 's decision, effort of all other workers in firm j , who are members of different households, is unaffected, $\mathcal{E}_{it} = \mathcal{E}_{ht}$ and $\mathcal{E}_{i't} = \mathcal{E}_t$ for $i' \neq i$. Thus, the effort condition becomes,

$$\mathcal{E}_{ht}^{1+\phi-} = \frac{(1+\zeta)}{(1+\phi)\zeta} \frac{Z_t}{\gamma C_{ht}^\eta} (1-\alpha) A_t \left(\mathcal{E}_t N_t \right)^{-\alpha} \quad (53)$$

and the elasticity of effort exerted by members of household h with respect to employment in that household, using equation (3), is given by,

$$\frac{N_{ht}}{\mathcal{E}_{ht}} \frac{\partial \mathcal{E}_{ht}}{\partial N_{ht}} = \frac{C_{ht}}{\mathcal{E}_{ht}} \frac{\partial \mathcal{E}_{ht}}{\partial C_{ht}} \cdot \frac{N_{ht}}{C_{ht}} \frac{\partial C_{ht}}{\partial N_{ht}} = -\frac{\eta}{1+\phi-} \frac{W_{ht} N_{ht}}{C_{ht}} = -\frac{\eta}{1+\phi-} \quad (54)$$

Substituting this elasticity into equation (49) above, gives expression (16) in the text.

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Table 1. The Vanishing Procyclicality of Labor Productivity

	Corr with output			Corr with labor input		
	Pre-84	Post-84	Change	Pre-84	Post-84	Change
Output per hour						
BP	0.60	0.25	−0.35	0.19	−0.40	−0.59
	[0.06]	[0.07]	[0.09]	[0.09]	[0.07]	[0.11]
4D	0.62	0.18	−0.44	0.09	−0.54	−0.62
	[0.05]	[0.09]	[0.10]	[0.07]	[0.08]	[0.11]
HP	0.61	0.04	−0.57	0.17	−0.56	−0.73
	[0.05]	[0.09]	[0.10]	[0.08]	[0.07]	[0.10]
Output per worker						
BP	0.78	0.60	−0.18	0.31	−0.15	−0.47
	[0.04]	[0.05]	[0.06]	[0.08]	[0.10]	[0.13]
4D	0.76	0.46	−0.30	0.13	−0.33	−0.46
	[0.03]	[0.09]	[0.09]	[0.08]	[0.11]	[0.14]
HP	0.77	0.40	−0.37	0.25	−0.32	−0.57
	[0.03]	[0.08]	[0.09]	[0.07]	[0.09]	[0.11]

Standard errors in brackets are calculated from the variance-covariance matrix of the second moments using the delta method. Data are from the BLS labor productivity and cost program (LPC) and refer to the private sector. Labor input is total hours worked in the first panel and employment in the second panel, consistent with the definition of labor productivity. The sample period is 1949-2007.

Table 2. The Rising Volatility of Labor Input

	Std. Dev.			Relative Std. Dev.		
	Pre-84	Post-84	Ratio	Pre-84	Post-84	Ratio
Employment (private sector)						
BP	1.57	0.91	0.58	0.66	0.81	1.23
	[0.08]	[0.05]	[0.04]	[0.03]	[0.05]	[0.09]
4D	2.44	1.54	0.63	0.66	0.94	1.43
	[0.13]	[0.12]	[0.06]	[0.03]	[0.08]	[0.15]
HP	1.62	1.16	0.72	0.66	0.97	1.47
	[0.08]	[0.07]	[0.06]	[0.03]	[0.06]	[0.12]
Hours (private sector)						
BP	1.93	1.18	0.61	0.81	1.06	1.30
	[0.09]	[0.06]	[0.04]	[0.03]	[0.05]	[0.08]
4D	2.94	1.92	0.65	0.79	1.17	1.48
	[0.15]	[0.13]	[0.06]	[0.04]	[0.09]	[0.13]
HP	1.96	1.44	0.73	0.80	1.20	1.50
	[0.10]	[0.08]	[0.05]	[0.03]	[0.07]	[0.10]
Hours (total economy)						
BP	1.68	0.85	0.51	0.71	0.76	1.07
	[0.08]	[0.04]	[0.04]	[0.03]	[0.04]	[0.07]
4D	2.56	1.47	0.57	0.69	0.89	1.30
	[0.14]	[0.10]	[0.05]	[0.03]	[0.07]	[0.11]
HP	1.71	1.06	0.62	0.70	0.89	1.27
	[0.09]	[0.07]	[0.05]	[0.03]	[0.05]	[0.09]

