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Imperfect Mobility and Heterogeneous Labour in CGE Modelling

Dorothee Flaig Universität Hohenheim

Presentation delivered at the 2013 Annual Meeting of the International Agricultural Trade Research Consortium (IATRC) Clearwater Beach, FL, December 15-17, 2013

THREE ESSAYS ON BLACK SEA GRAIN MARKETS

Kateryna G. Schroeder

IATRC 2013 Annual Meeting

December 16, 2013







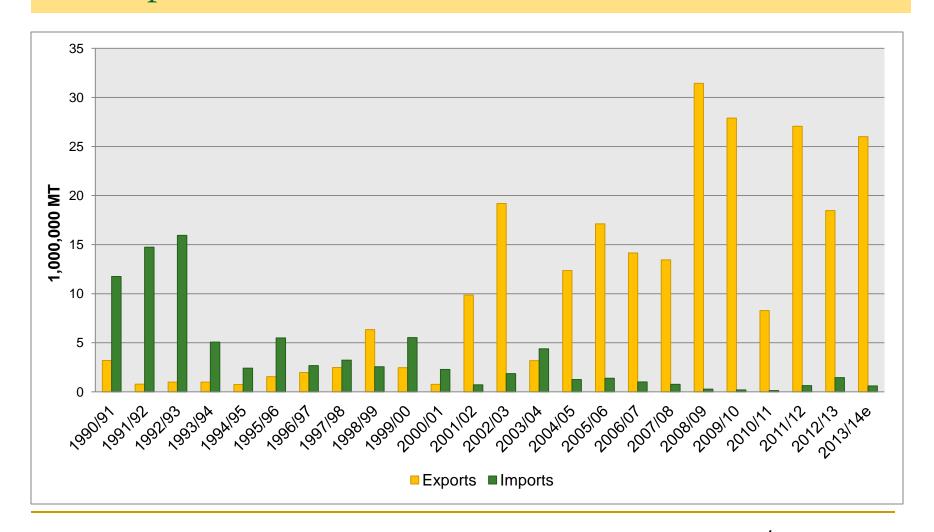
Black Sea Grain Market

In 2012/13, Ukraine and Russian were among top ten exporters of: Wheat (18.5 mln. tons) Corn (14.6 mln. tons) Barley (4.6 mln tons)





Dynamics of the Russian and Ukrainian wheat exports and imports



Source: USDA, 11/2013

Decision	Period	Wheat	Barley	Com	
10/11/2006	10/17/2006-12/31/2006	400	600	600	
12/08/2006	12/14/2006-06/30/2007	3	600	500	
02/13/2007	02/15/2007-06/30/2007	3	606	30	
02/22/2007	02/26/2007-06/07/2007	3	Quotas cancelled	Quotas cancelled	
05/22/2007	05/22/2007	Quotas cancelled	-	-	
06/20/2007	07/01/2007-10/31/2007	3	3	3	
09/26/2007	01/01/2008-03/31/2008	200	400	600	
03/28/2008	04/01/2008 - 04/30/2008	200	400	Automatic licensing	
04/23/2008	04/2008-07/01/2008	1,200	900	Automatic licensing	
05/21/2008	05/21/2008	Quotas and licenses are cancelled			
10/06/2010	10/20/2010-12/2010	500	200	2,000	
12/08/2010	12/2010 - 02/2011	1,000	200	3,000	
03/30/2011	04/04/2011 - 07/01/2011	1,000	200	5,000	
05/2011	05/2011		Quotas are cance	lled	
05/2011	05/2011-01/2012		Tariffs are introdu	ıced	
10/2011	10/2011	Tariffs can	celled, except for ba	arley (01/01/2012)	
10/10/2011	07/01/2011 - 06/30/2012	Ministry of Agricultural Policy and Food of Ukraine signed a Memorandum of Understanding with the grain exporters; amount of allowed exports established at 24.76 mmt of grain			
07/2012	07/01/2012 – 06/30/2013	Memorandum of Understanding was extended for 2012-13 marketing year, amount of allowed exports established at 24.76 mmt of grain			
11/2012	11/2012 - 06/30/2013	Export amounts under the Memorandum of Understanding were increased by 1.5 mln. tons (for corn only)			
06/2013	07/01/2013 - 06/30/2014	Memorandum of	f Understanding was marketing yea	s extended for 2013-14 ar	

Source: UkrAgroConsult (2013); Kobuta et al. (2012); FAO-EBRD (2010).

