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DETERMINANTS OF FOOD INDUSTRY PERFORMANCE – SURVEY DATA AND REGRESSIONS FOR DENMARK

Hartley Furtan and Johannes Sauer***

Abstract

This paper empirically investigates the determinants of firms' performance in the agri-food sector by using recent survey data for Denmark. Treating sales per employee as a proxy for value addition we estimate several bootstrapped regression models to draw conclusions on the marginal effects of potential performance determinants such as the form and nature of ownership, stage of the food chain and commodity sector, new product development, staff quality, firms' competitive stance, and elements of firms' strategy. To draw robust inferences we apply, besides the ordinary heteroscedasticity corrected Tobit ML-estimator, a nonparametric least absolute deviations estimator (LAD/CLAD) based on a quantile regression procedure. The results indicate that we cannot reject the hypothesis of no influence of dominant orientation on value added. Rather, firms' focus on human capital, stage and commodity sector better explains their value addition. We can reject the hypothesis that regional networks have no influence on value added. Location in Århus, emphasis on human capital and the negative influence of outsourcing on value added all provide supporting evidence. We reject the hypothesis of no influence of FDI, and moreover propose that FDI has targeted the domestic Danish market as a source of value added.

Keywords

value added, innovation, organizational type

1 Introduction

There is a perceived need to increase the rate at which food industry firms add value to products. For example, a recent paper by TRAILL and MUELENBERG (2002) lists a number of hypotheses as to the relationship between innovation and firm performance in the food industry. There has been surprisingly little research into the attributes of firms that can, do, or might deliver such value added. Moreover, such investigation often has used case studies with limited recourse to statistics and modelling of firm behaviour. The current paper uses a recent survey of Danish food industry firms¹ to identify relationships between value added and the attributes and behaviour of the firms. The models developed are used to test hypotheses drawn from the literature.

CHRISTY and CONNOR (1989) have proposed that changes in value addition are associated with structural and behavioural change in the food marketing system. Focusing on structure, ROGERS (2001) found that market shares of the largest firms in the US food industry were highly correlated with those large firms' shares of industry value added. CONVERSELY, GOULD and CARSON (1998) as well as BUHR (2004) both identified small firms' value-adding activities that are essentially defensive: adding value as an alternative to cost and logistics saving through economies of scale. BUHR's interviews with food industry firms revealed that

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they view product differentiation as being key to adding value. Branding (as a product differentiation tool) was investigated by BAKER et al. (2006) and BRESTER and SCHROEDER (1995) for linkages to value addition, and in both cases the relationship was found to be associated with vertical relationships in the marketing channels. BRESTER (1999) examined cases of vertical integration as a means of adding value, and found linkages to a broad range of management variables including human resource management, asset acquisition and the form and extent of sales growth.

COLTRAIN et al. (2000) examined a selection of firms with various differentiation strategies, and defined “innovation” and “co-ordination” as the two main sources of value addition. In that study, these two activities tended to focus on relations within the chain. BOSWORTH and LOUNDES (2002) also found strong links between several forms of innovation within the marketing channel and value added. Given the apparent significance of innovation in value addition, it is somewhat surprising that research and development (R&D) actions and expenditures have generally not been found to be closely associated with value addition (HESHMATI and PIETOLA, 2004; BOSWORTH and LOUNDES, 2002). JUNGnickel et al. (2004) and WALKENHORST (2001) examined relationships between international features of firms (foreign direct investment (FDI), export performance, staff located abroad) and measures of productivity and value added. Both found a positive relationship between value added and FDI, and between value added and exports. Those studies used publicly-available databases that did not contain information on some key variables on these themes, such as the characteristics of staff and the use of outsourcing. Strong links between value added and choice of marketing channel, and channel relationships, have been identified by BROWN (1995, in a Canadian study) and SONOBE et al. (2004, in a Chinese study across several sectors). BAKER et al. (2006) found evidence that value generated at one stage of the chain might be expropriated by other firms by applications of branding behaviour, particularly where such behaviour was related to interactions between retailers and other firms. LAWRENCE et al. (1997) found that inter-stage relationships were important in value addition, particularly when allocation of production, processing, distribution and retailing space were concerned.

