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*Department of Agricultural &
Resource Economics, UCB*
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Measuring the Poverty Reduction
Potential of Land in Rural Mexico

Frederico Finan
University of California, Berkeley

Elisabeth Sadoulet
University of California, Berkeley

Alain de Janvry
University of California, Berkeley

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Addresses: Frederico Finan, Giannini Hall 207, University of California, Berkeley, California 94720. Tel: 510-643-5418. Fax: 510-643-8911. E-mail: Finan@are.berkeley.edu.
Elisabeth Sadoulet, Giannini Hall 207, University of California, Berkeley, California 94720. Tel: 510-642-7225. Fax: 510-643-8911. E-mail: Sadoulet@are.berkeley.edu.
Alain de Janvry, Giannini Hall 207, University of California, Berkeley, California 94720. Tel: 510-642-3348. Fax: 510-643-8911. E-mail: Alain@are.berkeley.edu.

I. Introduction

Access to land has for many years been advocated as one of the potentially most effective approach to rural poverty reduction (Warriner, 1969; Thiesenhusen, 1989; Dorner, 1992; Binswanger, Deininger, and Feder, 1995). This can be seen in the extensive and sometimes dramatic land reforms programs that have occurred in numerous countries around the world, and in particular in Latin America (de Janvry, 1981). Justification for these programs is based on the assumption that land has a strong welfare generating potential for beneficiaries. Yet, land reform programs have frequently met with limited success in reducing poverty, and a number of recent empirical studies has called into question the importance of land as a poverty-reducing instrument (López and Valdés, 2000a). There is, however, a remarkable absence of solid empirical evidence measuring the potential of land in reducing poverty, particularly carefully taking into account the specific conditions under which land is used by beneficiaries, when we well know that these conditions are key determinants of success or failure.

There is little doubt that land is a source of income, based on the empirical evidence provided by numerous studies that have documented the existence of a positive association between land and incomes (see for example Scott (2000) for Chile; Gunning et al. (2000) for Zimbabwe; Grootaert, Kanbur, and Oh (1997) for Côte d'Ivoire; Bouis and Haddad (1990) for the Philippines; and Carter and May (1999) for South Africa). However, in many cases, this positive association actually translates into small income gains for the range of land endowments provided by land reform programs. For instance, McCulloch and Baulch (2000) simulated the impact of a policy giving two hectares of land to households in rural Pakistan with less than this amount to find that it has virtually no effect on income poverty. López and Valdés (2000b) and their co-authors find in empirical studies for eight Latin American countries that the income generating potential of land is also quite small. For example, their estimates of income to land elasticities suggest that landholdings in rural areas of Colombia would have to quadruple in order for the poorest 40% of farm households to reach just the poverty line. This led them to recommend looking into approaches other than access to land in order to attack rural poverty.

Although methodologies vary across studies, the empirical literature has in general ignored three important considerations when measuring the marginal poverty-reducing value of land. First, previous studies often assumed a linear specification for the estimation of an income equation. This assumption can be overly restrictive when constraints on a household's ability to effectively use its assets create a complex nonlinear relationship between land endowments and income. In this paper, we use nonparametric techniques to overcome this difficulty. Second, poverty is multidimensional, and an income representation of poverty is both very noisy and overly restrictive as a measure of household welfare. We consequently define an alternative welfare indicator that captures the multidimensionality of poverty as advocated in recent poverty studies (World Bank, 2001). Finally, differences across households in their vital characteristics, their overall asset positions, and their contextual circumstances demonstrate that return to land for them is significantly affected by these conditions. In measuring the return to land, it is consequently important to account for the high degree of heterogeneity across rural households and to use this to make conditional -- as opposed to the usual absolute -- statements about the poverty-reducing value of land.

In recent years, programs of access to land have returned high on the agenda of poverty reducing strategies pursued by governments, NGOs, and international development agencies (Deininger and Feder, 2002). This has led to large scale initiatives in countries such as Brazil (Cedula da Terra and INCRA) and Colombia (land market assisted reform), and to new initiatives in Mexico (Procede), Honduras (Land Bank), and Bolivia (validation of titles). Measuring the poverty reducing value of land and the conditions for successful use of this instrument has thus become all the more urgent. Additionally, this has to be done from a household (as opposed to a farm) perspective, where beneficiaries may engage in off-farm activities (Reardon et al., 2001, de Janvry and Sadoulet, 2001), and where access to even a small amount of land, given labor market failures, may provide an effective platform for livelihood strategies that allow to move out of poverty, something that was not considered in typical farm-oriented land reforms that only considered on-farm agricultural activities.

Since we are interested in measuring how land can reduce poverty in contexts of extensive rural poverty, we use the household data collected in 1997 by the Mexican Program for Education, Health, and Nutrition (PROGRESA) in poor rural communities. Results show that a small amount of land can create large income gains as it permits to mobilize family labor affected by labor market failures. In general, land has a high but decreasing marginal return for endowments of less than three hectares and a constant return for land sizes larger than three hectares. These results are consistent with the theory that small landowners with low opportunity costs of labor exhibit a higher reservation price for land (see Carter and Mesbah (1993) and Carter and May (1999)). Additionally, the marginal welfare value of land depends crucially on both the complementary assets (such as education) and the contextual settings (such as infrastructure) of the poor, prohibiting any absolute statement about the poverty reduction value of land. We also find that ethnicity is an important negative social asset as the marginal value of land for non-indigenous households is on average twice as high as it is for indigenous households. When complementary assets and a favorable context are in place, access to even a small amount of land can help households escape poverty. The paper thus helps establish the complex set of conditions under which land can be a valuable poverty-reducing instrument.

The rest of the paper is outlined as follows. Section 2 derives our estimative income equation from a microeconomic household model for conditions that imply a nonlinear relationship between welfare and land. In section 3, we explain the semiparametric estimation of the income equation. The data are described in section 4, including a discussion on our welfare measure and the structure of landholdings in Mexico. The empirical results follow in section 5, and we conclude the paper in section 6.

II. Theoretical Framework: Derivation of the income equation

In this section we derive the specification of our income equation from a standard agrarian household production model. We consider a situation of multiple market imperfections to investigate how these distortions affect the economic return to land.

The theoretical framework borrows from the work of Carter and Mesbah (1993) and it assumes three important frictions: 1) Land transactions are ignored and access to land is treated as exogenous, 2) households face the possibility of off-farm unemployment, and 3) access to credit increases with land size.

Under these assumptions, consider a household that generates income by cultivating agricultural land, in addition to possibly supplying labor at an exogenously determined market wage, w . The household is endowed with T hectares of land and \bar{L} hours of labor per year that are employed in on-farm agricultural work (L_f) and/or off-farm activities (L_s). The household cultivates a single crop using X units of input purchased at a per unit market price of q . The crop can be sold at an exogenous market price p . Let $F(L_f, X, T; z)$ be the production function, where z represents the set of household and contextual characteristics that affect the return on productive assets. Let $\Omega(L_s)$ denote the number of days employed as a function of labor supplied, L_s , where $\Omega' > 0, \Omega'' \leq 0$. Let $\Gamma(T)$ denote the amount of working capital available at an interest rate i to a household with land endowment, T . The cost of production, qX , must be financed by the sum of initial wealth K , wage income $w\Omega(L_s)$, and available capital $\Gamma(T)$.