Standard errors in brackets are calculated from the variance-covariance matrix of the second moments using the delta method. Data are from the BLS labor productivity and cost program (LPC) and refer to the private sector. Hours (total economy) is an unpublished series for economy-wide hours constructed by the BLS and used in Francis and Ramey (2009). The sample period is 1949-2007.

Table 3. The Rising Volatility of Wages

	Std. Dev.			Relative Std. Dev.		
	Pre-84	Post-84	Ratio	Pre-84	Post-84	Ratio
Compensation per hour (NIPA, private sector)						
BP	0.71	0.99	1.38	0.30	0.88	2.93
	[0.05]	[0.06]	[0.12]	[0.02]	[0.07]	[0.31]
4D	1.72	1.61	0.93	0.46	0.98	2.11
	[0.12]	[0.11]	[0.09]	[0.04]	[0.12]	[0.32]
HP	0.85	1.03	1.21	0.35	0.86	2.48
	[0.06]	[0.06]	[0.11]	[0.03]	[0.08]	[0.29]
Compensation per hour (NIPA, total economy)						
BP	0.78	0.86	1.10	0.33	0.76	2.32
	[0.05]	[0.05]	[0.10]	[0.02]	[0.07]	[0.24]
4D	1.85	1.57	0.85	0.50	0.95	1.92
	[0.10]	[0.10]	[0.07]	[0.03]	[0.12]	[0.26]
HP	0.84	0.95	1.14	0.34	0.80	2.33
	[0.05]	[0.08]	[0.11]	[0.02]	[0.09]	[0.30]
Earnings per hour (CES, private sector)						
BP	1.38	0.40	0.29	0.54	0.34	0.63
	[0.12]	[0.02]	[0.03]	[0.04]	[0.03]	[0.07]
4D	1.85	0.97	0.52	0.52	0.56	1.08
	[0.14]	[0.08]	[0.06]	[0.05]	[0.06]	[0.16]
HP	1.27	0.46	0.36	0.49	0.37	0.75
	[0.11]	[0.03]	[0.04]	[0.05]	[0.04]	[0.10]
Compensation per hour (NIPA, private sector, CPI deflator)						
BP	0.83	0.99	1.20	0.35	0.89	2.54
	[0.05]	[0.06]	[0.10]	[0.02]	[0.08]	[0.27]
4D	1.93	1.70	0.88	0.52	1.04	1.99
	[0.13]	[0.12]	[0.08]	[0.04]	[0.13]	[0.30]
HP	0.96	1.04	1.08	0.39	0.87	2.21
	[0.06]	[0.06]	[0.09]	[0.03]	[0.08]	[0.26]

Standard errors in brackets are calculated from the variance-covariance matrix of the second moments using the delta method. Wages are calculated as real compensation per hour. Compensation and hours data for the private sector are from the BLS labor productivity and cost program. Compensation data for NIPA compensation are combined with an unpublished economy-wide series for hours constructed by the BLS and used in Francis and Ramey (2009). Compensation from the establishment survey or Current Employment Statistics (CES) exclude non-wage payments. The sample period is 1949-2007 for the NIPA and 1965-2004 for the CES-based series.

