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Do Firms Interact Strategically?

A Structural Model of the Multi-Stage Investment Timing Game
in Offshore Petroleum ProductionC.-Y. Cynthia Lin¹

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Abstract

When individual petroleum-producing firms make their exploration and development investment timing decisions, positive information externalities and negative extraction externalities may lead them to interact strategically with their neighbors. This paper examines whether these inefficient strategic interactions take place in U.S. federal lands in the Gulf of Mexico. In particular, it analyzes whether a firm's production decisions and profits depend on the decisions of firms owning neighboring tracts of land. The empirical approach is to estimate a structural econometric model of the firms' multi-stage investment timing game. Although the model only permits the identification of the net effect of the two countervailing externalities, and not each individually, theory suggests that the importance of the extraction externality relative to the information externality should be greater on small tracts than on large tracts, and the data are consistent with this theory. Also as expected, the externalities intensify as the tract size decreases.

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1 Introduction

Petroleum production is a multi-stage process involving sequential investment decisions. The first stage is exploration: when a firm acquires a previously unexplored tract of land, it must first decide whether and when to invest in the drilling rigs needed to begin exploratory drilling. The second stage is development: after exploration has taken place, a firm must subsequently decide whether and when to invest in the production platforms needed to develop and extract the reserve. Because the profits from petroleum production depend on market conditions such as the oil price that vary stochastically over time, an individual firm producing in isolation that hopes to make dynamically optimal decisions would need to account for the option value to waiting before making either irreversible investment (Dixit & Pindyck, 1994).

The dynamic decision-making problem faced by a petroleum-producing firm is even more complicated when its profits are affected not only by exogenous market conditions, but also by the actions of other firms producing nearby. When firms own leases to neighboring tracts of land that may be located over a common pool of reserve, there are two types of externalities that add a strategic (or non-cooperative)² dimension to firms' investment timing decisions and may render these decisions socially inefficient.³ The first type of externality is an *information externality*: if tracts are located over a common pool or share common geological features so that their ex post values are correlated, then firms learn information about their own tracts when other firms drill exploratory wells or install production platforms on neighboring tracts (Hendricks & Porter, 1996). The information externality is a positive one, since a firm benefits from its neighbors' information.⁴ A second type of externality is an *extraction externality*: when firms have competing rights to a common-pool re-

²In this paper, I use the terms "strategic" and "non-cooperative" interchangeably.

³In my broad definition of an externality, I say that an externality is present whenever a non-coordinated decision by individual firms is not socially optimal.

⁴Although the information externality has a positive effect on a firm's profits, it is socially inefficient. For example, it may cause firms to play a non-cooperative timing game that leads them to inefficiently delay production, since the possibility of acquiring information from other firms may further enhance the option value to waiting. If firms are subject to a lease term by the end of which they must begin exploratory drilling, or else relinquish their lease, then the information externality would result in too little exploration at the beginning of the lease term and duplicative drilling in the final period of the lease (Hendricks & Porter, 1996; Porter, 1995). In contrast, the optimal coordinated plan would entail a sequential search in which one tract would be drilled in the first period and, if productive, a neighboring tract is drilled in the next (Porter, 1995).

source, strategic considerations may lead them to extract at an inefficiently high rate (Libecap & Smith, 1999a; Libecap & Wiggins, 1985). The extraction externality is a negative one, since it induces a firm to produce inefficiently. Owing to both information and extraction externalities, the dynamic decision-making problem faced by a petroleum-producing firm is not merely a single-agent problem, but rather can be viewed as a multi-agent, non-cooperative game in which firms behave strategically and base their exploration and development policies on those of their neighbors.

Both externalities lead to strategic interactions that are socially inefficient. The information externality leads to an inefficient delay in exploration. The extraction externality leads to excessively high extraction rates and less total oil extracted. Both types of strategic behavior lead to lower profits for the petroleum-producing firms and lower royalty revenue for the federal government. It is therefore important to analyze whether these strategic interactions place, and therefore whether policies that can mitigate the strategic interactions should be implemented.

Since 1954, the U.S. government has leased tracts from its federal lands in the Gulf of Mexico to firms interested in offshore petroleum production by means of a succession of lease sales. A lease sale is initiated when the government announces that an area is available for exploration, and nominations are invited from firms as to which tracts should be offered for sale. In a typical lease sale, over a hundred tracts are sold simultaneously in separate first-price, sealed-bid auctions. Many more tracts are nominated than are sold, and the nomination process probably conveys little or no information (Porter, 1995). A tract is typically a block of 5000 acres or 5760 acres (Marshall Rose, Minerals Management Service, personal communication, 9 November 2005). The size of a tract is often less than the acreage required to ensure exclusive ownership of any deposits that may be present (Hendricks & Kovenock, 1989), and tracts within the same area may be located over a common pool (Hendricks & Porter, 1993). To date, the largest petroleum field spanned 23 tracts. Depending on water depth, 57-67 percent of the fields spanned more than one tract and 70-79 percent spanned three or fewer tracts (Marshall Rose, Minerals Management Service, personal communication, 31 March 2005). Because neighboring tracts of land may share a common pool of petroleum reserve, information and extraction externalities that lead firms to interact strategically may be present. As a consequence, petroleum production on the federal leases may be inefficient.

In this paper, I analyze whether a firm's investment timing decisions and profits

depend on the decisions of firms owning neighboring tracts of land. Do the positive information externalities and negative extraction externalities have any net strategic effect that may cause petroleum production to be inefficient?

To answer this question, I develop and estimate a structural econometric model of the firms' multi-stage investment timing game, and apply the model to data on petroleum production on U.S. federal lands in the Gulf of Mexico.

The research presented in this paper is important for several reasons. First, an empirical analysis of investment timing decisions enables one to examine whether the strategic interactions that are predicted in theory actually occur in practice. Second, the estimation of strategic interactions, especially those that arise in dynamic decision-making, is of methodological interest. Third, my results have implications for leasing policy: if the strategic effects and externalities turn out to be large, then the program by which the U.S. government leases tracts to firms may be inefficient, and possible modifications should be considered.

The results from the structural econometric model do not indicate that externalities from exploration have any net strategic effect. A firm's profits from development does not depend significantly on the exploration decisions of its neighbor. In contrast, externalities from development do have a net strategic effect. A firm's real profits from developing increase when its neighbor develops, perhaps because this is a signal to the firm that the neighbor's exploratory efforts were successful, and therefore that there may be deposits present.

There are several possible explanations why the results reject strategic, non-cooperative behavior during exploration. One is that the tract sizes are large enough that cross-tract externalities are insignificant or even nonexistent. A second is that cross-tract externalities exist, but firms owning neighboring tracts cooperate to jointly internalize the inefficient externalities they impose on each other, for example by forming joint ventures in exploration⁵ or by consolidating their production rights through purchase or unitization.⁶ A third is that cross-tract externalities are significant, but

⁵Joint ventures in exploration occur less frequently than one might expect, however, because negotiations are contentious, because firms fear allegations of pre-sale anti-trust violations (Marshall Rose, Minerals Management Service, personal communication, 3 May 2005), and because prospective partners have an incentive to free ride on a firm's information gathering expenditures (Hendricks and Porter, 1992). In their theoretical model of the persuasion game, Hendricks and Kovenock (1989) find that, even with well-defined property rights, bargaining does not eliminate all the inefficiencies of decentralized drilling decisions. As a consequence, the information externality may not be fully internalized.

⁶Under a unitization agreement, a single firm is designated as the unit operator to develop the

the positive information externality exactly cancels the negative extraction externality, resulting in zero net strategic effects.

To distinguish among these three explanations for the lack of strategic interactions during exploration, I estimate the strategic interactions by tract size. If externalities are insignificant when tracts are large enough, then one would expect to see strategic, non-cooperative behavior only on small tracts. This is because the smaller the tract size, the more likely the tracts are located over a common pool, and therefore the more acute the information and extraction externalities faced by the firms. Evidence for significant strategic interactions on small tracts but not on large tracts would thus be consistent with the first explanation.

If externalities exist even for the largest tracts in the sample, but are eliminated through coordination, then, assuming firms would coordinate regardless of tract size, one would not expect to see any strategic, non-cooperative behavior even on small tracts. Insignificant strategic interactions regardless of tract size would thus be consistent with the second explanation.⁷

If strategic interactions do not occur on net because the positive information externality exactly cancels the negative extraction externality, then one may not expect the exact cancellation to still take place when the tract size is small. This is because the geographical span of the information externality is larger than that of the extraction externality: while the former only requires that tracts may share common geological features, the latter requires that tracts may be located over a common pool. As a consequence, it is possible for the information externality to be present on all the tracts in the sample, but for the extraction externality to be present on only the smaller tracts. Theory therefore suggests that the importance of the extraction externality relative to the information externality should be greater on small tracts than on large tracts. Thus, if the externalities cancel when all the

entire reservoir, while the other firms share in the profits according to negotiated formulas (Libecap & Smith, 1999b). There are many obstacles to consolidation, however, including contentious negotiations, the need to determine relative or absolute tract values, information costs, and oil migration problems (Libecap & Wiggins, 1984). In addition, another free rider problem that impedes coordination is that firms may fear that if they reveal to other firms their information or expertise, for example about how to interpret seismic data, then they may lose their advantage in future auctions (Hendricks & Porter, 1996). Thus, despite various means of coordination, firms may still behave strategically and non-cooperatively, and information and extraction externalities may not be fully internalized.

⁷It is also possible that the coefficients that arise when firms coordinate are significant, but are different from those that would arise under the non-cooperative outcome.

tract sizes are considered, one might expect that the negative extraction externality would dominate the positive information externality when the sample is limited to small tracts only. Strategic interactions that are more significantly negative on small tracts than on large tracts would therefore be consistent with the third explanation.

According to the structural estimation, the importance of strategic interactions depends on tract size. As expected, strategic interactions are more likely to take place on smaller tracts, where the externalities are more acute. When the tract size is large enough, the net strategic effects of the externalities from both exploration and development disappear. Also as predicted by theory, the relative importance of the extraction externality from exploration with respect to the information externality is greater on small tracts than on large tracts; on large tracts, the two externalities cancel each other out.

The results suggest that, by selling predominantly large tracts, the federal government has minimized inefficiencies in petroleum production that may have resulted from non-cooperative strategic interactions.

The balance of the paper proceeds as follows. I begin with a brief review of the relevant literature in Section 2. Section 3 presents the model and Section 4 explains the econometric estimation. The data are described in Section 5. Section 6 analyzes the results. Section 7 concludes.

2 Contributions to Existing Literature

The exploration timing game in offshore petroleum production in the Gulf of Mexico has been examined in a seminal series of papers by Kenneth Hendricks, Robert Porter and their co-authors (see e.g. Hendricks & Kovenock, 1989; Hendricks & Porter, 1993; Hendricks & Porter, 1996). These papers focus on the information externality associated with exploratory drilling. They analyze this externality and the learning and strategic delay that it causes by developing theoretical models of the exploration timing game. In addition, Hendricks and Porter (1993, 1996) calculate the empirical drilling hazard functions for cohorts in specific areas, and study the determinants of the exploration timing decision and of drilling outcomes. According to their results, equilibrium predictions of plausible non-cooperative models are reasonably accurate and more descriptive than those of cooperative models of drilling

timing.

The structural econometric analysis presented in this paper improves upon the existing literature on the exploration timing game in offshore petroleum production in several ways. First, unlike the theoretical models and reduced-form empirical analyses conducted by Hendricks, Porter and their co-authors, a structural approach yields estimates of the structural parameters of the discrete choice dynamic game. With these structural parameters, one can identify the effects of a neighbor's exploration and development decisions on the profits a firm would get from developing its tract.

A second way in which this paper contributes to the existing literature on the information externality in offshore petroleum production is that it combines the externality problem with real options theory. Oil production is a multi-stage process involving sequential investment decisions. Since the decision to explore a reserve entails an irreversible investment, the value of an unexplored reserve is the value of the option to invest in exploration. Similarly, the value of an explored but undeveloped reserve is the value of the option to invest in development. There is thus an option value to waiting before making either investment because the value of a developed reserve can change, either because exogenous conditions such as the oil price might change, or because there is a chance that neighboring firms might explore or develop first. Moreover, because these two types of investment are made sequentially, they act as compound options: completing one stage gives the firm an option to complete the next (Dixit & Pindyck, 1994).