Formally, the household chooses time allocation and purchased inputs to maximize its income:

$$\begin{aligned} & \text{Max}_{L_s, L_f, X} pF(L_f, X, T; z) - qX + w\Omega(L_s) - i(qX - K - w\Omega(L_s)) \\ & \text{s.t.} \\ & L_s + L_f \leq \bar{L} \quad (\text{a}) \\ & qX \leq K + w\Omega(L_s) + \Gamma(T) \quad (\text{b}) \\ & L_s \geq 0, L_f \geq 0. \quad (\text{c}) \end{aligned} \tag{1}$$

Assuming an interior solution for the labor allocation, the first-order conditions of this maximization problem can be written:

$$\begin{aligned}
pF_L &= w(1+i+\lambda)\Omega' \\
pF_X &= q(1+i+\lambda) \\
\lambda(K+w\Omega(L_s)+\Gamma(T)-qX) &= 0; \quad \lambda \geq 0; \quad K+w\Omega(L_s)+\Gamma(T)-qX \geq 0
\end{aligned} \tag{2}$$

where λ is the Lagrange multiplier associated with the liquidity constraint (1b).

If (*) denotes solution values of the choice variables that maximize the program above, then the income equation associated with income maximizing behavior can be specified as follows,

$$Y = pF(L_f^*, X^*, T; z) - q(1+i)X^* + w(1+i)\Omega(L_s^*) + iK = \Lambda(p, q, i, w, \bar{L}, K, T, z). \tag{3}$$

The income equation is a function of prices, household's endowment of productive assets, and any characteristic that affects the return to these assets. We can differentiate equation (3) to see how an increase in landholdings affects household income,

$$\begin{aligned}
\frac{dY}{dT} &= pF_T + (pF_L - w(1+i)\Omega') \frac{dL_f^*}{dT} + (pF_X - q(1+i)) \frac{dX^*}{dT} \\
&= pF_T + \lambda w \Omega' \frac{dL_f^*}{dT} + \lambda q \frac{dX^*}{dT}
\end{aligned} \tag{4}$$

If capital markets are perfect ($\lambda = 0$), then the terms in parentheses are identically equal to zero and the marginal value of land is simply the value of its marginal product, i.e. $\frac{dY}{dT} = pF_T$. Moreover, if households face the same opportunity costs of labor and inputs, and we assume constant returns to scale, then the marginal return to land is constant for all land endowments.

Conversely, with imperfections in labor and credit markets, the terms in parentheses are positive. An increase in landholding has both a direct and an indirect effect on income. In addition to directly increasing production, more land affects the distortions in the allocation of production inputs. As the land endowment increases, labor allocated to the farm increases, the under-employment rate Ω decreases, and the marginal productivity of labor Ω' increases. On the other hand, as the land endowment increases, the credit constraint may either tighten or loosen (λ can increase or decrease) depending on the relative increase in credit availability $\Gamma(T)$ and on the demand for inputs. For instance, in poor rural areas characterized by thin labor markets, an increase in household land will increase the marginal product value of household labor and reduce the difference between its shadow wage and the market wage. If, as Eswaran and Kotwal (1986) suggest, larger farms have better access to credit, then an increase in landholding will increase the use of variable inputs and reduce the distortion in the input markets as well. With market distortions, we consequently expect the marginal value of land to vary with the land endowment, and quite possibly in a nonlinear manner.

For better insight into how imperfections in the land, labor, and credit markets affect the land's potential to generate income, we simulate in Figure 1 the model above with specific functional forms and a rough parameterization, given our data.¹

¹ We use a Cobb-Douglas production function $F(\cdot) = L_f^{1/3} X^{1/3} T^{1/3}$, unemployment characterized by $\Omega(L_s) = L_s - 4 * L_s^2$, and credit availability by $\Gamma(T) = 1300 * \Phi\left(\frac{T-10}{3}\right)$. We take $p = 1500$ pesos/metric ton of corn, the average wage for a nonagricultural worker $w = 1280$ pesos per month, the average per unit cost of corn production $q = 460$ pesos per month, and the interest rate $i = .05$.

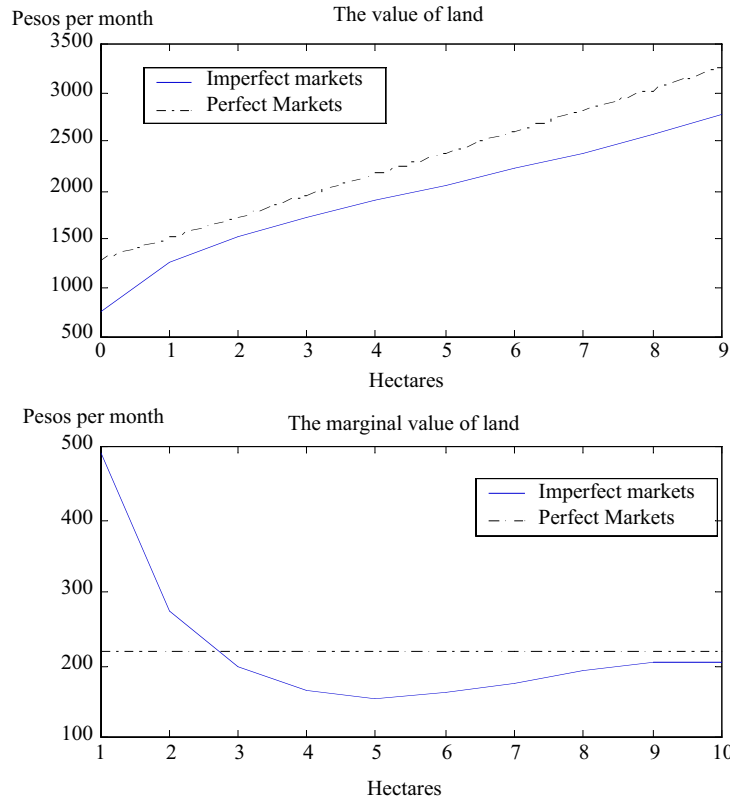


Figure 1. The income value of land under market imperfections

The top panel displays the relationship between income and land for both perfect and imperfect markets. When markets are perfect, landless households must work entirely off-farm and earn a monthly income of 1280 pesos. With access to a hectare of land, household income will increase by the marginal value of land, 228 pesos, as the household adjusts its labor allocation between working its land and off-farm employment. In the presence of market imperfections, the mapping of land to income shifts downward and is no longer linear. The curve intercepts the y-axis at 768 pesos: the expected wage rate when faced with possible unemployment. With distortions in the labor market, the marginal value of land is high for small farmers since another plot of land will also increase their shadow wage. As land endowments increase and the household can allocate its labor more effectively, the marginal return to land declines. Around four hectares, as access to credit improves, the marginal value of land begins to increase. A household's inability to generate sufficient economic livelihood depends not only on its land endowment but also on its ability to effectively utilize the amount of land it does possess.

To estimate the marginal value of land, as depicted in Figure 1, the typical linear approximation to an the income equation is an inadequate specification. Without knowing what are the underlying frictions of our environment, and hence the shape of the relationship between land and income, we relax completely the functional form for land and explore this mapping with nonparametric estimation techniques.

