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Effect of Parental Migration on the Academic Performance of Left-behind Children in Northwestern China

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Abstract:

China's rapid development and urbanization has induced large numbers of rural residents to migrate from their homes in the countryside to urban areas in search of higher wages. As a consequence, it is estimated that more than 60 million children in rural China are left behind and live with relatives, typically their paternal grandparents. These children are called Left Behind Children (or LBCs). There are concerns about the potential negative effects of parental migration on the academic performance of the LBCs that could be due to the absence of parental care. However, it might also be that when a child's parents work in the city away from home, their remittances can increase the household's income and provide more resources and that this can lead to better academic performance. Hence, the net impact of out-migration on the academic performance of LBCs is unclear.

This paper examines changes in academic performance before and after the parents of students out-migrate. We draw on a panel dataset collected by the authors of more than 13,000 students at 130 rural primary schools in ethnic minority areas of rural China. Using difference-in-difference and propensity score matching approaches, our results indicate that generally parental migration has significant, positive impacts on the academic performance of LBCs (which we measure using standardized English test scores). Heterogeneous analysis using our data demonstrates that the positive impact on LBCs is greater for poorer performing students.

Keywords: migration, academic performance, left-behind children, difference-in-difference, rural China

JEL classification: O12, O15

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Introduction

China's rapid development and urbanization has induced large numbers of rural residents to migrate from their homes in the countryside to urban areas (Hu et al., 2008; Wen and Lin, 2012; MHRSS, 2013). In the course of migration, it is common for migrants to leave their children behind in their home communities with a surrogate caregiver (Ye et al., 2006). As a consequence, in the past decade a new population has emerged in China known as Left Behind Children, henceforth LBCs (Duan and Zhou, 2005). Statistics from the Sixth Population Census show that there were more than 61 million LBCs in China (ACWF, 2013), of which one-third are still enrolled in compulsory education (MOE, 2014).

The education of LBCs has drawn attention from researchers, though the literature is mixed concerning the direction of the effect of parental migration on the academic performance of LBCs (Yang, 2008; Chen et al., 2009; Lahaie et al., 2009; Giannelli and Mangiavacchi, 2010; Chang et al., 2011; Antman, 2012; Lu, 2012; Wang, 2014; Xu and Xie, 2015; Roy et al., 2015). In some cases, researchers have found a positive relationship between parental migration and academic performance of LBCs (Yang, 2008; Chen et al., 2009; Roy et al., 2015). Research finds that this may occur through mechanisms such as relaxing household liquidity constraints (Du et al., 2005) and encouraging higher investments in LBCs (Edwards and Ureta, 2003; Yang, 2008; Lu and Treiman, 2011; Antman, 2012; Ambler et al., 2015; Malik, 2015). However, some researchers claim that they have identified negative effects of parental migration on the educational outcomes of LBCs (Meyerhoefer and Chen, 2011; Zhao et al., 2014; Zhou et al., 2014; Zhang et al., 2014). These researchers find that the

negative effects are mainly due to the absence of parental care (Lahaie et al., 2009; Ye and Lu, 2011) or to the increased time LBCs spend at home doing on-farm or in-home work (Chang et al., 2011; McKenzie and Rapoport, 2011). Additionally, other studies have found that there is no relationship between parental migration and the academic performance of LBCs (Zhou et al., 2015).

While many studies have examined the effect of being an LBC on learning and other educational outcomes, there are a number of systematic weaknesses in the literature that may account for the mixed impacts. First, some of the studies do not have a valid comparison group (e.g. Lahied, 2009; Meyerhoefer and Chen, 2011; Zhou et al., 2014). Second, many of the previous studies are based on samples that are quite small (Ye and Lu, 2011; Lu, 2012; Zhou et al., 2014). Third, some of the studies do not use careful measures of academic performance which may not serve as an objective measure of educational outcomes (Chen et al., 2009).

In addition, most studies only examine the *overall* effect of being an LBC on educational outcomes and do not consider the fact that there may be important heterogeneous effects which may account for the differences in findings among the studies. For example, only a limited number of previous studies distinguish between the effects of the paternal and maternal migrant status when estimating the impact on academic performance of LBCs (Chen et al., 2009; Antman, 2013; Wang, 2014). Also, some studies show that the effect of parental migration on academic performance varies by the gender of LBCs (Wang, 2014) or the mother's level of education (Sawyer, 2014).

The overall goal of this study is to examine the effects of parental migration on the academic performance of LBCs. To meet this goal, we have three specific objectives. First, we compare the distribution of children's scores across different

types of households. Second, we use difference-in-difference and propensity score matching approaches to examine whether parental migration affects the academic performance of LBCs. Third, we examine how the impact of parental migration varies by different sample characteristics, such as a student's gender, his/her starting academic performance, his/her household social economic status, and the level of his/her mother's education. These analyses will help us identify the heterogeneous impact of different types of household migration on the educational outcomes of LBCs.

Data

In order to achieve our objectives, we conducted two rounds of surveys: a baseline survey and an endline survey. A total of 13,055 students in 130 elementary schools participated in our study. In the following subsections, we present the study's sampling protocol and data collection approach.

Sampling

Our sampling frame was restricted to Haidong Prefecture, a poor minority area in Qinghai Province in northwest China. In order to create a sample with enough variation in household migration status to conduct our analysis, we chose to focus our study on poor, rural areas with high population densities and high rates of off-farm employment. A quarter of the population of Qinghai lives in Haidong Prefecture, even though it accounts for only about 2% of the province's total area. Additionally, of the six counties in the prefecture, five of the counties are nationally designated poor counties (National Bureau of Statistics of China, 2014). For these reasons, Haidong Prefecture was determined to be a suitable location to select our sample.

The next step in the sampling protocol was to choose the sample schools. We obtained a comprehensive list of schools in our six sample counties from each

county's education bureau. Based on these lists, we randomly selected 130 schools with classes in grades 1 to 6 in the six sample counties to be included in our sample.

We decided to focus on students in the fourth and fifth grades for two reasons. We believe that students of this age were old enough to be able to fill out their own survey forms and take a standardized examination, but also young enough that they could be followed for a sufficient period of time. In each grade, we randomly selected 2 classes (if there were more than 2 classes in the grade). On average there were 1.3 fourth grade classes and 1.4 fifth grade classes per school. All students in sample classes participated in our survey. In total, the sample included 13,055 students.

Descriptive statistics generated from our data show that the profile of sample students is fairly typical of students from rural areas. Approximately 48.2% of the sample students were girls. In the annual yearbook published by the Ministry of Education (2014), girls in rural China account for nearly the same percentage, 47 percent, of each of China's cohorts that are in rural schools.¹ The age of the students ranged between 9 and 18 years in 2003 when we conducted the baseline survey. However, 99% of the students were between the ages of 9 and 13 years.

Although at the time of the baseline survey the sample included a total of 130 schools and 13,055 students, there was some attrition (848 students) by the end of the study, primarily due to school transfers or absence due to illness/injury. This rate of attrition is low compared to other studies conducted with children in rural China (Mo et al., 2014; Lai et al., 2015) and unlikely to impact our findings. By the time of the endline survey in 2014, we were able to follow up with 12,207 students.

Data collection

¹ According to our calculation using data published by statistical yearbooks of Shaanxi, Ningxia, Qinghai, Gansu and Xinjiang, in 2013, girls in rural areas of northwest China also account for nearly the same percentage, namely, 47%, of the class.

The research group conducted two waves of surveys in the 130 sample schools. The first round of survey was a baseline survey conducted with all students in all sample schools in September 2013 at the beginning of the academic school year. The second wave of the survey was our endline survey, which was conducted at the end of June in 2014, a time that coincided with the end of the 2013-2014 academic year.

Academic performance

In each wave of the survey, the enumeration team visited all 130 schools and conducted a three-part survey. In the first part students were given a 30-minute standardized English test, the scores of which we used as our measure of student academic performance. Before each round of the survey, we tested the English test items with over two hundred 4th and 5th grade students to ensure the quality of the baseline and endline English examinations. All the questions in the endline test were different from those in the baseline test. We administered and printed the test ourselves to ensure that it was not possible for the students to prepare for the examination. Also, our enumeration team strictly proctored the test in order to minimize cheating. The team also enforced time limits for the examinations.