While literature on the financial theory of option valuation is abundant, structural models applying the theory to the oil production process that account for strategic considerations have yet to be developed. Hurn and Wright (1994) test reduced-form implications of the theory via a hazard model, but neither estimate a structural model nor account for possible strategic interactions. Paddock, Siegel and Smith (1998) compare the option valuation estimate of the market value of selected offshore petroleum tracts with estimates from other valuation methods and with the winning bids, but do not account for either information externalities or extraction externalities. Similarly, Pesaran (1990) estimates an intertemporal econometric model for the joint determination of extraction and exploration decisions of a "representative" profit-maximizing oil producer, but does not examine the case of multiple producers that may interact strategically.

The third innovation this paper makes to the existing literature on the informa-

tion externality in offshore petroleum production is that while the existing literature focuses exclusively on externalities that arise during exploratory drilling, my model allows for extraction externalities as well as information externalities that arise during both exploration and development. If firms do indeed learn about the value of their own tracts from the actions of their neighbors, then one would expect firms to update their own beliefs not only if their neighbors begin exploratory drilling, but also if, after having already begun exploring, the neighbors then decide to install a production platform. That a neighbor has decided to begin extracting after it explored should be at least as informative as the initiation of exploration in the first place. Furthermore, extraction externalities are another form of spillover that is not accounted for by previous studies of the investment timing game, and, unlike the information externality, is one that may have a negative effect on a firm's profits.

In addition to the literature on the information externality, a second branch of related literature is that on econometric models of discrete dynamic games (see e.g. Aguirregabiria & Mira, 2007; Bajari, Benkard & Levin, 2007 and references therein). In particular, my work applies a method developed by Pakes, Ostrovsky and Berry (2007) for estimating parameters of discrete dynamic games such as those involving firm entry and exit. This paper builds upon the work of Pakes et al. (2007) in several ways. First, unlike their paper, which uses simulated data, this paper estimates a discrete dynamic game using actual data. Second, while the entry and exit decisions they examine are two independent investments, the exploration and development decisions I examine are sequential investments: the decision to invest in development can only be made after exploration has already taken place. Thus, unlike the one-stage entry and exit games, the investment timing game is a two-stage game. The sequential nature of the investments is an added complexity that I address in my econometric model. Third, whereas the estimators Pakes et al. propose are for infinite-horizon dynamic games, the exploration stage of petroleum production is a finite-horizon dynamic optimization problem: firms must begin exploration before the end of the five-year lease term, or else relinquish their lease. As a consequence, an appropriate modification to the estimation algorithm is required. A fourth innovation I make is that, unlike Pakes et al., I do not assume that the profit function is a known function of the underlying state variables, but instead estimate its parameters from data. While Pakes et al. are able to appeal to economic theory to posit an exact form for profits as a function of state variables such as the number of firms in the industry,

I cannot. No economic theory predicts the relationship between profits and such state variables as whether or not a firm's neighbors explore or develop. Indeed, since the relationship between profits and the actions of one's neighbor is among the very parameters I hope to identify, I choose to estimate this relationship from the data rather than impose it *a priori*. The task of estimating these additional parameters requires the use of additional moment conditions.

There are several advantages to using a structural model. First, a structural model enables the estimation of all the structural parameters of the underlying dynamic game. These parameters include not only those governing the relationship between various state variables and the profits of firms, but also parameters governing the distribution of tract-specific private information. Second, the structural model addresses the endogeneity problems without the need for instruments. Measuring neighbors' effects is difficult owing to two sources of endogeneity. One source is the simultaneity of the strategic interaction: if tract i is affected by its neighbor j , then tract j is affected by its neighbor i . The other arises from spatially correlated unobservable variables (Manski, 1993; Manski, 1995; Robalino & Pfaff, 2005). Because the structural model is based on the equilibrium of the underlying dynamic game, however, it addresses the simultaneity problem directly by explicitly modeling the firms' strategies. Moreover, the problem of spatially correlated unobservables can be addressed by interpreting the profits in the model as expected profits conditional on observables. A third advantage to a structural model is that it enables one to estimate how a firm's profits are affected by the decisions of its neighbors; the sign of the effect indicates the net sign of the information and extraction externalities. Fourth, the structural model enables one to explicitly model each of the stages of the multi-stage dynamic decision-making problem faced by petroleum-producing firms. It is to this model that I now turn.

3 A Model of the Investment Timing Game

In my model of the investment timing game, each "market" k consists of an isolated neighborhood of adjacent tracts i that were each leased to a petroleum-producing firm on the same date. Time t denotes the number of years after the lease sale date. Firms must begin exploration before time T , the length of the lease term, or else relinquish their lease. Let the "lease term time" τ_{kt} of market k at time t be

given by:

$$\tau_{kt} = \begin{cases} t & \text{if } t = 0, 1, \dots, T-1 \\ T & \text{if } t \geq T \end{cases}$$

For each market k , the state of the market t years after the leases began is given by a vector Ω_{kt} of discrete and finite-valued state variables that are observed by all the firms in market k and as well as by the econometrician. Let θ denote the vector of parameters to be estimated.

At the beginning of each period t , the owner of each tract i must make one of two investment decisions. If tract i has not been explored before time t , its owner must decide whether to invest in exploration at time t . If tract i has been explored but has not been developed before time t , its owner must decide whether to invest in development at time t . For each period t , all firms make their time- t investment decisions simultaneously.

Each firm's time- t investment timing decision depends in part on the state of the market $\Omega_{kt} \equiv (N_{kt}, X_{kt}, \tau_{kt})$, which can be decomposed into endogenous state variables N_{kt} , exogenous profit-shifting state variables X_{kt} , and the lease term time τ_{kt} . Investment decisions depend on N_{kt} and X_{kt} because these state variables are assumed to affect profits. Because of the finite-horizon nature of the firm's exploration investment problem, the finite-valued and exogenous lease term time τ_{kt} affects investment decisions as well, as will be explained below. In the present model, there are two endogenous state variables N_{kt} : the total number of tracts in market k that have been explored before time t , and the total number of tracts in market k that have been developed before time t . These endogenous state variables capture the strategic component of the firms' investment timing decisions. The exogenous state variables X_{kt} include the drilling cost and the oil price and are assumed to evolve as a finite state first-order Markov process: $X_{k,t+1} \stackrel{iid}{\sim} F_X(\cdot | X_{kt})$. In other words, the next period's value $X_{k,t+1}$ of the exogenous state variables are assumed to be independently and identically distributed (i.i.d.) with a probability distribution that depends only on the time- t realization X_{kt} of the exogenous state variables, and not additionally on what happened before time t (Dixit & Pindyck, 1994).⁸

⁸The lease term time τ_{kt} evolves as a finite state first-order Markov process as well. I include this exogenous finite-valued variable τ_{kt} as a separate argument distinct from X_{kt} both because it does not affect profits and also to elucidate my later exposition of the finite-horizon nature of the

In addition to the publicly observable state variables Ω_{kt} , each firm's time- t investment timing decision also depends on two types of shocks that are private information to the firm and unobserved by either other firms or by the econometrician. The first source of private information is a pre-exploration shock μ_{it} to an unexplored tract i at time t . This pre-exploration shock, which is only observed by the firm owning tract i , represents any and all private information that affects the exploration investment decision made on tract i at time t . Such private information may include, for example, idiosyncratic shocks to exploration costs and the outcome of the post-sale, pre-exploration seismic study conducted on tract i at time $t - 1$.⁹ Following Pakes, Ostrovsky and Berry (2007), assume that the pre-exploration shock μ_{it} is an independently and identically distributed random variable with an exponential distribution and mean σ_μ . That is, $\mu_{it} \stackrel{iid}{\sim} \text{exponential}(\sigma_\mu)$.

The second source of private information is a pre-development shock ε_{it} to an explored but undeveloped tract i at time t . This pre-development shock, which is only observed by the firm owning tract i , represents any and all private information that affects the development investment decision made on tract i at time t . Such private information may include, for example, the outcome of the exploratory drilling conducted on tract i at time $t - 1$. Following Pakes, Ostrovsky and Berry (2007), assume that the pre-development shock ε_{it} is an independently and identically distributed random variable with an exponential distribution and mean σ_ε . That is, $\varepsilon_{it} \stackrel{iid}{\sim} \text{exponential}(\sigma_\varepsilon)$. In addition, assume that the pre-exploration shocks μ_{it} and the pre-development shocks ε_{it} are independent of each other.¹⁰ All distributions are

exploration stage.

⁹Firms conduct and analyze seismic studies in order to help them decide whether or not to begin exploratory drilling (John Shaw, personal communication, 18 April 2003; Bob Dye, Apache, personal communication, 21 January 2004; Jon Jeppesen, Apache, personal communication, 21 January 2004; Mark Bauer, Apache, personal communication, 21 January 2004; Billy Ebarb, Apache, personal communication, 22 January 2004).

¹⁰The assumptions that both types of shocks are i.i.d. and independent of each other, while restrictive, are needed in order for the estimation technique used in this paper to work. If either type of shock were serially correlated (or if, at the extreme, there were tract fixed effects), then firms would base their decisions not only on the current values of the state variables and of their shocks, but also on past values of the state variables and shocks as well. The state space would then be too large. If the distribution of the pre-development shock ε_{it} depended on the realization of the pre-exploration shock μ_{it} (e.g., the μ_{it} at the time of exploration), then μ_{it} would be a state variable in the development stage of production. As a consequence, the econometrician would need to observe μ_{it} , which she does not. The i.i.d. assumption is reasonable if the shocks are interpreted to encompass all idiosyncratic factors affecting investment decisions, including managerial shocks and technological shocks. Moreover, since one of my state variables is the average winning bid,

common knowledge.

In the absence of strategic considerations, the firm owning tract i would base its investment timing decisions on only the exogenous state variables X_{kt} , the lease term time τ_{kt} , and the private shocks μ_{it} and ε_{it} . To derive its dynamically optimal investment policy, it would solve a single-agent dynamic programming problem.

If information and extraction externalities were present, however, then strategic considerations would become important. As a consequence, the exploration and development investment decisions of the firm owning tract i in market k would depend on the exploration and development investment decisions of the firms owning the other tracts in market k . In other words, the firm owning tract i would base its investment timing decisions not only on the exogenous state variables X_{kt} , the lease term time τ_{kt} , and the private shocks μ_{it} and ε_{it} , but also on the endogenous state variables N_{kt} as well, namely the total number of tracts in its market k that have been explored before time t and the total number of tracts in market k that have been developed before time t . Each firm would then no longer solve merely a single-agent dynamic programming problem, but rather a multi-agent dynamic game.

The equilibrium concept used in the model is that of a Markov perfect equilibrium. Each firm is assumed to play a Markov "state-space" strategy: the past influences current play only through its effect on the state variables. A firm's dynamically optimal investment policy is then the Markov strategy that it plays in the Markov perfect equilibrium, which is a profile of Markov strategies that yields a Nash equilibrium in every proper subgame (Fudenberg & Tirole, 1998).

While each firm's time- t investment decision depends on both the publicly available endogenous and exogenous state variables Ω_{kt} as well as the firm's own private information μ_{it} or ε_{it} , its perception of its neighbor's time- t investment decisions depend only on the publicly observable state variables Ω_{kt} . This is because, owing to the above assumptions on the observable state variables and on the unobservable shocks, firms can take expectations over their neighbors' private information.¹¹ In equilibrium, firms' perceptions of their neighbors' investment probabilities should be

which is a measure of tract value, it is reasonable to assume that, conditional on tract value, shocks are i.i.d.

¹¹ While each firm plays a pure strategy, from the point of view of their neighbors, they appear to play mixed strategies. Thus, as with Harsanyi's (1973) purification theorem, a mixed distribution over actions is the result of unobserved payoff perturbations that sometimes lead firms to have a strict preference for one action, and sometimes a strict preference for another.

consistent with those that are actually realized (Starr & Ho, 1969).