III. Econometric specification of the income equation

This section outlines the semiparametric procedure for estimating the relationship between income and land endowments. Our production model implies that any characteristic that affects the return to the productive assets of the household should influence the household's income. This list includes household demographics, constraints on factor use, as well as village and regional factors that capture employment opportunities and market integration. Although a fully nonparametric specification of the income equation

would best capture its underlying shape, with several possible covariates the computation cost of this technique is prohibitively high.² A semiparametric procedure allows us to relax the functional form on land and still control for the other factors that determine household income. Following Robinson (1988), we estimate a model of the following form,

$$y = \alpha + x\beta + g(z) + \varepsilon \quad (5)$$

where x of dimension $n \times k$ is the set of controls and z of dimension $n \times l$ is the household's land endowment. The variable y represents some measure of household welfare. The constant term is denoted by α , and β is a $k \times l$ vector of our parameters of interest. The error term, ε , is distributed normally. We assume that the functional form of $g(\cdot)$ is unknown and $E(\varepsilon | x, z) = 0$. Taking expectations of equation (5) conditional on z , we get:

$$E(y | z) = \alpha + E(x | z)\beta + g(z). \quad (6)$$

Subtracting (6) from (5) yields:

$$y - E(y | z) = [x - E(x | z)]\beta + \varepsilon. \quad (7)$$

With nonparametric estimates of $E(y | z)$ and $E(x | z)$, we can then estimate (7) by ordinary least squares to get an estimate $\hat{\beta}$, that is \sqrt{n} -consistent, asymptotically normal, and robust to any unknown form of $g(z)$.³ The estimator for $g(z)$ is simply:

$$\hat{g}(z) = E(y | z) - (\alpha + E(x | z)\hat{\beta}). \quad (8)$$

There are two important features to note about this procedure. First, since $\hat{\beta}$ converges at a rate of \sqrt{n} and $\hat{g}(z)$ converges at a slower rate of \sqrt{nh} (where h is the bandwidth size), the estimation of β does not affect the asymptotic distribution of $\hat{g}(X)$. Second, a constant term cannot be identified independently of $\hat{g}(z)$.

IV. The data

The data for this study come from the 1997-1998 surveys conducted for PROGRESA. This program targets poor rural communities, and thus our analysis focuses on the value of land for this particular segment of the rural population that lives in marginal communities. The survey covers approximately 25,000 households over 500 localities and 7 states. A stratified randomization procedure selected the localities and every household within each locality was included in the sample. The data are available at the individual, household, and locality levels, with detailed information on schooling, consumption, and employment patterns of the household.

Land endowments

With data on both rainfed and irrigated types of land, we convert plot size into hectares of rainfed corn equivalence (RFE). The average yield of corn for both irrigated and rainfed land is calculated for each

² With k explanatory variables and a sample size of N observations, to evaluate the density on a k -dimensional grid of G points, requires NG^k evaluations (Deaton 1997). Even with the semiparametric approach, the income equation took over 8 hours to estimate.

³ Both $E(y | z)$ and $E(x | z)$ are estimated using Cleveland (1979) robust locally weighed regression (LOWESS) technique, using a bandwidth of 0.8.

locality and normalized by the sample average yield of rainfed corn.⁴ A household's endowment of land in hectares of rainfed corn equivalence is then the weighted sum of its rainfed and irrigated landholdings, where the normalized averages are used as the appropriate weights (de Janvry, Gordillo, and Sadoulet, 1997). By normalizing farm size by its yield, this adjustment incorporates a measure of land quality, a variable that is typically hard to observe.

The population of the survey is definitely poor, with 52% of households in our sample classified in poverty. Only 54% of the sample possesses land, and the average farm size is less than 2.7 hectares. Of the 97% of land that is used for agriculture, 90% is cultivated in corn.

The rental market is very inactive with 5% of farmers participating in the rental market for an average 1.7 ha of rainfed land, which only represents 3.3% of the rainfed area (see Table 1). These proportions are slightly higher with 6% participation and 6% of rainfed land among the small farm size of less than 2 RFE hectares, but this remains negligible. Unfortunately, the data is unclear as to the proportion of farmers who rent-out, but we suspect that this percentage is also small. An *ejido* community exists in 60% of the localities in our sample. This may explain some of the inactivity in the land rental market since, until recently, land transactions were prohibited in all *ejidos*. Although we have no data on land sale or purchase transactions, this market is likely very thin too. We thus retain for our analysis household owned land and make the assumption (that we will test) that land is exogenous.⁵

Table 1: Average farm sizes by land and ownership type

	All	Farm size in RFE ha			
		<1 ha	1-2 ha	2-5 ha	>5 ha
Distribution of farms (%)	100	30	28	28	14
Farm size (ha RFE)	3.09 (5.55)	.70 (.24)	1.40 (.32)	3.09 (.83)	11.09 (11.27)
Rainfed (ha)	2.68	0.62	1.14	2.47	8.10
Households participating in the rental market (%)	5.1	6.2	6.2	4.4	2.0
Rented/sharecropped area, among participants (ha)	1.71	0.68	1.21	2.63	7.3
Irrigated (ha)	0.14	0.002	0.017	0.074	0.67
Households participating in the rental market (%)	0.4	0.01	0.17	0.62	0.94
Rented/sharecropped area, among participants (ha)	1.8	0.25*	0.55*	1.06	3.38

* Denotes less than 10 observations.

Figure 2 provides evidence on the correlation between land and welfare. The association between land and the poverty indicator used by PROGRESA is quite striking. Among those households with less than one RFE hectare, 62% are poor. This proportion drops to only 38% when we consider households with more than 8 hectares of land.

⁴ If the number of observations in a community was less than 30, we averaged yield at the municipality level. Also these averages did not include crop failures, which for rainfed is a yield of less than 0.2 ton/ha and irrigated a yield of less than 0.8 ton/ha.

⁵ We test this assumption in our estimation of the welfare equation.

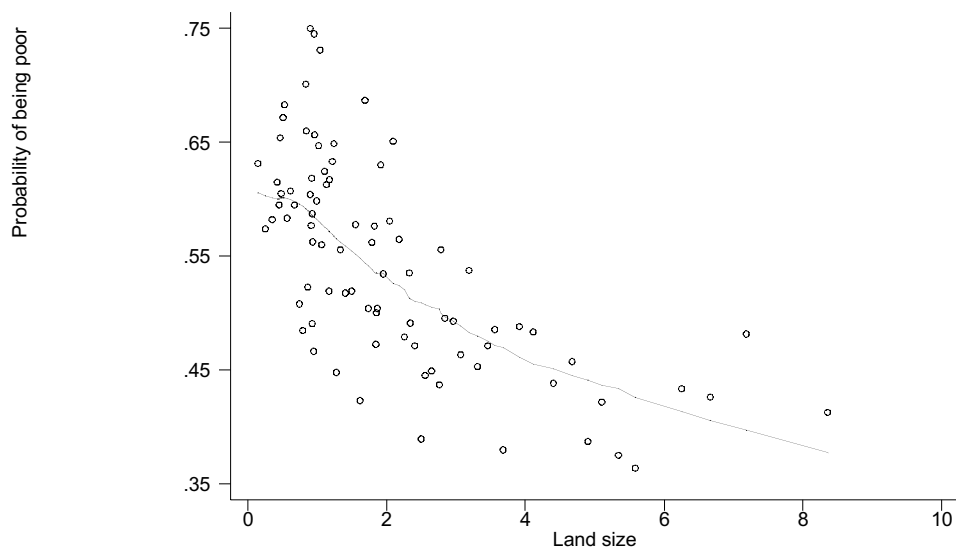


Figure 2: Poverty and land⁶

Welfare index

What is the best measure of welfare is a topic of much debate. Income, which is sensitive to volatile shocks, can be a poor indicator of long-term welfare for households predominately involved in agricultural and self-employed labor activities. With the possibility of households smoothing their consumption across time by borrowing, saving, and mutual insurance, consumption measures are considered much more reliable and more theoretically sound (Ravallion, 1996; Deaton 1997). Yet, both of these monetary-based measures fail to capture the multidimensionality of poverty, and in effect neglect important dimensions of welfare such as access to sanitation, access to water, and adequate shelter (World Bank, 2001).