We use the standardized English test scores as our measure of academic performance. English test scores were measured during the endline and baseline surveys using a 30 minute English tests. The English tests were constructed by trained psychometricians. Mathematics test items for the endline and baseline tests were first selected from the standardized English curricula for primary school students in China (and Qinghai provinces in particular) and the content validity of these test items was checked by multiple experts. The psychometric properties of the test were then checked using data from extensive pilot testing. We use standardized test scores rather than raw test scores to make student performance comparable across different grades

and classes, different periods, and different cohorts. Specifically, in order to standardize each individual observation we subtracted the mean of the comparison group and divided by the standard deviation (SD) of the distribution of the comparison group (the comparison group consists of the households that neither mother nor father out migrated between the two rounds of surveys—for more details of the group, please see the subsection below). Therefore, a standardized score of 0.2 represents someone who scored 0.2 standard deviations above the average of the comparison group. We standardized scores by the grades of the students separately. Figure 1 depicts the distribution of the standardized baseline English test scores.

We chose English as our subject of study interest for two reasons. First, English is one of the main subjects included as part of the competitive examination system in China that determines entrance into both senior high school and college (McKay, 2002; Bolton and Graddol, 2012). It is a fact that for the past decade or more English takes up nearly the same share as Math and Chinese in China's national high school and college entrance exams (*zhongkao* and *gaokao*). Specifically, the share of English in the overall exam ranges from 20 percent to 25 percent.

Second, English teaching and English learning are particularly weak in poor areas of rural China (Li, 2002; Zhao, 2003; Hu, 2005; Hu, 2009). Studies have shown that a low English score is one of the largest impediments against keeping rural students from attending senior high school in China (Loyalka, 2014). Because of this, it must certainly be true that low competency in English would seriously hinder the academic progress of rural students. Due to these reasons we believe that English is an appropriate subject that we can use in the our study to measure student academic performance.²

² Before each round of the survey, we tested the English test items with over 200 fourth and fifth grade students to construct baseline and endline English exams. In doing so, our test is with moderate difficulty and high distinction

Parental migration

In order to measure the key independent variable, parental migration status, we collected detailed information on the migration histories of each student's parents. The information came from the survey questionnaire that was filled by students under the supervision of enumerators. In the questionnaire, we included a section that asked for the migration status of each parent during the past several months. As a way of cross checking, the homeroom teacher was asked to verify the information on the parental migration status of each student. Based on the information of parental migration, there are two main types of households of interest in this study: migrant households (in which at least one parent out-migrated between our baseline and endline surveys) and non-migrant households (in which neither parent out-migrated between our baseline and endline surveys).

Recognizing that the effect of parental migration on student performance may be affected by which family member out-migrates (i.e., father, mother or both), we further subdivided the migrant households into six types of households: *Any Parent Migrated* households (father, mother or both parents out-migrated), *Father Migrated Only* households, *Father Migrated* households (unconditional on mother's migration status), *Mother Migrated Only* households, *Mother Migrated* (unconditional on father's migration status), and *Both Parents Migrated* households. It should be noted that the six types of households are not mutually exclusive. For brevity, when we talk about all of these households as a group, we call them *New Migrant* households to distinguish them from households that were already in the migrant labor force by the time of the baseline survey. In addition, we define *Never Migrant* households as those

level, as shown in Figure 1.

in which both parents stayed at home in both 2013 and 2014. Appendix Table 1 contains a list of the key independent variable names and definitions.

We use these types of parental migration variables to evaluate the effects of parental migration on the academic performance of LBCs. We make use of the variation in household migration status during the period of time between the baseline and endline surveys to evaluate the effect of migration status on school outcomes. In doing so, conceptually, our sample students are being divided into a treatment group (*New Migrant* households) and a comparison group (*Never Migrant* households). Sub-treatments in this framework are carried out using the six different types of migrant households.

Other covariates

In the third part of the survey we collected data on the characteristics of the sample students. From this part of the survey we were able to create a set of demographic and socioeconomic variables. The dataset includes measures of each student's characteristics, such as *female*, *age*, *ethnic minority*, *5th grade*, *boarding student*. We also created a number of variables measuring family characteristics, including *assets*,³ *father has junior high school or higher degrees*, *mother has junior high school or higher degrees*, and *number of siblings*. This information is beneficial to our research for two reasons: first, it allows us to explore whether the effects of parental migration on the school performance of LBCs are heterogeneous across children and households; second, these variables may directly affect school performance and by controlling for them we may more efficiently measure the effect of parental migration on school performance.

³ Asset is calculated by each account of family durable goods multiplying by their prices, then sum all index and take the logarithm.

Parental Migration and Academic Performance

In this section we seek to compare the distribution of children's scores across households of different migrant status. To do so, we first describe the prevalence of migrant households. Then, we present the correlations between migration and academic performance by comparing changes in academic performance of LBCs in the periods before and after their parents out-migrated with changes in migration status.

Prevalence of Migrant Households

Similar to the state of migration in many other rural areas in China (Rozelle et al., 1999), many households were already in the migrant labor force in 2013 when we conducted the baseline survey. Of the 12,207 households in our sample, there were 5,483 (44.9%) households in which at least one parent out-migrated (Table 1, column 1, rows 1-3). Within our sample of migrant households, we found differences in prevalence among types of migration. Of the 5,483 households with out-migrants, only the father out-migrated in 2,730 households, this accounts for 22.4% of the total number of households or 49.8% of the migrant households (column 1, row 1). In contrast, only the mother out-migrated in 560 households, which accounts for 4.6% of the total households or 10.2% of migrant households (column 1, row 2). According to our data, both parents out-migrated in 2,193 households, which is 18.0% of the total number of households or 40.0% of migrant households (column 1, row 3).

In addition, our study finds that the number of new migrant households in our sample rose rapidly during the study period. Among the 6,724 households that did not have any migrating parents in 2013 (column 1, row 4), at least one of the parents in 2,205 (32.8%) of these households entered the migrant labor force between the September 2013 baseline survey and the June 2014 endline survey (column 2, row 4).

After accounting for the 1,663 households that out-migrated at baseline and returned to the village at endline (column 8, rows 1-3), the total number of migrant households rose to 6,025 households (49.4% of the total sample) by June 2014 (column 2, row 5). This rise represents a 4.5 percentage point increase from the baseline migration level. Our sample also included a subset of households that did not send out any migrant during the study period. Specifically, 4,519 households (37.0% of the total sample) did not migrate in either period (column 8, row 4). This group of households provides the comparison group against which we can measure the impact of parental migration.

Correlation between Migration and Academic Performance

Our descriptive results suggest that an analysis of the *change* in English test scores in relation to a valid comparison group (i.e., the *Never Migrant* group) is necessary to evaluate the effect of parental migration on the academic achievement of LBCs. For example, Figure 2 shows that although students from *Both Parents Migrated*, *Mother Migrated Only*, and *Mother Migrated* households scored lower than those from the *Never Migrated* households in the endline survey, on average, the scores of the students from these households were already lower before their parents migrated. An analysis that does not take into consideration the performance of students over two periods may misattribute a student's initial performance as a product of parental migration.

When we compare the change in standardized English test score from the baseline to the endline survey between students of *New Migrant* households and those of *Never Migrant* households, the average standardized English test score of students increased, ranging from 0.02 – 0.12 SD (Figure 3).⁴ This suggests that, taking into

⁴ Since we normalized raw English test scores relative to the distribution of the *Never Migrant* households, the change of scores of students of *Never Migrant* households were 0.00 SD.

consideration the baseline test scores of students, parental migration actually may have had a positive effect on the test scores of their children.

The increases in English test scores of children from *New Migrant* households, however, may not be solely explained by parental migration. Further analysis of our data reveals that school performance may be explained by many factors other than migration activities that change over time and differ between migrant and non-migrant households. For example, higher income could have a positive effect on the grades of children from migrant households over time that might offset any other adverse effects. Therefore, further multivariate analysis is needed to explore the impact of parental migration on academic performance while holding other factors constant.

Methodology

In addition to the descriptive analysis (section above), in this section we use difference-in-difference and propensity score matching approaches to examine whether parental migration affects the academic performance of their LBCs. Firstly, we employ a difference-in-difference approach to test the impact. We also use a matching approach to check and see whether our results are robust to our choice of estimating approaches. Finally, we extend the cross-sectional matching estimator to a longitudinal setting and implement a difference-in-difference matching estimation approach in attempt to control for an additional part of unobservable factors. In the following subsections, we introduce the details of those approaches in sequence.