Price transmission

- Price transmission (or price cointegration) refers to the co-movement shown by prices of the same good in different location;
- Such models are used to assess competitive market equilibrium or market efficiency, a price based indicator that holds on the conditions of spatial equilibrium;
- Some benefits:
 - Use of price data only;
 - Allows for short and long run dynamics' analysis;
 - Allows for relaxing assumptions of linearity and symmetric adjustment.

Limitations

- Frequent reliance on the price data only;
- Transaction and transportation costs are assumed to be equal to zero or set as a fixed proportion of the prices used;
- Price vs. market integration;
- (!) Price transmission parameters summarize overall effects that might affect prices in different markets.
 Further research is needed to study separate role of each factor.

STRUCTURE OF THE DISSERTATION

- Essay 1a: Black Sea and World Wheat Market Price Integration Analysis
- Essay 1b: Short- and Long-Run Relationships Between Ukrainian Barley and World Feed Grain Export Prices
- Essay 2: Analysis of the Asymmetric Price Transmission in the Ukrainian Wheat Supply Chain
- Essay 3: Export Restrictions and Price Volatility in the Ukrainian Wheat Market: Evidence From a Dynamic Conditional Correlation GARCH Model

METHODS USED

Cointegration tests

Engle and Granger (1987) method

Step 1: Estimate the following regression model in

levels: $P_t^{RUS} = \beta_0 + \beta_1 P_t^{US} + \varepsilon_t$

Step 2: Test residuals ε_t for stationarity

Johansen Maximum Likelihood (ML) method

Goal: Estimate the rank of Π in the following model:

$$\Delta P_{t} = \Pi_{0} + \Pi P_{t-1} + \sum_{i=1}^{k-1} \Gamma_{i} \Delta P_{t-i} + \theta_{t}$$

TAR and M-TAR models

 Asymmetric price transmission - implies that the adjustment towards the equilibrium is of different magnitude regardless of the direction of the change.

$$\Delta \bar{\varepsilon_t} = \gamma_1 \, \bar{\varepsilon}_{t-1} + \sum_{i=1}^p \gamma_{i+1} \Delta \, \bar{\varepsilon}_{t-i} + \omega_t$$
, where
$$\Delta \bar{\varepsilon}_t = I_t \gamma_1 \, \bar{\varepsilon}_{t-1} + (1 - I_t) \gamma_2 \, \bar{\varepsilon}_{t-1} + \varphi_t$$

$$I_{t} = \begin{cases} 1 & \text{if } \varepsilon_{t-1} \ge 0 \\ 0 & \text{if } \bar{\varepsilon}_{t-1} < 0 \end{cases}$$

or

$$I_{t} = \begin{cases} 1 & \text{if } \Delta \bar{\varepsilon}_{t-1} \geq 0 \\ 0 & \text{if } \Delta \bar{\varepsilon}_{t-1} < 0 \end{cases}$$

TAR

Error-Correction Model (short-run dynamics)

 Error Correction Models (ECMs) estimate the speed at which a dependent variable returns to equilibrium after a change in an independent variable

$$\Delta P_{t}^{FL} = a_{0} + a_{1}\bar{\varepsilon}_{t-1} + \sum_{i=1}^{p} \delta_{i} \Delta P_{t-i}^{FL} + \sum_{j=1}^{n} \theta_{j} \Delta P_{t-j}^{FW} + \mu_{t}$$

Threshold Error Correction Models (TECMs)

$$\Delta P_{t}^{FL} = a_{0} + \rho_{1} I_{t} \bar{\varepsilon}_{t-1} + \rho_{2} (1 - I_{t}) \bar{\varepsilon}_{t-1} + \sum_{i=1}^{p} \delta_{i} \Delta P_{t-i}^{FL} + \sum_{j=1}^{n} \theta_{j} \Delta P_{t-j}^{FW} + \mu_{t}$$