Table 4. The Rising Volatility of Wages: Newly Hired Workers

	Std. Dev.			Relative Std. Dev.		
	Pre-84	Post-84	Ratio	Pre-84	Post-84	Ratio
Compensation per hour (NIPA, private sector)						
BP	0.46	1.01	2.20	0.18	0.88	4.78
	[0.06]	[0.06]	[0.29]	[0.04]	[0.08]	[1.01]
4D	1.71	1.63	0.95	0.37	0.97	2.62
	[0.13]	[0.11]	[0.10]	[0.05]	[0.13]	[0.50]
HP	0.59	1.05	1.78	0.20	0.86	4.22
	[0.08]	[0.07]	[0.27]	[0.05]	[0.08]	[1.09]
Mean usual hourly earnings (CPS)						
BP	0.49	0.54	1.10	0.20	0.47	2.39
	[0.05]	[0.03]	[0.13]	[0.04]	[0.04]	[0.48]
4D	1.48	1.27	0.86	0.32	0.76	2.36
	[0.14]	[0.11]	[0.11]	[0.05]	[0.09]	[0.46]
HP	0.66	0.83	1.27	0.23	0.68	3.02
	[0.11]	[0.06]	[0.24]	[0.05]	[0.07]	[0.78]
Mean usual hourly earnings newly hired workers (CPS)						
BP	1.42	0.67	0.47	0.57	0.58	1.03
	[0.16]	[0.05]	[0.07]	[0.08]	[0.05]	[0.17]
4D	4.49	2.77	0.62	0.97	1.65	1.69
	[0.60]	[0.26]	[0.10]	[0.17]	[0.24]	[0.38]
HP	3.54	2.31	0.65	1.22	1.88	1.54
	[0.47]	[0.18]	[0.10]	[0.25]	[0.18]	[0.35]

Standard errors in brackets are calculated from the variance-covariance matrix of the second moments using the delta method. CPS wage data are earnings per hour from the outgoing rotation groups, which limits the period for which quarterly data are available to after 1979. Wage series are constructed as in Haefke, Sonntag, and van Rens (2013) but are not corrected for composition bias for comparability with other data sources. However, keeping the composition of the labor force constant in terms of education, experience and demographic characteristics makes very little difference for the results presented here. The sample period is 1980-2005.

Table 5. Changes in Labor Market Dynamics in Major European Countries, 1960-2007

	Correlation Productivity				Relative Std. Dev.	
	with output		with employment		employment	
	Pre-84	Post-84	Pre-84	Post-84	Pre-84	Post-84
US 1949-2007 (baseline)	0.78	0.60	0.31	-0.15	0.66	0.81
US 1960-2007 (OR)	0.77	0.67	0.27	-0.03	0.67	0.74
France	0.86	0.68	0.06	-0.14	0.51	0.74
Germany	0.87	0.77	0.41	0.22	0.53	0.65
Italy	0.94	0.71	0.31	-0.54	0.37	0.84
UK	0.88	0.54	-0.02	-0.33	0.47	0.89

All data are bandpass filtered and refer to the private sector. Data for the baseline results for the US are from the BLS labor productivity and cost program (LPC), see Tables 1, 2 and 3 for details. Data for all other countries were collected by Ohanian and Raffo (2012) from the OECD Economic Outlook database and national statistics offices. For consistency with our baseline results, productivity is real output per worker and employment is in persons, although the Ohanian-Raffo data also allow to calculate output per hour and total hours.