Difference-in-Difference approach

We employ a Difference-in-Difference (hereafter, DID) approach to compare the outcomes (i.e. academic performance) of students in the treatment group before

and after the parent(s) out-migrated to students in the comparison group. This comparison produces what we call standard DID estimator. The model we estimated is restricted and unadjusted model:

$$\Delta Score_{is} = \alpha + \beta \cdot MIG_{is} + \lambda \cdot C_c + \varepsilon_{is} \quad (1)$$

where i denotes student in school s , $\Delta Score_{is}$ is the change in standardized English test score of student i in school s between baseline survey and endline survey (that is the standardized endline English test score (standard deviation) minus the standardized baseline English test score of the same student i in school s). MIG_{is} is the treatment variable which makes β the parameter of interest. In our analysis, we have six different treatments, as discussed above, namely: *Any Parent Migrated* households; *Father Migrated Only* households; *Father Migrated (Unconditional)* households; *Mother Migrated Only* households; *Mother Migrated (Unconditional)* households; *Both Parents Migrated* households. The county effect is captured by λ .

In addition to the standard DID estimator (Smith and Todd, 2005), we implemented three other DID estimators: an ‘unrestricted’ version that includes baseline outcomes as a right hand variable, an “adjusted” version that includes other covariates in addition to the treatment variable (in our case they are a series of control variables from the baseline survey), and an unrestricted/adjusted model that combines the features of both the “unrestricted” and “adjusted” model. The unrestricted and adjusted DID estimators relax the implicit restrictions in the standard DID estimator that the coefficient associated with baseline outcomes and covariates gathered from baseline survey equals one. The combination of unrestricted and adjusted DID estimators relaxes both of these assumptions.

In summary, the models to be estimated are:

The unrestricted and unadjusted model is:

$$\Delta Score_{is} = \alpha + \beta \cdot MIG_{is} + \delta \cdot Score_{is,baseline} + \lambda \cdot C_c + \varepsilon_{is} \quad (2)$$

The restricted and adjusted model is:

$$\Delta Score_{is} = \alpha + \beta \cdot MIG_{is} + \gamma \cdot X_{is} + \lambda \cdot C_c + \varepsilon_{is} \quad (3)$$

And, the unrestricted and adjusted model is:

$$\Delta Score_{is} = \alpha + \beta \cdot MIG_{is} + \delta \cdot Score_{is,baseline} + \gamma \cdot X_{is} + \lambda \cdot C_c + \varepsilon_{is} \quad (4)$$

where the term X_{is} is a vector of covariates that are included to capture the characteristics of students, their parents and households, such as *gender*, *age*, *ethnicity*, *grade*, and *number of siblings*. The data that were used to create all of the covariates were collected at the baseline survey (or before parental migration). $Score_{is,baseline}$ represents the standardized baseline English test score of student i in school s .

We also use a version of equation (4) to test for the heterogeneous effects of parental migration on the academic performance of LBCs. We do this by including an interaction term between the treatment dummy variables and the potential variables that may affect the outcome through the treatment heterogeneously. The model to test this is:

$$\Delta Score_{is} = \alpha + \beta_1 \cdot MIG_{is} + \beta_2 \cdot D_{is} + \beta_3 \cdot MIG_{is} \cdot D_{is} + \delta \cdot Score_{is,baseline} + \gamma \cdot X_{is} + \lambda \cdot C_c + \varepsilon_{is} \quad (5)$$

where the coefficients on the interaction term, β_3 , indicate the heterogeneous treatment effects. In our analysis, we include several different variables in the D matrix, including *standardized English test score in the baseline*, *female*, *ethnic minority*, *only child*, *assets*, *father has junior high school or higher degrees*, *mother has junior school or higher degrees*.

In all of the regressions, we accounted for the clustered design by constructing Huber-White standard errors clustered at the school level (relaxing the assumption that disturbance terms are independent and identically distributed within schools).

Propensity Score Matching Approach

In addition to the set of DID estimators, we also used a matching approach to check and see whether our results are robust to our choice of estimators. Rosenbaum and Rubin (1983) proposed Propensity Score Matching (henceforth, PSM) as a way to reduce the bias in the estimation of treatment effects with observational data sets. PSM allows the analyst to match a student in the treatment group with a similar student from the comparison group and interpret the difference in their academic performance as the effect of the parental migration activities when observable characteristics of *Never Migrant* and *New Migrant* households are continuous, or when the set of explanatory factors that determine parental migration contains multiple variables (Rosenbaum and Rubin, 1985). With the right data, it is possible to estimate the propensity scores of all households and compare the outcomes of *Never Migrant* and *New Migrant* households that have similar propensity scores.

In order to implement the matching estimator successfully, we follow a series of well-established steps (Caliendo and Kopeinig, 2008). First, since matching is only justified over the common support region, we check whether there is a large overlap in the support of the covariates between the *New Migrant* and *Never Migrant* households. Intuitively, wide common support means that there is a fairly large overlap in the propensity scores. In our study, the common support is fairly wide in our sample (Appendix Figure 1). This means that we can estimate the average treatment effect for the treated of a large portion of the sample.

In the second step, we choose the method of matching. In this study, we use the nearest neighbor matching method with replacement. The standard errors are bootstrapped using 1,000 replications. The last step is to assess the matching quality. Since we do not condition on all covariates but on the propensity score alone in PSM,

it has to be checked whether the matching procedure is able to balance the distribution of the relevant covariates in both the comparison and treatment group. To do so, we use balance tests described in Dehejia and Wahba (1999, 2002). The balancing tests were satisfied for all covariates.

In order to guard against the potential source of bias (shown by Abadie and Imbens, 2002), we also implemented the Bias-Corrected Matching (henceforth, BCM) estimator developed by Abadie and Imbens (2006). To minimize geographic mismatch, we enforce exact matching by county. Each treatment observation is matched to three control observations with replacement, which is few enough to enable exact matching by county for nearly all observations but enough to reduce the asymptotic efficiency loss significant (Abadie and Imbens, 2006). Matching is based on a set of 9 covariates, including *female*, *age*, *ethnic minority*, *5th grade*, *boarding student*, *assets*, *father has junior high school or higher degrees*, *mother has junior or higher degrees*, and *number of siblings*, which are time-invariant or were measured in the baseline survey (see Table 2). The weighting matrix uses the Mahalanobis metric, which is the inverse of the sample covariance matrix of the matching variables.

Finally, since all matching methods only match observations based upon observable covariates, they do not account for all unobservable covariates. To control for part of the unobservable factors, in particular, those factors that are time-invariant, we extended the cross-sectional matching estimator to a longitudinal setting and implemented a Difference-in-Difference Matching estimator (henceforth, DDM). When implementing DDM, we use both PSM and BCM methods.

Results of Multivariate Analysis

The results from the DID analysis using Models (1) - (4) for the version of model that uses the *Any Parent Migrated* household variable as the treatment demonstrate that the models perform fairly well and are consistent with our intuition. The coefficients on some of the control variables are also in accordance with our intuition (Table 3). For example, when we use the unrestricted and adjusted specification of the empirical model (column 4), the scores of older students drop relatively more than those of younger students (column 4, row 3). This finding might be considered reasonable since, *ceteris paribus*, students that enter elementary school at an older age may have an initial advantage (because they are relatively more mature) that gradually disappears as younger children catch up over the course of primary school, which is consistent with other findings (for example, Fredriksson and Öckert, 2005). Additionally, when a student is from a household that is part of a non-Han ethnic minority group, the student's score drops relatively more than Han students. There are many studies that show the academic performance of ethnic minority students in China lags substantially behind those of Han students (Gustafsson and Sai, 2014; Yang et al., 2015). In the rest of paper, we focus mainly on the results from the unrestricted and adjusted model. We do so because this regression has a higher goodness of fit (or *R*-square) statistic. In part, almost certainly, this better fit reflects the importance of capturing beginning scores, which embodies the unobserved ability of a student as well as other covariates.

One of the most important findings in Table 3 is that we reject the hypothesis that parental migration negatively affects the academic performance of children. In all four models, the coefficient of the *Any Parent Migrated* household dummy variable is not negative. In fact, the coefficients are all positive and significantly different from zero. The magnitudes of the coefficients range from 0.04 to 0.08 SD. This means that,

everything else held constant, after any parent in a household out-migrated between baseline and endline surveys, their child's standardized English test scores actually rose relative to the children of *Never Migrant* households. In other words, unlike claims made by some researchers (Meyerhoefer and Chen, 2011; Zhao et al., 2014; Zhou et al., 2014; Zhang et al., 2014), according to the results of the analysis in Table 3, parental migration did not hurt the academic performance of LBCs. At least in the migrant households in our sample area, migration led to improved school performance.