This paper is organized into seven sections. The second section is a brief overview of the Danish food sector. This section is included because our data comes from a survey of the Danish food sector. The third section is made up of the economic model and hypothesis while the fourth section reports the data and estimation procedures we follow. The fifth and sixth section reports the results and a discussion of the results. The results indicate support for our theoretical model of innovation and for innovation in the Danish food sector. The final section is the conclusions.

2 Danish Food Sector

There has been substantial consolidation in Danish food sector in the period 1995-2000 (BAKER, 2003). The reduction in number of food processing firms has been much faster in Denmark than in other parts of Europe. Similarly the reduction in wholesale firms in the food industry has reduced faster in Denmark than most other EU countries. In terms of concentration the Danish food sector is similar to other parts of Europe. Denmark has CR5 of about 56% while the CR4 in the US is about 27% at the national level. However when buying power is considered the Danish effective CR is around 70%. This figure puts Denmark pretty much in line with the rest of the EU. One strong trend that has occurred in the Danish food marketing chains is the increase in the share of the wholesale market controlled by non-specialized stores. This trend is apparent in most EU countries, but Denmark has the largest increase. In contrast the US has shown a strong increase in all types of food retailing outlets.

We use the Danish food sector as our case study because of the data available from BAKER (2003). Also the Danish economy has been rated as one of the most innovative in the world by agencies like the World Economic Forum (2007). Agriculture and food make up a significant portion of the economy in Denmark thus we expect the food sector to be innovative and yield robust conclusions to our tests for innovative activity.

3 Model and Hypotheses

A national ideas production function has been suggested by CHELLARAJ et al. (2005) as one means to track the influence of ideas on productivity. The CHELLARAJ et al. (2005) model builds on the path breaking work by ROMER (1990). In this paper we suggest that a production function which includes non-traditional inputs such as research expenditures, institutional forms, and foreign direct investment is a useful construct to model the question of how innovation activity affects the food sector. Most often the hypotheses that are tested in a paper such as this one are generated from the derived comparative statics of the theoretical model. In our case we do not have useful comparative statics thus we need to generate our hypothesis elsewhere. To deal with this concern we take hypothesis regarding innovation in the food sector from other research that is readily available in the literature (e.g. TRAIL and MUELENBERG, 2002). We assume that the firm wants to maximize expected profit subject to a technology constraint. We also assume that the firm has no pricing power in either the output or input market. We specify this as $MaxE(\Pi)$ subject to $NY = h[X, A(X, R, C, F)] + \varepsilon$ where expected profit is $E(\pi)$, N is a vector of the number of products, Y is a vector of output levels, X is a vector of traditional inputs, R is the level research expenditure, C is a vector of contractual arrangements, F is foreign direct investment, and ε is a random variable. In this specification the $A(\cdot)$ is an innovation function.

An increase in research expenditures will increase the level of innovation. This has been shown extensively in the agricultural economics literature and has been the main focus of numerous studies estimating rates of return. The impact of increasing the level of contractual arrangements is to make the market more efficient by reducing transaction costs. Finally the level of foreign direct investment is thought to increase the level of innovation. Foreign investment brings with it new techniques, opens up markets, and increases the level of business capital (contacts). A change in the level of $A(\cdot)$ shifts the function in a non parallel fashion. The large the level of traditional inputs used by the firm the greater is the impact of innovation on the total output value of the firm. The change in $A(\cdot)$ can increase the level of output of existing outputs (Y) or it can increase the number of products produced (N) or both (NY). For this reason the first order conditions for profit maximization are not tractable. We list the hypothesis which we have taken from the literature and we test using three different econometric models are shown in table 1. To the best of our knowledge none of these have been, to date, tested using firm-level data from the food industry.