Bai and Perron (1998, 2003) Structural Break Test

Allows for finding multiple breaks at unknown times

Uses dynamic programming algorithm to identify optimal number of breaks

 Decision is based on finding min Residual Sum of Squares (RSS) and Bayesian Information Criteria (BIC)

RESEARCH GOALS: Essays 1a & 1b

Horizontal Price Transmission

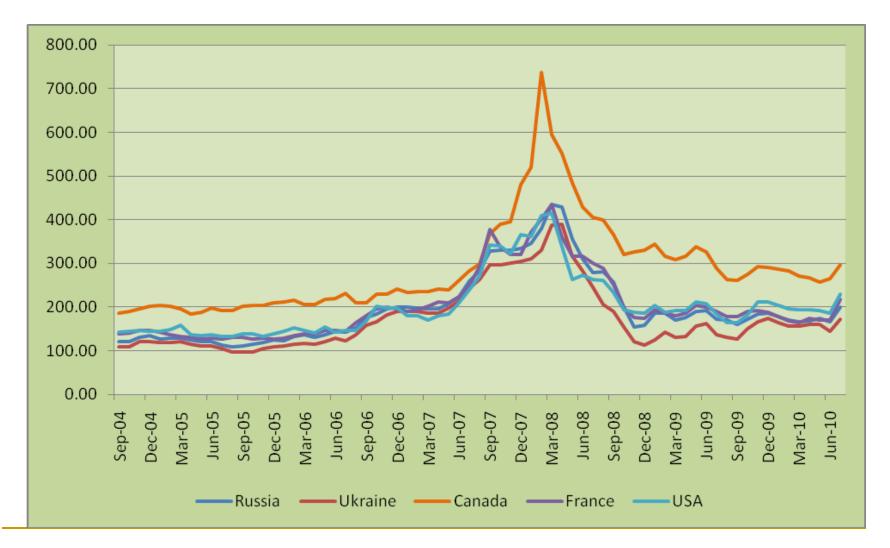
- To check whether Black sea grain markets are cointegrated with the world grain markets
 - Essay 1a studies price dynamics between Russia and Ukraine and other major wheat exporters U.S., EU, and Canada
 - Essay 1b investigates price transmission between Ukrainian barley and Australian barley, EU barley, Canadian barley, and U.S. corn
- To test for the presence of the structural change in the identified long-run price relationships
- To analyze if price transmission is symmetric (for the pairs of series that are cointegrated)
- To investigate the short run dynamics between cointegrated series

Essay 1a: Black Sea and World Wheat Market Price Integration Analysis





Comparison of the analyzed wheat price series, \$ per ton



Source: IGC, 2011



Cointegration tests' results – pairwise for Russia

	Engel and Granger procedure						
Pairs of series	# of lags ADF		ADF PP		Johansen method		
				Ho(H1)	Trace	5%CV	
Danie France	2	<i>5 20**</i>	5 0 4 * *	R=0((r>0)	25.98**	19.99	
Russia-France	2	-5.32**	-5.24**	R=1(r>1)	6.69	9.13	
Russia-Canada	1	2.20	-2.38	R=0((r>0)	13.23	19.99	
Russia-Callada	1	-2.30	-2.30 -2.38	R=1(r>1)	5.12	9.13	
Russia-USA	1	-3.79**	2 01**	r=0((r>0)	15.47	19.99	
Kussia-USA	1	-3./9***	-3.81**	R=1(r>1)	3.75	9.13	