The results hold when we examine other types of migrant households: we do not find any negative effects of parental migration on the academic performance of children (Table 4). For each of the four specifications, we look at the effect of parental migration on academic performance in all six types of migrant households.⁵ In 22 out of the 24 cases the coefficient is positive. The coefficients are only negative for *Mother Migrated Only* households (column 2, row 4). In each of these two cases, however, there is no statistically significant effect of parental migration on academic performance. Interestingly, when the father out migrated (column 4, row 3) or mother out migrated (unconditional) (column 4, row 5) or both parents out migrated (column 4, row 6), the standardized English test scores of LBCs improved significantly.

So why is it that parental migration does not appear to have a negative effect on the academic performance of LBCs and in some cases even appears to have a positive effect? Although we cannot answer this question from our analyses, one possible reason is that the income effect of remittances is relatively large compared to the adverse effect of less parental care. If parental migration leads to higher income,

⁵ In Table 3, we only report the coefficients on the treatment variable. The rest of the results are suppressed for brevity but are available from the authors upon request. We report the results for 24 different regressions. For completeness in Table 4, we include the results of the effect of Any Parent Migrated on school performance, but, in fact, this is a duplication of the results from Table 3, row 1.

as found in Du et al. (2005), the migrant households that experience rising incomes may be able to provide better nutrition, improved access to educational supplies and burden their children with less housework. With the addition of these household resources and the lessening of burdens, parental migration may have a positive effect on the academic performance of children. The positive income effect is probably behind our finding that the largest positive effects are found in the *Both Parents Migrated* households (Table 4, row 6). This result may arise since the family income would improve more when both parents out migrated comparing to other types of *New Migrant* households. This income effect also appears to be largely offsetting any negative effects of parental absence—such as the decline in parental care and oversight. Thus, on the whole, our results strongly suggest that parental migration is having net positive effect on the academic performance of LBCs when both parents out-migrate.

Results from Matching

The results of the cross-sectional matching analysis, regardless of the method of matching, also reveal that parental migration has no significant negative effect on the academic performance of LBCs. When propensity score matching is used to examine the effect of parental migration on academic performance for all six types of *New Migrant* households, there are no cases in which the coefficient on the treatment variable is negative and statistically significant (Table 4, column 1, rows 1a, 2a, 3a, 4a, 5a and 6a). The same is true when Bias-Corrected Matching is used (column 1, rows 1b, 2b, 3b, 4b, 5b and 6b).

In fact, results from matching are quite similar to those from the DID analyses. When we use Bias-Corrected Matching, which arguably is generating more reliable

estimates and standard errors, we find that the coefficients on the treatment variables in the *Any Parent Migrated* household, the *Father Migrated* (unconditional) household, the *Mother Migrated* household and the *Both Parents Migrated* household are positive and statistically significant. The magnitudes of the coefficients also are similar to those from the DID analyses.

In addition, and importantly, the findings remain largely the same when the DDM estimator is used (Table 4, column 2). Regardless of whether we use PSM or BCM, none of the coefficients of any of the treatment variables are significantly negative. In fact, most of them are positive and significant. Hence, whether using DID, PSM or DDM, we find no significant negative effects of parental migration on the academic performance of left-behind children.

Heterogeneous Effect of Parental Migration on Academic Performance

While we have found no significant negative impacts, mostly positive impacts, of parental migration on the academic performance of LBCs, all of these results have been for the average household (that is, for the typical migrant households). It is possible, however, that the impacts could vary for different subgroups, i.e., different types of migrant households, of our sample. In this section we use model (5) which is presented in Section 4 to test the heterogeneous effects of several variables.

Specifically, we will look at the heterogeneous effects of parental migration on: students who are poor and higher performing (using *standardized English test score in the baseline*); students of different gender (*female*); students who are Han and non-Han (*ethnic minority*); students who are an only child or who have siblings (*only child*); students from poorer and richer families (*assets*), students with parents who have lower and higher levels of education (using either *father has junior high school*

or higher degrees, or mother has junior high school or higher degrees). For brevity, we only report the results of the unrestricted and adjusted model, but the results are robust to this specification of the model.

The heterogeneous analysis shows that the positive impact of parental migration on LBCs is greater for poor performing students (Table 6, columns 4-6, row 1). These results mean that, everything else held constant, parental migration affects academic performance of LBCs with different starting academic performance in a heterogeneous way. Although it is beyond the scope of this paper to isolate the exact reason why parental out-migration helps poorer performing students more than higher performing students, it may be the additional resources that are available to households from newly out-migrating parents are able to overcome one or more of the educational barriers that were limiting the performance of the students (making them poorer performing). For example, it has been shown in a number of papers that when students are better nourished, their academic performance rises (Luo et al., 2012). If the newly available remittances were used to improve nutrition in the households where parents had recently left, this might lead to better academic performance by students that originally were not being provided enough nutrition and, hence, performed at a sub-par level. Remittances might also be used for other performance-enhancing investments of households with poorer performing students, such as, remedial tutoring or additional books or learning software and associated computer hardware.

The results of the heterogeneous analysis also demonstrate that the positive impact of parental migration on LBCs maybe offset if the mother of an LBC has at least junior high school degree. The coefficient on the interaction term between that variable indicating *Mother Migrated* (unconditional) households and mother's

education level – Mother has at least junior high school degree or not – is -0.10 SDs and is significant at the 5% level (Table 6, columns 5, row 6). Hence, if a student's mother does not have at least junior high school degree, the scores of LBCs would improve when mothers out-migrate or both parents migrate. In contrast, in the case of mothers with higher levels of education, the positive effects of out-migration (that are found for the average student) could be offset. While (again) we do not know exactly why, the results are consistent with the interpretation that there is a parental care-household resources trade-off when the mother of the student has the ability (from her higher level of education) to provide academic performance-enhancing care (e.g., from time spent tutoring her child). However, when a student's mother is poorly educated, she may not have the ability to help her child with his/her studies and so when she leaves and begins to earn an income providing the household with additional resources there is a net positive gain. Since our subject of interest is English, this interpretation is more reasonable. Compared to math or mandarin, higher education is probably needed for the mother of an LBC so that she could tutor her child in learning English. The results are similar to those found in previous studies which find that the impact of parental migration on academic performance of LBCs differs based on the background of parents, especially for the education levels of parents (e.g. Sawyer, 2014).

As is shown in Table 6, we find no significant evidence of heterogeneous effects for other student demographic and family characteristics, including gender, ethnic minority, only child, asset and father education level (Table 5, rows 2-5, 7). In other words, like the results for the average households reported in Tables 3 to 5, the results from DID analysis demonstrate that there are no significant effects of parental migration on the academic performance of LBCs and this is true in the case of either:

boys or girls; Han or non-Han ethnic minorities; only children or children with siblings; as well as children from families in which the father has at least a junior high school degree or not.

Interestingly, although ethnicity does not matter when we aggregate all ethnic minorities into a single group, the impacts do differ when we differentiate minorities by sub-population. In other words, we do find heterogeneous impacts when students are Tibetan (versus the impacts when they are non-Tibetan) and when students are members of the Tu minority. Compared to the students in the *Never Migrant* households, Tibetan students in the *Any Parent Migrated* households (+0.11 SDs), *Father Migrated* households (+0.10 SDs), *Mother Migrated Only* households (+0.26 SDs) and *Mother Migrated* households (+0.13 SDs) improved more in their standardized English test scores than Han students (Columns 1, 3-5, row 3a). In contrast, Tu students in *Mother Migrated Only* households (-0.62 SDs) and *Mother Migrated* (-0.22 SDs) lagged behind in their standardized English test scores than Han students in those treatment subgroups, respectively (Columns 4-5, row 3b).

So what is happening? From our data, we find that Tibetan students perform worse than other students ($t=10.49$, $p<0.01$). This result is consistent with results from above that the positive impact on LBCs is greater for poor performing students. Our data also show that the education level of parents of Tu students is significantly higher than that of other non-Han ethnic minorities ($t=4.70$, $p<0.01$). As what we have discussed above, when mothers of LBCs have the ability (from their higher level of education) to provide academic performance-enhancing care, the positive impact of parental migration on LBCs may be offset.

