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The Stata Journal is published quarterly by the Stata Press, College Station, Texas, USA.

Address changes should be sent to the Stata Journal, StataCorp, 4905 Lakeway Drive, College Station, TX 77845, USA, or emailed to sj@stata.com.





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Merger simulation with nested logit demand

Jonas Björnerstedt Swedish Competition Authority Stockholm, Sweden jonas@bjornerstedt.org Frank Verboven
University of Leuven
Leuven, Belgium
frank.verboven@kuleuven.be

Abstract. In this article, we show how to implement merger simulation in Stata as a postestimation command, that is, after estimating an aggregate nested logit demand system with a linear regression model. We also show how to implement merger simulation when the demand parameters are not estimated but instead calibrated to be consistent with outside information on average price elasticities and profit margins. We allow for a variety of extensions, including the role of (marginal) cost savings, remedies (divestiture), and conduct different from Bertrand–Nash behavior.

Keywords: st0349, mergersim, merger simulation, aggregate nested logit model, unit demand and constant expenditures demand

1 Introduction

Competition and antitrust authorities have long been concerned with the possible anticompetitive effects of mergers. This is in particular the case for horizontal mergers, which are mergers between firms selling substitute products. The traditional concern has been that such mergers raise market power, which may hurt consumers and reduce total welfare (the sum of producer and consumer surplus). At the same time, however, it has been recognized that mergers may also result in cost savings or other efficiencies. While such cost savings may often be insufficient to reduce prices and benefit consumers, it has been shown that even small cost savings can be sufficient to raise total welfare (see Williamson [1968] and Farrell and Shapiro [1990]). Despite the possible total welfare gains, most competition authorities in practice take a consumer surplus standard when evaluating proposed mergers.

Merger simulation is increasingly used as a tool to evaluate the effects of horizontal mergers. Consistent with policy practice, the focus is often on the price and consumer surplus effects, but various applications also evaluate the effects on total welfare.² Merger simulation aims to predict the merger effects in the following three steps.

^{1.} According to Williamson's (1968) analysis, the deadweight loss from the output reduction after the merger is a second-order effect that is easily compensated by the cost savings from the merger. However, Posner (1975) argues that there is another source of inefficiency from mergers because firms must spend wasteful resources to make a merger and maintain market power. In this alternative view, it may be more natural to use consumer surplus as a standard to evaluate mergers and to ignore the transfer from consumers to firms.

Early contributions to the merger simulation literature are Werden and Froeb (1994), Nevo (2000), Epstein and Rubinfeld (2002), and Ivaldi and Verboven (2005). For a recent survey, see Budzinski and Ruhmer (2010).

The first step specifies and estimates a demand system, usually one with differentiated products. The second step makes an assumption about the firms' equilibrium behavior, typically multiproduct Bertrand-Nash, to compute the products' current profit margins and their implied marginal costs. The third step usually assumes that marginal costs are constant and computes the postmerger price equilibrium, accounting for increased market power, cost efficiencies, and perhaps remedies (such as divestiture). This enables one to compute the merger's effect on prices, consumer surplus, producer surplus, and total welfare. Stata is often used to estimate the demand system (the first step) but not to implement a complete merger simulation (including the second and third steps). In this article, we show how to implement merger simulation in Stata as a postestimation command, that is, after estimating the parameters of a demand system for differentiated products. We also illustrate how to perform merger simulation when the demand parameters are not estimated but rather calibrated to be consistent with outside industry information on price elasticities and profit margins. We allow for a variety of extensions, including the role of (marginal) cost savings, remedies (divestiture), and conduct different from Bertrand-Nash behavior.

We consider an oligopoly model with multiproduct price-setting firms that may partially collude and have constant marginal cost. Following Berry (1994), we specify the demand system as an aggregate nested logit model, which can be estimated with market-level data using linear regression methods (as opposed to the individual-level nested logit model). We consider both a unit demand specification, as in Berry (1994) and Verboven (1996), and a constant expenditures specification, as in Björnerstedt and Verboven (2013). The model requires a dataset on products sold in one market, or in a panel of markets, with information on the products' prices, their quantities sold, firm and nest identifiers, and possibly other product characteristics.

In section 2, we discuss the merger simulation model, including the nested logit demand system. In section 3, we introduce the commands required to carry out the merger simulation. Section 4 provides examples and section 5 concludes.

2 Merger simulation with an aggregate nested logit demand system

2.1 Merger simulation

Suppose there are J products, indexed by j = 1, ..., J. The demand for product j is $q_j(\mathbf{p})$, where \mathbf{p} is a $J \times 1$ price vector, and its marginal cost is constant and equal to c_j . Each firm f owns a subset of products F_f and chooses the prices of its own products $j \in F_f$ to maximize

$$\Pi_f(\mathbf{p}) = \sum_{j \in F_f} (p_j - c_j) q_j(\mathbf{p}) + \phi \sum_{j \notin F_f} (p_j - c_j) q_j(\mathbf{p})$$

where $\phi \in (0,1)$ is a conduct parameter to allow for the possibility that firms partially coordinate. If $\phi = 0$, firms behave noncooperatively as multiproduct firms. If $\phi = 1$,

they behave as a perfect, joint-profit maximizing cartel. A Bertrand-Nash equilibrium is defined by the following system of first-order conditions:

$$q_j(\mathbf{p}) + \sum_{k \in F_f} (p_k - c_k) \frac{\partial q_k(\mathbf{p})}{\partial p_j} + \phi \sum_{k \notin F_f} (p_k - c_k) \frac{\partial q_k(\mathbf{p})}{\partial p_j} = 0, \qquad j = 1, \dots, J \quad (1)$$

Let $\boldsymbol{\theta}$ be a $J \times J$ product-ownership matrix, with $\theta(j,k)=1$ if products j and k are produced by the same firm and $\theta(j,k)=\phi$ otherwise. If $\phi=0$ (no collusion), $\boldsymbol{\theta}$ becomes the usual block diagonal matrix; if all firms own only one product, $\boldsymbol{\theta}$ becomes the identity matrix. Furthermore, let $\mathbf{q}(\mathbf{p})$ be the $J\times 1$ demand vector, $\boldsymbol{\Delta}(\mathbf{p})\equiv \partial \mathbf{q}(\mathbf{p})/\partial \mathbf{p}'$ be the $J\times J$ Jacobian of first derivatives, and \mathbf{c} be the $J\times 1$ marginal cost vector. We can then write (1) in vector notation as

$$\mathbf{q}(\mathbf{p}) + \{ \boldsymbol{\theta} \odot \boldsymbol{\Delta}(\mathbf{p}) \} (\mathbf{p} - \mathbf{c}) = 0$$

This can be inverted to write price as the sum of marginal cost and a markup, where the markup term (inversely) depends on the price elasticities and on the product-ownership matrix:

$$\mathbf{p} = \mathbf{c} - \{\boldsymbol{\theta} \odot \boldsymbol{\Delta}(\mathbf{p})\}^{-1} \mathbf{q}(\mathbf{p})$$
 (2)

For single-product firms with no collusion ($\phi = 0$), the markup term is price divided by the own-price elasticity of demand. With multiproduct-firms and partial collusion, the cross-price elasticities also matter, and this increases the markup term (if products are substitutes).

