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Market Integration and Price Leadership in the Low Quality Rice Export Market

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Introduction

Nearly 80 percent of market share in global rice exports are highly concentrated in top five countries (Rice yearbook, 2013). Hence, domestic agricultural and trade policies in these countries typically have profound effects on world rice prices.



There are only a few studies analyzing rice price relationships in the presence of different rice quality segments which focus on the low quality market and none have explored the role of India in this market.

Objective

To examine the dynamics of rice export prices and price leadership in the low quality rice sector (i.e., 25% broken rice). We particularly focus on four key exporting countries in the low quality market segment: Thailand, India, Vietnam, and Pakistan.

Data

The data set includes 902 weekly observations of free on board (FOB) prices from January 1996 to April 2013

Our study divides the data into 4 major cases: (1)Full Sample (Jan. 5, 1996 to April 12, 2013), (2) Subsample 1 (Jan. 5, 1996 to Oct. 5, 2007), (3)Subsample 2 (Oct. 12, 2007 to Sept. 9, 2011), (4)Subsample 3(Sept. 16, 2011 to April 12, 2013).

We investigate dynamic price relationships in three specific regimes – from 1996 to 2007 prior to India's export ban, from late 2007 to 2011 with India out of the market due to the export ban, and from 2011 to 2013 when India re-entered the market and Thailand started their domestic mortgage scheme.

Estimation strategies

1. Unit root tests

- Clemente-Moñés-Reyes tests with structural
- Augmented Dickey-Fuller (ADF) test without breaks



2. Johansen Procedure

Using trace and maximum eigenvalue tests to investigate cointegrating price relationships in both bivariate and multivariate analyses.



3. Vector Autoregressive Model (VAR)

$$p_{t} = \eta + A_{1}p_{t-1} + A_{2}p_{t-2} + ... + A_{k}p_{t-k} + \varepsilon_{t}$$

where \mathbf{p}_{t} is a $n \times I$ vector of price variables, η is a *n x 1* vector of constant, and $A_1, ..., A_n$ are $n \times n$ matrices of parameters. ξ_{k} is a $n \times 1$ vector of disturbances.

Vector Error Correction Model (VEC)

$$\begin{split} \Delta p_{t} &= \eta + \prod p_{t-l} + \sum_{i=l}^{k-l} \Gamma_{i} \Delta p_{t-i} + \epsilon_{t} \\ \text{where} &\quad \prod = \sum_{j=l}^{k} \!\! A_{j} \! - \!\! I \;, \quad \Gamma_{i} = \!\! - \!\! \sum_{j=i+l}^{k} \!\! A_{j} \end{split}$$

Note: $\prod = \alpha \beta'$ in which the adjustment coefficients (α) and the cointegration vector (β) .



4. Test Hypotheses of price leadership We test for **long-run exclusion** (i.e. $H_0: \beta_{ii} = 0$) and weak exogeneity (i.e. $H_0: \alpha_{ii} = 0$). Goulven (1999) defined price leadership status in accordance to hypothesis tests on the combination of parameter restrictions as follow:

(i) $\beta \neq 0, \alpha = 0$: price leader

(ii) $\beta \neq 0, \alpha \neq 0$: price follower

(iii) $\beta = 0, \alpha = 0$: segmented market

(iv) $\beta = 0, \alpha \neq 0$: regulator market

Results

I. Bivariate Cointegration

Variables	β	β̂ ₂	ôţ	ô ₂	Constant			
Full sample								
TH-VT	1	-1.20***	-0.036***	0.036***	1.014			
TH-PAK	1	-1.214***	-0.034***	0.017**	1.071			
VT-PAK	1	-1.038***	-0.054**	0.035***	0.195			
Sub-sample 1								
TH-VT	1	-1.089***	-0.046***	0.039***	0.425			
TH-PAK	1	-1.093***	-0.032***	0.024**	0.435			
TH-INDIA	1	-0.847***	-0.0003	0.022***	-0.821			
VT-PAK	1	-0.973***	-0.030***	0.018**	-0.146			
VT-INDIA	1	-0.679***	-0.006	0.019***	-1.662			
PAK-INDIA	1	-0.810***	-0.004	0.023***	-0.959			
Sub-sample 3								
PAK-INDIA	1	6.257***	-0.0005	-0.163***	-43.147			
NT - which this is 1 - 1 - 1 - 1 - 1 - 1 - 1 - 1 - 1 - 1								

Note: ***, **, * denotes significance at 1%, 5% and 10% level

he 25% broken rice prices of Thailand, Vietnam and stan are interdependent in the bivariate analysis. ote that there is no cointegrating relationships for price airs included India in the Full sample, even India is the

II. Multivariate Cointegration

In the multivariate analysis, we find that Indian low uality rice prices are cointegrated with low quality rice es of Thailand, Vietnam, and Pakis

III. Market price leadership Analysis

HYPOTHESES	THAI	VIETNAM	PAKISTAN	INDIA
Long-run exclusion	2.669	31.24***	11.42***	2.742*
Weak exogeneity	2.51	2.88*	0.73	0.25
Price leader status	segmente d	follower	leader	leader
Long-run exclusion	34.5***	54.52***	22.9***	6.736*
Weak exogeneity	11.51***	7.14*	5.88	1.6
Price leader status	follower	follower	leader	leader
Long-run exclusion	16.72***	46.5***	29.01***	5.352*
Weak exogeneity	2.80	7.54**	1.81	0.23
Price leader status	leader	follower	leader	leader
	Long-run exclusion Weak exogeneity Price leader status Long-run exclusion Weak exogeneity Price leader status Long-run exclusion Weak exogeneity Price leader	Long-run 2.669 exclusion Weak 2.51 exogeneity Price leader status Long-run 34.5"'' exclusion Weak 11.51"'' exclusion Wrice leader follower status Long-run 16.72"'' exclusion Weak 2.80 exogeneity Price leader leader	Long-run 2.669 31.24*** exclusion Weak 2.51 2.88* exogeneity Price leader status Long-run 34.5*** exclusion Weak 11.51*** 7.14* excogeneity Price leader follower status Long-run 16.72*** 46.5*** exclusion Weak 2.80 7.54** exogeneity Price leader leader follower	Long-run 2.669 31.24" 11.42" exclusion Weak 2.51 2.88' 0.73 exogeneity Price leader status Long-run 34.5" 54.52" 22.9" exclusion Weak 11.51" 7.14' 5.88 exogeneity Price leader follower follower leader status Long-run 16.72" 46.5" 29.01" exclusion Weak 2.80 7.54" 1.81 exogeneity Price leader leader follower leader

ietnam is consistently a price follower in the low quality ce market, while Pakistan and India are the leaders in

Main Conclusions

rovide the following insights

- India has been a price leader in the market for 25% broken rice for over 17 years.
- 1 Thailand is likely a segmented market in the 1996 to
- Thailand has emerged as a new price leader together with Pakistan and India since late 2011.

