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AN APPLICATION TO MEDICAL RESEARCH

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Muhammad I. Bhatti

In this note we outlined the procedure to obtain upper tail percentage points of the distribution of the average of k independent sample correlation coefficient (r) each of which is based on n pairs of observations. Null distribution of \bar{r} demonstrates its application to medical research.

Key words: Average correlation coefficient, Bessel function, characteristics function, critical values, Fubini's theorem, medical research.

1. INTRODUCTION

The α , β and δ cells of the pancreatic islets of Langerhans produce respectively the hormones glucagon, insulin and somatostatin which play central roles in the control of blood glucose concentration [Cooperstein and Dudley (1981, Chapter 18)]. Direct and indirect interactions of these cells can be seen by Bondy and Rosenberg (1977) and Kuhl, Faber, Hornnes and Jansen (1978), Gerich (1977), etc. In biomedical sciences biostatisticians are asked to assess the significance of average correlation coefficient (\bar{r}) between various sets of above-mentioned cells and determine its critical values. The details of such experiments are given in Bhatti (1987), Bergons, Tam, Molner, Rajotte and McGregor (1986). This note is an attempt to provide the distribution of \bar{r} and a table of such critical values for small sample size. These critical points are determined by using modern availability of computing power, with the hope of better and accurate results. In Section 2 of this note the derivation of the distribution of mean correlation coefficient (\bar{r}) is given. Section 3 demonstrates the table of critical values for selected values of K_1 (the number of correlations in average), N (the number of points from which each correlation is calculated), and α' (the upper tail area). If one is interested in different or higher values of K_1 , N and α' he can modify APL programs given at the end of the Appendix or he can obtain the extended version of the APL programs from the author on request.

2. DERIVATION OF THE DISTRIBUTION OF \bar{r}

The purpose of this section is to obtain upper tail percentage points of the distribution of the average of k independent sample correlation coefficients (r) each of which is based on n pairs of

independent observations. According to Hogg and Craig (1978, p.302) the null ($\rho=0$, where ρ is the population correlations coefficient) p.d.f. of r based on n pairs of observations is

$$g(r) = \begin{cases} \frac{\Gamma[(n-1)/2]}{\Gamma(1/2)\Gamma[(n-2)/2]} (1-r^2)^{(n-4)/2} & , \text{ for } -1 < r < 1 \\ 0 & , \text{ otherwise .} \end{cases}$$

where

$$r = \frac{\Sigma(x_i - \bar{x})(y_i - \bar{y})}{[\Sigma(x_i - \bar{x})^2 \Sigma(y_i - \bar{y})^2]^{1/2}}$$

The moment-generating function of the null distribution of r is therefore:

$$M_r(t) = E(e^{tr}) = \frac{\Gamma[(n-1)/2]}{\Gamma(1/2)\Gamma[(n-2)/2]} \int_{-1}^1 e^{tr} (1-r^2)^{\frac{n-3}{2} - \frac{1}{2}} dr .$$

$$M_r(t) = E(e^{tr}) = \Gamma[(n-1)/2] 2^{\frac{n-3}{2}} t^{-\frac{(n-3)}{2}} I_{\frac{(n-3)}{2}}(t), \text{ if } n > 2$$

Where

$$I_{\frac{(n-3)}{2}}(t) = \frac{(t/2)^{\frac{n-3}{2}}}{(\frac{n-3}{2})!} \left[1 + \frac{t^2}{2(n-1)} + \frac{t^4}{2^3(n-1)(n+1)} + \dots \right]$$

is the modified Bessel function of the second kind of order $(n-3)/2$. The corresponding characteristic function (c.f.) can be obtained by replacing t with it in $M_r(t)$. Also, note that since the p.d.f. of the distribution of r is symmetric, the c.f. will be real. Thus,

$$\phi_r(t) = M_r(it) = \Gamma[(n-1)/2] 2^{\frac{(n-3)}{2}} (it)^{-\frac{1}{2}(n-3)} I_{\frac{(n-3)}{2}}(it) .$$

Expressing $\phi_r(t)$ in terms of the Bessel function $J_{\frac{n-3}{2}}(t)$, and by using the relation

$$I_\nu(z) = i^{-\nu} J_\nu(iz) ,$$

we get the c.f. as,

$$\phi_r(t) = \Gamma[(n-1)/2] 2^{\frac{(n-3)}{2}} t^{-\frac{1}{2}(n-3)} J_{\frac{(n-3)}{2}}(t) .$$