Equation (2) serves two purposes. First, it can be rewritten to uncover the premerger marginal cost vector \mathbf{c} based on the premerger prices and estimated price elasticities of demand; that is,

$$\mathbf{c}^{\mathrm{pre}} = \mathbf{p}^{\mathrm{pre}} + \left\{ \boldsymbol{\theta}^{\mathrm{pre}} \odot \boldsymbol{\Delta}(\mathbf{p}^{\mathrm{pre}}) \right\}^{-1} \mathbf{q}(\mathbf{p}^{\mathrm{pre}})$$

Second, (2) can be used to predict the postmerger equilibrium. The merger involves two possible changes: a change in the product ownership matrix from $\boldsymbol{\theta}^{\text{pre}}$ to $\boldsymbol{\theta}^{\text{post}}$ and, if there are efficiencies, a change in the marginal cost vector from \mathbf{c}^{pre} to \mathbf{c}^{post} . To simulate the new price equilibrium, one may use fixed point iteration on (2), possibly with a dampening parameter in the markup term, or another algorithm such as the Newton method (see, for example, Judd [1998, 633]).

2.2 Nested logit demand system

The demand system $\mathbf{q} = \mathbf{q}(\mathbf{p})$ for the J products, $j = 1, \ldots, J$, is specified as a nested logit model with two levels of nests, referred to as groups and subgroups. This model belongs to McFadden's (1978) generalized extreme value discrete choice model. Consumers choose the alternative that maximizes random utility, which results in a specification for choice probabilities for each alternative. The nested logit model relaxes the independence of an irrelevant alternative property of the simple logit model and allows consumers to have correlated preferences for products that belong to the same subgroup or group. While discrete choice models were initially developed to analyze individual-level

data (see Train [2009] for an overview), Berry (1994) and Berry, Levinsohn, and Pakes (1995) show how to estimate the models with aggregate data. The dataset consists of $J \times 1$ vectors of the products' quantities \mathbf{q} , prices \mathbf{p} , and a $J \times K$ matrix of product characteristics \mathbf{x} , including indicator variables for the products' subgroup and group and their firm affiliation. The dataset is for either one market or a panel of markets, for example, different years or different regions and countries. The panel is not necessarily balanced, because new products may be introduced over time, or old products may be eliminated, and not all products may be for sale in all regions.

In addition to each product j's quantity sold q_j , its price p_j , and the vector of product characteristics x_j , it is necessary to observe (or estimate) the potential market size for the differentiated products. In the common unit demand specification of the nested logit, consumers have inelastic conditional demands: they buy either a single unit of their most preferred product $j=1,\ldots,J$ or the outside good j=0. The potential market size is then the potential number of consumers I, for example, an assumed fraction γ of the observed population in the market, $I=\gamma L$. An alternative is the constant expenditures specification, where consumers have unit elastic conditional demand: they buy a constant expenditure of their preferred product or the outside good. Here the potential market size is the potential total budget B, for example, an assumed fraction γ of total gross domestic product in the market, $B=\gamma Y$.

As shown by Berry (1994) and the extensions by Verboven (1996) and Björnerstedt and Verboven (2013), the aggregate two-level nested logit model gives rise to the following linear estimating equation for a cross section of products j = 1, ..., J:

$$\ln(s_i/s_0) = x_i\beta + \alpha \widetilde{p}_i + \sigma_1 \ln(s_{i|hg}) + \sigma_2 \ln(s_{h|g}) + \xi_i \tag{3}$$

A subscript t can be added to consider multiple markets or time periods, as in most empirical applications. The price variable is $\widetilde{p}_j = p_j$ in the unit demand specification, and $\widetilde{p}_j = \ln(p_j)$ in the constant expenditures specification. The variable s_j is the market share of product j in the potential market, $s_{j|hg}$ is the market share of product j in its subgroup h of group g, and $s_{h|g}$ is the market share of subgroup h in group g. More precisely, as discussed in more detail in Björnerstedt and Verboven (2013), the market shares are quantity shares in the unit demand specification

$$s_j = \frac{q_j}{I}, \quad s_{j|hg} = \frac{q_j}{\sum_{j \in H_{hg}} q_j}, \quad s_{h|g} = \frac{\sum_{j \in H_{hg}} q_j}{\sum_{h=1}^{H_{hg}} \sum_{j \in H_{hg}} q_j}$$

and they are expenditure shares in the constant expenditures specification

$$s_j = \frac{p_j q_j}{B}, \quad s_{j|hg} = \frac{p_j q_j}{\sum_{j \in H_{hg}} p_j q_j}, \quad s_{h|g} = \frac{\sum_{j \in H_{hg}} p_j q_j}{\sum_{h=1}^{H_{hg}} \sum_{j \in H_{hg}} p_j q_j}$$

where H_{hg} is the set (or number) of products of subgroup h of group g.

Furthermore, in (3), x_j is a vector of observed product characteristics, and ξ_j is the error term, which captures the product's quality that is unobserved to the econometrician. Equation (3) has the following parameters to be estimated: a vector of

mean valuations β for the observed product characteristics, a price parameter $\alpha < 0$, and two nesting parameters σ_1 and σ_2 , which measure the consumers' preference correlation for products in the same subgroup and group. The model reduces to a one-level nested logit model with only subgroups as nests if $\sigma_2 = 0$, to a one-level nested logit model with only groups as nests if $\sigma_1 = \sigma_2$, and to a simple logit model without nests if $\sigma_1 = \sigma_2 = 0$. The mean gross valuation for product j is defined as $\delta_j \equiv x_j \beta + \xi_j = \ln(s_j/s_0) - \alpha \tilde{p}_j - \sigma_1 \ln(s_{j|hg}) - \sigma_2 \ln(s_{h|g})$, so it can be computed from the product's market share, price, and the parameters α , σ_1 , and σ_2 .

In sum, the aggregate nested logit model is essentially a linear regression of the products' market shares on price, product characteristics, and (sub)group shares. In the unit demand specification, price enters linearly and market shares are in volumes; in the constant expenditures specification, price enters logarithmically and market shares are in values. In both cases, the unobserved product characteristics term, ξ_j , may be correlated with price and market shares, so instrumental variables should be used. Cost shifters would qualify as instruments, but these are typically not available at the product level. Berry, Levinsohn, and Pakes (1995) suggest using sums of the other products' characteristics (over the firm and the entire market). For the nested logit model, Verboven (1996) adds sums of the other product characteristics by subgroup and group.

3 The mergersim command

Various mergersim subcommands implement merger simulation as either commands before and after a linear nested logit regression to estimate α , σ_1 , and σ_2 or stand-alone commands where α , σ_1 , and σ_2 are specified by the user. With a panel dataset, one must time set the dataset before invoking the mergersim commands by using xtset id time or tsset id time, where id is the unique product identifier within the market, and time is the market identifier (time and region). Time setting is not required with a dataset for one market.

3.1 Syntax

```
mergersim init [if] [in], marketsize(varname) 
 \{\underline{q}uantity(varname) | \underline{p}rice(varname) | \underline{r}evenue(varname)\} [nests(varlist) \underline{unit}demand \underline{ces}demand alpha(\#) sigmas(\# [\#]) name(<math>string)]
```

mergersim market [if] [in], $\underline{\mathbf{f}}$ irm(varname) [conduct(#) name(<math>string)]

```
mergersim simulate [if][in], \underline{\text{firm}}(varname) {\underline{\text{buyer}}(\#) \underline{\text{seller}}(\#) | \underline{\text{newfirm}}(varname)} [conduct(\#) | \underline{\text{name}}(string) | \underline{\text{buyer}}(\#) sellereff(\#) \underline{\text{eff}}(iciencies(varname) | \underline{\text{newc}}(varname) | \underline{\text{newc}}(varname) | \underline{\text{newc}}(varname) | \underline{\text{newc}}(\#) | \underline{\text{method}}(\#) | \underline{\text
```

3.2 Options

Demand and market specification

The demand and market specification are set in mergersim init and mergersim market (and in mergersim simulate if mergersim market is not explicitly invoked by the user).

marketsize(varname) specifies the potential size of market (total number of potential buyers in unit demand specification, total potential budget in constant expenditures specification). marketsize() is required with mergersim init.