Table 6. Model Calibration

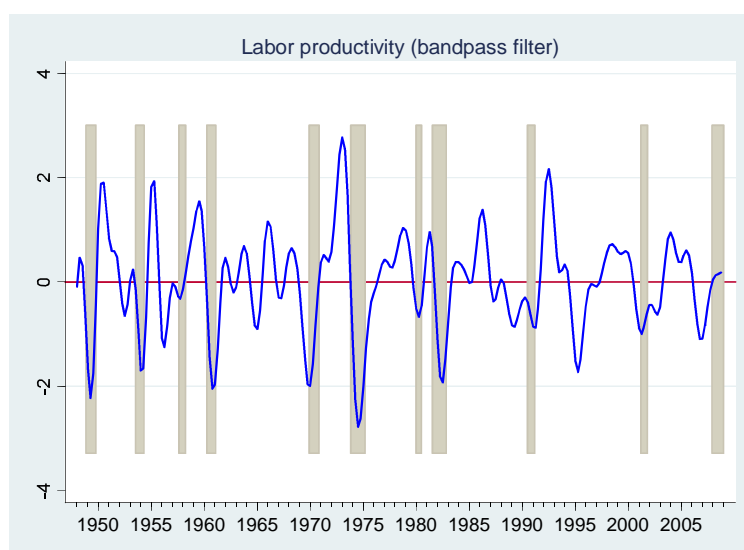
	Parameter	Target
Utility:	$\beta = 0.99$	quarterly data
	$\eta = 1$	log utility over consumption
	$\gamma = 1.24$	frictionless employment population ratio $\bar{N} = 0.7$
Production:	$f(N) = N^{1-\alpha}, \alpha = 1/3$	capital share
Effort:	$\zeta = 0.299$	normalization: frictionless $\bar{\mathcal{E}} = 1$
	$\phi = 0$	normalization so that \mathcal{E} is in utils
	$= 0.3$	total curvature $\phi +$ is a free parameter
Frictions:	$\delta = 0.35 - 0.20$	gross quarterly separations, decline in turnover
	$g(H) = \frac{\kappa}{1+\mu} H^{1+\mu}, \mu = 1.5$	convex adjustment costs
	$\kappa = 3.19$	frictions 3% of output pre-84
Shocks:	$\rho_A = 0.97, \sigma_A = 0.186$	normalization: $\text{sd}(y) = 1\%$
	$\rho_z = 0.97, \sigma_z = 0.173$	$\text{sd}(n) / \text{sd}(y) = 0.66$

Table 7. Simulation results

	frictions (% GDP)	empl/pop ratio \tilde{N}	correlation with output	productivity with empl	relative empl n_t	std.dev. wage w_t	std.dev. output y_t
<i>Data</i>							
Pre-84			0.78	0.31	0.66	0.30	
Post-84			0.60	-0.15	0.81	0.88	
<i>Model with flexible wages</i>							
$\delta = 0.40$	3.69	0.52	0.79	0.12	0.62	0.86	1.00
$\delta = 0.35$ (Pre)	3.00	0.57	0.75	0.03	0.66	0.86	1.00
$\delta = 0.30$	2.30	0.59	0.71	-0.07	0.71	0.86	1.00
$\delta = 0.25$	1.63	0.62	0.66	-0.15	0.76	0.86	1.00
$\delta = 0.20$ (Post)	1.02	0.65	0.61	-0.22	0.81	0.86	1.01
$\delta = 0.15$	0.53	0.67	0.57	-0.28	0.86	0.85	1.02
$\delta = 0$	0.00	0.70	0.48	-0.37	0.94	0.83	1.04
<i>Model with endogenous wage rigidity ($\rho = 1, \bar{R} = 0.95$)</i>							
$\delta = 0.40$	3.69	0.52	0.80	0.31	0.63	0.68	1.00
$\delta = 0.35$ (Pre)	3.00	0.57	0.76	0.22	0.66	0.69	1.00
$\delta = 0.30$	2.30	0.59	0.73	0.14	0.69	0.70	1.00
$\delta = 0.25$	1.63	0.62	0.69	0.05	0.72	0.70	1.00
$\delta = 0.20$ (Post)	1.02	0.65	0.65	-0.03	0.76	0.70	1.00
$\delta = 0.15$	0.53	0.67	0.62	-0.09	0.79	0.70	1.00