Cointegration tests' results – pairwise for Ukraine

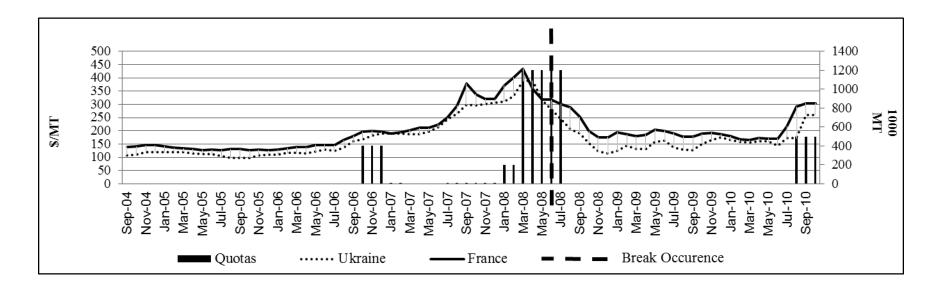
	Engel and Granger procedure					
Pairs of series	# of lags	ADF	PP	Johansen method		hod
				Ho(H1)	Trace	5%CV
III . L	2	2.22	2 (14	R=0((r>0)	24.66**	19.99
Ukraine-France	3	-2.33	-3.64*	R=1(r>1)	5.28	9.13
Illradina Canada	1	1.00	1.00	R=0((r>0)	12.48	19.99
Ukraine - Canada	1	-1.90	-1.99	R=1(r>1)	4.70	9.13
Illradia o IIC A	1154 2 201 2 201*		r=0((r>0)	12.48	19.99	
Ukraine-USA	2	-2.91	-3.24*	R=1(r>1)	4.56	9.13



Long-run dynamics

- The long-run elasticities are equal to
 - □ 1.04 (case of Russia-France)
 - 1.16 (case of Russia-USA)
 - □ 1.05 (case of Ukraine-France)

The story of a structural break



	# of lags	ADF	PP	KPSS
Pre-break period	4	-2.61	-3.14*	0.17
Post-break period	2	-1.75	-2.74	0.35*

Source: IGC, 2011

Short-run Dynamics

Dependent variable	Independent variable	# of lags	Speed of adjustment, α ₁	LM test	F _{test}
Russia	France	2;2	-0.18 (-0.82)	1.87 [0.76]	12.78**
France	Russia	1;1	-0.59 (-2.82)**	1.73 [0.78]	10.08**
Russia	USA	1;1	-0.09 (-0.98)	3.28 [0.51]	14.78**
USA	Russia	1;1	-0.30 (-2.11)**	1.66 [0.79]	4.03**
Ukraine	France	2;2	-0.20 (-1.99)*	0.92 [0.63]	11.38**
France	Ukraine	1;1	-0.25(-2.63)**	2.76 [0.25]	9.13**

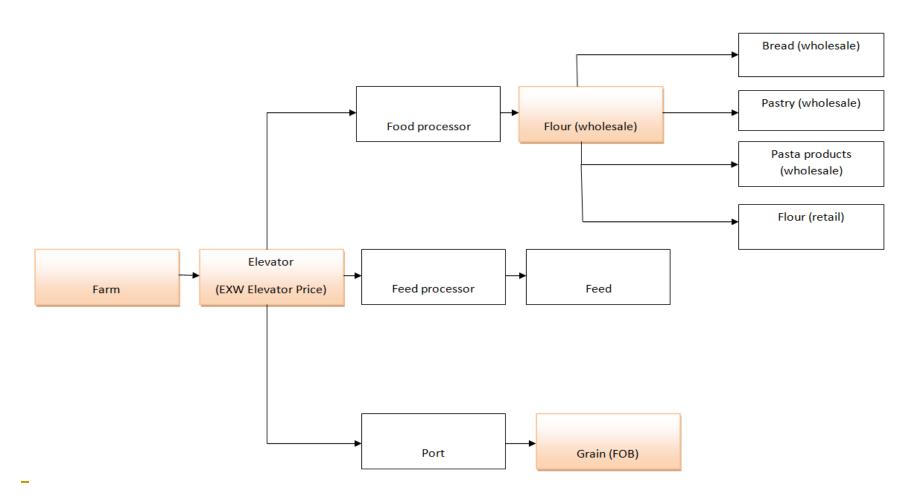
Essay 2: Analysis of Asymmetric Price Transmission in the Ukrainian Wheat Supply Chain