In testing these hypotheses we use three different econometric models. We approach hypothesis 1-4 by specifying the following functional relationship (model 1 and model 2). In model 1:

$$VA_{it} = f(emp1_{it}, emp2_{it}, dFDI_{it}, dR_{ijt}, dCO_{it}, dMS_{ikt}, dFS_{ikt}, dO_{int}) \quad (1)$$

where VA_{it} is value added or a proxy for profits measured by sales per employee, $emp1_{it}$ is the percentage of employees outside Denmark, $emp2_{it}$ is the percentage of employees with university education, $dFDI_{it}$ is a dummy variable for foreign direct investment in the firm, dR_{ijt} is a dummy variable for location in region j .² The variable dCO_{it} is a dummy variable for

² $j \in \left\{ \begin{array}{l} \text{København, Århus, Sønderjylland, Fyn, Viborg, Nordjylland, Vejle, Storstrøms, Frederiksborg,} \\ \text{Ringkøbing, Ribe, Roskilde, Vestsjælland} \end{array} \right\}$.

cooperative form, dMS_{it} a dummy variable for marketing stage k^3 and the dummy variable dFS_{it} denotes firm i 's operation in sector l .⁴ Finally we include a dummy variable to denote firm i 's dominant orientation m .⁵

We do not measure the innovation activity of the firm directly rather we anticipate that more innovative firms (i.e. introducing more new products) will generate more value added. Intuition concerning human capital leads us to expect that the percentage of employees with a university education has a positive influence on value added. The percentage of employees outside Denmark is (a proxy for off-shoring), as a cost-lowering activity, is expected to increase value added. We expect that firms with some foreign direct investment will be more productive, and so exhibit higher value added than firms without foreign direct investment. We anticipate a cluster or agglomeration effect, which will be indicated by the location dummy variable, and following (ASHEIM and COENEN 2005) expect that the co-operative form will have a negative influence on value added. We expect retail and processing firms to show greater value-added than other firms. We have no particular expectations about different levels of value addition by separate commodity sectors. Our examination of co-ordination within the marketing chain centres on vertical integration, which is expected to be positively associated with value added. This effect may, however, be difficult to detect due to its association with other factors. Model 2 uses a different estimator but the identical variables as model 1. To test hypothesis 5 we propose a second model (model 3)

$$\Delta VA_{it,t-1} = f(\Delta emp1_{it,t-1}, \Delta emp2_{it,t-1}, \Delta dFDI_{it,t-1}, \Delta dO_{im,t-1}, \Delta dCO_{it,t-1}, R_i, MS_i, FS_i,) \quad (2)$$

where Δ denotes the difference of (t)-(t-1) for each variable, with $t = 2005$ and $t-1 = 2000$. Model 2 allows us to test for the variables' marginal contribution to growth in value added for each of the firms. In addition to the variables defined above, a selection of other variables was included in the models at various stages of specification. The variable $\Delta dCO_{it,t-i}$ drops out because there was no change in the number of firms of the cooperative type over the time period. Only the most statistically significant variables, as well as those important for testing the hypotheses listed in table 1, remained in the final models' specification.

Table 1: Hypotheses tested

	Hypothesis	Reference
1	There is no relationship between company size and innovation	Traill and Muelenberg
2	Successful firms do not have a single dominant orientation to product, process, nor market	Traill and Muelenberg
3	Regional innovation systems do not have an impact on a firm's performance	Asheim and Coenen
4	Foreign Direct Investment has as no influence on a firm's performance	Aitken and Harrison
5	In product and market oriented firms, new product development does not drive firm growth.	Traill and Muelenberg

4 Data and Estimation Procedure

The data used in this study are drawn from a survey of 444 Danish food industry (non-farm) firms.⁶ To identify target firms, a commercial database of firms' contact details was

³ $k \in \{\text{primary, processing, wholesale, retail, ingredients}\}$

⁴ $l \in \{\text{fruits and vegetables, dairy, pork, poultry, meat, unspecialised}\}$

⁵ $m \in \{\text{market, process, product}\}$

⁶ Further details of the survey and data are available from the authors.

purchased, with stratified sampling based on size (across size groups but excluding firms with less than 5 employees) and sector (just 8 sectors included), and across three stages of the marketing chain (retail, wholesale and processing). After eliminating defunct firms, incorrect contact details and subsidiaries of other firms in the sample, telephone contacts were made with 444 firms. The survey procedure yielded 131 valid responses (a 30% response rate on 444 firms). The survey questionnaire addressed several elements of strategy and behaviour, and sought responses for 2005 (the current year) and 2000 (in retrospective). The telephone interview-based survey was conducted between November 2005 and March 2006. Descriptive statistics for the data set employed in the models can be obtained from the authors.