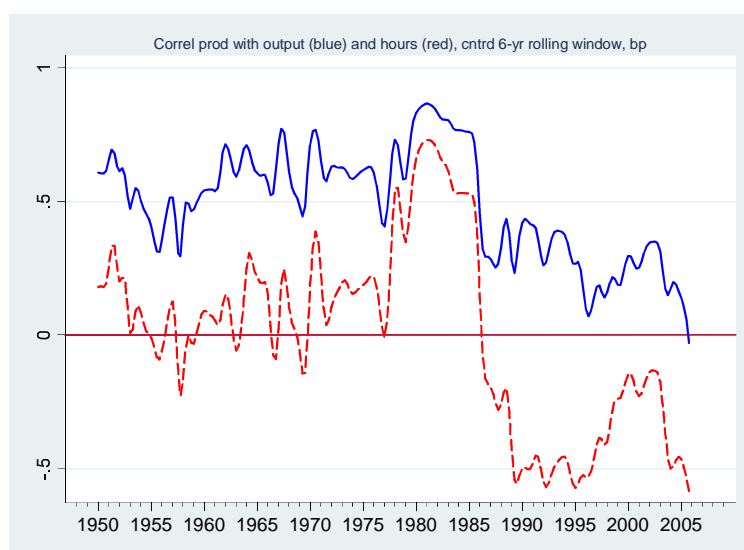
Moments for the model are based on simulated time series of 200,000 quarters. We simulate the model for 201,000 quarters but ignore the first 1,000 quarters to eliminate the effect of the initial conditions. Numbers in bold are calibration targets.

Figure 1. The Vanishing Procyclicality of Labor Productivity



Output per hour in the US private sector. Shaded areas are NBER recessions.

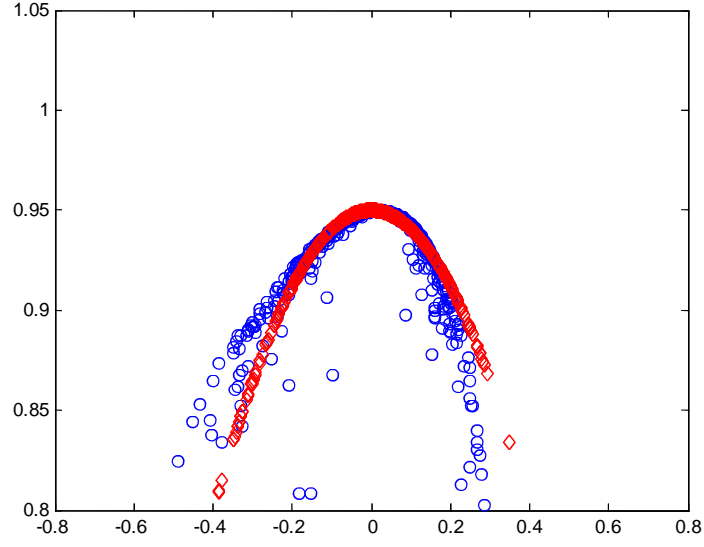
Figure 2. The Vanishing Procyclicality of Labor Productivity: Rolling Correlations



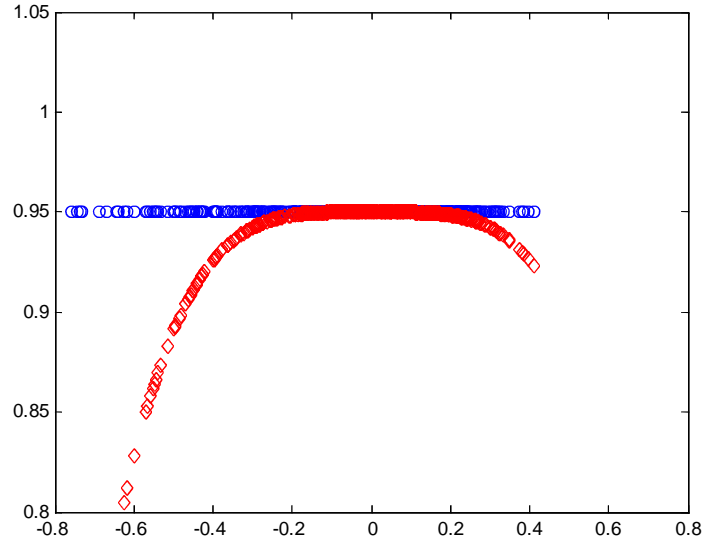
Correlations are calculated in a centered 6-year rolling window of quarterly bandpass-filtered data.