Conclusions

In this paper, we have tried to understand whether or not the academic performance of LBCs suffers when their father, mother or both parents migrate from their home communities to the city (or at least away from home). Despite the perception that is commonly found in the literature and the popular press, our results – somewhat surprisingly – show that there seem to be no significant negative effects of parental migration on the academic performance of LBCs. Comparing the change in the standardized English test scores before and after parents out-migrated between children from migrant households and those from non-migrant households, we can reject the hypothesis that parental migration negatively affects the academic performance of LBCs. In fact, in the analysis of most migrant households, especially in those in which any parent migrated, father migrated, mother migrated or both parents migrated, migration is shown to have a statistically significant and positive effects on the academic performance of LBCs.

In addition, we also sought to explain how the impacts of parental migration vary for different subgroups, i.e., different types of migrant households, of our sample. Results from our data show that the positive impact of parental migration on LBCs is greater for poor performing students. However, the positive impact maybe offset if the mother of an LBC has at least junior high school degree. In contrast, we find no evidence of heterogeneous effects by other student demographic and family characteristics, including gender, ethnic minority, only child, asset and father education level.

Based on these results, it might be tempting to conclude that policy makers do not need to take any action (to help LBCs) since there are no measurable negative effects of migration on school performance. If there were, education officials might want to reduce class sizes or hire more qualified teachers to improve the mentoring

program in schools in which there were many LBCs. Boarding schools might offer some of the services that parents originally carried out before they entered the migrant labor force. Ultimately, measures can be promoted to offer the children of migrants who lived in China's cities better access to urban schools so parents would not have to leave their children behind. However, all of these programs are costly. Although there might be good reason to implement such policies anyhow, according to our results, they should not be carried out on the ground of the negative effect of migration on school performance. In addition, since both groups of children (children from migrant households and non-migrant households) perform poorly on academic performance, we recommend that special programs designed by policy makers to improve education among left-behind children be expanded to cover all children in rural China.

Although we have tried a number of alternative approaches to identify the effect of migration, and although the findings are largely robust, if the assumptions underlying our methodologies were not valid, our estimates could be biased. Even though we controlled for many observed and time-invariant unobserved factors, there still may be factors that are known to the parents of migrants and potential migrants but are not observable to the econometrician. For example, it may be that all parents who were in the village with their children in the baseline survey worry about whether or not their migration decision would negatively affect the academic performance of their children. If it is the case that those parents who – though having an opportunity to migrate – believed that the grades of their children would suffer decided not to migrate, while those that believed their children's grades would not suffer decided to migrate, then our results would be subject to selection bias.

If there was, in fact, such a selection bias and we did not account for it (as we were unable to – due to the absence of any effective instrumental variable), would our results be useless? We believe not. We believe even if there was a selection bias our results are showing that when rural parents out-migrate, the academic performance of their children do not suffer. It is true that part of the reason for the non-negative effect may be exactly this selection effect – parents do not go when they believe the scores of the children would suffer. But, from society's point of view, there is less cost in terms of academic performance of its children due to parental migration.

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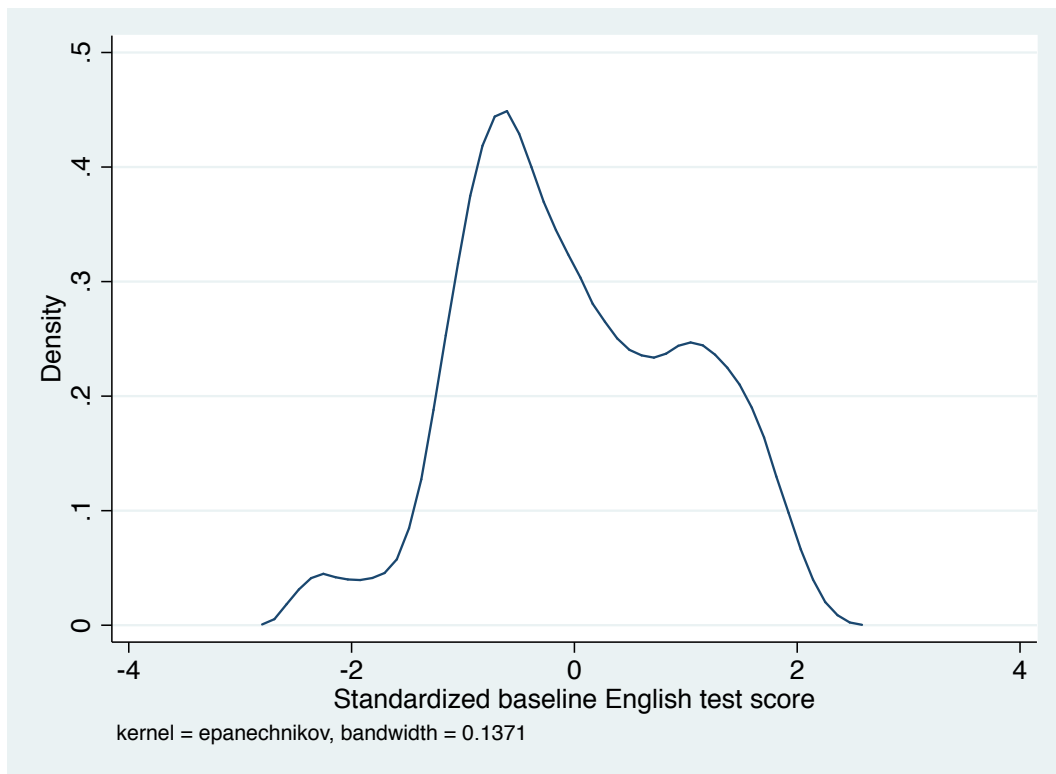


Fig 1. Kernal distribution of Standardized baseline English test score.

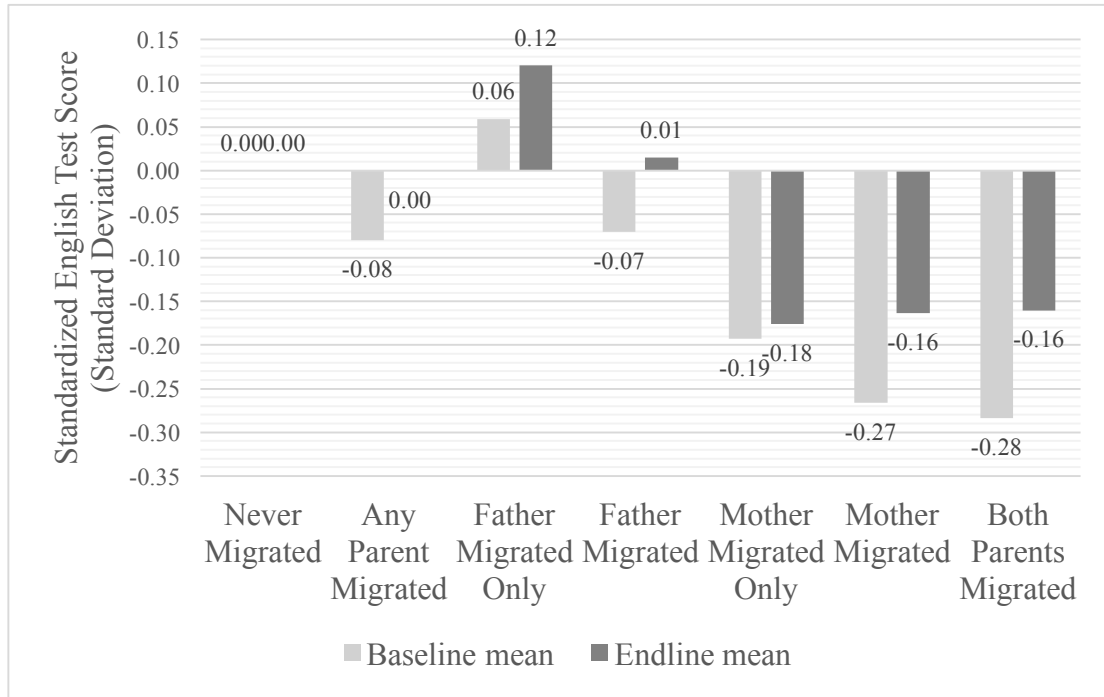


Fig 2. The standardized English test score of baseline and endline survey in different migrant households.