Table 2: Descriptive Statistics

<i>Variable</i>	<i>Mean (Median)</i>	<i>Stdev</i>	<i>Min</i>	<i>Max</i>
sales per employee (mill DKK)	4.16 (1.41)	10.67	0.01	133.33
number of new products introduced (n)	93.15 (50)	417.86	0	5000
percentage of employees outside Denmark (%)	2.88 (2.12)	11.36	0	76
percentage of employees with university degree (%)	5.67 (2.67)	14.66	0	100
foreign direct investment in the firm (dummy: 1 - yes, 0 - no)	0.12	0.33	0	1
regional location of the firm (dummy: 1 - yes, 0 - no)				
- københavn (12 obs)	0.16	0.37	0	1
- århus (24 obs)	0.09	0.29	0	1
- sønderjylland (22 obs)	0.08	0.28	0	1
- fyn (36 obs)	0.14	0.34	0	1
- viborg (16 obs)	0.06	0.24	0	1
- nordjylland (26 obs)	0.10	0.29	0	1
- vejle (32 obs)	0.12	0.33	0	1
- storstrøms (6 obs)	0.02	0.15	0	1
- frederiksborg (8 obs)	0.03	0.17	0	1
- ringkøbing (6 obs)	0.02	0.15	0	1
- ribe (6 obs)	0.02	0.14	0	1
- roskilde (10 obs)	0.04	0.19	0	1
- vestsjælland (8 obs)	0.03	0.17	0	1
ownership of/by the firm* (dummy: 1 - yes, 0 - no)				
- owned by a farmer/farmer cooperative (22 obs)	0.08	0.28	0	1
- owned by a non-food firm (41 obs)	0.16	0.37	0	1
- ownership of a retail outlet (5 obs)	0.02	0.14	0	1
- owned by distributor/wholesaler (6 obs)	0.04	0.19	0	1
- ownership of a distributor/wholesaler (10 obs)	0.06	0.25	0	1
- ownership of a processor (11 obs)	0.04	0.20	0	1
- owned by services (6 obs)	0.02	0.15	0	1
- owned by ingredients (1 obs)	0.01	0.06	0	1
marketing stage of the firm (dummy: 1 - yes, 0 - no)				
- primary (8 obs)	0.03	0.17	0	1
- processing (108 obs)	0.41	0.49	0	1
- wholesale (78 obs)	0.29	0.46	0	1
- retail (54 obs)	0.21	0.41	0	1
- ingredients (10 obs)	0.04	0.19	0	1
sector of the firm (dummy: 1 - yes, 0 - no)				
- feeding (2 obs)	0.01	0.09	0	1
- fruits & vegetables (24 obs)	0.09	0.29	0	1
- dairy (46 obs)	0.18	0.38	0	1
- beef (10 obs)	0.04	0.19	0	1
- pork (6 obs)	0.02	0.15	0	1
- poultry (12 obs)	0.05	0.21	0	1
- meat (50 obs)	0.19	0.39	0	1
- unspecified (112 obs)	0.43	0.49	0	1
dominant orientation of the firm (dummy: 1 - yes, 0 - no)				
- market (85 obs)	0.32	0.47	0	1
- process (15 obs)	0.06	0.23	0	1
- product (64 obs)	0.24	0.43	0	1

*"ownership" refers to ownership of firms (or by firms) outside the stage of the marketing chain occupied by the firm.

Econometric Model 1 - Bootstrapped Random-Effects Tobit Estimation

The dependent variable in model 1, VA_{it} - sales per employee, is censored at zero and thus violates a classical assumption of the linear regression model. Consequently, we use a censored regression (also known as Tobit) model (see MADDALA, 1983). By choosing a random-effects (RE) approach the unobservable factors that differentiate the two cross-section units (2000, 2005) are assumed to be best characterized as randomly distributed variables. The cross-sectional units of our analysis - agri-food companies - vary quite a lot with respect to size, business focus and management style as well as strategy, risk aversion etc. By assuming that these differences are randomly distributed the general form of a RE model is given as