Figure 3. Endogenous Wage Rigidity: Wage Rule

$$\rho = 1 \text{ (quadratic)}$$

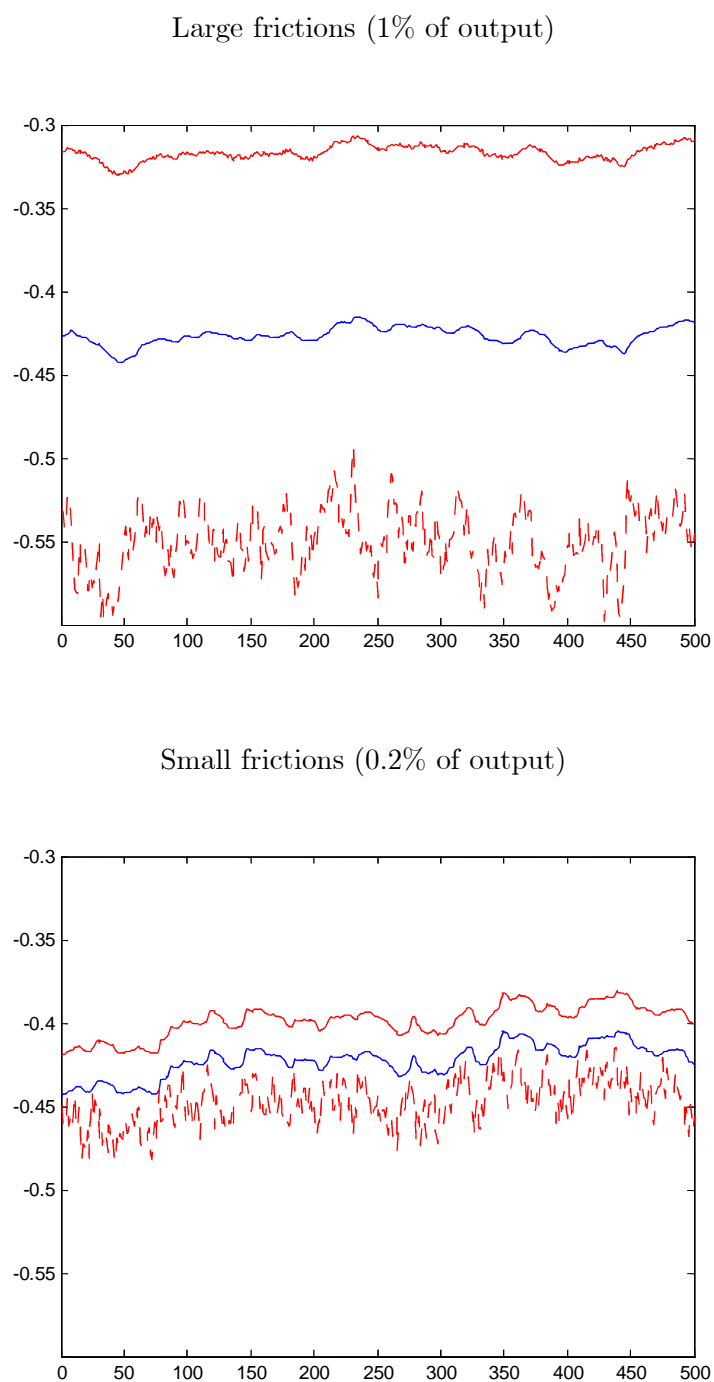


$$\rho = 2 \text{ (quartic)}$$



Wage rigidity R_t as a function of the relative distance of the wage from the center of the bargaining set $(W_t - W_t^*) / \frac{1}{2} (W_t^{UB} - W_t^{LB})$, see equation (23). The red diamonds represent the theoretical non-linear relation. The blue circles are simulated data from a second-order approximation of the model. In these graphs, frictions are set to 0.3% of output.

Figure 4. Endogenous Wage Rigidity: Simulated Wage Data



Simulated wage paths for the model with large frictions and small frictions. In this calibration, the wage is more volatile if frictions are smaller.