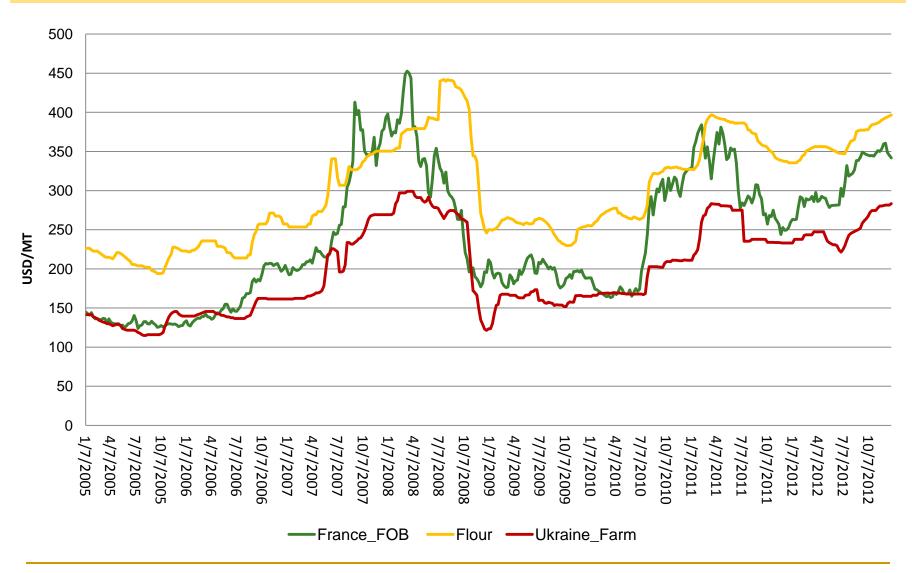
Objective:

 To investigate short- and long-run vertical price transmission behavior in the Ukrainian wheat supply chain

Simplified Ukrainian wheat supply chain (without the retail level, except for flour)



Data



Source: APK-Inform

Long-run Dynamics

$$P_{t}^{FL} = \beta_{0} + \beta_{1}P_{t}^{FW} + \epsilon_{t}$$

Or Dynamic OLS

$$P_t^{FL} = \ \theta_0 + \theta_1 P_t^{FW} + \sum_{i=-m}^{+m} \Delta P_{t-i}^{FW} + \epsilon_t$$

- The long-run elasticities are equal to
 - 0.69 (case of Farm-France)
 - 0.74 (case of Flour-Farm)

Structural breaks

	Break date	Confidence interval	BIC	RSS
Break 1	8/17/2007	07/20/07 - 08/24/07	2.03	-947.44
Break 2	10/24/08	10/17/08 - 10/31/08		
Break 3	01/01/10	12/18/09 - 3/05/11		
Break 4	03/18/11	3/11/11 – 4/1/11		

Period 1 (January 2005 till August 2007)

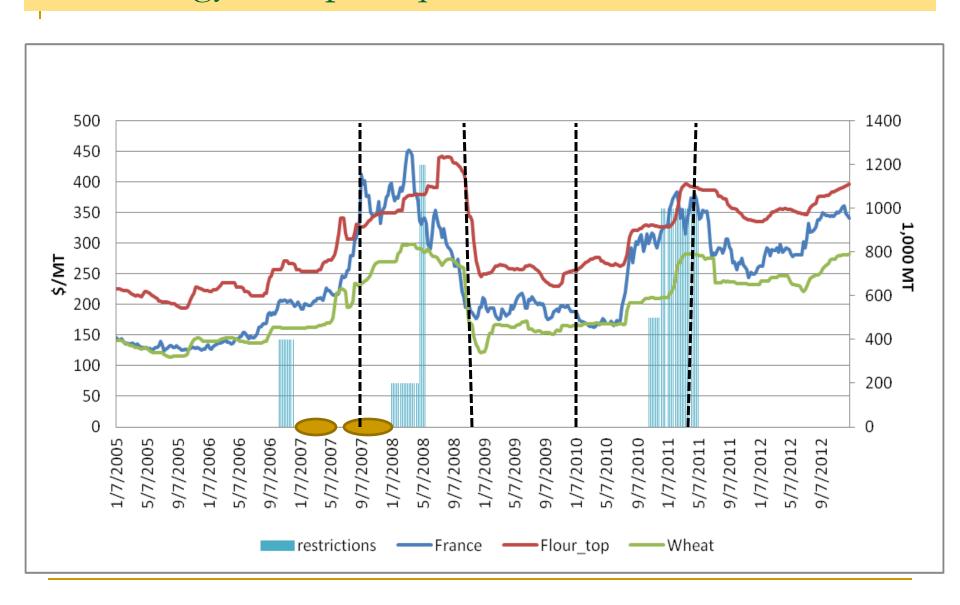
Period 2 (September 2007 – October 2008)

