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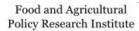
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Presentation delivered at the 2013 Annual Meeting of the International Agricultural Trade Research Consortium (IATRC) Clearwater Beach, FL, December 15-17, 2013

### THREE ESSAYS ON BLACK SEA GRAIN MARKETS

#### Kateryna G. Schroeder

### IATRC 2013 Annual Meeting December 16, 2013





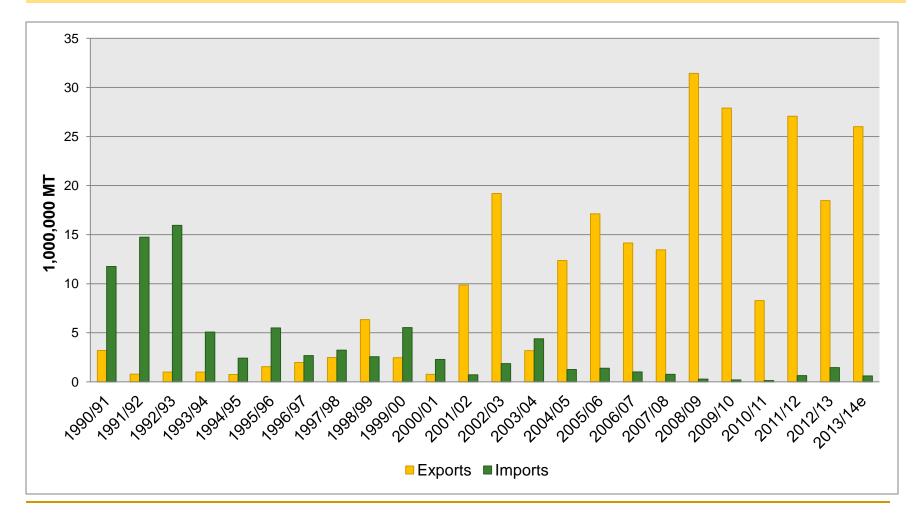


### Black Sea Grain Market

In 2012/13, Ukraine and Russian were among top ten exporters of: Wheat (18.5 mln. tons) Corn (14.6 mln. tons) Barley (4.6 mln tons)



# Dynamics of the Russian and Ukrainian wheat exports and imports



**Source:** USDA, 11/2013

Decision	Period	Wheat	Deuteu	Com	
Decision	Period	wneat	Barley	Com	
10/11/2006	10/17/2006-12/31/2006	400	600	600	
12/08/2006	12/14/2006-06/30/2007	3	600	500	
02/13/2007	02/15/2007-06/30/2007	3	606	30	
02/22/2007	02/26/2007-06/07/2007	3	Quotas cancelled	Quotas cancelled	
05/22/2007	05/22/2007	Quotas cancelled	-	-	
06/20/2007	07/01/2007-10/31/2007	3	3	3	
09/26/2007	01/01/2008-03/31/2008	200	400	600	
03/28/2008	04/01/2008 - 04/30/2008	200	400	Automatic licensing	
04/23/2008	04/2008-07/01/2008	1,200	900	Automatic licensing	
05/21/2008	05/21/2008	Quotas and licenses are cancelled			
10/06/2010	10/20/2010-12/2010	500	200	2,000	
12/08/2010	12/2010 - 02/2011	1,000	200	3,000	
03/30/2011	04/04/2011 - 07/01/2011	1,000	200	5,000	
05/2011	05/2011		Quotas are cance	lled	
05/2011	05/2011-01/2012		Tariffs are introdu	ıced	
10/2011	10/2011	Tariffs can	celled, except for ba	arley (01/01/2012)	
10/10/2011	07/01/2011-06/30/2012	Ministry of Agricultural Policy and Food of Ukraine signed a Memorandum of Understanding with the grain exporters; amount of allowed exports established at 24.76 mmt of grain			
07/2012	07/01/2012 - 06/30/2013	Memorandum of Understanding was extended for 2012-13 marketing year, amount of allowed exports established at 24.76 mmt of grain			
11/2012	11/2012 - 06/30/2013	Export amounts under the Memorandum of Understanding were increased by 1.5 mln. tons (for corn only)			
06/2013	07/01/2013 - 06/30/2014	Memorandum of	f Understanding was marketing yea	s extended for 2013-14 ar	

## Price transmission

- Price transmission (or price cointegration) refers to the co-movement shown by prices of the same good in different location;
- Such models are used to assess competitive market equilibrium or market efficiency, a price based indicator that holds on the conditions of spatial equilibrium;

Some benefits:

- Use of price data only;
- Allows for short and long run dynamics' analysis;
- Allows for relaxing assumptions of linearity and symmetric adjustment.

### Limitations

Frequent reliance on the price data only;

- Transaction and transportation costs are assumed to be equal to zero or set as a fixed proportion of the prices used;
- Price vs. market integration;
- (!) Price transmission parameters summarize overall effects that might affect prices in different markets. Further research is needed to study separate role of each factor.