Thus, the c.f. of the distribution of the mean \bar{r} of k independent values of r (i.e., the c.f. of the distribution of \bar{r}) is given by

$$\begin{aligned} \phi_{\bar{r}}(t) &= M_{\bar{r}}(it) = [M_r(it/k)]^k \\ \phi_{\bar{r}}(t) &= \left[\Gamma[(n-1)/2] 2^{\frac{(n-3)}{2}} (t/k)^{-\frac{1}{2}(n-3)} J_{\frac{(n-3)}{2}}(t/k) \right]^k . \end{aligned}$$

Note, that $\phi_{\bar{r}}(t)$ is an even function of t , since $\phi_{\bar{r}}(t) = \phi_{\bar{r}}(-t)$ and thus the distribution of \bar{r} is symmetric. Using the inversion formula for characteristic function, see Feller (1968), we can obtain the p.d.f. of the distribution of \bar{r} as,

$$\begin{aligned} f(\bar{r}) &= \frac{1}{2\pi} \int_{-\infty}^{\infty} e^{-it\bar{r}} \phi_{\bar{r}}(t) dt \\ &= \frac{1}{\pi} \int_0^{\infty} \cos(t\bar{r}) \phi_{\bar{r}}(t) dt . \end{aligned}$$

The upper tail area above r_0 is,

$$P(\bar{r} > r_0) = \frac{1}{2} - P(0 \leq \bar{r} \leq r_0) = \frac{1}{2} - \int_0^{r_0} f(\bar{r}) d\bar{r}$$

$$P(\bar{r} > r_0) = \frac{1}{2} - \int_0^{r_0} \left\{ \frac{1}{\pi} \int_0^{\infty} \cos(t\bar{r}) \phi_{\bar{r}}(t) dt \right\} d\bar{r} .$$

Using Fubini's theorem,

$$\begin{aligned} P(\bar{r} > r_0) &= \frac{1}{2} - \frac{1}{\pi} \int_0^{\infty} \phi_{\bar{r}}(t) \left[\int_0^{r_0} \cos(t\bar{r}) d\bar{r} \right] dt \\ &= \frac{1}{2} - \frac{1}{\pi} \int_0^{\infty} \phi_{\bar{r}}(t) \frac{1}{t} \sin(tr_0) dt . \end{aligned}$$

An APL program to evaluate the critical values is given in Appendix I.

3. COMPUTATION OF CRITICAL VALUES

The results of these computed critical values are given in Table 1, for selected values of N , K_1 and α' . Where N is the number of points from which each correlation is calculated (3, 4, 5, ..., 10), K_1 is the number of correlations in average (1, 2, 3, ..., 10) and α' is the upper tail area (0.05, 0.025, 0.01, 0.005).

TABLE OF CRITICAL VALUES FOR SELECTED VALUES OF N, K_1 AND α

K_1		1	2	3	4	5	6	7	8	9	10
3	α'										
	.05	.988	.8458	.6775	.5852	.5227	.4766	.4409	.4123	.3887	.3686
	.025	.997	.9211	.7952	.6836	.6133	.5610	.5200	.4868	.4592	.4360
	.01	.9995	.9761	.8828	.7891	.7119	.6533	.6074	.5703	.5391	.5117
	.005	.9998	.9956	.9297	.8535	.7734	.7129	.6641	.625	.5908	.5625
4	.05	.9121	.6821	.5518	.4768	.4260	.3884	.3594	.3362	.3169	.3005
	.025	.9511	.7739	.6455	.5591	.5015	.4583	.4246	.3975	.3750	.3557
	.01	.9870	.8608	.7383	.6484	.5840	.5356	.4971	.4663	.4404	.4179
	.005	.9999	.9082	.7988	.7051	.6367	.5859	.5449	.5117	.4834	.4590
5	.05	.8030	.5881	.4773	.4126	.3688	.3364	.3114	.2913	.2745	.2604
	.025	.8823	.6802	.5591	.4853	.4351	.3977	.3684	.3449	.3254	.3088
	.01	.9744	.7656	.6465	.5654	.5088	.4663	.4331	.4058	.3833	.3643
	.005	.9743	.8164	.7012	.6172	.5566	.5117	.4756	.4463	.4219	.4004
6	.05	.7292	.5239	.4263	.3689	.3297	.3008	.2784	.2605	.2455	.2329
	.025	.8120	.6094	.5005	.4346	.3894	.3559	.3298	.3086	.2913	.2764
	.01	.9409	.6982	.5811	.5078	.4570	.4185	.3882	.3638	.3433	.3262
	.005	.9547	.7500	.6328	.5557	.5010	.4590	.4268	.4004	.3779	.3594
7	.05	.6694	.4775	.3889	.3365	.3009	.2745	.2541	.2377	.2241	.2125
	.025	.7588	.5571	.4575	.3972	.3560	.3252	.3013	.2820	.2659	.2524
	.01	.8447	.6416	.5332	.4648	.4180	.3828	.3550	.3325	.3140	.2979
	.005	.8887	.6924	.5810	.5098	.4590	.4209	.3906	.3662	.3457	.3286
8	.05	.6216	.4416	.3599	.3115	.2786	.2451	.2352	.2200	.2074	.1968
	.025	.7065	.5166	.4238	.3682	.3296	.3013	.2791	.2611	.2463	.2336
	.01	.7910	.5977	.4951	.4316	.3877	.3545	.3291	.3081	.2910	.2761
	.005	.8389	.6475	.5410	.4746	.4258	.3906	.3623	.3398	.3208	.3047
9	.05	.5884	.4128	.3367	.2913	.2605	.2377	.2200	.2058	.1940	.1841
	.025	.6680	.4839	.3970	.3445	.3086	.2817	.2610	.2444	.2305	.2186
	.01	.7529	.5615	.4644	.4048	.3633	.3320	.3081	.2886	.2725	.2583
	.005	.8008	.6094	.5078	.4443	.3994	.3662	.3398	.3184	.3008	.2852
10	.05	.5493	.3889	.3173	.2747	.2455	.2241	.2074	.1940	.1829	.1735
	.025	.6328	.4568	.3745	.3250	.2910	.2659	.2462	.2305	.2173	.2062
	.01	.7178	.5313	.4373	.3818	.3428	.3135	.2905	.2722	.2568	.2437
	.005	.7666	.5781	.4373	.4199	.3789	.3457	.3203	.3003	.2832	.2695