$$VA_{it}^* = \beta' x_{it} + u_{it} + \varepsilon_{it} \quad (3)$$

where VA_{it}^* denotes the latent variable (value added) for firm i at time t , x_{it} as a vector of the observable explanatory variables for firm i in period t , u_{it} as a vector capturing the effects of relevant unobservable variables and time-invariant factors characterizing firm i in period t , and ε_{it} as the stochastic disturbances of the model for firm i and period t . The two randomly distributed stochastic elements of [3] form the composite error term as

$$\xi_{it} = u_{it} + \varepsilon_{it} \quad (4)$$

which is assumed to be normally distributed with the following characteristics

$$\xi_{it} \sim N(0, \Sigma), \quad \Sigma = \begin{bmatrix} \sigma_u^2 & \sigma_u \sigma_\varepsilon \\ \sigma_u \sigma_\varepsilon & \sigma_\varepsilon^2 \end{bmatrix} \quad (5)$$

with σ denoting the standard deviation as usual. The dependent variable in [3] VA_{it}^* denotes the latent variable and VA_{it} as the proxy sales per employee for firm i at time t . Hence, we construct the left-censored variable VA_{it} used in estimation as

$$VA_{it} = \begin{cases} VA_{it}^* & \text{if } VA_{it}^* > 0 \\ L & \text{if } VA_{it}^* \leq 0 \end{cases} \quad (6)$$

where L denotes the lower censoring bound, and use a RE Tobit ML estimation procedure to obtain estimates of the parameters of the vector β' in (3) by maximizing the log-likelihood function

$$L(\beta, \sigma) = \sum_{i=1}^n \left(\ln \frac{1}{\sigma} \phi \left(\frac{VA - x\beta}{\sigma} \right) + \ln \Phi \left(\frac{x\beta - \tau}{\sigma} \right) \right) \quad (7)$$

where ϕ and Φ are the probability density function and the cumulative density function, respectively, for the standard normal distribution, σ as the standard deviation for ξ , and τ as the threshold of censoring, here zero. As is common econometric knowledge, robust inference requires that the distribution of the error terms follow a homoscedastic pattern. Hence, we use the heteroscedasticity-consistent covariance matrix estimator proposed by White (1980) and report the corrected standard errors and t-statistics. To test for small-sample bias we further investigate the robustness of our estimates obtained by (1) by applying a simple stochastic re-sampling procedure based on bootstrapping techniques (see EFRON and TIBSHIRANI 1993). As is extensively discussed by HOROWITZ (2001), the bias of the bootstrap as an estimator of $\hat{\Psi}_n$, $B_{\hat{\Psi}} = \hat{\Psi}_n - \hat{\Psi}_n^*$, is itself a feasible estimator of the bias of the asymptotic estimator of the true population parameter Ψ_n .

Model 2 - Censored Least Absolute Deviations Estimator (CLAD)

Both violations - heteroscedastic error terms and a non-normal error distribution - lead to highly inconsistent Tobit regression results. However, there are alternative estimation procedures which do not require the adherence to these error related assumptions. Beside the method of Least SQUARES GAUSS and LAPLACE suggested the estimation method of Least Absolutes based on the minimization of the sum of the absolute residuals (TAYLOR, 1974). This estimator consequently measures the error term as the absolute distance of the estimated values from the true values and belongs to the median family of estimators. POWELL (1984 and 1986) generalized the Least Absolute Deviations estimation for the censored regression model based on a non-negative dependent variable. We choose this nonparametric censored least absolute deviations estimator (CLAD) as a second estimation approach for our pooled sample. Different contributions (ARABMAZAR and SCHMIDT, 1981; VIJVERBERG, 1987; ROGERS, 1993) show that the CLAD estimator is robust to heteroscedasticity and is consistent and asymptotically normal for a wide class of error distributions. The CLAD estimator is more robust to outliers, which arise frequently in the case of survey data due to erroneous responses. If we re-write equations 3 and 6 as follows: $VA_{it} = \max(x_{it}\beta' + \varepsilon_i, L)$ (8)