Any two of price(), quantity(), or revenue() are required.

quantity (varname) specifies the quantity variable.

price(varname) specifies the price variable.

revenue (varname) specifies the revenue variable.

nests(varlist) specifies one or two nesting variables. The outer nest is specified first. If only one variable is specified, a one-level nested logit model applies. If the option is not specified, a simple logit model applies.

unitdemand specifies the unit demand specification (default).

- cesdemand specifies the constant expenditure specification rather than the default unit demand specification.
- alpha(#) specifies a value for the alpha parameter rather than using an estimate. Note that this option has no effect if mergersim market has been run.
- sigmas(# [#]) specifies a value for the sigma parameters rather than using an estimate. In the two-level nested logit, the first sigma corresponds to the log share of the product in the subgroup, and the second corresponds to the log share of the subgroup in the group.
- name(string) specifies a name for the simulation. Variables created will have the specified name followed by an underscore character rather than the default M₋. This option can be used with all the mergersim subcommands.

- firm(varname) specifies the integer variable, indexing the firm owning the product. firm() is required with mergersim market and mergersim simulate.
- conduct(#) measures the fraction of the competitors' profits that firms account for when setting their own prices. It gives the degree of joint profit maximization between firms before the merger in percentage terms (number between 0 and 1).

Merger specification

The merger specification is set in mergersim simulate or in mergersim mre.

Either the identity of buyer and seller firms or the new ownership structure are required. The identity corresponds to the value in the variable specified with the firm() option.

buyer (#) specifies the buyer ID in the firm variable.

seller(#) specifies the seller ID in the firm variable.

newfirm(varname) specifies postmerger ownership in more detail than the buyer and seller options. For example, it can be used to simulate divestitures or two cumulative mergers by manually constructing a new firm ownership variable that differs from the firm variable specified with the firm() option.

Efficiency gains, in terms of percentage reduction in marginal costs, can be specified by either all seller and buyer products using the buyereff() and sellereff() options or product by product with the efficiencies() option.

- buyereff(#) specifies the efficiency gain of all products of the buyer firm after the merger. A value of 0 indicates no efficiency gain. The default is buyereff(0). For example, to incorporate a 10% efficiency gain, specify the buyereff(0.1) option.
- sellereff(#) specifies the efficiency gain of all products of the seller firm after the merger.
- efficiencies (varname) specifies a variable for efficiency gains more generally (that is, product by product), where, for example, 0.2 is a 20% decrease in marginal costs, and 0 is no change.

newcosts(varname) specifies a variable for postmerger costs.

newconduct(#) specifies the degree of joint profit maximization between firms after the merger, in percentage terms. With a conduct value of 1, the profits of other firms are as important as own profits.

Computation

The computation options can be set in mergersim simulate, where the postmerger Nash equilibrium is computed.

method(fixedpoint | newton) specifies the method used to find postmerger Nash equilibrium. The option can be specified as fixedpoint or newton. The default is method(newton). The Newton method starts with one iteration of the fixedpoint method.

maxit(#) specifies the maximum number of iterations in the solver methods.

dampen(#) specifies an initial dampening factor lower than the default dampen(1) in the fixed-point method. If fixedpoint does not converge, the method automatically tries a dampening factor of half the initial dampening.

Display and results

keepvars specifies that all generated variables should be kept after simulation, calculation of elasticities, or minimal required efficiencies.

detail shows market shares in mergersim simulate. These market shares are relative to total sales (excluding the outside good). Market shares are in terms of volumes for the unit demand specification and in terms of value for the constant expenditure specification. Changes in consumer and producer surplus and in the Herfindahl–Hirshman index are also displayed.

3.3 Description

mergersim performs a merger simulation with the subcommands init, market, and simulate. mergersim init must be invoked first to initialize the settings. mergersim market calculates the price elasticities and marginal costs. mergersim simulate performs a merger simulation, automatically invoking mergersim market if the command has not been called by the user. In addition to displaying results, mergersim creates various variables at each step. By default, the names of these variables begin with M.

First, mergersim init initializes the settings for the merger simulation. It is required before estimation and before a first merger simulation. It defines the upper and lower nests; the specification (unit demand or constant expenditures demand); the price, quantity, and revenue variables (two out of three); the potential market size variable; and the firm identifier (numerical variable). It also generates the variables necessary to estimate the demand parameters (alpha and sigmas) using a linear (nested) logit regression, similar to Berry (1994) and the extensions of Björnerstedt and Verboven (2013). The names of the market share and price variables to use in the regression will depend on the demand specification and are shown in the display output of mergersim init. Alternatively, the demand parameters can be calibrated with the alpha() and sigmas() options rather than being estimated.

Second, mergersim market computes the premerger conditions—the gross valuations δ_j and marginal costs c_j of each product j—under assumptions regarding the degree of coordination. The computations are based on the last estimates of α , σ_1 , and σ_2 unless they are overruled by values specified by the user in the alpha() and

sigmas() options. mergersim market is required after mergersim init and before the first mergersim simulate. It is not necessary to specify mergersim market before additional mergersim simulates (unless one wants to specify new premerger values of δ_i and c_i).

Third, mergersim simulate computes the postmerger prices and quantities under assumptions regarding the identity of the merged firms, their cost efficiencies, and the degree of collusion (the same as before the merger). It is possible to repeat the command multiple times after estimation.

In addition to these three main subcommands, several other subcommands can provide useful information. For example, mergersim mre computes the minimum required efficiencies per product for the price not to increase after the merger. It can be invoked after mergersim init.

4 Examples

4.1 Preparing the data

To demonstrate mergersim, we use the dataset on the European car market, collected by Goldberg and Verboven (2001) and maintained on their webpages.³ We take a reduced version of that dataset with fewer variables and a slightly more aggregate firm definition; the dataset is called cars1.dta. Each observation comprises a car model, year, and country. The total number of observations is 11,483: there are 30 years (1970–1999) and 5 countries (Belgium, France, Germany, Italy, and the United Kingdom), which implies an average of 77 car models per year and country. The car market is divided into five upper nests (groups) according to the segments: subcompact, compact, intermediate, standard, and luxury. Each segment is further subdivided into lower nests (subgroups) according to the origin: domestic or foreign (for example, Fiat is domestic in Italy and foreign in the other countries). Sales are new car registrations (qu). Price is measured in 1,000 Euro (in 1999 purchasing power). The product characteristics are horsepower (in kilowatts), fuel efficiency (in liter/100 kilometers), width (in centimeters), and height (in centimeters). The commands below are provided in a script called example.do.

^{3.} See http://www.econ.kuleuven.be/public/ndbad83/frank/cars.htm.

- . use cars1
- . summarize year country co segment domestic firm qu price horsepower fuel
- > width height pop ngdp

Variable	Obs	Mean	Std. Dev.	Min	Max
year	11483	1985.43	8.540344	1970	1999
country	11483	2.918488	1.443221	1	5
co	11483	223.0364	206.6172	1	980
segment	11483	2.559087	1.289577	1	5
domestic	11483	.1886267	.3912288	0	1
firm	11483	14.49769	8.567491	1	34
qu	11483	19911.44	37803.6	51	433694
price	11483	18.49683	8.922665	5.260726	150.3351
horsepower	11483	57.26393	23.89019	13	169.5
fuel	11483	6.728904	1.709702	3.8	18.6
width	11483	164.4574	9.567716	122	188
height	11483	140.4434	4.631175	117.5	173.5
pop	11483	4.81e+07	2.18e+07	9660000	8.21e+07
ngdp	11483	1.76e+14	4.73e+14	5.18e+10	2.13e+15

A first key preparatory task is to define the two dimensions of the panel and to time set the data (unless there is only one cross-section). The first dimension is the "product", that is, the car model (for example, Volkswagen [vw] Golf). The second dimension is the "market", which can be defined as the country and year (for example, France in 1995).