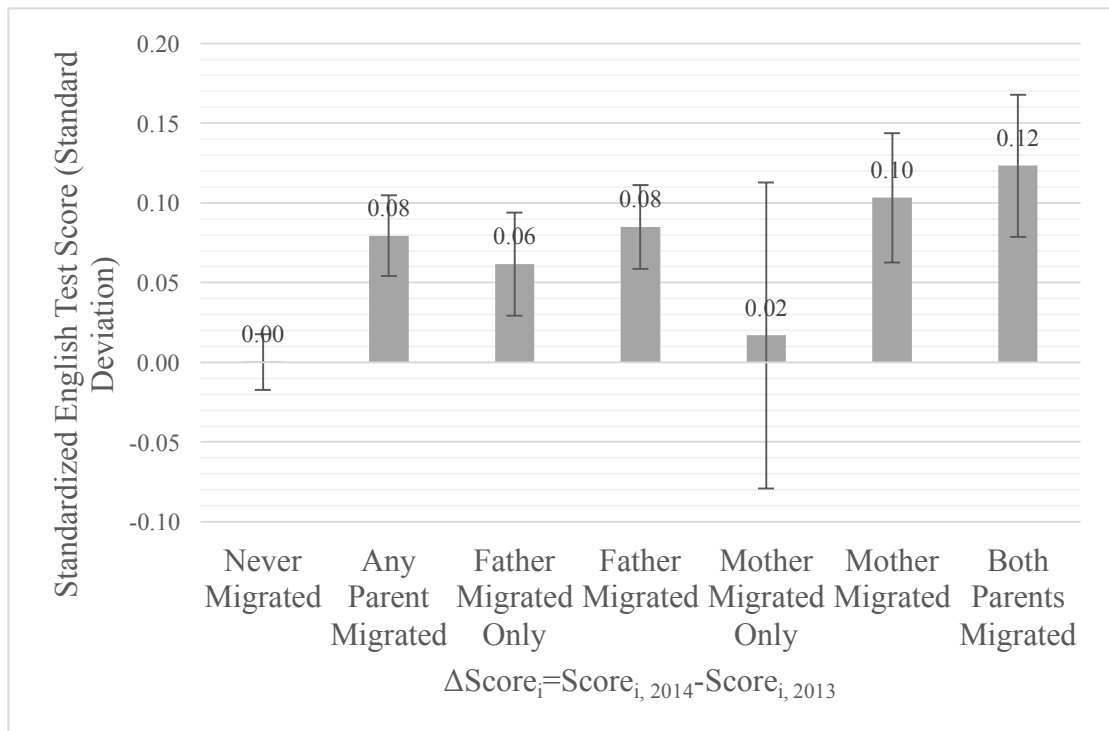


Fig 3. Change in standardized English test score before and after the parents of students out-migrate with 95% Confidence Interval (*CI*) in different migrant households.

Table 1. Patterns of migration in sample households in 2013 and 2014, Qinghai Province, China.

	Migration status in 2013	Migration status in 2014						
	(1)	(2)	(3)	(4)	(5)	(6)	(7)	(8)
	Number of households in 2013	Any Parent Migrated in 2014	Father Migrated Only in 2014	Father Migrated in 2014	Mother Migrated Only in 2014	Mother Migrated in 2014	Both Parents Migrated in 2014	Neither parent migrated in 2014
[1] Father migrated only	2730	1779	1358	1728	51	421	370	951
[2] Mother migrated only	560	353	63	221	132	290	158	207
[3] Both parents migrated	2193	1688	305	1579	109	1383	1274	505
[4] Neither parent migrated	6724	2205	1264	2028	177	941	764	4519
[5] Total number of households	12207	6025	2990	5556	469	3035	2566	6182

Data source: Authors' survey.

Column (1)=Column (3)+Column (5)+Column (7)+Column (8); Column (2)=Column (3)+Column (5)+Column (7); Column (4)=Column (3)+Column(7); Column (6)=Column (5)+Column (7).

The households in column 8, rows 1, 2 and 3 are return migrants (or those households in which households had a migrant in 2013 and in 2014 had returned home). These households are dropped from the multivariate analysis.

The households in row 1, columns 2-7; row 2, columns 2-7; row 3, columns 2-7 are always migrant households. These households are dropped from the multivariate analysis.

Total new migrants (or those households in which the parents did not migrate in 2013 and migrated in 2014) is found in column 2, row 4. Never migrants is found in column 7, row 4.

Table 2. Descriptive statistics of control variables used in the multivariate analysis.

	Total	Never migrant	Any Parent migrated		Father migrated only		Father migrated		Mother migrated only		Mother migrated		Both parents migrated	
Control variables	mean (s.e.)	mean (s.e.)	mean (s.e.)	H0: (2)=(3) Difference	mean (s.e.)	H0: (2)=(5) Difference	mean (s.e.)	H0: (2)=(7) Difference	mean (s.e.)	H0: (2)=(9) Difference	mean (s.e.)	H0: (2)=(11) Difference	mean (s.e.)	H0: (2)=(13) Difference
	(1)	(2)	(3)	(4)	(5)	(6)	(7)	(8)	(9)	(10)	(11)	(12)	(13)	(14)
Characteristics of the students														
[1] Female (1=female; 0=male)	0.49 (0.50)	0.49 (0.50)	0.48 (0.50)	-0.01 (0.01)	0.49 (0.50)	0.00 (0.02)	0.48 (0.50)	-0.01 (0.01)	0.45 (0.50)	-0.04 (0.04)	0.45 (0.50)	-0.03* (0.02)	0.46 (0.50)	-0.03 (0.02)
[2] Age (years)	10.83 (1.11)	10.77 (1.10)	10.97 (1.13)	0.20*** (0.04)	10.92 (1.12)	0.18*** (0.04)	10.97 (1.13)	0.20*** (0.04)	10.95 (1.09)	0.15* (0.08)	11.03 (1.13)	0.21*** (0.06)	11.05 (1.14)	0.23*** (0.06)
[3] Ethnic minority (1=yes; 0=no)	0.51 (0.50)	0.50 (0.50)	0.54 (0.50)	0.04* (0.02)	0.49 (0.50)	0.03* (0.02)	0.54 (0.50)	0.04* (0.02)	0.54 (0.50)	0.02 (0.04)	0.61 (0.49)	0.04 (0.03)	0.62 (0.48)	0.05 (0.03)
[3a] Hui Minority (1=yes; 0=no)	0.30 (0.46)	0.32 (0.47)	0.26 (0.44)	-0.05* (0.03)	0.24 (0.43)	-0.03** (0.02)	0.26 (0.44)	-0.05* (0.03)	0.26 (0.44)	-0.08** (0.04)	0.29 (0.45)	-0.07 (0.04)	0.30 (0.46)	-0.07 (0.05)
[3b] Tibetan (1=yes; 0=no)	0.11 (0.31)	0.08 (0.27)	0.17 (0.37)	0.08*** (0.02)	0.12 (0.32)	0.05*** (0.01)	0.16 (0.37)	0.08*** (0.02)	0.20 (0.40)	0.11*** (0.04)	0.23 (0.42)	0.13*** (0.04)	0.23 (0.42)	0.13*** (0.04)
[3c] Tu minority (1=yes; 0=no)	0.05 (0.21)	0.04 (0.19)	0.06 (0.24)	0.02** (0.01)	0.08 (0.27)	0.03*** (0.01)	0.06 (0.24)	0.02** (0.01)	0.05 (0.22)	0.01 (0.02)	0.04 (0.20)	0.01 (0.01)	0.04 (0.19)	0.01 (0.01)
[4] 5th grade (1=yes; 0=no)	0.52 (0.50)	0.53 (0.50)	0.51 (0.50)	-0.02 (0.02)	0.52 (0.50)	-0.01 (0.02)	0.51 (0.50)	-0.02 (0.02)	0.53 (0.50)	0.00 (0.04)	0.50 (0.50)	-0.03 (0.02)	0.49 (0.50)	-0.04* (0.02)
[5] Boarding student (1=yes; 0=no)	0.15 (0.36)	0.12 (0.32)	0.22 (0.41)	0.10*** (0.02)	0.20 (0.40)	0.07*** (0.02)	0.22 (0.41)	0.10*** (0.02)	0.19 (0.39)	0.07** (0.03)	0.24 (0.43)	0.13*** (0.03)	0.26 (0.44)	0.14*** (0.03)
Characteristics of the parents and the households														
[6] Log (asset)	9.59 (0.47)	9.61 (0.39)	9.55 (0.60)	-0.06*** (0.02)	9.55 (0.50)	-0.06*** (0.02)	9.56 (0.54)	-0.05*** (0.02)	9.52 (1.06)	-0.08 (0.08)	9.55 (0.70)	-0.04 (0.03)	9.56 (0.59)	-0.04 (0.02)
[7] Father has at least junior high school degree (1=yes; 0=no)	0.49 (0.50)	0.52 (0.50)	0.44 (0.50)	-0.07*** (0.01)	0.49 (0.50)	-0.05*** (0.02)	0.43 (0.50)	-0.08*** (0.02)	0.48 (0.50)	-0.02 (0.04)	0.37 (0.48)	-0.10*** (0.02)	0.35 (0.48)	-0.12*** (0.02)
[8] Mother has at least junior high school degree (1=yes; 0=no)	0.36 (0.48)	0.39 (0.49)	0.30 (0.46)	-0.08*** (0.01)	0.32 (0.47)	-0.09*** (0.02)	0.29 (0.45)	-0.09*** (0.02)	0.36 (0.48)	-0.00 (0.03)	0.26 (0.44)	-0.08*** (0.02)	0.24 (0.43)	-0.10*** (0.02)
[9] Number of siblings	1.57 (1.51)	1.54 (1.45)	1.65 (1.61)	0.10** (0.05)	1.53 (1.41)	0.07 (0.05)	1.64 (1.61)	0.10** (0.05)	1.67 (1.67)	0.07 (0.11)	1.81 (1.84)	0.13* (0.07)	1.84 (1.87)	0.14* (0.08)
[10] Number of Observations	6724	4519	2205	6724	1264	5783	2028	6547	177	4696	941	5460	764	5283