Period 3 (November 2008 – January 2010)

Period 4 (February 2010 – March 2011)

Period 5 (April 2011 – December 2012)

Chronology of export quotas and structural breaks



Dynamic OLS results for different regime dummies

Variable	Coefficient
Const	2.17 (0.35)***
Trend	0.0005(0.00)***
P _t FR	0.55 (0.06)***
Regime 1	0.01 (0.07)
Regime 2	-0.34 (0.08)***
Regime 3	-0.33 (0.17)*
Regime 4	-0.14 (0.07)**
Const_r1	-0.08 (0.41)
Const_r2	2.08 (0.48)***
Const_r3	1.6 (0.92)*
Const_r4	0.66 (0.38)*
ΔP_{t-1}^{FR}	-0.43(0.09)***
ΔP_{t-2}^{FR}	-0.31 (0.09)***
ΔP_{t+1}^{FR}	-0.04 (0.09)
ΔP_{t+2}^{FR}	-0.02 (0.08)
AIC	-1033.76

[&]quot;The Newey-West heteroskedasticity and autocorrelation robust standard errors are reported in parenthesis. Asterisks denote levels of significance (* for 10 percent, ** for 5 percent, *** for 1 percent).

Testing for asymmetric price transmission: M-TAR model

	France-Wheat	Wheat-Flour
Variable	Parameter estimate	Parameter estimate
γ_1	-0.06 (-2.01)**	-0.03 (-0.99)
γ_2	-0.09(-3.44)**	-0.10 (-4.33)**
$H_0: \gamma_1 = \gamma_2 = O(\Phi)$	7.72**	9.75**
$H_0: \gamma_1 = \gamma_2(F)$	0.81[0.37]	4.61 [0.03]**
τ	0	0

Short-run Dynamics

Dependent variable	Independent variable	# of lags	Speed of adjustment, $ ho_1$	Speed of adjustment, $ ho_1$	Speed of adjustment, $ ho_2$
FrenchFOB	Wheat	1;1	-0.0240 (-1.33)	-	-
Wheat	FrenchFOB	2;2	-0.0638 (-5.70)**	-	-
Flour	Wheat	1;1		-0.02 (-1.23)	-0.11(-6.08)**

Essay 3: Export Restrictions and Price Volatility in the Ukrainian Wheat Market: Evidence from a Dynamic Conditional Correlation GARCH Model



Objective

to investigate the development of Ukrainian wheat farm price volatility during the export interventions compared to open trade

to study interdependence of the world (French) and Ukrainian domestic wheat price volatility in the analyzed period

GARCH-DCC

- DCC is a type of multivariate GARCH model.
 See Engle(2002).
- It approximates a dynamics conditional correlation matrix and permits to evaluate the level of interdependence between markets across time

GARCH-DCC (cont.)

$$\boldsymbol{H}_{t} = \boldsymbol{D}_{t} \boldsymbol{R}_{t} \boldsymbol{D}_{t}$$

- H_t is the conditional covariance matrix of the price returns' vector
- $D_t = diag\{\sqrt{h_{jj}}\}, j = 1,...,J$, is a 3x3 matrix of the standardized disturbance variances from the univariate GARCH models
- $R_t = (diag(Q_t))^{-1/2}Q_t(diag(Q_t))^{-1/2} a$ symmetric dynamic correlations matrix of standardized residuals, where

$$Q_{t} = \{\rho_{ij,t}\} = (1 - \alpha - \beta)\overline{Q} + \beta Q_{t-1} + \alpha(u_{t-1}u'_{t-1})$$