- . egen yearcountry=group(year country), label
- . xtset co yearcountry

panel variable: co (unbalanced)

time variable: yearcountry, 1 to 150, but with gaps

delta: 1 unit

Note that the panel is unbalanced because most models are not available throughout the entire period or in all countries.

A second key preparatory task is to define the potential market size. For the car market, it is sensible to adopt a unit demand specification. We specify the potential market size as total population divided by 4, a crude proxy for the number of households. In practice, the potential market size in a given year may be lower because cars are durable and consumers who just purchased a car may not consider buying a new one immediately.

. generate MSIZE=pop/4

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4.2 Performing a merger simulation

Merger simulation can now proceed in three steps.

Initializing the merger simulation settings

The first step initializes the settings for the merger simulation using the command mergersim init. The next example specifies a two-level nested logit model where the groups are the segments and the subgroups are domestic or foreign with the segments. This requires the option nests(segment domestic). The specification is the default unit demand specification. The price, quantity, market size, and firm variables are also specified.

```
. mergersim init, nests(segment domestic) price(price) quantity(qu)
> marketsize(MSIZE) firm(firm)

MERGERSIM: Merger Simulation Program
Version 1.0, Revision: 218

Unit demand two-level nested logit

Depvar Price Group shares
M_ls price M_lsjh M_lshg

Variables generated: M_ls M_lsjh M_lshg
```

merger init creates market share and price variables labeled with an M_ prefix (the default prefix). The variable M_ls is the dependent variable $\ln(s_j/s_0)$, M_lsjh is the log of the subgroup share $\ln(s_{j|hg})$, and M_lshg is the log of the group share $\ln(s_{h|g})$.

We can estimate the nested logit model with a linear regression estimator using instrumental variables to account for the endogeneity of the price and market share variables. As a simplification to illustrate the approach, we consider a fixed-effects regression without instruments.

```
. xtreg M_ls price M_lsjh M_lshg horsepower fuel width height domestic year
> country2-country5, fe
Fixed-effects (within) regression
                                                   Number of obs
                                                                             11483
Group variable: co
                                                   Number of groups
                                                                                351
R-sq: within = 0.8948
                                                   Obs per group: min =
                                                                                  1
       between = 0.7576
                                                                   avg =
                                                                               32.7
       overall = 0.8427
                                                                   max =
                                                                                146
                                                   F(13,11119)
                                                                           7271.50
corr(u_i, Xb) = -0.0147
                                                   Prob > F
                                                                            0.0000
        M_ls
                     Coef.
                             Std. Err.
                                             t
                                                   P>|t|
                                                              [95% Conf. Interval]
                 -.0468375
                              .0013002
                                         -36.02
                                                   0.000
                                                             -.0493861
                                                                         -.0442888
       price
                  .9047371
                              .0041489
                                         218.07
                                                   0.000
                                                             .8966045
                                                                          .9128696
      M_lsjh
                              .0085109
      M_lshg
                  .5677968
                                          66.71
                                                   0.000
                                                                          .5844796
                                                               .551114
  horsepower
                  .0038279
                              .0005921
                                           6.46
                                                   0.000
                                                              .0026672
                                                                          .0049886
                 -.0270919
                                          -5.97
                                                   0.000
                                                             -.0359892
        fuel
                              .004539
                                                                         -.0181946
                                                   0.000
       width
                  .0103757
                              .0016768
                                           6.19
                                                              .0070889
                                                                           .0136625
      height
                  .0004322
                              .0022161
                                           0.20
                                                   0.845
                                                             -.0039117
                                                                           .0047761
    domestic
                  .5230743
                              .0124205
                                          42.11
                                                   0.000
                                                             .4987279
                                                                           .5474206
        year
                  .0017336
                              .0012022
                                           1.44
                                                   0.149
                                                             -.000623
                                                                           .0040902
    country2
                 -.6621749
                                .01399
                                         -47.33
                                                   0.000
                                                            -.6895977
                                                                         -.6347521
    country3
                 -.5883123
                              .0147382
                                         -39.92
                                                   0.000
                                                             -.6172017
                                                                         -.5594229
    country4
                 -.7129762
                              .0137524
                                         -51.84
                                                   0.000
                                                             -.7399333
                                                                          -.686019
    country5
                 -.4155907
                               .016715
                                         -24.86
                                                   0.000
                                                             -.448355
                                                                          -.3828265
       _cons
                 -8.193457
                             2.246407
                                          -3.65
                                                   0.000
                                                            -12.59681
                                                                         -3.790101
     sigma_u
                 .52455749
     sigma_e
                 .36374004
                  .6752947
                              (fraction of variance due to u_i)
F test that all u_i=0:
                            F(350, 11119) =
                                                 22.69
                                                                 Prob > F = 0.0000
```

The parameters that will influence the merger simulations are the price parameter $\alpha = -0.0468$ and the nesting parameters $\sigma_1 = 0.905$ and $\sigma_2 = 0.568$ (the coefficients of, respectively, M_lsjh and M_lshg). These estimates satisfy the following restrictions from economic theory: $\alpha < 0$ and $1 > \sigma_1 \ge \sigma_2 \ge 0$. However, it is important to stress that the fixed-effects estimator is inconsistent because price and the subgroup and group market share variables are endogenous. As discussed in Berry (1994), an instrumental-variable estimator is required (for example, using ivreg or xtivreg with appropriate instruments). We therefore use only the results from the fixed-effects estimator for illustration.

Analyzing premerger market conditions

The second step in the merger simulation calculates the premerger market conditions (the products' gross valuations and their marginal costs and the price elasticities of demand) using the command mergersim market. In the example below, these calculations are done for only the five countries in 1998. Because no values for α , σ_1 , and σ_2 are specified, mergersim market uses the parameters in the last available Stata estimation, that is, the ones from a fixed-effects regression.

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. mergersim market if year == 1998

Supply: Bertrand competition

Demand: Unit demand two-level nested logit

Demand estimate

xtreg M_ls price M_lsjh M_lshg horsepower fuel width height domestic year

> country2-country5, fe
Dependent variable: M_ls

Parameters

alpha = -0.047

sigma1 = 0.905

sigma2 = 0.568

Own- and Cross-Price Elasticities: unweighted market averages

variable	mean	sd	min	max
M_ejj M_ejk M_ejl M_ejm	-7.488 0.766 0.068 0.001	3.761 1.276 0.120 0.002	-30.454 0.003 0.000 0.000	-1.710 10.908 0.768 0.011

Observations: 449

Pre-merger Market Conditions Unweighted averages by firm

firm code	price	Marginal costs	Pre-merger Lerner
BMW	20.194	17.499	0.146
Fiat	15.277	10.553	0.372
Ford	14.557	11.923	0.207
Honda	20.094	17.941	0.128
Hyundai	12.915	10.849	0.179
Kia	10.814	8.772	0.207
Mazda	14.651	12.557	0.156
Mercedes	25.598	21.569	0.162
Mitsubishi	15.955	13.825	0.145
Nissan	15.438	13.259	0.159
GM	21.054	18.633	0.135
PSA	16.243	13.533	0.194
Renault	15.518	12.837	0.203
Suzuki	9.289	7.226	0.234
Toyota	14.560	12.430	0.172
VW	18.990	16.388	0.181
Volvo	23.167	20.912	0.099
Daewoo	13.871	11.789	0.170

Variables generated: M_costs M_delta

These results imply fairly high own-price elasticities for the products in 1998, -7.488 on average. The cross-price elasticities are higher for products within the same subgroup (0.766) than for products of a different subgroup (0.068) and especially for products of a different group (0.001). The Lerner index or percentage markup over marginal cost varies from 9.9% to 37.2%, with a tendency of higher percentage markups for firms with lower-priced models (a feature of most unit demand-logit models).