With these concerns, we construct a welfare index consisting of various dwelling characteristics (running water, electricity, has a bathroom, number of rooms, and dirt floors), household durables (ownership of a blender, refrigerator, television, and truck), in addition to the short-term measures of consumption expenditure and non-farm labor income. We aggregate these various indicators of poverty using principal component techniques.⁷

Table 2 displays the average of various household characteristics by quintiles of the welfare index. There is a consistent positive relationship between assets (land, farm animals, household education level), income, and welfare. The difference in land assets between the highest and lowest quintile is only 2.2 RFE hectares, which is not entirely surprising given our sample of poor communities. The average education level of the households in the highest quintile is almost two full years higher than in the lowest quintile. Poverty is strongly associated with ethnicity, as fifty percent of households in the lowest quintile are

⁶ To make the graph more readable, each point represents 137 observations.

⁷ The principal components were computed for each of the seven regions separately. For every region, the first Eigen value captures more than 25% of the total variance and is about three times as high as the next Eigen value. We, therefore, use this first component as our welfare index. It is a weighted average of the normalized variables retained for the analysis. The coefficients of the different variables are very similar across regions. We report the average of the coefficients in the following welfare index formula:

$$W = -.30 (\text{dirt floor}) + .25 (\text{\# rooms}) + .17 (\text{potable water}) + .18 (\text{bathroom}) + .12 (\text{sewage}) + .30 (\text{electricity}) + .38 (\text{blender}) + .32 (\text{refrigerator}) + .36 (\text{oven}) + .24 (\text{radio}) + .35 (\text{television}) + .23 (\text{truck}) + .11 (\text{labor income}) + .19 (\text{consumption}).$$

indigenous compared to only 25% in the highest. Also, while the number of agricultural workers per household varies little across quintiles, households in the highest quintile have almost three times the number of nonagricultural workers as the lowest quintile. Households in the highest quintile earn approximately 2.2 times as much as the households in the lowest quintile. Non-labor transfers represent almost the same share of income in all quintiles. Off-farm earnings are an important source of income, especially for those in the lowest quintile where 68% of their total income is derived from off-farm activities.

Table 2. Mean characteristics of the poor across welfare levels

	Quintiles of the welfare index					Total
	1	2	3	4	5	
Number of households	2,740	2,740	2,740	2,740	2,740	13,700
Land assets in hectares						
Rainfed	2.28	2.28	2.52	2.74	3.78	2.72
Irrigated	0.06	0.08	0.12	0.18	0.32	0.15
Total in rainfed equivalence	2.46	2.53	2.87	3.21	4.63	3.14
Farm animals						
Horses, mules, and cattle	1.10	1.47	2.19	2.84	4.45	2.42
Household characteristics						
Number of adult equivalent persons	4.84	5.38	5.72	5.93	5.96	5.56
Mean level of education in household	1.96	2.45	2.86	3.25	3.89	2.88
Head of household variables						
Education	1.91	2.31	2.46	2.64	3.04	2.47
Age	47.0	48.1	48.7	49.5	51.8	49.0
Indigenous head	0.59	0.56	0.42	0.28	0.17	0.4
Occupation type within the household						
Agricultural workers	0.88	0.94	0.91	0.82	0.63	0.84
Nonagricultural workers	0.10	0.14	0.20	0.29	0.39	0.22
Seller	0.20	0.25	0.26	0.27	0.32	0.26
Family-related business	0.38	0.42	0.45	0.56	0.70	0.50
Household income						
Total (pesos per month)	799	933	1126	1343	1848	1210
Off-farm wage income (%)	68.2	66.1	63.4	61.2	55.9	61.7
Non-labor transfers (%)	19.4	19.9	21.1	22.0	22.9	21.4

Thus far, descriptive statistics suggest a positive correlation between land and welfare. In the following section, we explore this relationship further and quantify the economic value of land.

V. Estimation Results: The welfare generating potential of land

In this section, we estimate a household welfare equation to explore the relationship between land and welfare. There are three main findings. First, compared to the semiparametric approach, a linear specification estimates a much lower marginal value of land. Second, in these poor communities, small

plots of land have a large welfare value. Finally, the marginal value of land differs according to a household's complementary assets and contextual setting.

Our final sample is restricted to the 12,034 landowners with complete household and village level data. Excluding the landless reduces the possibility of selection bias and ensures that the estimation is across a sample of households that do not pursue wildly different livelihood strategies. The final specification of the welfare equation consists in land, household characteristics, social and institutional assets, village characteristics, and states dummies. We also include the number of working-age adults in the household by their education level. Since members of households with smaller farms may tend to migrate more, we incorporate all the household members that have migrated within the last five years into our definition of the number of working-age adults. This will help to mitigate any potential bias in household labor force due to endogenous migration strategies.

Parametric estimation

Regression (A) in Table 3 assumes a linear specification. With our dependent variable measured in welfare units, we divide each of the estimated coefficients by the monetary transfer coefficient, 0.00025, to obtain results that are easier to interpret.⁸ This calculation is reported in the fourth column along with its standard error. The marginal welfare value of land is highly significant. An increase of one RFE hectare will increase monthly welfare by 125 pesos. For a better sense of the order of magnitude, the average monthly wage of an agricultural worker is 700 pesos, and the price of a metric ton of corn, in 1998, was on average 1460 pesos. With the average yield for rainfed corn at 1 ton/ha, 125 pesos per month or \$1500 pesos per year is a meaningful amount. To compare these results with other analyses in the literature, we also estimate this equation for income and consumption. In the income regression, the estimated return to land is not significantly different from zero. In the consumption regression, the estimated return to land is significantly different from zero, but quite small. An increase of one RFE hectare increases monthly consumption by 11.5 pesos, which is only 1.5% of average monthly consumption in the sample.⁹ This result is of an order of magnitude similar to estimates reported in López and Valdés (2000a). Choice of an adequate indicator of welfare is thus important in assessing the poverty-reduction value of land.

Human capital assets are associated with large welfare effects. A marginal increase in the household head's education level raises welfare by 374 monthly pesos, an amount which is three times the return to one hectare of land. While addition of an adult male with primary education increases welfare by a significant 925 pesos, the return of an adult male with more than secondary education is 2,795 pesos; a differential gain of 202%. Uneducated adults do not contribute to household welfare at the margin, and children and elderly men are negatively associated with household welfare. There is a high cost associated with ethnicity as being indigenous reduces welfare by 3,117 monthly pesos. Psacharopoulos and Patrinos (1993) documented that 80% of the indigenous population in Mexico lives in poverty, compared to 17.9% for the non-indigenous population, and in our sample 65% of the indigenous households are in fact classified as poor compared to 44.5% of the non-indigenous. With much of rural poverty associated with ethnicity, it is not surprising that indigenous households have on average a much lower welfare than non-indigenous households.

To capture the role of networks in migration (Winters et al., 2001), migration assets are defined as the number of village members who have migrated out of the state over the last 5 years, separating domestic migration in other states of Mexico from foreign migration. Foreign migration assets contribute 170 more monthly pesos to welfare than domestic migration assets.

⁸ Monetary transfers are the monthly amount of pesos a household receives from non-labor sources. This includes government transfers other than PROCAMPO and scholarships.

⁹ Several studies also use a log-linear specification. With this specification, we estimate a land to income elasticity of .105.