Data source: Author's survey.

* significant at 10%; ** significant at 5%; *** significant at 1%. Mean values are reported in the table with robust standard errors in parentheses clustered at school level. County dummies are controlled.

The within-school difference between the column (2) and column (3) is calculated by regressions of each of row variables on the dummy variable that represent *Any Parent Migrated* households. The within-school difference between the column (2) and column (5) is calculated by regressions of each of row variables on the dummy variable that represent *Father Migrated Only* households. The within-school difference between the column (2) and column (7) is calculated by regressions of each of row variables on the dummy variable that represent *Father Migrated* households. The within-school difference between the column (2) and column (9) is calculated by regressions of each of row variables on the dummy variable that represent *Mother Migrated Only* households. The within-school difference between the column (2) and column (11) is calculated by regressions of each of row variables on the dummy variable that represent *Mother Migrated* households. The within-school difference between the column (2) and column (13) is calculated by regressions of each of row variables on the dummy variable that represent *Both Parents Migrated* households.

Table 3. Difference in difference regression results analyzing the effects of migration activities of parents on school performance of students, Qinghai Province, China.

Dependent variable: $\Delta \text{Score}_i = \text{Score}_{i, 2014} - \text{Score}_{i, 2013}$	Restricted & Unadjusted	Unrestricted & Unadjusted	Restricted & Adjusted	Unrestricted & Adjusted
VARIABLES	(1)	(2)	(3)	(4)
<i>Treatment variable (MIG)</i>				
[1] Any Parent Migrated (1=yes; 0=no)	0.08*** (0.02)	0.04** (0.02)	0.06*** (0.02)	0.04** (0.02)
<i>Characteristics of the students</i>				
[2] Female (1=female; 0=male)			0.07*** (0.02)	0.15*** (0.02)
[3] Age (years)			-0.01 (0.01)	-0.05*** (0.01)
[4] Ethnic minority (1=yes; 0=no)			-0.00 (0.03)	-0.05* (0.02)
[5] 5th grade (1=yes; 0=no)			0.04 (0.04)	0.07** (0.03)
[6] Boarding student (1=yes; 0=no)			0.17*** (0.04)	0.08** (0.04)
<i>Characteristics of the parents and the households</i>				
[7] Log (asset)			-0.01 (0.02)	-0.00 (0.01)
[8] Father has at least junior high school degree (1=yes; 0=no)			-0.02 (0.02)	0.03** (0.01)
[9] Mother has at least junior high school degree (1=yes; 0=no)			-0.07*** (0.02)	-0.03 (0.02)
[10] Number of siblings			0.00 (0.01)	-0.01 (0.01)
[11] Standardized pre English test score (standard deviation)		-0.35*** (0.03)		0.29* (0.16)
[12] County dummy	YES	YES	YES	YES
[13] Constant	0.02 (0.03)	-0.19*** (0.04)	0.24 (0.18)	6,724 0.20
[13] Number of observations	6,724	6,724	6,724	6,724
[14] R-squared	0.01	0.17	0.03	0.20

Data source: Author's survey.

* significant at 10%; ** significant at 5%; *** significant at 1%. Robust standard errors in parentheses clustered at school level.

Table 4. Difference in difference regression results analyzing the effects of migration activities of parents on school performance of students in all six types of migrant households, Qinghai Province, China.

Dependent variable: $\Delta \text{Score}_i = \text{Score}_{i, 2014} - \text{Score}_{i, 2013}$

	Restricted & Unadjusted	Unrestricted & Unadjusted	Restricted & Adjusted	Unrestricted & Adjusted
Treatment variables	(1)	(2)	(3)	(4)
[1] Any Parent Migrated	0.08*** (0.02)	0.04** (0.02)	0.06*** (0.02)	0.04** (0.02)
No. of observations	6,724	6,724	6,724	6,724
R ²	0.01	0.17	0.03	0.20
[2] Father Migrated Only	0.06*** (0.02)	0.03 (0.02)	0.04** (0.02)	0.03 (0.02)
No. of observations	5,783	5,783	5,783	5,783
R ²	0.01	0.16	0.02	0.19
[3] Father Migrated (unconditional)	0.08*** (0.02)	0.04** (0.02)	0.06*** (0.02)	0.04** (0.02)
No. of observations	6,547	6,547	6,547	6,547
R ²	0.01	0.17	0.03	0.20
[4] Mother Migrated Only	0.01 (0.05)	-0.04 (0.04)	0.00 (0.04)	-0.04 (0.04)
No. of observations	4,696	4,696	4,696	4,696
R ²	0.00	0.17	0.03	0.20
[5] Mother Migrated (unconditional)	0.10*** (0.02)	0.04* (0.02)	0.07*** (0.02)	0.04* (0.02)
No. of observations	5,460	5,460	5,460	5,460
R ²	0.01	0.18	0.03	0.21
[6] Both Parents Migrated	0.12*** (0.03)	0.06** (0.03)	0.09*** (0.02)	0.07*** (0.02)
No. of observations	5,283	5,283	5,283	5,283
R ²	0.01	0.18	0.03	0.21

Data source: Author's survey.

* significant at 10%; ** significant at 5%; *** significant at 1%. Robust standard errors in parentheses clustered at school level. County dummies are controlled.

The full version of the regressions from models (1) – (4) is not reported for brevity purpose but is available from the authors upon request.

Table 5. Evaluating the effects of migration activities of parents on school performance of students in all six types of migrant households using matching and difference-in-difference matching, Qinghai Province, China.

Treatment variables	(1)			(2)		
	Matching			Difference-in-difference matching		
	Average treatment effect for the treated	Std. Err.	t-stat/z-value	Average treatment effect for the treated	Std. Err.	t-stat/z-value
Any Parent Migrated						
[1a] Propensity score matching	0.11***	0.04	3.00	0.08***	0.02	3.58
[1b] Bias corrected matching	0.07***	0.02	4.37	0.07***	0.02	4.05
Father Migrated Only						
[2a] Propensity score matching	0.01	0.05	0.25	0.03	0.03	1.19
[2b] Bias corrected matching	0.06***	0.02	3.10	0.07***	0.02	3.02
Father Migrated (unconditional)						
[3a] Propensity score matching	0.09***	0.04	2.35	0.06***	0.02	2.77
[3b] Bias corrected matching	0.08***	0.02	4.51	0.08***	0.02	4.05
Mother Migrated Only						
[4a] Propensity score matching	0.01	0.11	0.11	-0.02	0.07	-0.26
[4b] Bias corrected matching	0.02	0.05	0.46	0.05	0.05	0.91
Mother Migrated (unconditional)						
[5a] Propensity score matching	0.08*	0.05	1.82	0.02	0.03	0.55
[5b] Bias corrected matching	0.09***	0.02	3.73	0.08***	0.03	3.25
Both Parents Migrated						
[6a] Propensity score matching	0.10**	0.05	2.01	0.08***	0.03	2.35
[6b] Bias corrected matching	0.10***	0.02	4.02	0.09***	0.03	3.26

Data source: Author's survey.