Simulating the merger effects

The third step performs the actual merger simulation using the mergersim simulate command. The example below considers a merger where General Motors (GM) (firm = 15) sells its operations to VW (firm = 26). Note that the merger simulations would be the same if VW sold its operations to GM. We first carry out the merger simulations for Germany in 1998, where it can be considered a domestic merger (because GM sells the Opel brands, which are produced in Germany). It is assumed that there are no marginal cost savings to the seller or the buyer and that there is no partial coordination (neither before nor after the merger).

```
. mergersim simulate if year == 1998 & country == 3, seller(15) buyer(26)
> detail
```

Merger Simulation			
			Simulation method: Newton
	Buyer	Seller	Periods/markets: 1
Firm	26	15	Number of iterations: 6
Marginal cost savings			Max price change in last it: 4.5e-06

Prices
Unweighted averages by firm

Relative change	Post-merger	Pre-merger	firm code
0.003	18.002	17.946	BMW
0.000	15.341	15.338	Fiat
0.023	13.362	13.093	Ford
0.000	15.780	15.778	Honda
0.000	12.912	12.912	Hyundai
0.000	11.276	11.276	Kia
0.000	14.231	14.229	Mazda
0.003	20.155	20.114	Mercedes
0.000	15.834	15.832	Mitsubishi
0.000	15.103	15.101	Nissan
0.076	21.054	19.921	GM
0.000	16.399	16.397	PSA
0.000	15.295	15.292	Renault
0.000	9.225	9.225	Suzuki
0.000	13.020	13.019	Toyota
0.036	17.739	17.182	VW
0.000	22.154	22.149	Volvo
0.000	13.484	13.483	Daewoo

Variables generated: M_price2 M_quantity2 M_price_ch (Other M_ variables > dropped)

(output omitted)

The results show prices before and after the merger (in 1,000 Euro) and the percentage price change averaged by firm. This information is provided standard, even without the detail option at the end. The merger simulations predict that GM will on average raise its prices by 7.6%, while VW will on average raise its prices by 3.6%. The rivals respond with only very small price increases (with the exception of Ford).⁴

Because the new price vector is saved, one can use Stata's graphics to plot these results. Consider the following commands:

```
. generate perc_price_ch=M_price_ch*100
(11386 missing values generated)
. graph bar (mean) perc_price_ch if country==3&year==1998,
> over(firm, sort(perc_price_ch) descending label(angle(vertical)))
> ytitle(Percentage) title(Average percentage price increase per firm)
```

This produces the following plot:

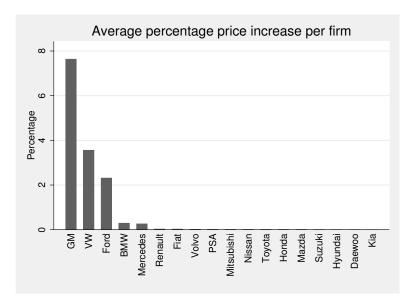


Figure 1. Average percentage price increase per firm after merger of GM and VW

^{4.} Note that one can also specify the detail option to display the market shares before and after the merger and the percentage point difference. If one is interested to see more detailed results, one can use additional options under mergersim results. One can also use standard Stata commands, such as table, based on the variables M_price (premerger price) and M_price2 (postmerger price).

Without the detail option after the mergersim simulate command, the output reports only the price information. The detail option produces additional results on the following variables (premerger, postmerger, and changes): market shares by firm, the Herfindahl index, C4 and C8 ratios (market share of 4 and 8 largest firms), and consumer and producer surplus.⁵

Market shares by quantity Unweighted averages by firm

firm code	Pre-merger	Post-merger	Difference
BMW	0.074	0.079	0.005
Fiat	0.043	0.045	0.003
Ford	0.095	0.132	0.037
Honda	0.012	0.012	0.001
Hyundai	0.006	0.006	0.000
Kia	0.003	0.003	0.000
Mazda	0.025	0.027	0.002
Mercedes	0.100	0.116	0.017
Mitsubishi	0.015	0.017	0.001
Nissan	0.025	0.027	0.002
GM	0.166	0.108	-0.058
PSA	0.034	0.037	0.003
Renault	0.051	0.054	0.003
Suzuki	0.006	0.006	0.000
Toyota	0.027	0.029	0.002
VW	0.300	0.280	-0.020
Volvo	0.012	0.013	0.001
Daewoo	0.006	0.007	0.001

	Pre-merger	Post-merger
HHS:	1501	1972
C4:	66.07	71.50
C8:	86.21	88.01

	Change
Consumer surplus:	-1,839,750
Producer surplus:	1,303,353

For example, the Herfindahl index increases from 1,501 to 1,972. Consumer surplus (in Germany) drops by 1.8 billion Euro or 586 Euro per car (because 3.1 million cars were sold in Germany in 1998). This is partly compensated by an increase in producer surplus of 1.3 billion Euro.

^{5.} In logit and nested logit models, consumer surplus (up to a constant) is given by the well-known log(sum) expression divided by the marginal utility of income. Caution is warranted in the constant expenditure specification because marginal utility is not constant. See Train (2009).

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4.3 Accounting for efficiencies, remedies, and partial collusion

It is possible to account for several specific features of the merger.

Efficiencies

First, one may account for the possibility that the buying or the selling firm benefits from a marginal cost saving, which may be passed on to consumer prices. The cost saving is expressed as a percentage of current marginal cost. In the command below, the options sellereff(0.2) and buyereff(0.2) mean that the seller and the buyer each have a marginal cost saving of 20% on all of their products.

```
. mergersim simulate if year == 1998 & country == 3, seller(15) buyer(26)
> sellereff(0.20) buyereff(0.20) method(fixedpoint) maxit(40) dampen(0.5)
```

Merger Simulation			
	Buyer	Seller	Simulation method: Dampened Fixed point Periods/markets: 1
Firm	26	15	Number of iterations: 19
Marginal cost savings	.2	.2	Max price change in last it: .

Prices			
${\tt Unweighted}$	averages	by	firm

firm code	Pre-merger	Post-merger	Relative change
BMW	17.946	17.703	-0.011
Fiat	15.338	15.265	-0.004
Ford	13.093	13.125	0.003
Honda	15.778	15.737	-0.002
Hyundai	12.912	12.908	-0.000
Kia	11.276	11.274	-0.000
Mazda	14.229	14.212	-0.001
Mercedes	20.114	19.259	-0.031
Mitsubishi	15.832	15.810	-0.001
Nissan	15.101	14.981	-0.005
GM	19.921	18.980	-0.022
PSA	16.397	16.372	-0.002
Renault	15.292	15.261	-0.003
Suzuki	9.225	9.219	-0.001
Toyota	13.019	13.005	-0.001
VW	17.182	15.717	-0.075
Volvo	22.149	22.036	-0.005
Daewoo	13.483	13.477	-0.000

Variables generated: M_price2 M_quantity2 M_price_ch (Other M_ variables > dropped)

There is now a predicted price decrease in Germany of -2.2% for GM and -7.5% for VW. This implies that the 20% cost savings are sufficiently passed to consumers. To obtain convergence, we used a fixed-point iteration with a dampening factor of 0.5 because the default Newton method did not converge. sellereff() and buyereff() assume the same percentage cost saving for all products of the seller and buyer. A

more flexible option is efficiencies(), which enables one to have product-specific percentage cost saving based on the variable that enters in efficiencies().