The model has at least one Markov perfect equilibrium, and each equilibrium generates a finite state Markov chain in Ω_{kt} tuples (Pakes, Ostrovsky & Berry, 2007).¹² Although model assumptions do not guarantee a unique equilibrium, they do insure that there is only one set of equilibrium policies that is consistent with the data generating process. It is thus possible to use the data itself to pick out the equilibrium that is played. For large enough samples, the data will pick out the correct equilibrium and the estimators for the parameters in the model will be consistent (Pakes, Ostrovsky & Berry, 2007).¹³

The firm's dynamic decision-making problem is as follows. The first-stage problem is to determine the optimal policy for investment in exploration. Because firms must begin exploration before the end of their lease term, or else relinquish their lease, this is a finite-horizon problem. As a consequence, firms' decisions will depend not only on the profit-shifting state variables N_{kt} and X_{kt} , but also on time t . However, since firms can only make exploration decisions at the beginning of periods $t = 0, \dots, T - 1$, the time dependence only applies until time $t = T - 1$, after which exploration can no longer begin and the endogenous variable that counts the total number of tracts in the market that have been explored stays constant. It is for this reason that the exogenous and finite-valued state variable "lease term time" τ_{kt} captures the entire time dependence of the problem.

The second stage of the firm's dynamic decision-making problem is to determine the optimal timing for investing in the development of a tract that has already been explored. This second-stage problem has both a finite-horizon component and an infinite-horizon component. A firm's development strategy depends in part on its perceptions of the future exploration policies of the firms in the market. Since exploration policies depend on time until time $t = T - 1$, this means that perceptions, and therefore development strategies, will depend on time for $t < T$. As a consequence, the dynamic programming problem for time $t < T$ is a finite-horizon problem. However, because the lease term only applies to the exploration stage of production, and because the endogenous variable that counts the total number of tracts in the market that have been explored – a variable that depends on the time-dependent exploration

¹²A Markov chain is a Markov process on a finite state space (Stokey, Lucas & Prescott, 1989).

¹³This assumes that the same equilibrium is played in each market. If a mixed strategy equilibrium is played, then it is assumed that the same mixed strategy equilibrium is played in each market.

policies of the firms in the market – stays constant after the lease term expires, the dynamic programming problem for the development stage from time T onwards is an infinite-horizon problem that does not depend on time. Thus, once again, the lease term time τ_{kt} sufficiently captures the entire time dependence of the problem.

The firm's sequential investment problem is a two-stage optimization problem, and can be solved backwards using dynamic programming (Dixit & Pindyck, 1994). In the second, or development, stage of oil production, a firm with an explored but undeveloped tract i must decide if and when to invest in a production platform. Assume that the profit $\pi^d(\Omega_{kt}, \varepsilon_{it}; \theta)$ that a firm will get after developing tract i at time t can be separated into a deterministic component and a stochastic component as follows:

$$\pi^d(\Omega_{kt}, \varepsilon_{it}; \theta) = \pi_0^d(\Omega_{kt}; \theta) + \varepsilon_{it} , \quad (1)$$

where the deterministic component of profit is linear in the publicly observable state variables:

$$\pi_0^d(\Omega_{kt}; \theta) \equiv N'_{kt} \gamma_N + X'_{kt} \gamma_X , \quad (2)$$

and where the stochastic component is the privately observed pre-development shock ε_{it} .¹⁴ The development profit is therefore independent of time (and lease term time) except through the state variables (N_{kt}, X_{kt}) and the shock ε_{it} .

Let $\gamma \equiv (\gamma_N, \gamma_X)$ denote the vector of all the coefficients in the development profit function. The coefficients γ_N in the profit function on the endogenous state variables N_{kt} – the total number of tracts in the market that have been explored and the total number of tracts in the market that have been developed – indicate whether and how one firm's profits depend on the production decisions of its neighbors. If a neighbor explores, then the state variable counting the total number of tracts in the market that have been explored increases by one and the value of the development profits increase by the value of its coefficient. Similarly, if a neighbor develops, then the state variable counting the total number of tracts in the market that have been developed increases by one and the value of the development profits increase by the

¹⁴If there were additional market state variables that affected profits but were unobserved by the econometrician, then $\pi^d(\Omega_{kt}, \varepsilon_{it}; \theta)$ can be interpreted as the expected profits conditional on the available information Ω_{kt} (Pakes, Ostrovsky & Berry, 2007). Under this interpretation, spatially correlated unobservables do not pose a concern.

value of its coefficient. The coefficients γ_N on the endogenous variables thus measure the net effects of the information and extraction externalities, and therefore indicate whether firms interact strategically on net. Positive values of the coefficients γ_N would indicate that the information and extraction externalities were positive on net, and therefore that the information externality was dominant. Negative values would indicate that the externalities were negative on net, and therefore that the extraction externality was dominant.

The value V^e of an explored but undeveloped tract i in market k at time t is given by:

$$V^e(\Omega_{kt}, \varepsilon_{it}; \theta) = \max\{\pi^d(\Omega_{kt}, \varepsilon_{it}; \theta), \beta V^{ce}(\Omega_{kt}; \theta)\}, \quad (3)$$

where $\beta \in (0, 1)$ is the discount factor and $V^{ce}(\Omega_{kt}; \theta)$ is the continuation value to waiting instead of developing at time t . For the structural estimation, I set the discount factor β to 0.9. The continuation value to waiting is the expectation over the state variables and shocks of next period's value function, conditional on not developing this period:

$$V^{ce}(\Omega_{kt}; \theta) = E[V^e(\Omega_{k,t+1}, \varepsilon_{i,t+1}; \theta) | \Omega_{kt}, I_{it}^d = 0], \quad (4)$$

where I_{it}^d is an indicator for whether development began on tract i at time t .

Let $g^d(\Omega_{kt}; \theta)$ denote the probability of developing an explored but undeveloped tract i at time t conditional on the publicly available information Ω_{kt} on time t , but not on the private information ε_{it} . The development probability $g^d(\Omega_{kt}; \theta)$ function represents a firm's perceptions of the probability that a neighbor owning an explored but undeveloped tract will decide to develop its tract in period t , given that the state of their market at time t is Ω_{kt} . Moreover, a firm's expectation of its own probability of development in the next period is simply the expected value of the next period's development probability, conditional on this period's state variables: $E[g^d(\Omega_{k,t+1}; \theta) | \Omega_{kt}]$.

Using the exponential distribution for ε_{it} and equation (1) for development profits as shown in Appendix A, the continuation value $V^{ce}(\cdot)$ can be reduced to:

$$V^{ce}(\Omega_{kt}; \theta) = E[\beta V^{ce}(\Omega_{k,t+1}; \theta) + \sigma_\varepsilon g^d(\Omega_{k,t+1}; \theta) | \Omega_{kt}, I_{it}^d = 0], \quad (5)$$

and the development probability $g^d(\cdot)$ can be reduced to the following function of the

continuation value, the state variables and the parameters:

$$g^d(\Omega_{kt}; \theta) = \exp \left(-\frac{\beta V^{ce}(\Omega_{kt}; \theta) - \pi_0^d(\Omega_{kt}; \theta)}{\sigma_\varepsilon} \right). \quad (6)$$

In the first, or exploration, stage of oil production, a firm with an unexplored tract i must decide if and when to invest in exploratory drilling. Owing to the sequential nature of the investments, the publicly observable deterministic component of the payoff $\pi_0^e(\cdot)$ to exploring in the first stage is equal to the expected value of having an explored but undeveloped tract in the second stage, net the cost of exploration $c^e(\cdot)$:

$$\pi_0^e(\Omega_{kt}; \theta) \equiv E_\varepsilon [V^e(\Omega_{kt}, \varepsilon_{it}; \theta) | \Omega_{kt}] - c^e(\Omega_{kt}; \theta), \quad (7)$$

where the exploration cost is assumed to be linear in the exogenous cost-shifting state variables:¹⁵

$$c^e(\Omega_{kt}; \theta) = -X'_{kt} \alpha. \quad (8)$$

Assume that the actual payoff $\pi^e(\cdot)$ to exploring tract i at time t also includes a privately observed stochastic component as well:

$$\pi^e(\Omega_{kt}, \mu_{it}; \theta) = \pi_0^e(\Omega_{kt}; \theta) + \mu_{it}, \quad (9)$$

where the stochastic component is the pre-exploration shock μ_{it} .

The value V^n of an unexplored tract i in market k and time t is given by:

$$V^n(\Omega_{kt}, \mu_{it}; \theta) = \max \{ \pi^e(\Omega_{kt}, \mu_{it}; \theta), \beta V^{cn}(\Omega_{kt}; \theta) \}, \quad (10)$$

where $V^{cn}(\Omega_{kt}; \theta)$ is the continuation value to waiting instead of exploring at time t . The continuation value to waiting is the expectation over the state variables and shocks of next period's value function, conditional on not exploring this period:

$$V^{cn}(\Omega_{kt}; \theta) = E [V^n(\Omega_{k,t+1}, \mu_{i,t+1}; \theta) | \Omega_{kt}, I_{it}^e = 0], \quad (11)$$

where I_{it}^e is an indicator for whether exploration began on tract i at time t . The lease term imposes the following boundary condition:

¹⁵I define costs with a negative sign so that the coefficients can be interpreted as coefficients in the exploration profit function. Variables that increase cost will decrease profit, and vice versa.

$$V^n((N, X, T), \mu; \theta) = 0 \quad \forall N, X, \mu. \quad (12)$$

Let $g^e(\Omega_{kt}; \theta)$ denote the probability of exploring an unexplored tract i at time t conditional on the publicly available information Ω_{kt} on time t , but not on the private information μ_{it} . As with the development probability, the current value of the exploration probability represents a firm's perceptions of the probability that a neighbor owning an unexplored tract will decide to explore its tract in period t , given that the state of their market at time t is Ω_{kt} ; its expected value at time $t+1$ represents a firm's expectation of its own probability of exploration in the next period.

Using the exponential distribution for μ_{it} and equation (9) for exploration profits as shown in Appendix A, the continuation value $V^{cn}(\cdot)$ to waiting instead of exploring can be reduced to:

$$V^{cn}(\Omega_{kt}; \theta) = E[\beta V^{cn}(\Omega_{k,t+1}; \theta) + \sigma_\mu g^e(\Omega_{k,t+1}; \theta) | \Omega_{kt}, I_{it}^e = 0], \quad (13)$$

and the exploration policy function $g^e(\cdot)$ can be reduced to the following function of the continuation values, state variables and parameters:

$$g^e(\Omega_{kt}; \theta) = \exp \left(- \frac{\beta V^{cn}(\Omega_{kt}; \theta) - (\beta V^{ce}(\Omega_{kt}; \theta) + \sigma_\varepsilon g^d(\Omega_{kt}; \theta)) + c^e(\Omega_{kt}; \theta)}{\sigma_u} \right). \quad (14)$$

Owing to the sequential nature of the investment decisions, the continuation value $V^{ce}(\cdot)$ and the investment probability $g^d(\cdot)$ from the development stage appear in the expression for the investment probability $g^e(\cdot)$ in the exploration stage.

The ex ante expected value of an unexplored tract at time $t = 0$, where expectations are taken over the pre-exploration shock μ , is given by:

$$E_\mu[V^n(\Omega_{k0}, \mu_{i0}; \theta) | \Omega_{k0}] = \beta V^{cn}(\Omega_{k0}; \theta) + \sigma_\mu g^e(\Omega_{k0}; \theta). \quad (15)$$

4 The Structural Econometric Model

The econometric estimation technique I use employs a two-step semi-parametric estimation procedure. It is an extension of the estimator proposed by Pakes, Os-

trovsky and Berry (2007) to finite-horizon, multi-stage games. In the first step, the continuation values are estimated non-parametrically and these estimates are used to compute the predicted probabilities of exploration and development. In the second step, the parameters $\theta \equiv (\sigma_\mu, \sigma_\varepsilon, \gamma, \alpha)'$ are estimated by matching the predicted probabilities with the actual probabilities in the data. I will now describe each step in turn.