Table 3. Parametric and semiparametric estimation of the welfare equation

	Average value	Household welfare index							
		Parametric (A)				Semiparametric (B)			
		Coef.	t-stat	Peso value	st. err.	Coef.	t-stat	Peso value	st. err.
Land assets (RFE Hectares)	3.14	0.03	8.5	125	32				
Monetary transfers (*1000)	39.56	0.00025	4.5	1	0	0.00024	5.1	1	0
Head of household characteristics									
Gender (dummy)	0.92	0.30	4.8	1223	370	0.29	4.7	1219	349
Age (years)	49.2	0.01	9.3	61	15	0.01	8.7	58	13
Education level	2.49	0.09	11.2	374	89	0.09	11.6	379	81
Labor force (number of individuals)									
Male adults with education = 0	0.19	-0.04	-0.9	-149	163	-0.05	-1.2	-210	179
Male adults with 0 < education < 6	0.47	0.06	2.4	261	125	0.06	2.0	233	126
Male adults with education = 6	0.31	0.23	7.9	925	236	0.22	8.0	920	215
Male adults with 6 < education ≤ 9	0.16	0.47	12.5	1921	451	0.45	12.9	1905	404
Male adults with education > 9	0.05	0.68	9.5	2795	685	0.66	10.8	2793	608
Female adults with education = 0	0.32	0.03	0.8	102	133	0.02	0.5	76	140
Female adults with 0 < education < 6	0.44	0.30	10.0	1235	296	0.29	10.0	1245	273
Female adults with education = 6	0.28	0.54	17.8	2213	503	0.53	18.1	2256	460
Female adults with 6 < education ≤ 9	0.12	0.66	15.7	2686	616	0.64	16.8	2731	563
Female adults with education > 9	0.03	1.05	11.1	4267	1021	1.03	14.4	4355	916
Children (under 17 years old)	2.59	-0.01	-1.0	-31	33	-0.01	-1.3	-44	34
Males, at least 55 years old	0.35	-0.12	-2.4	-480	227	-0.13	-2.8	-543	221
Females, at least 55 years old	0.31	0.24	6.2	973	267	0.23	6.2	968	246
Social and institutional assets									
Indigenous household (dummy)	0.40	-0.76	-19.8	-3117	710	-0.72	-18.9	-3060	627
Access to agricultural cooperative	0.04	0.01	0.2	56	312	-0.03	-0.3	-106	327
Church present (dummy)	0.40	0.02	0.6	78	133	0.01	0.4	48	134
Migration Assets (number of individuals in the village)									
Out of the state	3.28	0.01	3.0	60	24	0.02	3.1	70	27
Out of Mexico	1.21	0.06	7.3	230	59	0.06	9.5	241	53
Locality Characteristics									
State Road (dummy)	0.17	0.22	4.8	882	268	0.22	5.2	952	265
Federal Road (dummy)	0.26	0.22	5.7	898	255	0.23	6.4	992	251
Health center (dummy)	0.14	0.24	4.8	964	292	0.21	4.6	910	268
Min. distance to an urban center (km)	107	-0.0044	-7.4	-18.0	4.7	-0.0043	-7.2	-18.2	4.4
Min. distance to the state capital (km)	155	-0.0016	-3.9	-6.5	2.2	-0.0016	-4.2	-6.9	2.2
Male agricultural wage (daily)	32.6	0.0194	12.9	79.1	18.4	0.0185	13.1	78.5	16.6
Male non-agricultural wage (daily)	3.82	-0.0016	-1.1	-6.4	5.9	-0.0014	-1.1	-5.9	5.4
Male self-employed wage (daily)	1.69	0.0051	2.5	20.8	9.5	0.0053	2.8	22.4	9.2
Female agricultural wage (daily)	13.0	0.0044	4.3	18.0	5.7	0.0042	4.6	17.7	5.2
Female non-agricultural wage (daily)	1.14	0.0074	2.6	30.0	13.2	0.0077	3.0	32.6	12.6
Female self-employed wage (daily)	1.14	-0.0002	-0.1	-0.7	10.6	-0.0005	-0.2	-2.0	10.5
Population	393	0.0004	5.5	1.4	0.4	0.0003	4.9	1.4	0.4
State dummies (reference is Guerrero)									
Hidalgo	0.17	-0.22	-3.5	-893	323	-0.21	-3.4	-899	316
Michoacan	0.12	-1.56	-18.3	-6359	1438	-1.63	-20.7	-6914	1397
Puebla	0.15	-0.44	-6.4	-1797	486	-0.45	-6.5	-1889	473
Queretaro	0.04	-1.18	-10.6	-4826	1157	-1.23	-12.4	-5199	1104
San Luis Potosi	0.15	-0.50	-6.9	-2035	527	-0.58	-8.4	-2466	556
Veracruz	0.28	-0.60	-8.3	-2438	609	-0.63	-9.0	-2686	605
Intercept	1.0	-1.34	-9.8	-	-	-	-	-	-
Endogenous variable (mean welfare)	0.079								
Number of observations	12034								
F(43, 11990) / F(42, 11992)		133				118			
R ²		0.32				0.29			

* Standard errors were computed with the delta method and are displayed in parentheses.
The t-statistics are based on Eicker-White corrected standard errors.

Several contextual variables are important determinants of household welfare. Available in only 14% of the villages, the presence of a health center increases welfare quite substantially by 964 monthly pesos. Access to state and federal roads, which helps reduce transactions costs, contribute greatly to welfare, as does proximity to an urban center. After controlling for differential asset positions and village characteristics, state effects remain significant. Relative to Guerrero, all states are poorer.

Before proceeding with the semiparametric analysis, we explore the robustness of our results and consider the possibility of endogeneity bias. Despite the dearth of data on land market activities, our assumption of treating land as exogenous depends on the absence of intergenerational transmission of unobserved characteristics that determine land endowments as well as household welfare. Given our cross-sectional data, it is difficult to imagine a meaningful household level variable that is correlated with land endowments but does not affect welfare. Consequently, we use as instruments variables that are statistically valid, but admittedly ad hoc. On that basis, we instrument land with average village farm size, the number of uneducated males, and the number of uneducated females. The first-stage regression (not shown) indicates that the instruments are strong predictors of household landholdings (F-statistic $F(3,11990) = 65.1$). The overidentification tests fail to reject the null hypothesis that the instruments are statistically valid (Davidson and MacKinnon, 1993). The IV estimations of the welfare equation produced estimates that were not statistically different from the OLS estimations (P-value = 0.18).

For another check of robustness, we estimate a village fixed-effects model to test whether any unobserved village-level variable affects the estimated return to land. The fixed-effects estimates of the return to land is slightly higher at 206 pesos but not significantly different from the OLS regression result of 125 pesos (F-statistic $F(1,11526) = 1.05$; P-value = 0.31). As long as land quality remains constant within villages, it appears that adjusting land area by its yield has done a reasonable job of controlling for land quality differences across villages. Finally, we estimate a log linear specification of landholdings to discover that, evaluated at the average land size in the sample, the marginal welfare value of land more than doubles to 361 monthly pesos. This result suggests that the choice of functional form does matter and that a semiparametric approach to the data is warranted.

Semiparametric estimation

Regression (B) displayed in Table 3 corresponds to the second step of the semiparametric procedure in estimating equation (7) above. Interestingly, the semiparametric procedure estimates coefficients very similar to the estimates of the OLS regression, suggesting that land is orthogonal to the other covariates. The resulting semiparametric estimate of the welfare value of land, $g(X)$, is shown in Figure 3, where welfare appears as an increasing concave function of land. The vertical axis refers to welfare units and includes the constant term. This result suggests that a linear specification would in fact provide a poor approximation.