Instead of simulating the prices in the postmerger equilibrium with efficiencies, one can compute the minimum required efficiency (percentage cost saving by product) for the prices to remain unchanged after the merger; see Froeb and Werden (1998) or Röller, Stennek, and Verboven (2001). This can be done with the mergersim mre command:

. mergersim mre if year == 1998 & country == 3, seller(15) buyer(26)

max						
43.649						
43.620						
0.401						
Weighted average MRE: 0.221 Observations: 19						

The generated variable M_mre refers to the minimum required efficiency per product owned by the merging firms and is set to a missing value for the products of the non-merging firms. According to the results, the minimum required efficiencies for the 19 products of the merging firms are on average 12.3% (unweighted) and 22.1% (weighted by sales).

Divestiture as a remedy

Second, one may account for divestiture as a remedy to mitigate the price effects of a merger. Under such a remedy, the competition authority accepts the merger on the condition that the firms sell some of their products or brands. To simulate the effects of a merger with divestiture, one can replace the options buyer (#) and seller (#) with the option newfirm(varname), which specifies a variable for the new ownership structure after the merger. To illustrate, we consider a merger between Renault (firm = 18) and PSA (firm = 16), where PSA sells the brands Peugeot and Citroën. This merger would substantially raise average prices in France: 59.8% for the Renault products and 63.1% for the PSA products (ignoring entry and substitution to other countries). To mitigate the anticompetitive effects, the competition authority may request that PSA sell one of its brands, Citroën (brand = 4), to Fiat (firm = 4). The commands below show how to simulate the effects of such a merger with divestiture after creating the appropriate variable firm_rem for the new ownership structure.

^{6.} Note that this example starts with mergersim init and moves to mergersim simulate without performing a regression to obtain the price and nesting parameters. In this case, mergersim continues to use the most recent results.

Variable name

firm_rem

Simulation method: Newton Periods/markets: 1

Number of iterations: 7

Marginal cost savings Max price change in last it: 9.7e-08

Prices
Unweighted averages by firm

Ownership from:

firm code	Pre-merger	Post-merger	Relative change
BMW	18.342	18.347	0.000
Fiat	12.688	12.749	0.006
Ford	11.995	12.001	0.001
Honda	15.742	15.744	0.000
Hyundai	9.862	9.863	0.000
Kia	7.040	7.040	0.000
Mazda	12.536	12.536	0.000
Mercedes	25.239	25.240	0.000
Mitsubishi	14.880	14.880	0.000
Nissan	12.371	12.372	0.000
GM	18.963	18.966	0.000
PSA	15.303	16.317	0.089
Renault	14.996	17.114	0.162
Suzuki	7.824	7.824	0.000
Toyota	12.638	12.638	0.000
VW	17.735	17.744	0.001
Volvo	22.641	22.642	0.000
Daewoo	13.939	13.940	0.000

Variables generated: M_price2 M_quantity2 M_price_ch (Other M_ variables > dropped)

The results show that the merger with divestiture raises the average price only by 16.2% for Renault and by 8.9% for the Peugeot brand, whereas the price of Fiat (now including the Citroën brand) increases by 0.6%. The option newfirm(varname) can also be used for other applications, for example, to assess the impact of two consecutive mergers.

Conduct

Third, one may account for the possibility that firms partially coordinate, that is, take into account a fraction of the competitors' profits when setting prices. Assume, for example, that firms maintain the same degree of coordination before and after the merger: one can set the conduct parameter such that the markups are in line with outside estimates. Performing mergersim market before mergersim simulate enables one to verify whether the conduct parameter results in premerger markups in line with outside estimates. This is shown in the following example (which returns to the earlier merger between GM and VW in Germany).

```
. mergersim market if year == 1998 & country == 3, conduct(0.5)
```

Supply: Partial collusion, conduct = .5
Demand: Unit demand two-level nested logit

Demand estimate

xtreg M_ls price M_lsjh M_lshg horsepower fuel width height domestic year

> country2-country5, fe
Dependent variable: M_ls

Parameters

alpha = -0.047

sigma1 = 0.905

sigma2 = 0.568

Own- and Cross-Price Elasticities: unweighted market averages

variable	mean	sd	min	max
M_ejj	-6.907 0.781	2.876	-22.039	-3.339
M_ejk M_ejl	0.781	1.141 0.123	0.007 0.001	4.920 0.637
${ t M_ejm}$	0.001	0.002	0.000	0.011

Observations: 97

Pre-merger Market Conditions Unweighted averages by firm

firm code	price	Marginal costs	Pre-merger Lerner
BMW	17.946	13.079	0.290
Fiat	15.338	10.845	0.334
Ford	13.093	8.114	0.419
Honda	15.778	11.433	0.286
Hyundai	12.912	8.818	0.349
Kia	11.276	7.196	0.391
Mazda	14.229	10.012	0.315
Mercedes	20.114	13.753	0.348
Mitsubishi	15.832	11.612	0.280
Nissan	15.101	10.651	0.316
GM	19.921	14.862	0.297
PSA	16.397	12.106	0.299
Renault	15.292	10.893	0.340
Suzuki	9.225	5.084	0.461
Toyota	13.019	8.794	0.379
VW	17.182	12.104	0.352
Volvo	22.149	17.596	0.208
Daewoo	13.483	9.339	0.346

Variables generated: M_costs M_delta

The results show that if firms coordinate by taking into account 50% of the competitors' profits, then the Lerner index becomes almost twice as high as when there is no coordination. The predicted price effects after the merger can now be computed.

- . mergersim simulate if year == 1998 & country == 3, seller(15) buyer(26) > conduct(0.5)

Firm Marginal cost savings	Buyer 26	Seller 15	Simulation method: Newton Periods/markets: 1 Number of iterations: 6 Max price change in last it: 2.1e-07
Conduct:	Pre .5	Post .5	nar price change in race iv. 2.10 of

Unweighted averages by firm

firm code	Pre-merger	Post-merger	Relative change
BMW	17.946	18.125	0.011
Fiat	15.338	15.434	0.007
Ford	13.093	13.881	0.063
Honda	15.778	15.889	0.008
Hyundai	12.912	13.019	0.009
Kia	11.276	11.379	0.009
Mazda	14.229	14.334	0.008
Mercedes	20.114	20.427	0.025
Mitsubishi	15.832	15.956	0.008
Nissan	15.101	15.194	0.007
GM	19.921	21.171	0.084
PSA	16.397	16.503	0.007
Renault	15.292	15.395	0.008
Suzuki	9.225	9.314	0.010
Toyota	13.019	13.115	0.008
VW	17.182	17.947	0.049
Volvo	22.149	22.265	0.005
Daewoo	13.483	13.584	0.008
- I			

Variables generated: M_price2 M_quantity2 M_price_ch (Other M_ variables > dropped)

Under partial coordination, the merger simulation predicts larger price increases. On one hand, there is a larger predicted price increase for the merging firms: this feature does not hold generally, because the merging firms already partially coordinate before the merger. On the other hand, there is also a larger predicted price increase for the outsider firms: this feature may hold more generally because it reflects that outsiders have more cooperative responses to price changes by the merging firms.