Propensity scores are estimated using the same set of covariates as in Table 2.

* significant at 10%; ** significant at 5%; *** significant at 1%. t-stats are reported for propensity score matching and z-values are reported for bias-corrected matching in parentheses.

We use propensity scores as a tool to enforce a common support. We use the nearest neighbor matching with replacement. Following Smith and Todd (2005), we match students based on the log odds ratio and standard errors are bootstrapped using 1,000 replications.

To minimize geographic mismatch, we enforce exact matching by county. Each treatment observation is matched to 3 control observations with replacement. The weighting matrix uses the Mahalanobis metric, which is the inverse of the sample covariance matrix of the matching variables.

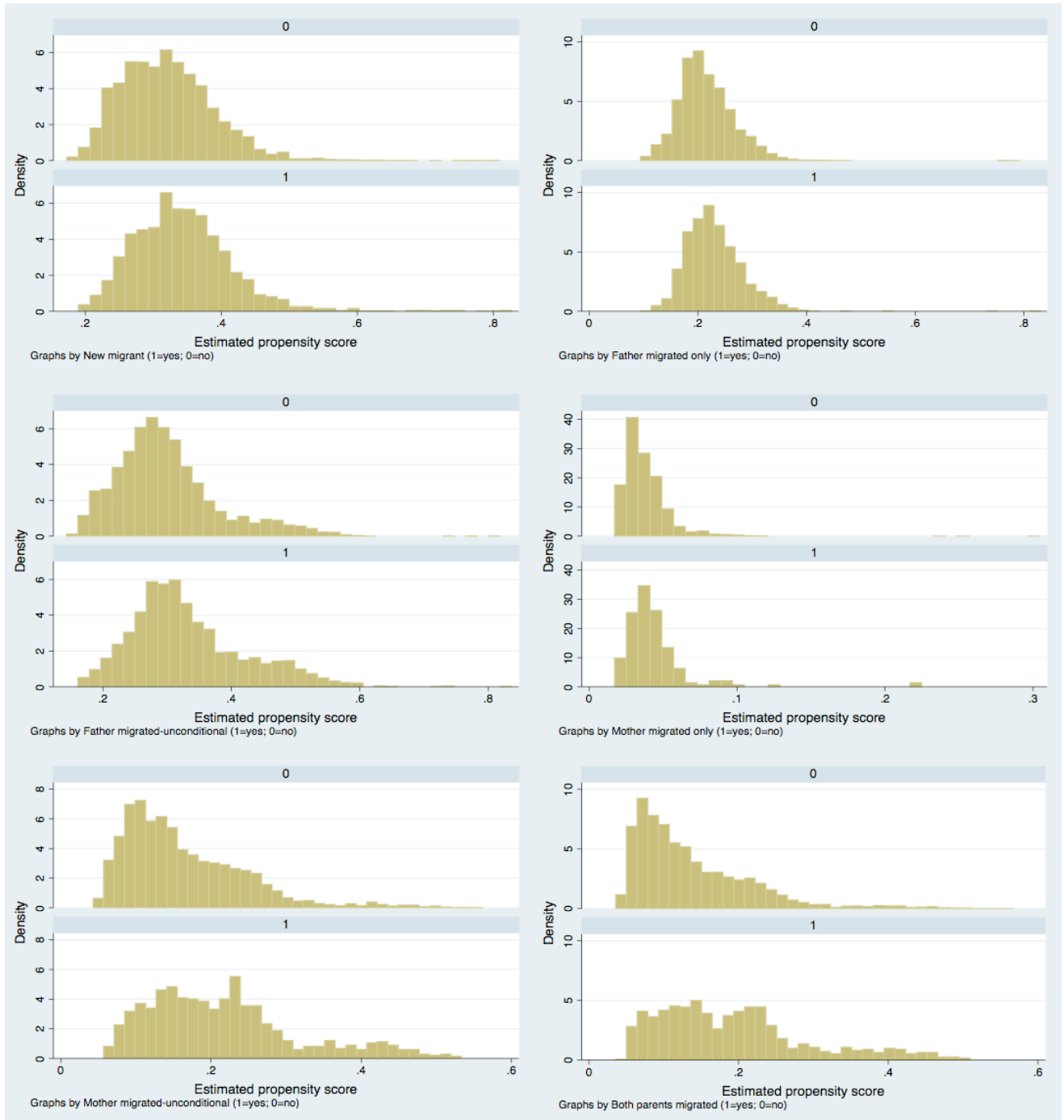
Table 6. Heterogeneous effect.

Dependent variable: $\Delta \text{Score}_i = \text{Score}_{i, 2014} - \text{Score}_{i, 2013}$	(1)	(2)	(3)	(4)	(5)	(6)
VARIABLES	Any Parent Migrated	Father Migrated Only	Father Migrated	Mother Migrated Only	Mother Migrated	Both Parents Migrated
<i>Characteristics of the students</i>						
[1] MIG * Standardized pre English test score (standard deviation)	-0.02 (0.01)	0.02 (0.02)	-0.01 (0.01)	-0.08* (0.04)	-0.07*** (0.02)	-0.06*** (0.02)
[2] MIG * Female	-0.02 (0.03)	-0.04 (0.03)	-0.03 (0.03)	0.01 (0.09)	0.01 (0.05)	0.01 (0.06)
[3] MIG * Ethnic Minority	0.03 (0.03)	0.02 (0.04)	0.02 (0.03)	0.08 (0.09)	0.03 (0.05)	0.01 (0.05)
[3a] MIG * Tibetan	0.11** (0.05)	0.08 (0.06)	0.10* (0.05)	0.26** (0.10)	0.13** (0.05)	0.10 (0.06)
[3b] MIG * Tu minority	-0.08 (0.07)	-0.02 (0.08)	-0.04 (0.08)	-0.62** (0.24)	-0.22** (0.09)	-0.09 (0.10)
[4] MIG * Only child	-0.04 (0.04)	0.00 (0.05)	-0.04 (0.05)	-0.05 (0.09)	-0.08 (0.06)	-0.09 (0.07)
[5] MIG * Asset	-0.01 (0.04)	-0.06 (0.07)	-0.02 (0.05)	0.02 (0.02)	0.02 (0.02)	0.02 (0.03)
<i>Characteristics of the parents and the households</i>						
[6] MIG * Mother has at least junior high school degree	-0.02 (0.04)	0.04 (0.04)	-0.01 (0.04)	-0.10 (0.09)	-0.10** (0.05)	-0.08 (0.05)
[7] MIG * Father has at least junior high school degree	0.00 (0.03)	0.03 (0.04)	0.00 (0.03)	0.02 (0.09)	-0.03 (0.04)	-0.03 (0.05)
[8] Standardized per English test score (standard deviation)	YES	YES	YES	YES	YES	YES
[9] Control variables	YES	YES	YES	YES	YES	YES
[10] School Fixed effects	YES	YES	YES	YES	YES	YES
[11] Number of Observations	6,724	5,783	6,547	4,696	5,460	5,283

* significant at 10%; ** significant at 5%; *** significant at 1%. Robust standard errors in parentheses clustered at school level. County dummies are controlled.

Appendix Table 1. Definition of different types of migrant households.

Migration status (key independent variable name)	Definition
[1] Any Parent Migrated	Households in which both parents lived at home by September 2013 and at least on parent – either the father, mother or both parents – out-migrated by June 2014
[2] Father Migrated Only	Households in which only the father out-migrated by June 2014 but was at home by September 2013
[3] Father Migrated (Unconditional)	Households in which the father was at home by September 2013 but out-migrated by June 2014 (including households in which the mother was either at home or not at home in 2014)
[4] Mother Migrated Only	Households in which only the mother out-migrated by June 2014 but was at home by September 2013
[5] Mother Migrated (Unconditional)	Households in which the mother was at home by September 2013 but out-migrated by June 2014 (including households in which the father was either at home or not at home in 2014)
[6] Both Parents Migrated	Households in which both parents were at home by September 2013, but out-migrated by June 2014



Appendix Fig. 1 Overlap in the support of the covariates between the six different types of *New Migrant* households and *Never Migrant* households.