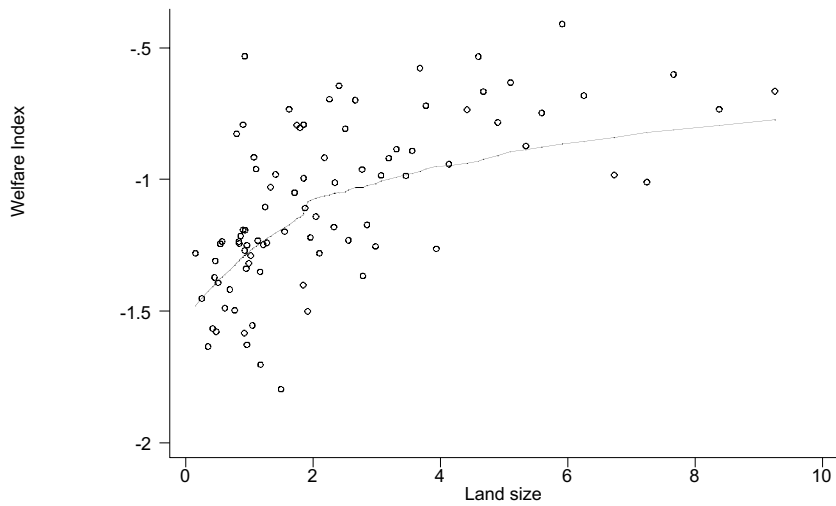


Figure 3: Welfare value of land¹⁰

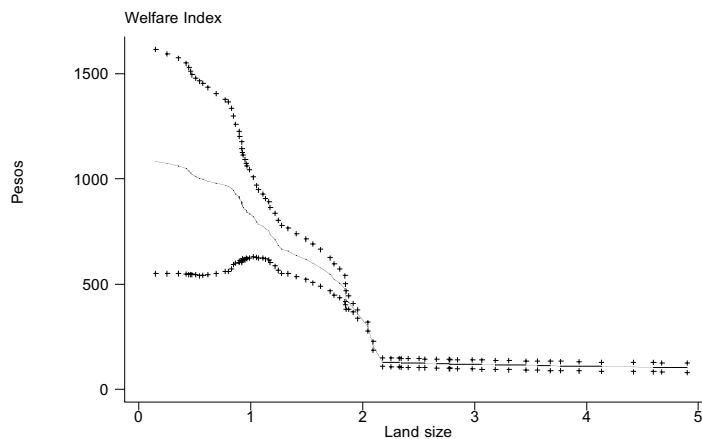


Figure 4: Marginal welfare value of land

Figure 4 displays the marginal value of land for the welfare index along with the 95% point-wise confidence intervals.¹¹ The y-axis has been normalized to the appropriate peso value. An additional hectare of land for farmers with at most one hectare is associated with an increase in their monthly welfare of 880 pesos (see Table 4). This is an important contribution to welfare as it represents 26% more than an average agricultural worker's income. The curve appears to asymptote at an estimated value that is comparable to the parametric fit, suggesting that a linear specification can grossly underestimate the marginal value of land. A simple average of the return to land measured in monthly expenditure over the range of observed farm sizes is 414 pesos, which is more than three times the benchmark case of 125 pesos.

¹⁰ To make the graph more readable, each point represents 120 households.

¹¹ It should be noted that pointwise confidence intervals do not suggest that all the estimated values jointly fall within these bounds (Cleveland and Devlin (1988)).

Table 4. Marginal value of land in monthly pesos

Farm size in RFE hectares	Number of households	Welfare index
		Average marginal value of land
≤ 1	3,557	880
1 – 3	5,180	301
> 3	3,452	102
All	12,189	414

Recalling Figure 1, the shape of Figure 4 is remarkably similar to our theoretical predictions of the impact of a labor market constraint on the return to land. Figure 4 appears to capture the fact that, for small farmers, additional land garners a return that is higher than the simple production value of the extra plot of land. Additional land will increase the marginal product value of household labor and other productive assets, thus reducing the impact of any distortions that may exist. The shape suggests that lack of rural employment opportunities is a possible constraint that these household face.

Heterogeneity in the marginal value of land

An important element of the debate on the welfare enhancing value of land is to account for complementary asset endowments and the existence of market failures. The high degree of heterogeneity in both asset positions and exposure to market imperfections suggests that the marginal value of land should not be constant across rural households. While our previous parametric and semiparametric specifications permit the household’s demographic and contextual characteristics to shift the welfare equation, we still restrict the coefficient to be the same across a very diverse sample.

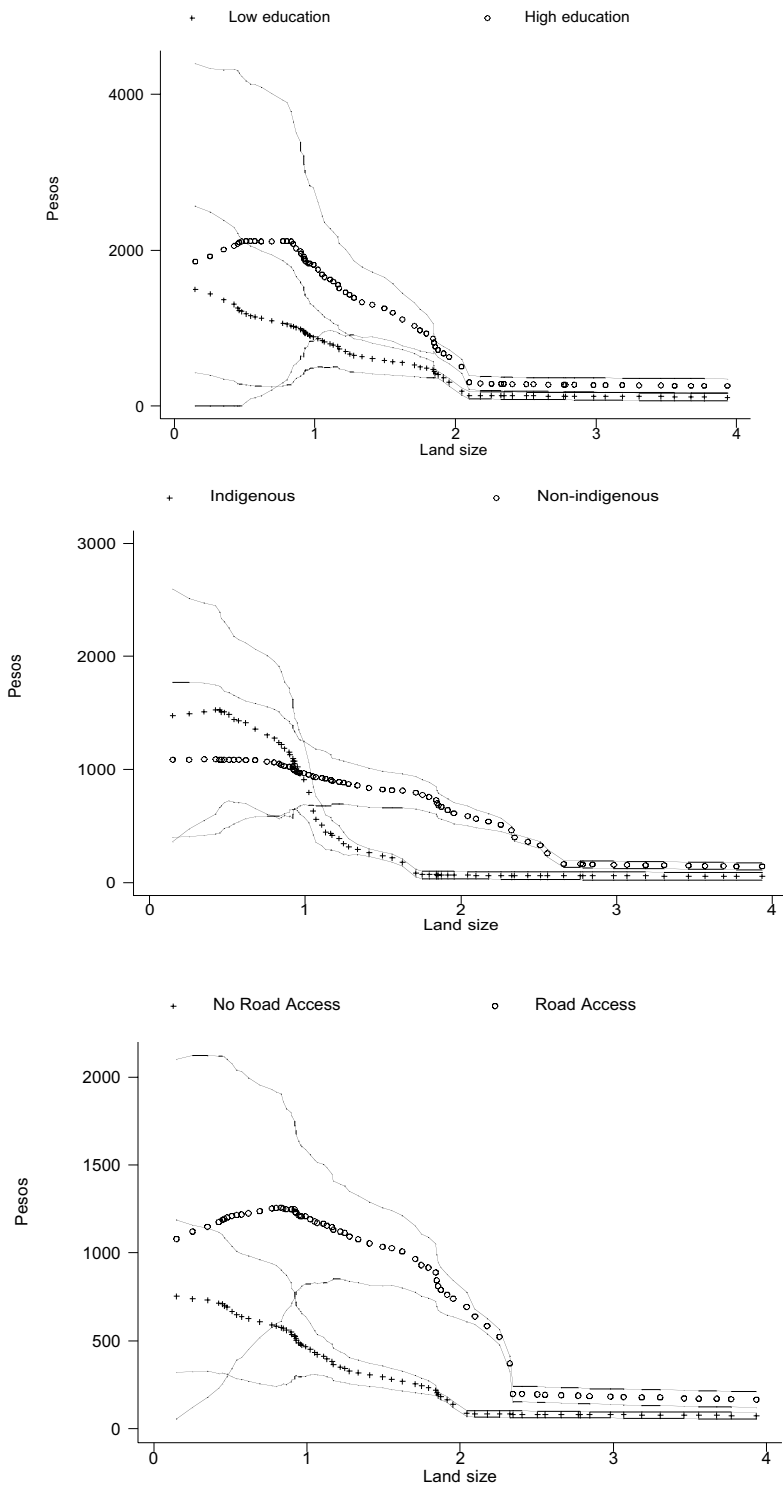
The pooling of data across different subgroups may also cause spurious curvature in the semiparametric estimation of land (Bhalotra and Attfield, 1998). Suppose, for example, that the relationship between welfare and land is actually linear, but that indigenous land is systematically smaller and produces a higher return. The marginal value of land for non-indigenous households will be lower and lie predominantly to the right of the indigenous sample. While the relationship is in fact linear for each of these subgroups, the combination of these data will result in a nonlinear representation.