4.4 Calibrating instead of estimating the price and nesting parameters

Calibration

The merger simulation results depend on the values of three parameters: α , σ_1 , and σ_2 (and on the price and quantity data per product). A practitioner may not want to rely too heavily on the econometric estimates of these parameters and may want to verify whether the elasticities and markups are consistent with external industry information. Here a practitioner would not estimate but "calibrate" the parameters such that they result in price elasticities and markups that are equal to external estimates. Such calibration is possible by specifying the options alpha() and sigmas() to mergersim market. The selected values overrule the values in memory, for example, the ones from a previous estimation. In the lines below, we specify $\alpha = -0.035$ (closer to 0 as compared with the econometric estimate of $\alpha = -0.047$), and we keep σ_1 and σ_2 to the previous values. Hence, we calibrate α such that demand would be less elastic. The results from this calibration indeed imply lower price elasticities (on average -5.5):

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. mergersim market if year == 1998 & country == 3

Supply: Bertrand competition

Demand: Unit demand two-level nested logit

Demand calibration

Parameters

alpha = -0.035

sigma1 = 0.910

sigma2 = 0.570

Own- and Cross-Price Elasticities: unweighted market averages

variable	mean	sd	min	max
M_ejj	-5.457	2.273	-17.430	-2.640
M_ejk	0.624	0.911	0.006	3.946
M_ejl	0.045	0.093	0.000	0.480
M_ejm	0.001	0.001	0.000	0.008

Observations: 97

Pre-merger Market Conditions Unweighted averages by firm

Pre-merger Lerner	Marginal costs	price	firm code
0.193	14.738	17.946	BMW
0.229	12.297	15.338	Fiat
0.287	9.765	13.093	Ford
0.189	12.921	15.778	Honda
0.223	10.294	12.912	Hyundai
0.248	8.681	11.276	Kia
0.206	11.455	14.229	Mazda
0.255	15.030	20.114	Mercedes
0.186	13.019	15.832	Mitsubishi
0.209	12.155	15.101	Nissan
0.199	16.573	19.921	GM
0.197	13.576	16.397	PSA
0.236	12.302	15.292	Renault
0.294	6.586	9.225	Suzuki
0.246	10.280	13.019	Toyota
0.254	13.540	17.182	VW
0.144	18.974	22.149	Volvo
0.220	10.860	13.483	Daewoo

Variables generated: M_costs M_delta

The next lines show what this calibration implies for merger simulation.

. mergersim simulate if year == 1998 & country == 3, seller(15) buyer(26)

Merger Simulation

Simulation method: Newton

Buyer Seller Periods/markets: 1

Firm 26 15 Number of iterations: 6

Marginal cost savings Max price change in last it: 5.9e-06

Prices
Unweighted averages by firm

firm code	Pre-merger	Post-merger	Relative change
BMW	17.946	18.018	0.004
Fiat	15.338	15.342	0.000
Ford	13.093	13.443	0.030
Honda	15.778	15.781	0.000
Hyundai	12.912	12.912	0.000
Kia	11.276	11.276	0.000
Mazda	14.229	14.231	0.000
Mercedes	20.114	20.167	0.003
Mitsubishi	15.832	15.835	0.000
Nissan	15.101	15.103	0.000
GM	19.921	21.372	0.098
PSA	16.397	16.399	0.000
Renault	15.292	15.296	0.000
Suzuki	9.225	9.226	0.000
Toyota	13.019	13.020	0.000
VW	17.182	17.892	0.045
Volvo	22.149	22.155	0.000
Daewoo	13.483	13.484	0.000

Variables generated: M_price2 M_quantity2 M_price_ch (Other M_ variables > dropped)

These results show that the predicted price increase is larger when demand is less elastic.

Application: Bootstrapping confidence intervals

One can also use the calibration options alpha() and sigmas() to implement a parametric bootstrap for constructing confidence intervals of the computed merger effects. The following lines perform three steps. First, we take 100 draws for α , σ_1 , and σ_2 assuming the parameters are normally distributed. Second, we perform 100 merger simulations for each draw. Third, we save the results for the average price increase of the buying firm and the selling firm, and we compute summary statistics.

```
. quietly mergersim init, nests(segment domestic) price(price) quantity(qu)
> marketsize(MSIZE) firm(firm)
. matrix b=e(b)
. matrix V=e(V)
. matrix bsub = (b[1,1], b[1,2], b[1,3])
. matrix Vsub = (V[1,1], V[1,2], V[1,3] \setminus V[2,1], V[2,2], V[2,3] \setminus V[3,1],
> V[3,2], V[3,3])
. local ndraws 100
. set seed 1
. preserve
. drawnorm alpha sigma1 sigma2, n(`ndraws´) cov(Vsub) means(bsub) clear
(obs 100)
. mkmat alpha sigma1 sigma2, matrix(params)
. restore
. matrix pr_ch = J(`ndraws´,2,0)
. forvalues i = 1 2 to `ndraws' {
  2. local alpha = params[`i´,1]
  3. local sigma1 = params[`i´,2]
  4. local sigma2 = params['i',3]
  5. quietly mergersim init, nests(segment domestic) price(price) quantity(qu)
> marketsize(MSIZE) firm(firm) alpha(`alpha´) sigmas(`sigma1´ `sigma2´)
  6. quietly mergersim simulate if year == 1998 & country == 3, seller(15)
> buyer(26)
  7. sum M_price_ch if year == 1998 & country == 3&firm==15, meanonly
  8. matrix pr_ch[`i´,1] = r(mean)
  9. sum M_price_ch if year == 1998 & country == 3&firm==26, meanonly
 10. matrix pr_ch[`i´,2] = r(mean)
 11. }
. clear
. quietly svmat pr_ch , names(pr_ch)
. sum pr_ch1 pr_ch2
    Variable
                     Obs
                                 Mean
                                         Std. Dev.
                                                          Min
                                                                     Max
                      100
                             .0763034
                                          .0031552
                                                     .0667844
                                                                .0844396
      pr_ch1
                      100
                             .0355618
                                         .0015778
                                                     .0307121
                                                                .0394875
      pr_ch2
```

Earlier, we obtained point estimates for the percentage price increase of 7.6% for GM and 3.6% for VW (for the base scenario). The 95% confidence intervals for these price increases are [6.7–8.4]% for GM and [3.1–4.0]% for VW.

4.5 Constant expenditures demand

We can finally illustrate how to do merger simulation based on a constant expenditures demand instead of a unit demand specification. For cars, this may not be a realistic option, because consumers typically buy one unit or no unit rather than constant expenditures. Nevertheless, we can use the constant expenditures specification to see how functional form affects the predictions from merger simulation.

First, we need to define the potential market size.

. generate MSIZE1=ngdpe/5

This assumes the potential expenditures on cars in a country and year are 20% of total gross domestic product.