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APPENDIX-I

APL - PROGRAM USED TO CALCULATE CRITICAL VALUES OF THE
DISTRIBUTION OF THE AVERAGE SAMPLE CORRELATION.

The program consists of the following input:

- (1) $N_1 = 100$; the number of subintervals in the numerical integration.
- (2) N ; Number of points from which each correlation is calculated (3, 4, 5, 6, ..., 10).
- (3) K_1 ; Number of correlations in average (1, 2, 3, ..., 10).
- (4) S.G; Upper tail area (0.05, 0.025, 0.01, 0.005).

The output consists of the requisite critical values given below:

CRITICAL VALUES FOR DISTRIBUTION OF AVERAGE CORRELATION

VRBAR[□]▽
▽ RBAR

- [1] $X1 \leftarrow (0, N1) \div N1$
 - [2] $X2 \leftarrow (X1[1+ N1] + X1[N1]) \div 2$
 - [3] $K \leftarrow ((01) * 0.5) \times (!((N-3) \div 2)) \div (!((N-4) \div 2))$
 - [4] $I2 \leftarrow 0$
 - [5] $A1 \leftarrow 0$
 - [6] $X3 \leftarrow A1 + X2$
 - [7] $BV \leftarrow N1 \rho 0$
 - [8] $L \leftarrow 1$
 - [9] $U \leftarrow X3[L]$
 - [10] $BV[L] \leftarrow Kx + (1 \div N1) \times (20(Ux(20(0X2)))) \times ((100x2) * (N-3))$
 - [11] $L \leftarrow L+1$
 - [12] $\rightarrow (9x(L \leq N1)) + (13x(L > N1))$
 - [13] $I3 \leftarrow + / ((10RxK1xX3) \div X3) \times (BV * K1) \times (1 \div N1)$
 - [14] $I2 \leftarrow I2 + I3$
 - [15] $A1 \leftarrow A1 + 1$
 - [16] $J1 \leftarrow ((1 \div 01) \times I3) * 2$
 - [17] $L1 \leftarrow 10 * 8$
 - [18] $\rightarrow (6x(J1 \geq L1) \times (A1 < 100)) + (19x((J1 < L1) + (A1 \geq 100)))$
 - [19] $I1 \leftarrow 0.5 - (1 \div 01) \times I2$
- ▽

VRNDF[□]▽
▽ RNDF

- [1] $C \leftarrow ((L11 \times 10000) + (((10000 \times I1) - (L10000 \times I1)) > 0.5)) \div 10000$
- ▽

VRSRCH[□]▽
▽ RSRCH

- [1] $A \leftarrow 0$
 - [2] $B \leftarrow 1$
 - [3] $R \leftarrow (A+B) \div 2$
 - [4] R
 - [5] $RBAR$
 - [6] $RNDF$
 - [7] $\rightarrow (8x(C > SG)) + (10x(C < SG)) + (12x(C = SG))$
 - [8] $A \leftarrow R$
 - [9] $\rightarrow 3$
 - [10] $B \leftarrow R$
 - [11] $\rightarrow 3$
 - [12] R
- ▽

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