Table 5 compares the return to land across subgroups of the population using parametric and semiparametric estimations of the welfare index. Comparing subgroups in the linear specification, only the access/no-access to road split records a significant difference in the return to land. However, when we relax the functional specification, the marginal value of land, as displayed in Figure 5, does vary across these subgroups, at least beyond one hectare.

Households with a higher average adult education level achieve a higher return to their land. The income generating potential of land used by indigenous households is less than half that used by non-indigenous households for landholdings greater than one hectare. Access to a paved road is important for reducing transactions costs on product and factor markets, and households residing in a village with a road garner a higher return to land.

The conditions under which land is accessed is thus fundamental in deriving welfare from land. For non-indigenous small farmers (less than one ha. of land), one additional hectare adds 393 pesos/month for those without education and without access to road, compared to 2,819 pesos for those with at least primary education and access to a road (Table 5). For those between 1 and 3 ha., the gain rises from 210 pesos/month to 2,200 pesos/month. When access to land is combined with complementary assets (education) and public goods (infrastructure), the gain in welfare from one additional hectare for these two categories of households is thus three to four times more than an average agricultural worker’s income. These results vindicate those who have been arguing for a long time that only programs of access to land accompanied by complementary reforms to secure the competitiveness of beneficiaries can achieve the desired welfare objectives (Warriner, 1969).

Figure 5: Marginal value of land across subgroups



Note that, in all the graphs, the general curvature of the land function remains consistent. This provides evidence that the curvature is not spurious.

Table 5. Marginal welfare value of land in monthly pesos

	Land Size	Sample Size	Parametric	Semiparametric		
				Farm size		
				<1 ha	1–3 ha	>3 ha
Head of household education						
No education	2.61	4108	130 (81)	1092 [1262]	462 [1764]	105 [1082]
At least primary school	3.79	2033	229 (179)	2011 [588]	1006 [878]	245 [567]
Ethnicity						
Indigenous	2.52	4800	76.7 (48)	1277 [1778]	302 [2090]	55 [932]
Non-indigenous	3.55	7234	157.5 (44)	1045 [1729]	647 [3021]	134 [2484]
Infrastructure						
No access to a road	3.09	6875	70.9** (24)	594 [2044]	253 [2983]	68 [1848]
Access to a road	3.21	5159	174 (57)	1213 [1463]	758 [2128]	155 [1568]
Education and infrastructure for non-indigenous households						
No education and no access to road		1157		393 [289]	210 [487]	33 [38]
At least primary school and access to road		619		2819 [147]	2200 [269]	345 [203]

All these regressions include the same covariates displayed in Table 3. The standard errors shown in parentheses are computed by using the delta method. The figures in brackets are the number of households with the corresponding amount of land.

(**) denotes difference across subgroups is significant at 10%.

Heterogeneity in the level of welfare

To better understand the potential of access to land as an instrument to reduce poverty, we investigate how differences in asset position and contextual setting translate into differences in welfare level. In order to decompose these differences, we re-estimate our welfare equation using a spline function for land. In this situation, a flexible parametric specification is needed since the nonparametric approach came at the expense of not being able to identify the constant term. However, the nonparametric exploration of the land to welfare contour does guide us in our choice of the number and location of the knot points, which in a spline regression can be fairly arbitrary.

In the spirit of Oaxaca's (1973) wage-gap decomposition, we explore the welfare differential between subgroups. The welfare gap, $\bar{W}_A - \bar{W}_B$, between group A and group B owning the same amount of land z can be decomposed as,

$$\bar{W}_A(z) - \bar{W}_B(z) = g_A(z) - g_B(z) + \bar{x}_A(\beta_A - \beta_B) + (\bar{x}_A - \bar{x}_B)\beta_B, \quad (9)$$

where \bar{x}_A, \bar{x}_B denote the average endowments of the determinants of household welfare, β_A, β_B the estimated parameters, and g_A, g_B the spline functions for the two subgroups of households. The first term in equation (9) represents, at a given farm size, the difference in the return to land between the two groups.

The second term captures differences in the estimated return to the other determinants x of welfare due to omitted factors. The last term is the portion of the welfare gap attributed to differences in endowments.

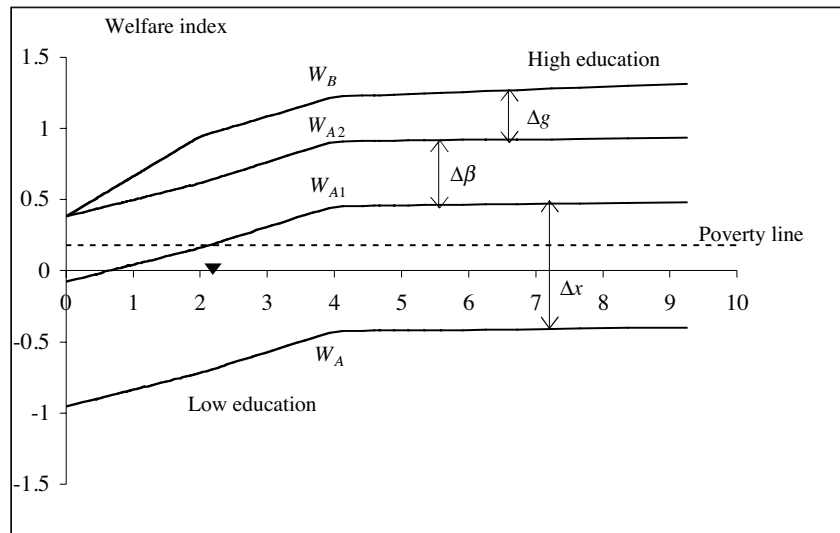
Role of complementary assets: The decomposition of the difference in predicted welfare levels between households with a low and a high average education level is depicted by landholding size in Figure 6. The different curves represent the following path in the decomposition:

$W_A = g_A(\text{land}) + \bar{x}_A \beta_A$	Welfare of group A.
$W_{A1} = g_A(\text{land}) + \bar{x}_B \beta_A$	Welfare of group A with non-land assets of group B.
$W_{A2} = g_A(\text{land}) + \bar{x}_B \beta_B$	Welfare of group A with non-land assets and return to non-land assets of group B.
$W_B = g_B(\text{land}) + \bar{x}_B \beta_B$	Welfare of group A with non-land assets, return to non-land assets, and return to land of group B = Welfare of group B.