The CLAD estimator of β minimizes the sum of absolute deviations, $|\varepsilon|$, assuming a conditional median restriction on the error term. The objective function can thus be specified as:

$$S_n(\beta) = \min \left\{ \frac{1}{n} \sum_{i=1}^n |VA_i - \max\{L, \beta' x'_i\}| \right\} \quad (9)$$

whereby the estimator uses the observations so that the median is preserved by monotonic functions. Hence, the CLAD estimator involves the minimization of an objective function that is not necessarily convex in β . Thus, obtaining a global minimum of (10) implies the usage of numerical minimization algorithms based on the approximations of the first derivative. The optimization procedure follows JONSTON and DINARDO (1997). Since the estimator's asymptotic variance-covariance matrix involves the estimation of the density function of the error term, we use bootstrap estimates of the standard errors with about 1000 draws following the re/sampling procedure outlined above.

Model 3 - Nonparametric Quantile Regression

Model 3 is based on the differences of the variables' values between the two time periods. Hence, the dependent variable ΔVA_{it} no longer has a censored distribution. However, due to the small sample size as well as the survey related frequency of outliers, we use again a nonparametric quantile regression procedure based on a least absolute deviation estimation (LAD). Equations (2) and (3) are estimated by following the procedure outlined in the previous section. Hence, (8) is adjusted to account for an uncensored dependent variable

$$\Delta VA_{it,t-1} = \max(\Delta x_{it,t-1}\beta' + \varepsilon_i) \quad (10)$$

The LAD estimator of β minimizes again the sum of absolute deviations, $|\varepsilon|$, assuming a conditional median restriction on the error term. The objective function is now

$$S_n(\beta) = \min \left\{ \frac{1}{n} \sum_{i=1}^n |\Delta VA_{it,t-1} - \max\{L, \beta' \Delta x'_{it,t-1}\}| \right\} \quad (11)$$

Obtaining a global minimum of (11) implies again the usage of numerical minimization algorithms based on the approximations of the first derivative. We finally also bootstrap the quartile regression models for the differenced sample following the re-sampling procedure outlined above and obtaining the sampling characteristics of our vector of parameters as described by (8) after re-sampling the 110 observations with replacement.

5 Estimation Results

The diagnostic tests conducted for the Tobit regression indicate no serial correlation, no rejection of the normality hypothesis with respect to the residual, and a rejection of the hypothesis of no joint parameter significance at the 5% level of test (see LR chi-squared value). The overall model significance is satisfactory (see adjusted McFadden's R^2 , McKelvey-Zavonias R^2 , and the Akaike information criteria value AIC), all given the modest sample size and the use of survey data. This conclusion is backed up by the bootstrapped bias-corrected standard errors confirming the robustness of the various estimations (the parameter estimates can be obtained from the authors).

Overall, there is strong agreement between the results generated in models 1 and 2 in that the parameters estimated have in general, the same signs and pattern of significance. In particular, both models deliver significant and positive parameter estimates for the influence of FDI on value added. Both models identify the wholesale stage of the chain and the dairy sector as having significant and positive influences on value added. Both models deliver a highly significant, and positive, parameter estimate for the influence of staff education levels. Neither model identifies firms' dominant orientation (product or market) as a significant influence on value added, nor the percentage of employees located outside Denmark. We find no relationship between the innovative activity of a firm, as measured by the number of new products introduced, and the size of the firm. Notably, both models indicate no influence of new product introductions on value addition. The two models disagree on the influence of regional location, with model 1 delivering a strongly positive influence of location in Århus. The role of sector is also ambiguous, with model 1's result indicating a significant negative influence of "unspecialised" firms and model 2 indicating the opposite.

The results for model 3 (two specifications, see table 4) generally support those of models 1 and 2. FDI is a significant positive influence on growth in value added, as is firms' operation at the wholesale stage of the chain. Both specifications of model 3 indicate a negative influence on growth in value added of the number of employees outside Denmark. The non-significance of new product introduction in value addition (from models 1 and 2) was also found in both specifications of model 3: neither the change in new product introduction nor the average for the two years is a significant driver of growth in value addition. Although both specifications of model 3 deliver similar R^2 values, the second specification (using average number of new products rather than its first difference) delivers two more significant explanatory variables: a product orientation of the firm (a positive influence on growth in value added) and a market orientation (a negative influence). Neither of the two specifications deliver a significant parameter estimate for location in Århus, although in both cases the *t*-value indicates that the (positive) estimate is close to being significant at the 5% level of test.