4.1 Step 1: Estimating continuation values and predicted probabilities

The first step entails computing the non-parametric¹⁶ estimators $V^{ce}(\widehat{\Omega}_{kt}; \theta)$ and $V^{cn}(\widehat{\Omega}_{kt}; \theta)$ for the continuation values $V^{ce}(\Omega_{kt}; \theta)$ and $V^{cn}(\Omega_{kt}; \theta)$, respectively, given θ . To do so, historical empirical frequencies are used to estimate the elements of the Markov transition matrix governing the evolution of the finite-valued state variables from one period to the next. Estimators for the continuation values are subsequently derived from equations (5) and (13) using dynamic programming. These estimators are then substituted into equations (6) and (14) to obtain predicted probabilities for development and exploration, respectively.¹⁷

Formally, the non-parametric estimator $V^{ce}(\widehat{\Omega}_{kt}; \theta)$ for $V^{ce}(\Omega_{kt}; \theta)$ is derived from equation (5) and is computed as follows. For each period t , let each component of the vector \vec{V}_t^{ce} be $V^{ce}(\Omega_{kt}; \theta)$ evaluated at a different tuple of state variables. Similarly, for each period t , let each component of the vector \vec{g}_t^d be $g^d(\Omega_{kt}; \theta)$ evaluated at a different tuple of state variables. Finally, for each lease term time period τ , let M_τ^e be a transition matrix from the point of view of an owner of an explored but undeveloped tract who decides not to develop at time t . The element in the i^{th} row and j^{th} column is the probability that the state tuple next period will be the j^{th} tuple, given that the state tuple this period is the i^{th} tuple, given that the tract has already been explored but not yet developed at time t , and conditional on not developing at time t .

¹⁶The continuation values are non-parametric functions of the state variables Ω_{kt} conditional on the parameters θ .

¹⁷Rather than use historical empirical frequencies to estimate the Markov transition matrix, it is possible to compute an estimator for the matrix using the estimators for the exploration and development probabilities. However, because the latter, more complicated approach imposes a computational burden and because Pakes, Ostrovsky and Berry (2007) find that it did not improve the performance of their estimator, I choose the former, simpler approach.

The estimator $V^{ce}(\widehat{\Omega}_{kt}; \theta)$ is obtained from rewriting equation (5) in vector form:

$$\overrightarrow{V_t^{ce}} = M_\tau^e \left(\beta \overrightarrow{V_{t+1}^{ce}} + \sigma_\varepsilon \overrightarrow{g_{t+1}^d} \right). \quad (16)$$

The estimator is obtained after further substituting in the empirical average \widehat{M}_τ^e for M_τ^e . For $t \geq T$, since $\overrightarrow{V_t^{ce}} = \overrightarrow{V_{t+1}^{ce}} \forall t \geq T$, we can solve for a fixed point $\widehat{V}^{ce}((N_{kt}, X_{kt}, T; \theta))$, which, from Blackwell's Theorem, is unique. To obtain the estimator of the value function for $t < T$, we then iterate backwards in time from $t = T$ using $\overrightarrow{V_T^{ce}} = \widehat{V}^{ce}((N_{kt}, X_{kt}, T; \theta))$ as a boundary condition. The predicted probability of development is then given by:¹⁸

$$g^d(\widehat{\Omega}_{kt}; \theta) = \exp \left(- \frac{\beta V^{ce}(\widehat{\Omega}_{kt}; \theta) - \pi_0^d(\Omega_{kt}; \theta)}{\sigma_\varepsilon} \right). \quad (17)$$

The non-parametric estimator $V^{cn}(\widehat{\Omega}_{kt}; \theta)$ for the continuation value to waiting instead of exploring is derived from equation (13) and is computed in a similar fashion as the estimator of $V^{ce}(\cdot)$. For each period t , let each component of the vector $\overrightarrow{V_t^{cn}}$ be $V^{cn}(\cdot)$ evaluated at a different tuple of state variables. For each period t , let each component of the vector $\overrightarrow{g_t^e}$ be $g^e(\cdot)$ evaluated at a different tuple of state variables. Finally, for each lease term time period τ , let M_τ^n be a transition matrix from the point of view of an owner of an unexplored tract who decides not to explore at time t . The element in the i^{th} row and j^{th} column is the probability that the state tuple next period will be the j^{th} tuple, given that the state tuple this period is the i^{th} tuple, given that the tract has yet to be explored, and conditional on not exploring at time t .

The estimator $V^{cn}(\widehat{\Omega}_{kt}; \theta)$ is obtained from rewriting equation (13) in vector form:

$$\overrightarrow{V_t^{cn}} = M_\tau^n \left(\beta \overrightarrow{V_{t+1}^{cn}} + \sigma_\mu \overrightarrow{g_{t+1}^e} \right). \quad (18)$$

Substituting in the empirical average \widehat{M}_τ^n for M_τ^n , we can solve backwards in time from the boundary condition $\overrightarrow{V_{T-1}^{cn}} = 0$ implied by equation (12) to obtain $\widehat{V}^{cn}(\Omega_{kt}; \theta)$ for all $t \leq T - 1$. The predicted probability of exploration is then given by:¹⁹

¹⁸In practice, I use: $g^d(\widehat{\Omega}_{kt}; \theta) = \min \left\{ \exp \left(- \frac{\beta V^{ce}(\widehat{\Omega}_{kt}; \theta) - \pi_0^d(\Omega_{kt}; \theta)}{\sigma_\varepsilon} \right), 1 \right\}$.

¹⁹In practice, I use: $g^e(\widehat{\Omega}_{kt}; \theta) = \min \left\{ \exp \left(- \frac{\beta V^{cn}(\widehat{\Omega}_{kt}; \theta) - (\beta V^{ce}(\widehat{\Omega}_{kt}; \theta) + g^d(\widehat{\Omega}_{kt}; \theta) \cdot \sigma_\varepsilon) + c^e(\Omega_{kt}; \theta)}{\sigma_\mu} \right), 1 \right\}$.

$$g^e(\widehat{\Omega}_{kt}; \theta) = \exp \left(- \frac{\beta V^{cn}(\widehat{\Omega}_{kt}; \theta) - \left(\beta V^{ce}(\widehat{\Omega}_{kt}; \theta) + g^d(\widehat{\Omega}_{kt}; \theta) \cdot \sigma_\varepsilon \right) + c^e(\Omega_{kt}; \theta)}{\sigma_u} \right). \quad (19)$$

Owing to the sequential nature of the investment decisions, the estimators $V^{ce}(\widehat{\Omega}_{kt}; \theta)$ and $g^d(\widehat{\Omega}_{kt}; \theta)$ of the continuation value and the investment probability, respectively, from the development stage are needed to form the estimator of the investment probability $g^e(\widehat{\Omega}_{kt}; \theta)$ in the exploration stage.

4.2 Step 2: Generalized Method of Moments

After obtaining estimates of the continuation values and predicted probabilities as functions of the state variables Ω_{kt} and the parameters θ in the first step, I estimate the parameters θ in the second step using generalized method of moments (GMM). The moments I construct involve matching the probabilities of exploration and development predicted by the model, as given by equations (19) and (17), with the respective empirical probabilities $\overline{g^e(\Omega_{kt})}$ and $\overline{g^d(\Omega_{kt})}$ in the data. I also form moments that match, for those tracts that developed, the expected development profits conditional on development predicted by the model with the actual average realized profits $\overline{\pi^d(\Omega_{kt})}$ in the data. Additional moments are constructed by interacting the above moments with the state variables. The moment function $\Psi(\Omega_{kt}, \theta)$ is therefore:

$$\begin{aligned} & \left(g^e(\widehat{\Omega}_{kt}; \theta) - \overline{g^e(\Omega_{kt})} \right) \cdot n^{\text{not-yet-e}}(\Omega_{kt}) \\ & \left(g^d(\widehat{\Omega}_{kt}; \theta) - \overline{g^d(\Omega_{kt})} \right) \cdot n^{\text{e-not-yet-d}}(\Omega_{kt}) \\ & \left(\left(E \left[\pi^d(\Omega_{kt}, \varepsilon_{it}; \theta) | \pi^d(\Omega_{kt}, \varepsilon_{it}; \theta) > \beta V^{ce}(\Omega_{kt}; \theta) \right] - \overline{\pi^d(\Omega_{kt})} \right) \cdot n^d(\Omega_{kt}) \right) \\ & \Omega'_{kt} \left(\left(g^e(\widehat{\Omega}_{kt}; \theta) - \overline{g^e(\Omega_{kt})} \right) \cdot n^{\text{not-yet-e}}(\Omega_{kt}) \right) \\ & \Omega'_{kt} \left(\left(g^d(\widehat{\Omega}_{kt}; \theta) - \overline{g^d(\Omega_{kt})} \right) \cdot n^{\text{e-not-yet-d}}(\Omega_{kt}) \right) \end{aligned}$$

where, for each state of the market Ω_{kt} , $n^{\text{not-yet-e}}(\Omega_{kt})$ is the number tracts that have yet to be explored, $n^{\text{e-not-yet-d}}(\Omega_{kt})$ is the number of tracts that have been explored but not developed, and $n^d(\Omega_{kt})$ is the number of tracts that have been developed.

The GMM estimator $\hat{\theta}$ is then the solution to the problem:

$$\min_{\theta} \left[\frac{1}{n} \sum_{k,t} \Psi(\Omega_{kt}, \theta) \right]' W_n^{-1} \left[\frac{1}{n} \sum_{k,t} \Psi(\Omega_{kt}, \theta) \right], \quad (20)$$

where n is the number of observations (i.e., the number of market-time pairs) and W_n is a weight matrix. In the present estimation, the number of moments that involved interactions with state variables is chosen so that the number of moments is equal to the number of parameters. Because the system is therefore exactly identified, the weight matrix used is the identity matrix.²⁰

Standard errors are formed by a parametric bootstrap (Pakes, Ostrovsky & Berry, 2007), as follows. Estimates of the continuation values and the parameters are used to solve for the exploration and development probabilities arising in the Markov perfect equilibrium. These probability functions are then used in conjunction with the empirical transition matrix for the exogenous variables and with random draws, with replacement, from the empirical distribution of initial conditions to simulate 100 independent panels of size equal to the actual sample size. Finally, the structural econometric model is run on each of the simulated panels. The standard error is then formed by taking the standard deviation of the estimates from each of the pseudo random samples.

To assess the finite sample distribution of the estimators, I ran two Monte Carlo experiments. The results from both experiments indicate that the estimators recover the actual parameter values fairly well. Details and results of these experiments are provided in Appendix B.

5 Data

I use a data set on federal lease sales in the Gulf of Mexico between 1954 and 1990 compiled by Kenneth Hendricks and Robert Porter from U.S. Department of Interior data. There are three types of tracts that can be offered in an oil and gas lease sale: wildcat, drainage, and developmental. Wildcat tracts are located in

²⁰If the system is overidentified, a two-step GMM estimator can be used (Hansen, 1982; Graham, 2005). In the first step, a preliminary estimate of θ is obtained using the identity matrix as the weight matrix. In the second step, this preliminary estimate of θ is used to construct the optimal weighting matrix as specified by Chamberlain (1987), which is then used to obtain the final estimate of θ .

regions where no exploratory drilling has occurred previously and therefore where the geology is not well known. Exploration on wildcat tracts entails searching for a new deposit. In contrast, both drainage and developmental tracts are adjacent to tracts on which deposits have already been discovered; developmental tracts, in addition, are tracts that have been previously offered in an earlier lease sale but either whose previous bids were rejected as inadequate or whose leases were relinquished because no exploratory drilling was done (Porter, 1995).

I focus my attention on wildcat tracts offshore of Louisiana and Texas that were auctioned between 1954 and 1979, inclusive. I do so for several reasons. First, my restrictions are similar to those made by Hendricks and Porter (1996), thus enabling me to best compare my results with theirs. Second, since wildcat tracts are tracts on which no exploratory drilling has occurred previously, information externalities are likely to be most acute. Third, because the data set only contains production data up until 1990, the restriction to tracts sold before 1980 eliminates any censoring of either drilling or production.²¹ Additional restrictions I impose for a tract to be included in my data set are that it must be a tract for which location data is available, for which the first exploration occurred neither before the sale date nor after the lease term,²² and for which production did not occur before exploration.