Next we calibrate (rather than estimate) the parameters to $\alpha = -0.5$, $\sigma_1 = 0.9$, and $\sigma_2 = 0.6$.

```
. mergersim init, nests(segment domestic) ces price(price) quantity(qu)
> marketsize(MSIZE1) firm(firm) alpha(-0.5) sigmas(0.9 .6)
  (output omitted)
```

We can verify the premerger elasticities and markups at these calibrated parameters:

```
. mergersim market if year == 1998 & country == 3
```

```
Supply: Bertrand competition
```

Demand: Constant expenditure two-level nested logit

Demand calibration

Parameters

alpha = -0.500

sigma1 = 0.900

sigma2 = 0.600

Own- and Cross-Price Elasticities: unweighted market averages

variable	mean	sd	min	max
M_ejj	-5.574	0.493	-5.995	-4.054
M_ejk	0.426	0.493	0.005	1.946
M_ejl	0.039	0.065	0.000	0.283
M_ejm	0.001	0.001	0.000	0.006

Observations: 97

Pre-merger Market Conditions Unweighted averages by firm

firm code	price	Marginal costs	Pre-merger Lerner
BMW	17.946	14.375	0.194
Fiat	15.338	12.451	0.189
Ford	13.093	10.502	0.202
Honda	15.778	12.938	0.180
Hyundai	12.912	10.732	0.169
Kia	11.276	9.384	0.168
Mazda	14.229	11.684	0.177
Mercedes	20.114	14.228	0.260
Mitsubishi	15.832	12.978	0.180
Nissan	15.101	12.281	0.183
GM	19.921	15.784	0.206
PSA	16.397	13.473	0.179
Renault	15.292	12.504	0.188
Suzuki	9.225	7.661	0.170
Toyota	13.019	10.739	0.175
VW	17.182	13.395	0.221
Volvo	22.149	17.606	0.201
Daewoo	13.483	11.201	0.169

Variables generated: M_costs M_delta

The premerger elasticities and markups are roughly comparable with the ones of the estimated unit demand model (with less variation between firms). However, as shown below, the merger simulation results in a larger predicted price increase: +10.1% for GM and +4.4% for VW. This follows from the different functional form: the constant expenditures specification has the property of quasi-constant price elasticity, whereas the unit demand specification has the property that consumers become more price sensitive as firms raise prices. For this same reason, efficiencies in the form of marginal cost savings would also be passed more to consumers under this specification.

. mergersim simulate if year == 1998 & country == 3, seller(15) buyer(26)
> detail

Merger Simulation

Buyer Seller Periods/markets: 1
Firm 26 15 Number of iterations: 7

Marginal cost savings Max price change in last it: 4.7e-09

Prices
Unweighted averages by firm

firm code	Pre-merger	Post-merger	Relative change
BMW	17.946	18.021	0.004
Fiat	15.338	15.342	0.000
Ford	13.093	13.302	0.017
Honda	15.778	15.781	0.000
Hyundai	12.912	12.912	0.000
Kia	11.276	11.276	0.000
Mazda	14.229	14.231	0.000
Mercedes	20.114	20.155	0.003
Mitsubishi	15.832	15.835	0.000
Nissan	15.101	15.103	0.000
GM	19.921	21.581	0.101
PSA	16.397	16.399	0.000
Renault	15.292	15.295	0.000
Suzuki	9.225	9.225	0.000
Toyota	13.019	13.020	0.000
VW	17.182	17.933	0.044
Volvo	22.149	22.159	0.000
Daewoo	13.483	13.484	0.000

Variables generated: M_price2 $M_quantity2$ M_price_ch (Other $M_variables > dropped)$

(output omitted)

Because the detail option was added, mergersim simulate reports additional results. Consumer surplus now drops by -2.2 billion Euro (versus -1.8 billion Euro in the unit demand specification), and producer surplus increases by 1.1 billion Euro (versus 1.3 billion Euro before).

	Pre-merger	Post-merger
HHS:	1501	1906
C4:	66.07	70.52
C8:	86.21	87.61

	Change
Consumer surplus:	-2,190,399
Producer surplus:	1,140,647

5 Conclusions

This overview has shown how to apply two specifications of the two-level nested logit demand system to merger simulation. We show that merger simulation can be applied as a postestimation command based on estimated parameter values, or it can be implemented without estimation but with calibrated parameters. The merger simulation results yield intuitive predictions given the assumed demand parameters. The set of merger simulation commands can be used to simulate the effects of horizontal mergers in a standard setting (differentiated products, multiproduct Bertrand price setting). One can also incorporate various extensions, including efficiencies in the form of cost savings, remedies through partial divestiture, and alternative behavioral assumptions (partial collusive behavior).

Other applications and extensions could be considered. For example, for the car market, it could be interesting to generalize the demand model to allow consumers to substitute between countries by introducing an upper nest for the choice of country instead of assuming such substitution is not possible. These additional substitution possibilities would limit the market power effects of mergers. Other demand models may also be considered, such as a random coefficients logit model or the almost ideal demand system.

6 References

- Berry, S., J. Levinsohn, and A. Pakes. 1995. Automobile prices in market equilibrium. *Econometrica* 63: 841–890.
- Berry, S. T. 1994. Estimating discrete-choice models of product differentiation. *RAND Journal of Economics* 25: 242–262.
- Björnerstedt, J., and F. Verboven. 2013. Does merger simulation work? Evidence from the Swedish analgesics market.
 - http://www.econ.kuleuven.be/public/ndbad83/Frank/Papers/Bjornerstedt%20&%20Verboven,%202013.pdf.
- Budzinski, O., and I. Ruhmer. 2010. Merger simulation in competition policy: A survey. Journal of Competition Law & Economics 6: 277–319.
- Epstein, R. J., and D. L. Rubinfeld. 2002. Merger simulation: A simplified approach with new applications. *Antitrust Law Journal* 69: 883–919.
- Farrell, J., and C. Shapiro. 1990. Horizontal mergers: An equilibrium analysis. *American Economic Review* 80: 107–126.
- Froeb, L. M., and G. J. Werden. 1998. A robust test for consumer welfare enhancing mergers among sellers of a homogeneous product. *Economics Letters* 58: 367–369.

^{7.} We stress, however, that the estimated parameters were based on an inconsistent fixed-effects estimator. In practice, one should use instrumental variables to estimate the parameters consistently.

- Goldberg, P. K., and F. Verboven. 2001. The evolution of price dispersion in the European car market. Review of Economic Studies 68: 811–848.
- Ivaldi, M., and F. Verboven. 2005. Quantifying the effects from horizontal mergers in European competition policy. *International Journal of Industrial Organization* 23: 669–691.
- Judd, K. L. 1998. Numerical Methods in Economics. Cambridge, MA: MIT Press.
- McFadden, D. 1978. Modelling the choice of residential location. In *Spatial Interaction Theory and Planning Models*, ed. A. Karlqvist, L. Lundqvist, F. Snickars, and J. Weibull, 75–96. Amsterdam: North-Holland.
- Nevo, A. 2000. Mergers with differentiated products: The case of the ready-to-eat cereal industry. RAND Journal of Economics 31: 395–421.
- Posner, R. A. 1975. The social costs of monopoly and regulation. *Journal of Political Economy* 83: 807–828.
- Röller, L.-H., J. Stennek, and F. Verboven. 2001. Efficiency gains from mergers. *European Economy* 5: 31–128.
- Train, K. E. 2009. Discrete Choice Methods with Simulation. 2nd ed. Cambridge: Cambridge University Press.
- Verboven, F. 1996. International price discrimination in the European car market. RAND Journal of Economics 27: 240–268.
- Werden, G. J., and L. Froeb. 1994. The effects of mergers in differentiated products industries: Logit demand and merger policy. *Journal of Law, Economics, and Organization* 10: 407–426.
- Williamson, O. E. 1968. Economies as an antitrust defense: The welfare tradeoffs. *American Economic Review* 58: 18–36.

About the authors

Jonas Björnerstedt is a researcher at the Swedish Competition Authority and at the University of Leuven (Belgium). His current research focuses on the empirical analysis of competition policy and industrial organization.

Frank Verboven is a professor of economics and industrial organization at the University of Leuven (Belgium) and research fellow at the Centre for Economic Policy Research. His current research focuses on the empirical analysis of industries with market power, with applications to issues in competition policy and regulation.