The primary focus of the GARCH-DCC model is on obtaining conditional correlations in R_t:

$$r_{ij,t} = \rho_{ij,t} / \sqrt{\rho_{ii,t}} \sqrt{\rho_{jj,t}}$$



Data

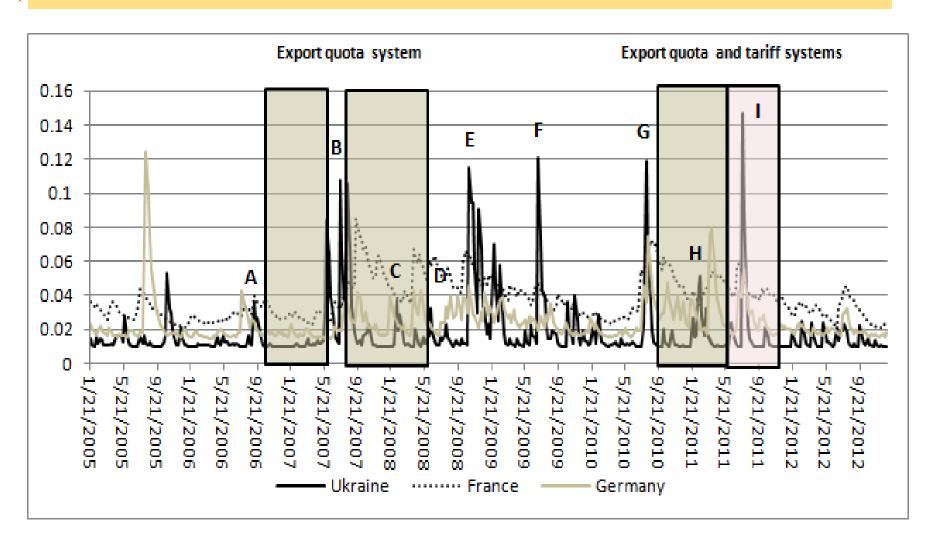
- Weekly prices for:
 - Ukrainian ex warehouse prices of milling wheat of Class III
 - German warehouse delivery price of bread wheat
 - French FOB soft wheat price (Rouen)

- Time span: from January 2005 till December 2012
- Sources: APK-Inform, AMI, HGCA

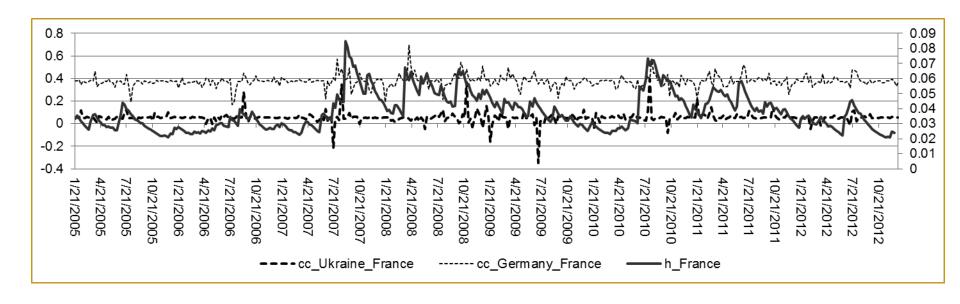
Univariate GARCH models' results

	France	Ukraine	Germany			
Volatility equations						
α_0	-0.00 (0.95)	-0.00 (0.75)	0.00 (0.91)			
α_1	0.14 (0.02)**	0.74 (0.00)***	0.37 (0.01)**			
β_1	0.83 (0.00)***	0.25 (0.00)***	0.56 (0.00)***			
$\alpha_1 + \beta_1$	0.97	0.99	0.93			
Log	809	1270	1041			
Likelihood						

Conditional variances



Dynamic correlations



	Ukraine-France	Germany-France
$r_{ij,t}$ min	-0.34	0.17
$r_{ij,t}max$	0.39	0.69
$r_{ij,t}$ mean	0.05	0.38
χ^2 -test: $R_t = R$	305.5***	442.9***

Volatility Transmission

	Nakatani & Teräsvirta test		Hafner & Herwartz test	
	π	p-value	τ	p-value
France- Ukraine	6.04	0.19	8.67	0.07*
France- Germany	9.15	0.06*	9.13	0.06*



THANK YOU!