For my structural estimation, I focus on two-tract markets.²³ In order for two tracts to qualify as a market, the two tracts must be within 6 miles north and south of each other or 6 miles east and west of each other,²⁴ both the sale dates and the

²¹Another reason to focus on the earlier lease sales is that post-auction lease transfers occurred less frequently in the past (Porter, 1995; John Rodi, Minerals Management Service, personal communication, 8 May 2003; Robert Porter, personal communication, 21 May 2003).

²²It is possible for a lease to receive a suspension of production (SOP) or suspension of operations (SOO) which will extend the life of the lease beyond its primary term (Jane Johnson, Minerals Management Service, personal communication, October 29, 2003). Exploratory drilling first occurred after the lease term on 77 (or 3.1 %) of the 2481 wildcat tracts sold before 1980.

²³I restrict the size of the market to two tracts for two reasons. First, limiting the market size to two minimizes the state space. This is because the number of possible combinations of state variables, which is the product of the cardinality of the supports of each of the state variables, is quadratic in the market size. Second, when there are only two tracts in the market, each tract is equidistant to all its neighbors. Since the federal tracts in the Gulf of Mexico form a grid, the next sensible size of a market is four. With four tracts, however, diagonal tracts are not as close together as tracts that share a side are; as a consequence, the distance between each pair of neighbors in the market is not the same. It is plausible that firms may weight the behavior of their neighbors by their distance; when neighbors are no longer symmetric, the econometric model becomes more complicated.

²⁴I choose 6 miles because the maximum tract size is 5760 acres, or 3 miles by 3 miles (Marshall Rose, Minerals Management Service, personal communication, 17 April 2003; Larry Slaski,

lease terms must be the same for both tracts, the tracts must be owned by different firms, and no other wildcat tracts from the same sale date can be within 6 miles of either tract. There are 87 such markets in my data set. Figure 1 maps out the tracts used.

Each discrete period t represents a year. The time of exploration is the year of the tract's first spud date. The time of development is the year of the tract's first production date.²⁵ It is possible that development begins in the same year as exploration does.

The panel spans the years 1954 to 1990. A market enters the panel when its tracts are sold. If both tracts are eventually developed, the market exits the panel when the second tract to develop first begins development. If neither tract has explored by the end of the five-year lease term, the market exits the panel when the leases expire.

There are two endogenous state variables N_{kt} : the total number of tracts in the market that have been explored and the total number of tracts in the market that have been developed. Because there are two tracts in each market, there are three possible values for the the total number of tracts that have been explored: 0, 1 and 2. As for the total number of tracts that have been developed, because a market ends once both tracts have been developed, there are only two possible values: 0 and 1.

The coefficients γ_N in the development profit function on the endogenous state variables indicate whether and how one firm's profits depend on the production decisions of its neighbors. Positive values of the coefficients γ_N would indicate that the information and extraction externalities were positive on net, and therefore that the information externality was dominant. Negative values would indicate that the externalities were negative on net, and therefore that the extraction externality was dominant. Moreover, one would expect each coefficient γ_N to be less than the cost of development. Otherwise, if the coefficients were greater than or equal to the development cost, this would mean that having a neighbor explore or develop would offset the cost of development, making development essentially costless. The effects

Minerals Management Service, personal communication, 25 April 2003). I convert latitude and longitude to miles using the following following factors from the Louisiana Sea Grant web site (<http://lamer.lsu.edu/classroom/deadzone/changedistance.htm>): 1 minute longitude in Louisiana offshore = 60.5 miles; 1 minute latitude = 69.1 miles.

²⁵More specifically, for both the time of exploration and the time of development, I take the floor of the number of years since the lease began.

of strategic interaction are unlikely to be that large.

I use three exogenous state variables X_{kt} ; these variables were chosen based on considerations of state space and data availability. The first exogenous state variable is the discretized average winning bid per acre over the two tracts in market k at time t . Because the winning bid is a measure of the value of the tract, the average winning bid over the tracts in the market captures any fixed market-specific variables such as geological structures that may affect profits. To construct this variable I average the winning bids per acre over the two tracts in the market, and then discretize the average into three bins: 0 = low (0 to 1 thousand 1982 \$/acre), 1 = medium (1 thousand to 5 thousand 1982 \$/acre), and 2 = high (over 5 thousand 1982 \$/acre). One expects that profits would increase in the value of the tract, and therefore that the coefficient on the winning bid in the development profit function is positive.

The second exogenous state variable is the discretized real drilling cost at time t . I use data on annual drilling costs from the American Petroleum Institute's *Joint Association Survey of the U.S. Oil & Gas Producing Industry* for the 1969-1975 data and its *Joint Association Survey on Drilling Costs* for the 1976-1990 data. The cost is average cost per well over all offshore wells (oil wells, gas wells, dry holes), in nominal dollars. I convert the nominal costs to real costs in 1982-1984 dollars using the consumer price index (CPI). I discretize the real drilling cost into two bins: 0 = low (0 to 2.5 million 1982-1984 \$/well), 1 = high (over 2.5 million 1982-1984 \$/well). The drilling cost is a measure of both exploration costs and development costs, and therefore enters into both the extraction profit function and the development profit function. In particular, from (8), the exploration cost is assumed to be the following function of the discretized drilling cost $drill_cost_t$:

$$c^e(\Omega_{kt}; \theta) = -\alpha \cdot (drill_cost_t + 1), \quad (21)$$

where α is now a scalar, so that exploration profits are:

$$\pi^e(\Omega_{kt}, \mu_{it}; \theta) = E_\varepsilon [V^e(\Omega_{kt}, \varepsilon_{it}; \theta) | \Omega_{kt}] + \alpha \cdot (drill_cost_t + 1) + \mu_{it}. \quad (22)$$

The discretized drilling cost is incremented by one so that costs are non-zero even when they fall in the low bin. Figure 2 plots the real drilling cost data, along with the bins.²⁶ The expected sign of the coefficient α on costs in the exploration profit

²⁶Before 1969, the real drilling cost is assumed to fall into the low bin. The 1982 real drilling

function is negative; higher costs should lower exploration profits. Similarly, the expected sign of the coefficient on costs in the development profit function is negative as well.

The third exogenous state variable is the discretized real oil price. I use the U.S. average crude oil domestic first purchase price from the EIA Annual Energy Review and deflate the time series to 1982-1984 dollars per barrel using the CPI. I discretize the real oil price into three bins: 0 = low (0 to 13 1982-1984 \$/barrel), 1 = medium (13 to 25 1982-1984 \$/barrel), and 2 = high (over 25 1982-1984 \$/barrel). Figure 3 plots the real oil price, along with the bins. The expected sign on oil price in the development profit function is positive: higher oil prices should increase revenues.

Table 1 presents summary statistics for the panel data used for the structural estimation. There are 1646 observations spanning 174 tracts and 87 markets. The markets range in duration from 2 years to 36 years, with an average length of 17.92 years (s.d. = 11.17). Of the 174 tracts, 122 were eventually explored and 66 were eventually developed. The average number of years to exploration, conditional on exploring, is 1.21 (s.d. = 1.42). The average number of years to development, conditional on developing, is 5.79 (s.d. = 3.51). For the 66 tracts that developed, the predicted ex post revenues, as calculated by Hendricks, Porter and Boudreau (1987), range from \$22,000 to \$298 million, with an average of \$49.34 million (s.d. = 65.53 million). The real gross profits from development, which are the predicted ex post revenues times the government royalty rate minus costs, but not net of the bid, also as calculated by Hendricks, Porter and Boudreau (1987), range from -\$38.10 million to \$18.80 million, with an average of -\$10.83 million (s.d. = 9.57 million). Table 6 also provides the summary statistics for the state variables in the panel. The number of possible combinations of state variables is the product of the cardinality of the supports of each of the state variables, or $3 \times 2 \times 3 \times 2 \times 3 = 108$.

The maximum tract size, as stipulated by a provision in section 8(b) of the Outer Continental Shelf Lands Act (OCSLA), 43 U.S.C. 1337(b)(1), is 5760 acres, or 3 miles by 3 miles (Marshall Rose, Minerals Management Service, personal communication, 17 April 2003). The distribution of tract sizes is presented in Figure 4. Most tracts are either 2500 acres, 5000 acres or 5760 acres in size. The mean tract size is 4460 acres (s.d. = 1300), and the median tract size is 5000 acres. Table 2

cost, which was unavailable because the 1982 issue of the *Joint Association Survey on Drilling Costs* was out of print, is assumed to fall in the high bin.

presents summary statistics by tract size. Large tracts are defined as tracts that are greater than or equal to 5000, 4000 or 3000 acres in size, respectively; small tracts are defined as tracts that are less than 5000, 4000 or 3000 acres in size, respectively. The distributions of the variables appear to be similar across tract sizes. Small tracts make up a small percentage of the tracts sold: of the tracts used in the structural estimation, only 37%, 27%, and 24% of tracts were less than 5000 acres, 4000 acres and 3000 acres in size, respectively.

6 Results

There are three different types of parameters to be estimated: the parameters $(\sigma_\mu, \sigma_\varepsilon)$ governing the distribution of the private information, the coefficient α on drilling costs in the exploration profit function, and the coefficients γ on the state variables in the development profit function.

The results from running the structural model on all tracts in the panel regardless of tract size are shown in Table 3. The coefficients $\gamma_N \equiv (\gamma_{tote}, \gamma_{told})$ on the two endogenous state variables can be interpreted as follows. Since the total number of tracts in the market that have been explored increases by one if a neighbor explores, the coefficient γ_{tote} on this variable measures how the profits from development change when a neighbor explores. Similarly, since the total number of tracts in the market that have been developed can only take values of 0 and 1, with 0 indicating that the neighbor has not developed and 1 indicating that the neighbor has developed, the coefficient γ_{told} on this variable measures how the profits from development change when a neighbor develops first. An important result is that the coefficient on the total number of tracts in the market that have been explored is statistically insignificant, which means that firms do not interact strategically on net during exploration.²⁷ In contrast, the coefficient on the total number of tracts in the market that have been developed is statistically significant and positive, which means that a firm's profits increase when its neighbor develops. This seems reasonable, because when a neighbor develops following exploration, this is a signal to the firm that the neighbor's

²⁷This result is consistent with the results from reduced-form analyses consisting of a discrete response model of a firm's exploration timing decision using variables based on the timing of a neighbor's lease term as instruments for the neighbor's decision, as these analyses also do not indicate that strategic, non-cooperative behavior occurs during the first stage of petroleum production (Lin, forthcoming).

exploratory efforts were successful, and therefore that there may be deposits present.

To assess the economic significance of the strategic interactions, I compare the coefficients γ_N on the endogenous state variables with the development cost $|\gamma_{drill} \cdot drill_cost_t|$, where γ_{drill} denotes the coefficient on the discretized real drilling cost $drill_cost_t$ in the development profit function. As noted above, one would expect each coefficient γ_N to be less than the cost of development. Since the maximum development cost is given by $|\gamma_{drill} \cdot \max(drill_cost_t)| = |\gamma_{drill}|$, one would therefore expect $\gamma_{tote} < |\gamma_{drill}|$ and $\gamma_{told} < |\gamma_{drill}|$. The results are consistent with the expectations. Further comparison of the coefficients γ_N with the mean development cost $|\gamma_{drill} \cdot \overline{drill_cost_t}|$ can give a measure of the economic importance of a neighbor's decisions to a firm's profits. In particular, the relative importance of a neighbor's exploration decision as a fraction of a firm's costs is given by $\frac{\gamma_{tote}}{|\gamma_{drill} \cdot \overline{drill_cost_t}|} = \frac{\gamma_{tote}}{0.49|\gamma_{drill}|} = 0.06$. Similarly, the relative importance of a neighbor's development decision as a fraction of a firm's costs is given by $\frac{\gamma_{told}}{|\gamma_{drill} \cdot \overline{drill_cost_t}|} = \frac{\gamma_{told}}{0.49|\gamma_{drill}|} = 0.19$. The small values of both of these numbers indicate that the effects of neighbors' decisions are second-order compared to costs.