6 Discussion of Results

Firms that introduced new products into the market did not have a significantly higher value added than those that did not introduce new products. Thus we can not reject our first hypothesis (hypothesis 1 table 1). (This result did not change when we looked at the growth in value added by firms.) There are at least three ways one can interpret this result. First, new product innovation is not always profitable, especially in the short run. Firms that have successful products in the market try to maintain the market for these products through advertising and driving down production and marketing costs. In such cases new products may be associated with higher costs of production at least in the short run. Second, food firms may be more likely to innovate through process innovation and thus lower the costs of production. Our data does not capture process innovation. Finally, innovation may occur through strategic alliances, such as with foreign direct investment. This may open up foreign markets allowing firms to drive down costs. The nature of the connection between new

product introductions and other activities of the firms is the subject of on-going research, but at the current study cannot associate it strongly with value addition.

There is an on going debate in the profession regarding the relationship between the size of firms and their innovativeness. In the food sector many of the new innovations are not as expensive to adopt as they are in very large capital intensive manufacturing industries, like the aerospace sector. Perhaps capital constraints are more important in determining which firms innovate, however we did not have data on the financial status of the firms. The strong positive influence of staff education levels on value addition indicates an important role for high quality human capital. This is highly consistent with the negative influence of outsourced labour, and indicates a commitment to “knowledge-based” industry. It also appears to be consistent with regional networking amongst firms and with our inability to reject our second hypothesis (hypothesis 2 table 1): the insignificance of dominant orientation. Our results indicate that instead of such orientations, Danish firms employ sector (i.e. dairy), stage (i.e. wholesale) and educated employees to deliver value added. This is, of course, not to say that Danish firms do not have a dominant orientation: rather we claim that it is not an important determinant of value addition across a range of firms. This interpretation is reinforced by the model of growth in value added, where dominant orientations do play a role, but in this case they explain changes within a single firm, rather than between firms. The lack of significance of the co-operative form as an explanation of value added is expected to some extent. However, we exercise caution in interpreting this result because the dummy variable used may be masking other effects. In particular, many of the processing firms in the sample are co-operatives and so the impact of both dummies may be diluted by co-occurrence.

We are able to reject the third hypothesis (hypothesis 3 table 1) that regional innovation systems do not have an impact on a firm’s performance. ASHEIM and COENEN (2005) examined the functional food ‘cluster’ in Scania and found that the location of a university and research organizations provided the seedbed for innovation. BRAADLAND (2003) and AVERMAETE and VIAENNE (2000) identified regional networks as sources of innovation in agro-industry, the bulk of it being organisational or involving the strengthening of existing brands and market positions. A major advantage of Århus is the University of Århus and the numerous food industry research facilities that have been built by both the public and private sector in the area. It is the spillovers between research individuals together with presence of university research that lead to new product and process innovation that increase the productivity of firms. These interactions support the successful food processing firms located in the Århus region. Our results support such an explanation of value addition, specifically by identifying Århus as a centre for innovative networks. In addition to being an innovation centre, Århus’ geographic location favours relatively lower cost access to export markets and the presence of an agro-industrial cluster (and the majority of Danish livestock produce) is likely to contribute to superior added value.

The positive influence of FDI on value addition implies we can reject our fourth hypothesis (hypothesis 4 table 1). Thus we can not rule out the importance that FDI has in the Danish food sector. AITKEN and HARRISON (1999) have outlined a number of targets of FDI in the manufacturing sector in Venezuela, including access to markets, general investment considerations and the introduction of specific skills, experience and capital. Each one of these explanations appears to have good application to the data used here, and reinforces the impression gained from the strong influence of FDI in the growth of value added (model 3). This is a particularly important result given the pressure for Danish food companies to globalize because of reduced tariff and non-tariff barriers. Our final hypothesis (hypothesis 5 table 1) is built on the relationship between the growth of a firm (as measured by change in value added) and the number of new products introduced to the market. We found no relationship between innovation and the growth in company size. This was an expected result. A lot of product and process innovation is done in smaller firms. Small firms without capital

limitations are often more flexible and adaptable to new ideas. We point out that the inclusion of average new product introduction levels in model 3 are associated with significance of dominant orientations, which do not feature strongly elsewhere in the results.