Here, we see that access to land does little to alleviate poverty among households with low education levels.¹² Differences in endowments explain on average 57% of the welfare gap. This graph emphasizes the importance of an integrated poverty alleviation strategy. A specialized program of access to land for the poor, such as land reform, without providing access to complementary assets will, consequently, do little to reduce poverty.

Role of household characteristics: Figure 7 shows the decomposition between indigenous and non-indigenous households. In this case, eliminating the endowment differential between indigenous and non-indigenous households would be insufficient to raise the indigenous households with less than 15 hectares of land out of poverty. Non-indigenous households receive an unexplained premium that is on average 55% of the welfare differential. If indigenous households were to receive the same return to their assets as non-indigenous households, they would only require access to less than three hectares of land to reach the poverty line.

Role of the context: Figure 8 plots the difference in welfare between households with and without access to a paved road. With access to a road, households only need less than 1 hectare of land to reach the poverty line, compared to households without access who need 8. Again, we see that if there were no difference either in endowments or in this case, in the return to land, less than 3 hectares of land would be required to lift households out of poverty.



¹² The poverty line was determined by estimating the PROGRESA poverty indicator on our welfare index. The threshold value maximizes the number of correctly classified households.

Figure 6: Welfare as a function of land assets:
Role of complementary asset

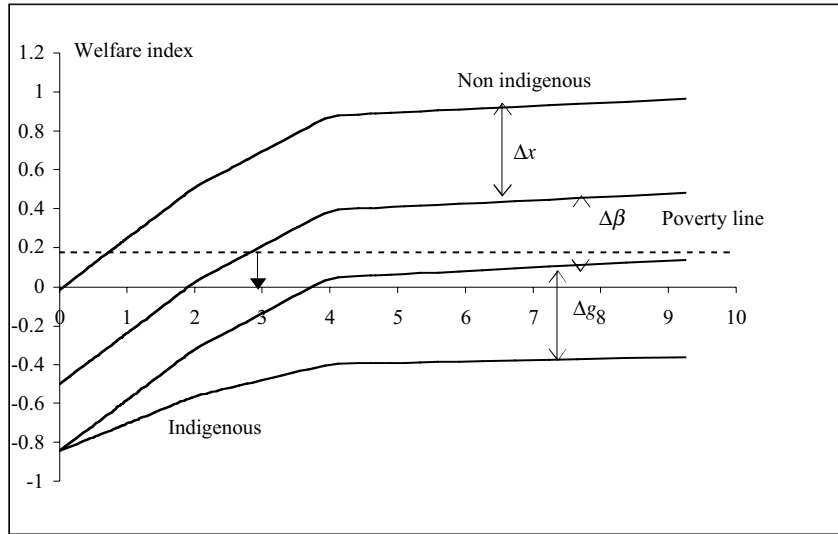


Figure 7: Welfare as a function of land assets:
Role of household characteristics

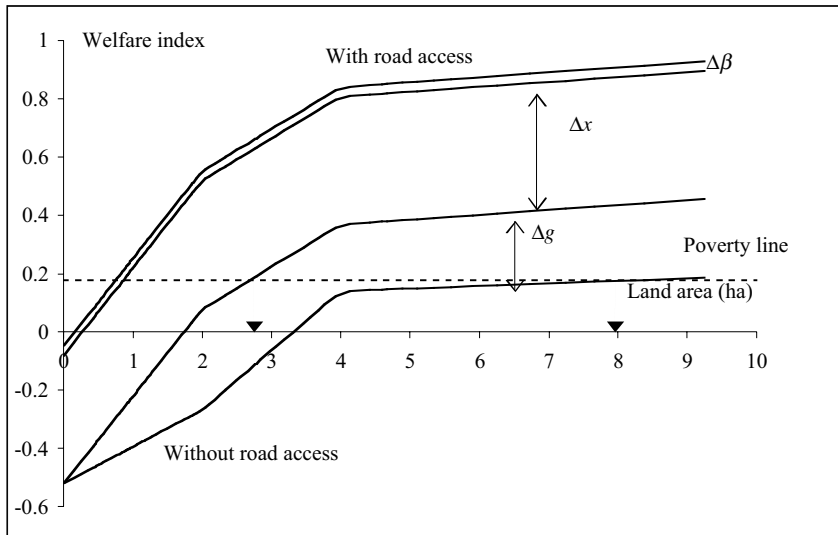


Figure 8: Welfare as a function of land assets:
Role of the context

VI. Conclusion

Access to land has for a long time been advocated as a strategy to reduce rural poverty (Warriner, 1969; Thiesenhusen, 1989; Dorner, 1992; de Janvry, 1981) and it has recently returned prominently on the policy agenda (Deininger and Feder, 2002). This is evidenced by the prevalence and magnitude of land reforms programs across the world, both past and present. However, there is an active controversy about the value of access to land in reducing poverty. Fair to say is that the conditions under which land can be effective for this purpose have not been satisfactorily established. Lopéz and Valdés (2000a) have, in particular, argued that land contributes little to income and that it is consequently better to look at other

instruments if rural poverty is a concern. Yet, the methodology that has been used in these studies has several limitations which we proceed to remedy in this paper. We use in particular a semiparametric approach to avoid restrictive specification of an income equation that theory predicts to be non-linear and complex, and a definition of poverty that goes beyond income alone as suggested by recent interdisciplinary studies (World Bank, 2001). In addition, we believe that any statement about the poverty value of land needs to be qualified in order to recognize the considerable heterogeneity of conditions under which land is accessed. This includes, most particularly, the role of household characteristics, the availability of complementary assets, and the context where land is used. To do this, we estimate the potential of land in reducing poverty in poor rural communities of Mexico in an effort to provide more accurate measurements and to identify the conditions under which access to land can help reduce poverty.

In general, we find that the marginal welfare value of land is quite high for households with less than one hectare of rainfed corn equivalence. For these farmers, an additional plot of land can increase welfare by as much as 880 monthly pesos, which represents 26% more than the average monthly income of an agricultural worker in the poor communities surveyed. This high return to land captures the increase it creates in the value of the marginal product of other household assets due to market failures, in addition to the increase in the direct production value of the land. We find that the shape of the mapping between land and welfare is consistent with a theory of labor rationing in off-farm employment.

Given the diversity of livelihood strategies across households, we investigate the role that heterogeneity plays in the ability of land to generate welfare. We observe that household characteristics, complementary assets, and contextual circumstances greatly influence the welfare generating potential of land. Social assets such as ethnicity lower the marginal value of land, whereas households with more education receive a higher return to land. Households that face lower transactions costs as measured by access to roads, garner a return to land that is two to three times as high as those without access to a road. Moreover, access to only one hectare of land can be sufficient to escape poverty for households living in villages with access to a paved road, in large part because Mexican farm households are engaged in off-farm activities that complement incomes derived from land. This suggests that land reform programs that seek massive land redistribution are not only politically difficult to achieve, but may not even be necessary, as better access to even small amounts of land can translate into significant welfare gains if this is done under conditions that allow effective use of the land and off-farm employment opportunities also exist.

These findings suggest that land can indeed be an important element of a poverty reduction strategy, but that there are specific conditions that must hold for this to be the case, calling on complementary interventions. Besides better access to land, it is important to improve control by beneficiaries over complementary assets such as education, and to improve the provision of public goods such as infrastructure needed for them to make an effective use of the land. Programs of access to land must consequently be packaged as elements of more comprehensive programs that secure the poverty reduction potential of land. We have established the nature of these complementary interventions for marginal rural communities in Mexico. They need to be systematically established for every particular population of poor for whom programs of access to land are being considered. It is only once this will have been done that the debate on the poverty reduction value of programs of access to land can be put to rest.

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