In absolute terms, strategic interactions in exploration are not only statistically insignificant, but economically insignificant as well: changes in profits (in 1982 \$) from a neighbor's exploration less than \$80,000 and greater than \$640,000 can be rejected at a 5% level. These values are small relative to predicted ex post revenues, which average \$49.34 million. Strategic interactions in development are statistically significant but only moderately economically significant: changes in profits from a neighbor's development less than \$700,000 and greater than \$980,000 can be rejected at a 5% level.

As for the values of the parameters governing the distribution of private information, both the parameter σ_μ from the distribution of the pre-exploration shock μ_{it} and the parameter σ_ε from the distribution of the pre-development shock ε_{it} are statistically significant. This suggests that private information has a statistically significant impact on both the exploration decision and the development decision.

In terms of economic significance, one way to interpret the mean σ_μ of the pre-exploration shock μ_{it} is to compare it with exploration costs. Since both the pre-exploration shock μ_{it} and the exploration cost function $c^e(\Omega_{kt}; \theta)$ enter linearly into the exploration profit function (9), the importance of private information in the exploration decision can be measured by comparing the mean σ_μ of the shock

with the mean $\overline{c^e(\Omega_{kt}; \theta)}$ of the costs. Expressed as a fraction of the mean exploration costs, where the mean exploration costs are computed by substituting the mean value of the discretized real drilling cost into the equation (21) for the exploration costs, the relative importance of private information is therefore given by: $\frac{\frac{\sigma_\mu}{c^e(\Omega_{kt}; \theta)}}{\frac{\sigma_\mu}{-\alpha \cdot (\overline{drill_cost_t} + 1)}} = \frac{\sigma_\mu}{-1.49\alpha}$. A large value of $\frac{\sigma_\mu}{-1.49\alpha}$ would indicate that private information plays a large role relative to costs in the first-stage exploration decision; a small value would indicate that costs are more relatively more important. In this case, the value is 0.33. Private information is about a third as important as costs. Thus, the role of private information in the exploration decision is both economically and statistically significant.

The mean σ_ε of the pre-development shock ε_{it} can be similarly compared with development costs to assess the importance of private information in the second-stage development decision. Since both the pre-development shock ε_{it} and the development cost $|\gamma_{drill} \cdot \overline{drill_cost_t}|$ enter linearly into the development profit function (1), the importance of private information in the development decision can be measured by comparing the mean σ_ε of the shock with the mean $|\gamma_{drill} \cdot \overline{drill_cost_t}|$ of the costs. Expressed as a fraction of the mean development costs, the relative importance of private information is therefore given by: $\frac{\frac{\sigma_\varepsilon}{|\gamma_{drill} \cdot \overline{drill_cost_t}|}}{\frac{\sigma_\varepsilon}{0.49|\gamma_{drill}|}} = \frac{\sigma_\varepsilon}{0.49|\gamma_{drill}|}$. A large value of $\frac{\sigma_\varepsilon}{0.49|\gamma_{drill}|}$ would indicate that private information plays a large role relative to costs in the second-stage development decision; a small value would indicate that costs are more relatively more important. In this case, the value is 0.91. Private information is almost as important as costs in the development decision.

The coefficients on the other covariates are all significant and have the expected sign. As expected, the coefficient on the discretized drilling cost is negative in both the exploration profit equation and the development profit equation. Also as expected, the profits from development increase in both the average winning bid and in the real oil price.

There are several possible explanations why the results do not provide evidence for strategic, non-cooperative behavior during exploration. One is that the tract sizes are large enough that cross-tract externalities are insignificant or even nonexistent. A second is that cross-tract externalities exist, but firms owning neighboring tracts cooperate to jointly internalize the externalities they impose on each other, for example through joint ventures or unitization. A third is that cross-tract externalities are significant, but the positive information externality exactly cancels the negative

extraction externality, resulting in zero net strategic effects.

To distinguish among these three explanations for the lack of strategic interactions during exploration, I estimate the strategic interactions by tract size. If externalities are insignificant when tracts are large enough, then one would expect to see strategic, non-cooperative behavior only on small tracts. This is because the smaller the tract size, the more likely the tracts are located over a common pool, and therefore the more acute the information and extraction externalities faced by the firms. Evidence for significant strategic interactions on small tracts but not on large tracts would thus be consistent with the first explanation.

If externalities exist even for the largest tracts in the sample, but are eliminated through coordination, then, assuming firms would coordinate regardless of tract size, one would not expect to see any strategic, non-cooperative behavior even on small tracts. Insignificant strategic interactions regardless of tract size would thus be consistent with the second explanation.

If strategic interactions do not occur on net because the positive information externality exactly cancels the negative extraction externality, then one may not expect the exact cancellation to still take place when the tract size is small. This is because the geographical span of the information externality is larger than that of the extraction externality: while the former only requires that tracts may share common geological features, the latter requires that tracts may be located over a common pool. As a consequence, it is possible for the information externality to be present on all the tracts in the sample, but for the extraction externality to be present on only the smaller tracts. Theory therefore suggests that the importance of the extraction externality relative to the information externality should be greater on small tracts than on large tracts. Thus, if the externalities cancel when all the tract sizes are considered, one might expect that the negative extraction externality would dominate the positive information externality when the sample is limited to small tracts only. Strategic interactions that are more significantly negative on small tracts than on large tracts would therefore be consistent with the third explanation.

Running the structural model on subsamples of the data set that differed in the acreage of the tracts in the panel would therefore enable one to distinguish among these three explanations for the lack of strategic interactions in the pooled sample. Assuming that the tract sizes differ for exogenous reasons, the results by acreage will also give a sense of whether or not the government can change the extent the which

firms behave strategically by changing the size of the tracts.

Table 4a presents the results from running the structural model on three subsamples consisting of larger tracts, defined as tracts that are greater than or equal to 5000 acres, 4000 acres, and 3000 acres in size, respectively. As the coefficients on the endogenous variables indicate, strategic interactions are neither economically nor statistically significant for any of these subsamples. For tracts greater than or equal to 5000 acres in size, the 95% confidence interval for the effect of a neighbor's exploration on profits (in 1982 \$) ranges from -\$0.18 million to \$0.10 million, and the 95% confidence interval for the effect of a neighbor's development on profits ranges from -\$0.07 million to \$0.05 million. These values are small relative to predicted ex post revenues, which average \$49.34 million in the pooled sample. Investment decisions and profits are instead driven primarily by private information, the average winning bid, the real drilling cost, and the real oil price.

Table 4b presents the results from running the structural model on three subsamples consisting of smaller tracts, defined as tracts that are less than 5000 acres, 4000 acres, and 3000 acres in size, respectively. Strategic interactions are statistically and economically significant in all three subsamples. The coefficients in the three specifications on the number of tracts in the market that have been explored indicate that real development profits decrease by a statistically significant \$10.73 million to a statistically significant \$15.08 million when a neighbor explores; a neighbor's exploration is roughly as important to profits as maximum development costs. The negative extraction externality thus appears to dominate: when a neighbor explores, a firm's profits decrease because the neighbor has begun production and is likely to eventually compete with the firm for the same common pool. The coefficients in the three specifications on the number of tracts in the market that have been developed indicate that real development profits increase by a statistically significant \$1.09 million to a statistically significant \$3.29 million when a neighbor develops. In this case, the positive information externality dominates: a firm benefits when its neighbor develops after it explores because this is a signal to the firm that the neighbor's exploratory efforts were successful, and therefore that there may be deposits present. The magnitude of the positive net strategic effect that results from a neighbor's development is one order of magnitude smaller than the negative net strategic effect that results from a neighbor's exploration. As expected, the magnitude of the strategic interactions in both exploration and development increase monotonically as the tract

sizes get smaller.

For the smaller tract sizes, investment decisions and profits are also driven by private information, the average winning bid, the real drilling cost, and the real oil price, as before, with two differences. First, pre-development private information plays a larger role in decision-making and profits for the smaller tracts than it does for the larger tracts. Second, the coefficient on the average winning bid is now negative, instead of positive, and is of smaller magnitude. On the smaller tracts, investment decisions and profits are governed more by post-auction private and public information, including that from the actions of one's neighbors, than by the initial estimate of tract value.

The results therefore indicate that when a neighbor explores, there are no strategic effects on larger tracts even though information and extraction externalities exist because these externalities cancel each other out. The cancelling no longer occurs on smaller tracts, however, where the negative extraction externality dominates. When a neighbor develops, there are no strategic effects on larger tracts because the tract size is large enough to enable a firm to internalize the externalities on its own. On small tracts, however, externalities are acute and the positive information externality dominates. Small tracts make up a small percentage of the tracts sold: of the tracts used in the structural estimation, only 37%, 27%, and 24% of tracts were less than 5000 acres, 4000 acres and 3000 acres in size, respectively. Thus, for the majority of tracts, externalities do not cause inefficient strategic interactions on net. The results suggest that, by making most of the tracts at least 5000 acres in size, the federal government has minimized the net effects of any externalities that may be present, and has thus eliminated any potential inefficiencies in petroleum production.

7 Conclusion

When individual petroleum-producing firms make their exploration and development investment timing decisions, information externalities and extraction externalities may lead them to interact strategically with their neighbors. A positive information externality arises if tracts are located over a common pool or share common geological features so that their ex post values are correlated, since firms learn information about their own tracts when other firms drill exploratory wells or install

production platforms on neighboring tracts. A negative extraction externality arises when tracts are located over a common pool, since firms are competing for the same stock of petroleum. Owing to both information and extraction externalities, the dynamic decision-making problem faced by a petroleum-producing firm is not merely a single-agent problem, but rather can be viewed as a multi-agent, non-cooperative game in which firms behave strategically and base their exploration and development policies on those of their neighbors.

This paper examines whether strategic considerations arising from information and extraction externalities are present. In particular, it analyzes whether a firm's investment timing decisions and profits depend on the decisions of firms owning neighboring tracts of land. The econometric approach employed is a structural econometric model of the firms' multi-stage investment timing game.

The research presented in this paper is important for several reasons. First, an empirical analysis of investment timing decisions enables one to examine whether the strategic interactions that are predicted in theory actually occur in practice. Second, the estimation of strategic interactions, especially those that arise in dynamic decision-making, is of methodological interest. The structural econometric methodology employed can be used to analyze externalities in a variety of contexts, including spillovers that arise during research and development. Third, my results have implications for leasing policy: if the strategic effects and externalities turn out to be large, then the program by which the U.S. government leases tracts to firms may be inefficient, and possible modifications should be considered.

Do the positive information externalities and negative extraction externalities have any net strategic effect that may cause petroleum production to be inefficient? The answer depends on tract size. When the tract sizes are large, firms do not impose externalities on each other on net when choosing to explore or develop, and, as a consequence, strategic considerations are second-order. This is the case with most of the tracts in the federal leasing program. However, in the few cases where the tract size is small, externalities do matter, and they cause firms to interact strategically with their neighbors. As expected, these externalities intensify as the tract size decreases. Also as predicted by theory, the relative importance of the extraction externality from exploration with respect to the information externality is greater on small tracts than on large tracts. The results suggest that, in making most tracts at least 5000 acres in size, the federal government has minimized inefficiencies in petroleum production

that may have resulted from non-cooperative strategic interactions.

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8 Appendix A: Derivations

In this Appendix, I provide the details for deriving the equations for the continuation values and investment probabilities used in the structural model of the investment timing game.