What then does explain the growth of firm size? We found two variables to be significant, FDI and if the firm is wholesale marketing stage of the industry. The most important of these is FDI. Business networking is extremely important in a globalizing economy; however it is also very expensive in terms of management time. One way to achieve the benefits business networks is through FDI. This result is consistent with the observation that more a liberalized trading environment provides potential benefits to those firms that can increase exports or imports through strategic alliances. The significance of the “wholesale” stage of the chain in value addition is likely to be associated with its rapid consolidation since 1995 (see BAKER, 2003). Consolidation, *ceteris paribus*, is likely to raise the sales per employee for remaining firms. The dynamic nature of this result is further supported by its strong significance in model 3. In practice, value added is strongly contested between retailers and wholesalers, both of which have experienced consolidation since 1995, although it has been most pronounced at wholesale level (BAKER, 2003). A similar within- versus between-firms argument explains the lack of significance of university education in model 3: although value added has increased for most firms between 2000 and 2005, the numbers of employees with a university education at any one firm probably has not.

The non-significance of several variables that is not reported in the results are worthy of note. First, export orientation of firms was dropped from the model for this reason. This unexpected result indicates that firms serving export markets face fewer opportunities for value addition than those concerned with domestic markets. More importantly, the result provides a clue about the purpose of FDI in the Danish food industry: it is unlikely to be motivated by access to EU and foreign markets, but rather addresses the Danish market. Although the survey provided data on competitiveness of markets (for products and inputs), these also dropped out during specification. Value addition is likely to be influenced by the structure of markets, but it appears that the effects have been captured in the models by variables such as sector and stage of chain. This result requires examination in future work.

Our ration for using value-added per employee is that it would track net value-added per employee among the firms and over time. One reviewer suggests this may be a problem when some firms outsource low cost labour and retain the high cost labour in house. Firms may do this (in fact we know they do) when out sourcing to lower labour cost countries. Higher cost labour, such as design engineers and accounts for example, are usually retained in house. This should not bias our results as the value added will reflect this type of innovation in business-organization. What is a problem is that we do not identify it in our regression model nor can we because of the data. This would be very interesting for further study especially when examining innovation in developed countries like Denmark.

7 Conclusions

This study examines some of the factors that influence innovation in the value-adding food sector. Our theoretical model is based upon the early work of ROMER (1990). A number of hypotheses were tested using data from the Danish food sector for the years 2000 and 2005. The hypotheses were drawn from a literature that has been dominated by general theoretical relationships and case studies but short on econometric modeling. The unique and new data set employed requires specific econometric techniques due to truncation of the distribution of the main dependent variable (a proxy for value addition) and the distribution of both the explanatory variables and the models' error terms. One of the contributions of this paper is the use of a new and interesting theoretical model. This model holds some considerable promise

in analyzing the impact of innovation activity in the food sector. Overall, the results of the econometric analysis are not inconsistent with our theoretical specification.

Specifically we find no relationship between company size and innovation (hypothesis 1 table 1). we cannot reject the null hypothesis of no influence of dominant orientation on value added (hypothesis 2 table 1). Firms' that focus on human capital, stage and commodity sector have a higher value addition. We can reject the hypothesis (hypothesis 3 table 1) that regional networks have no influence on value added. Location in Århus, emphasis on human capital and the negative influence of outsourcing on value added all provide supporting evidence. We reject the hypothesis (hypothesis 3 table 1) of no influence of FDI, and moreover propose that the FDI has targeted the domestic Danish market as a source of value added. Evidence on the importance of firms' dominant orientation is mixed at best, and we cannot reject the null hypothesis of no influence on value added (hypothesis 5 table 1).

This paper represents a beginning in the quantification of the impact of innovation in the food sector. More studies need to be done and reported before any strong conclusions can be drawn. However, given the importance of innovation and the interest of many governments in the topic this field promises to be a rich area of study for applied economists.

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