8.1 Stage 2: Development

Equation (5) for the continuation value V^{ce} to waiting instead of developing is derived as follows. Substituting in equation (1) for development profits, the expected truncated profits from development conditional on development can be written as:

$$\begin{aligned}
& E [\pi^d(\Omega_{kt}, \varepsilon_{it}; \theta) | \pi^d(\Omega_{kt}, \varepsilon_{it}; \theta) > \beta V^{ce}(\Omega_{kt}; \theta)] \\
&= E [\pi_0^d(\Omega_{kt}; \theta) + \varepsilon_{it} | \pi_0^d(\Omega_{kt}; \theta) + \varepsilon_{it} > \beta V^{ce}(\Omega_{kt}; \theta)] \\
&= \pi_0^d(\Omega_{kt}; \theta) + E [\varepsilon_{it} | \varepsilon_{it} > \beta V^{ce}(\Omega_{kt}; \theta) - \pi_0^d(\Omega_{kt}; \theta)] \\
&= \beta V^{ce}(\Omega_{kt}; \theta) + \sigma_\varepsilon,
\end{aligned} \tag{23}$$

where the final step comes from the exponential distribution for ε_{it} .²⁸ Using equation (23) for the expected truncated profits conditional on development, the continuation value can thus be written as:

$$\begin{aligned}
& V^{ce}(\Omega_{kt}; \theta) \\
&= E \left[g^d(\Omega_{k,t+1}; \theta) \cdot E [\pi^d(\Omega_{k,t+1}, \varepsilon_{i,t+1}; \theta) | \pi^d(\Omega_{k,t+1}, \varepsilon_{i,t+1}; \theta) > \beta V^{ce}(\Omega_{k,t+1}; \theta)] \mid \Omega_{k,t}, I_{it}^d = 0 \right] \\
&\quad + (1 - g^d(\Omega_{k,t+1}; \theta)) \cdot \beta V^{ce}(\Omega_{k,t+1}; \theta) \\
&= E [g^d(\Omega_{k,t+1}; \theta) \cdot (\beta V^{ce}(\Omega_{k,t+1}; \theta) + \sigma_\varepsilon) + (1 - g^d(\Omega_{k,t+1}; \theta)) \cdot \beta V^{ce}(\Omega_{k,t+1}; \theta) \mid \Omega_{k,t}, I_{it}^d = 0] \\
&= E[\beta V^{ce}(\Omega_{k,t+1}; \theta) + \sigma_\varepsilon g^d(\Omega_{k,t+1}; \theta) \mid \Omega_{k,t}, I_{it}^d = 0],
\end{aligned}$$

which yields equation (5) as desired.

Equation (6) for the development probability $g^d(\Omega_{kt}; \theta)$ is derived as follows. The probability $g^d(\Omega_{kt}; \theta)$ of developing tract i at time t given that development of tract i has not occurred previously can be expressed as:

$$g^d(\Omega_{kt}; \theta) = \Pr(\pi^d(\Omega_{kt}, \varepsilon_{it}; \theta) > \beta V^{ce}(\Omega_{kt}; \theta)),$$

where $\Pr(\cdot)$ denotes probability.

Substituting in equation (1) for development profits, we get:

$$\begin{aligned}
g^d(\Omega_{kt}; \theta) &= \Pr(\pi_0^d(\Omega_{kt}; \theta) + \varepsilon_{it} > \beta V^{ce}(\Omega_{kt}; \theta)) \\
&= \Pr(\varepsilon_{it} > \beta V^{ce}(\Omega_{kt}; \theta) - \pi_0^d(\Omega_{kt}; \theta)).
\end{aligned}$$

²⁸If $\varepsilon \sim \text{exponential}(\sigma_\varepsilon)$, then its pdf is $f(\varepsilon) = \frac{1}{\sigma_\varepsilon} \exp(-\frac{\varepsilon}{\sigma_\varepsilon})$ (Casella & Berger, 1990).

Substituting in the exponential distributional assumption for ε_{it} , we get:

$$g^d(\Omega_{kt}; \theta) = \exp \left(-\frac{\beta V^{ce}(\Omega_{kt}; \theta) - \pi_0^d(\Omega_{kt}; \theta)}{\sigma_\varepsilon} \right),$$

which yields equation (6) as desired.

8.2 Stage 1: Exploration

Equation (13) for the continuation value $V^{cn}(\cdot)$ to waiting instead of exploring is derived as follows.

$$\begin{aligned} & V^{cn}(\Omega_{kt}; \theta) \\ = & E \left[g^e(\Omega_{k,t+1}; \theta) \cdot E \left[\pi^e(\Omega_{k,t+1}, \mu_{i,t+1}; \theta) | \pi^e(\Omega_{k,t+1}, \mu_{i,t+1}; \theta) > \beta V^{cn}(\Omega_{k,t+1}; \theta) \right] | \Omega_{kt}, I_{it}^e = 0 \right] \\ & + (1 - g^e(\Omega_{k,t+1}; \theta)) \cdot \beta V^{cn}(\Omega_{k,t+1}; \theta) \\ = & E \left[g^e(\Omega_{k,t+1}; \theta) \cdot (\pi_0^e(\Omega_{kt}; \theta) + E[\mu_{it} | \mu_{it} > \beta V^{cn}(\Omega_{k,t+1}; \theta) - \pi_0^e(\Omega_{kt}; \theta)]) | \Omega_{kt}, I_{it}^e = 0 \right] \\ & + (1 - g^e(\Omega_{k,t+1}; \theta)) \cdot \beta V^{cn}(\Omega_{k,t+1}; \theta) \\ = & E[g^e(\Omega_{k,t+1}; \theta) \cdot (\beta V^{cn}(\Omega_{k,t+1}; \theta) + \sigma_\mu) + (1 - g^e(\Omega_{k,t+1}; \theta)) \cdot \beta V^{cn}(\Omega_{k,t+1}; \theta) | \Omega_{kt}, I_{it}^e = 0] \\ = & E[\beta V^{cn}(\Omega_{k,t+1}; \theta) + \sigma_\mu g^e(\Omega_{k,t+1}; \theta) | \Omega_{kt}, I_{it}^e = 0]. \end{aligned}$$

The derivation of equation (14) for the exploration probability $g^e(\Omega_{kt}; \theta)$ proceeds as follows. The exploration policy function $g^e(\Omega_{kt}^n; \theta)$, which is the probability of exploring tract i at time t given that exploration of tract i has not occurred previously, is given by:

$$g^e(\Omega_{kt}; \theta) = \Pr(\pi^e(\Omega_{kt}, \mu_{it}; \theta) > \beta V^{cn}(\Omega_{kt}; \theta)). \quad (24)$$

Substituting in equations (9) and (7) for exploration profits, we get:

$$\begin{aligned}
& g^e(\Omega_t^n; \theta) \\
&= \Pr(\pi_0^e(\Omega_{kt}; \theta) + \mu_{it} > \beta V^{cn}(\Omega_{kt}; \theta)) \\
&= \Pr(\mu_{it} > \beta V^{cn}(\Omega_{kt}; \theta) - \pi_0^e(\Omega_{kt}; \theta)) \\
&= \exp\left(-\frac{\beta V^{cn}(\Omega_{kt}; \theta) - \pi_0^e(\Omega_{kt}; \theta)}{\sigma_u}\right) \\
&= \exp\left(-\frac{\beta V^{cn}(\Omega_{kt}; \theta) - E_\varepsilon[V^e(\Omega_{kt}, \varepsilon_{it}; \theta)|\Omega_{kt}] + c^e(\Omega_{kt}; \theta)}{\sigma_u}\right) \\
&= \exp\left(-\frac{\beta V^{cn}(\Omega_{kt}; \theta) - (\beta V^{ce}(\Omega_{kt}; \theta) + \sigma_\varepsilon g^d(\Omega_{kt}; \theta)) + c^e(\Omega_{kt}; \theta)}{\sigma_u}\right),
\end{aligned}$$

which yields equation (14).

9 Appendix B: Monte Carlo Experiments

To assess the finite sample distribution of the estimators, Table 5 presents the results from two Monte Carlo experiments. In each experiment, I run 100 simulations of 87 markets each for a given set of fixed parameters.²⁹ To generate the simulated data, the Markov perfect equilibrium is first computed for the given set of parameters using dynamic programming. The exploration and development policy functions arising from the equilibrium are then used in conjunction with the empirical transition matrix for the exogenous variables and with random draws, with replacement, from the empirical distribution of initial conditions to simulate sample paths for each of 87 markets to form a simulated panel data set. Finally, the structural econometric model is run on each of the 100 simulated panels to obtain the finite sample distribution of the estimators.

The table reports the results for two different experiments. In experiment (1), the parameters are chosen so that neighbors have a large, negative effect and so that private information plays a small role in the exploration decision and a moderate role in the development decision.³⁰ In particular, the relative importance of a neighbor's exploration decision as a fraction of a firm's average development cost

²⁹I chose to simulate 87 markets because there are 87 markets in my actual data set.

³⁰See text for an explanation of how to interpret the relative values of the parameters.

is given by $\frac{\gamma_{\text{total}}}{\gamma_{\text{drill}} \cdot \text{drill_cost}_t} = \frac{\gamma_{\text{total}}}{0.49|\gamma_{\text{drill}}|} = -1.36$. Similarly, the relative importance of a neighbor's development decision as a fraction of a firm's average development cost is given by $\frac{\gamma_{\text{totald}}}{\gamma_{\text{drill}} \cdot \text{drill_cost}_t} = \frac{\gamma_{\text{totald}}}{0.49|\gamma_{\text{drill}}|} = -1.59$. Thus, when a neighbor explores or develops, this decreases profits by more than mean development costs. In terms of private information, the mean pre-exploration shock is only a small fraction of the mean exploration costs: $\frac{\sigma_{\mu}}{c^e(\Omega_{kt};\theta)} = \frac{\sigma_{\mu}}{-\alpha \cdot (\text{drill_cost}_t + 1)} = \frac{\sigma_{\mu}}{-1.49\alpha} = 0.08$. The mean pre-development shock is only a moderate fraction of mean development costs: $\frac{\sigma_{\epsilon}}{\gamma_{\text{drill}} \cdot \text{drill_cost}_t} = \frac{\sigma_{\epsilon}}{0.49|\gamma_{\text{drill}}|} = 0.91$. According to results of the first experiment, the estimators appear to recover the actual parameter values fairly well.

In experiment (2), the parameters are chosen so that neighbors have a small, positive effect and so that private information plays a large role in both the exploration and development decisions. In particular, the relative importance of a neighbor's exploration decision is a small fraction of a firm's development costs: $\frac{\gamma_{\text{total}}}{\gamma_{\text{drill}} \cdot \text{drill_cost}_t} = \frac{\gamma_{\text{total}}}{0.49|\gamma_{\text{drill}}|} = 0.51$. The relative importance of a neighbor's development decision is also a small fraction of a firm's costs: $\frac{\gamma_{\text{totald}}}{\gamma_{\text{drill}} \cdot \text{drill_cost}_t} = \frac{\gamma_{\text{totald}}}{0.49|\gamma_{\text{drill}}|} = 0.26$. In terms of private information, the mean pre-exploration shock is over two and a half times the mean exploration costs: $\frac{\sigma_{\mu}}{c^e(\Omega_{kt};\theta)} = \frac{\sigma_{\mu}}{-\alpha \cdot (\text{drill_cost}_t + 1)} = \frac{\sigma_{\mu}}{-1.49\alpha} = 2.68$. The mean pre-development shock is nearly four times mean development costs: $\frac{\sigma_{\epsilon}}{\gamma_{\text{drill}} \cdot \text{drill_cost}_t} = \frac{\sigma_{\epsilon}}{0.49|\gamma_{\text{drill}}|} = 3.82$. According to results of the second experiment, the estimators appear to recover the actual parameter values fairly well. Because the variances σ_{μ}^2 and σ_{ϵ}^2 of the distribution of the stochastic shocks are larger than in the first experiment, the standard deviations are larger as well, as expected.

Thus, results from both Monte Carlo experiments indicate that the estimators recover the actual parameter values fairly well.

FIGURE 1.

Wildcat tracts used (2-tract markets)

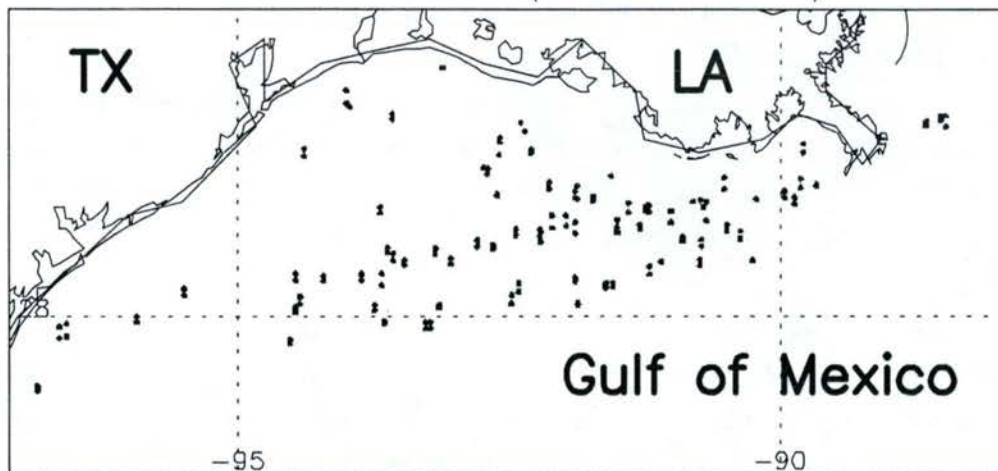


FIGURE 2.

U.S. offshore costs per well

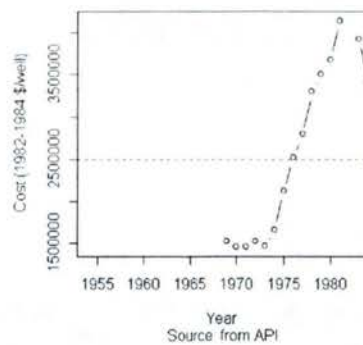


FIGURE 3.

U.S. average crude oil price

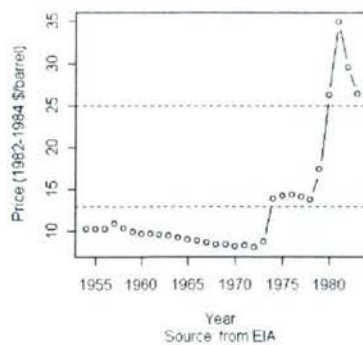


TABLE 1. Summary statistics

	# obs	mean	s.d.	min	max
<i>by tract</i>					
acreage (1000 acres)	174	4.46	1.30	0.94	5.76
number of years to exploration, conditional on exploring	122	1.21	1.42	0	4
number of years to development, conditional on developing	66	5.79	3.51	0	15
revenue (million 1982 \$), conditional on developing	66	49.34	65.53	0.02	298.0
gross profits (million 1982 \$), conditional on developing	66	-10.83	9.57	-38.10	18.80
<i>by market</i>					
number of time observations	87	17.92	11.17	2	36
<i>by market-year</i>					
# tracts in market that have been explored	1646	1.31	0.68	0	2
# tracts in market that have been developed	1646	0.31	0.46	0	1
discretized average winning bid per acre	1646	0.70	0.69	0	2
discretized real drilling cost	1646	0.49	0.50	0	1
discretized real oil price	1646	0.65	0.72	0	2

FIGURE 4.

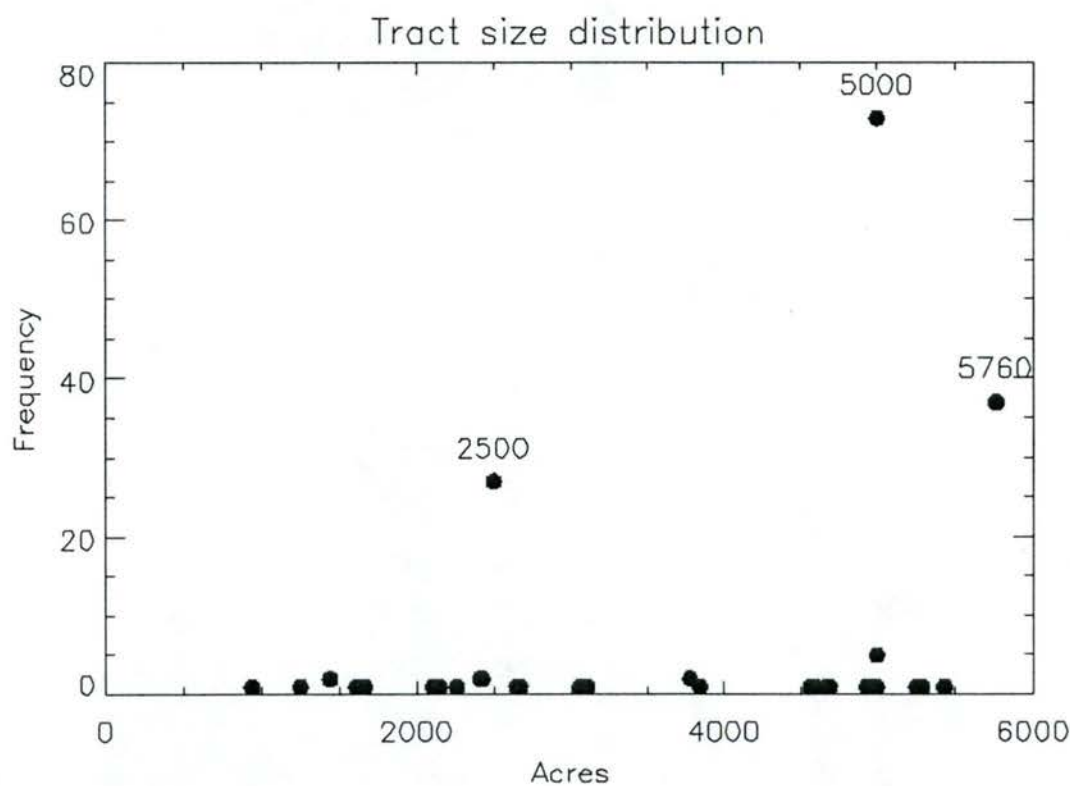


TABLE 2. Summary statistics by tract size

	Acreage						
	all tracts	≥ 5000	≥ 4000	≥ 3000	< 5000	< 4000	< 3000
<i>by tract</i>							
fraction of tracts that explore	0.70	0.75	0.75	0.76	0.60	0.60	0.55
fraction of tracts that develop	0.38	0.42	0.41	0.43	0.29	0.29	0.24
number of years to exploration, conditional on exploring	1.21 (1.42)	0.98 (1.24)	0.97 (1.25)	0.97 (1.25)	1.74 (1.75)	2.24 (1.69)	2.38 (1.75)
number of years to development, conditional on developing	5.79 (3.51)	5.72 (3.10)	5.72 (3.17)	5.58 (3.13)	5.35 (4.51)	5.83 (4.90)	6.78 (5.29)
revenue (million 1982 \$), conditional on developing	49.34 (65.53)	46.53 (62.34)	51.88 (69.85)	50.55 (67.19)	56.53 (79.46)	41.15 (56.89)	48.55 (64.42)
gross profits (million 1982 \$), conditional on developing	-10.83 (9.57)	-11.61 (8.77)	-11.23 (9.95)	-11.10 (9.62)	-7.27 (9.62)	-8.65 (7.85)	-9.08 (9.12)
<i>by market</i>							
number of time observations	17.92 (11.17)	17.38 (10.12)	17.34 (10.14)	17.20 (10.21)	19.34 (13.10)	20.47 (14.07)	20.79 (14.20)
<i>by market-year</i>							
# tracts in market that have been explored	1.31 (0.68)	1.39 (0.66)	1.38 (0.66)	1.39 (0.67)	1.16 (0.66)	1.18 (0.68)	1.13 (0.67)
# tracts in market that have been developed	0.31 (0.46)	0.36 (0.48)	0.35 (0.48)	0.37 (0.48)	0.24 (0.42)	0.25 (0.43)	0.20 (0.40)
discretized average winning bid per acre	0.70 (0.69)	0.81 (0.68)	0.81 (0.68)	0.79 (0.68)	0.57 (0.59)	0.55 (0.60)	0.58 (0.60)
discretized real drilling cost	0.49 (0.50)	0.54 (0.50)	0.54 (0.50)	0.53 (0.50)	0.43 (0.50)	0.40 (0.49)	0.39 (0.49)
discretized real oil price	0.65 (0.72)	0.72 (0.72)	0.71 (0.72)	0.70 (0.72)	0.56 (0.70)	0.52 (0.69)	0.51 (0.69)
# markets	87	55	61	64	29	21	19
# observations	1646	1012	1119	1165	590	451	414

TABLE 3. Pooled results

Parameter	Estimate	Standard Error
σ_{μ}	4.96	0.00
σ_{ε}	4.08	0.27
<i>coefficient α in the exploration profit function on:</i>		
discretized real drilling cost + 1	-10.00	0.00
<i>coefficients γ in the development profit function on:</i>		
# tracts in market that have been explored	0.28	0.18
# tracts in market that have been developed	0.84	0.07
discretized average winning bid per acre	5.56	0.09
discretized real drilling cost	-9.11	0.03
discretized real oil price	6.92	0.10
constant	5.00	0.09

Notes: There are 1646 observations spanning 87 markets. Standard errors are formed by bootstrapping 100 simulated panels of 87 markets each.

TABLE 4a. Results for large tracts

	Acreage		
	≥ 5000	≥ 4000	≥ 3000
σ_{μ}	5.00 (0.01)	4.99 (0.00)	5.00 (0.01)
σ_{ε}	4.96 (0.05)	4.96 (0.04)	4.96 (0.04)
<i>coefficient α in the exploration profit function on:</i>			
discretized real drilling cost + 1	-10.00 (0.00)	-10.00 (0.00)	-10.00 (0.00)
<i>coefficients γ in the development profit function on:</i>			
# tracts in market that have been explored	-0.04 (0.07)	-0.04 (0.07)	-0.04 (0.05)
# tracts in market that have been developed	-0.01 (0.03)	-0.01 (0.02)	-0.01 (0.03)
discretized average winning bid per acre	5.03 (0.03)	5.03 (0.03)	5.03 (0.03)
discretized real drilling cost	-9.98 (0.02)	-9.98 (0.02)	-9.98 (0.01)
discretized real oil price	5.08 (0.02)	5.08 (0.03)	5.08 (0.02)
constant	4.96 (0.04)	4.96 (0.04)	4.96 (0.03)
 # markets	 55	 61	 64
# observations	1012	1119	1165

Notes: The acreage is the acreage of each tract in the market. Standard errors are formed by bootstrapping 100 simulated panels of size equal to the actual sample size.

TABLE 4b. Results for small tracts

	Acreage		
	< 5000	< 4000	< 3000
σ_{μ}	4.90 (0.00)	4.85 (0.00)	4.80 (0.00)
σ_{ε}	16.37 (0.03)	12.96 (0.01)	17.80 (0.02)
<i>coefficient α in the exploration profit function on:</i>			
discretized real drilling cost + 1	-9.96 (0.00)	-9.91 (0.00)	-9.86 (0.00)
<i>coefficients γ in the development profit function on:</i>			
# tracts in market that have been explored	-10.73 (0.01)	-11.14 (0.01)	-15.08 (0.01)
# tracts in market that have been developed	1.09 (0.00)	1.51 (0.00)	3.29 (0.00)
discretized average winning bid per acre	-1.14 (0.00)	-0.03 (0.00)	-3.82 (0.00)
discretized real drilling cost	-10.30 (0.00)	-9.97 (0.00)	-9.94 (0.00)
discretized real oil price	4.82 (0.00)	4.80 (0.00)	4.85 (0.00)
constant	-1.79 (0.01)	-1.17 (0.00)	-4.38 (0.00)
# markets	29	21	19
# observations	590	451	414

Notes: The acreage is the acreage of each tract in the market. Standard errors are formed by bootstrapping 100 simulated panels of size equal to the actual sample size.

TABLE 5. Monte Carlo results

	True value	(1) Mean	Standard Deviation	True value	(2) Mean	Standard Deviation
σ_μ	1	2.17	0.14	40	40.00	0.00
σ_ϵ	4	4.25	0.03	15	14.01	1.41
<i>coefficient α in the exploration profit function on: discretized real drilling cost + 1</i>	-8	-7.70	0.04	-10	-10.00	0.01
<i>coefficients γ in the development profit function on:</i>						
# tracts in market that have been explored	-6	-6.00	0.00	2	1.81	0.22
# tracts in market that have been developed	-7	-7.00	0.00	1	0.93	0.08
discretized average winning bid per acre	8	8.00	0.00	5	4.97	0.02
discretized real drilling cost	-9	-9.00	0.00	-8	-8.00	0.01
discretized real oil price	10	10.00	0.00	5	4.98	0.01
constant	11	11.00	0.00	-15	-15.06	0.03

Notes: Results are from 100 simulations of 87 markets each.