## STRUCTURE OF THE DISSERTATION

- Essay 1a: Black Sea and World Wheat Market Price Integration Analysis
- Essay 1b: Short- and Long-Run Relationships Between Ukrainian Barley and World Feed Grain Export Prices
- Essay 2: Analysis of the Asymmetric Price Transmission in the Ukrainian Wheat Supply Chain
- Essay 3: Export Restrictions and Price Volatility in the Ukrainian Wheat Market: Evidence From a Dynamic Conditional Correlation GARCH Model

### **METHODS USED**

## Cointegration tests

• Engle and Granger (1987) method <u>Step 1</u>: Estimate the following regression model in levels:  $P_t^{RUS} = \beta_0 + \beta_1 P_t^{US} + \varepsilon_t$ 

#### **<u>Step 2</u>**: Test residuals $\varepsilon_t$ for stationarity

Johansen Maximum Likelihood (ML) method Goal: Estimate the rank of Π in the following model:

$$\Delta P_t = \Pi_0 + \Pi P_{t-1} + \sum_{i=1}^{k-1} \Gamma_i \Delta P_{t-i} + \theta_t$$

#### TAR and M-TAR models

 <u>Asymmetric price transmission</u> - implies that the adjustment towards the equilibrium is of different magnitude regardless of the direction of the change.

$$\Delta \overline{\varepsilon}_{t} = \gamma_{1} \overline{\varepsilon}_{t-1} + \sum_{i=1}^{p} \gamma_{i+1} \Delta \overline{\varepsilon}_{t-i} + \omega_{t}$$
, where  

$$\Delta \overline{\varepsilon}_{t} = I_{t} \gamma_{1} \overline{\varepsilon}_{t-1} + (1 - I_{t}) \gamma_{2} \overline{\varepsilon}_{t-1} + \varphi_{t}$$

$$I_{t} = \begin{cases} 1 \text{ if } \overline{\varepsilon}_{t-1} \ge 0\\ 0 \text{ if } \overline{\varepsilon}_{t-1} < 0 \end{cases} \quad \text{or} \qquad I_{t} = \begin{cases} 1 \text{ if } \Delta \overline{\varepsilon}_{t-1} \ge 0\\ 0 \text{ if } \Delta \overline{\varepsilon}_{t-1} < 0 \end{cases}$$

$$\text{TAR}$$

$$M\text{-TAR}$$

#### Error-Correction Model (short-run dynamics)

• <u>Error Correction Models (ECMs)</u> estimate the speed at which a dependent variable returns to equilibrium after a change in an independent variable  $\Delta P_t^{FL} = a_0 + a_1 \bar{\varepsilon}_{t-1} + \sum_{i=1}^p \delta_i \Delta P_{t-i}^{FL} + \sum_{i=1}^n \theta_j \Delta P_{t-j}^{FW} + \mu_t$ 

 $\frac{\text{Threshold Error Correction Models (TECMs)}}{\Delta P_t^{FL} = a_0 + \rho_1 I_t \bar{\varepsilon}_{t-1} + \rho_2 (1 - I_t) \bar{\varepsilon}_{t-1} + \sum_{i=1}^p \delta_i \Delta P_{t-i}^{FL} + \sum_{j=1}^n \theta_j \Delta P_{t-j}^{FW} + \mu_t}$ 

# Bai and Perron (1998, 2003) Structural Break Test

- Allows for finding multiple breaks at unknown times
- Uses dynamic programming algorithm to identify optimal number of breaks
- Decision is based on finding min Residual Sum of Squares (RSS) and Bayesian Information Criteria (BIC)

#### **RESEARCH GOALS:** Essays 1a & 1b

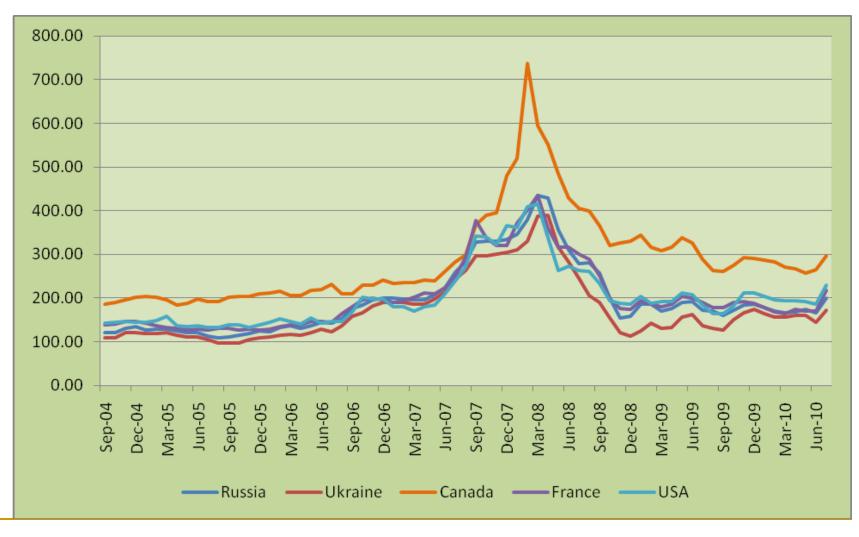
#### Horizontal Price Transmission

- To check whether Black sea grain markets are cointegrated with the world grain markets
  - Essay 1a studies price dynamics between Russia and Ukraine and other major wheat exporters – U.S., EU, and Canada
  - Essay 1b investigates price transmission between Ukrainian barley and Australian barley, EU barley, Canadian barley, and U.S. corn
- To test for the presence of the structural change in the identified long-run price relationships
- To analyze if price transmission is symmetric (for the pairs of series that are cointegrated)
- To investigate the short run dynamics between cointegrated series

### **Essay 1a: Black Sea and World Wheat** Market Price Integration Analysis



### Comparison of the analyzed wheat price series, \$ per ton



Source: IGC, 2011

#### Cointegration tests' results – pairwise for Russia

 $\sum$ 

	Engel	and Granger				
Pairs of series	# of lags	ADF	PP	Johansen method		
				Ho(H1)	Trace	5%CV
Durasia Engrada	2	5 20**	5 7 1 * *	R=0((r>0)	25.98**	19.99
Russia-France	2	-5.32**	-5.24**	R=1(r>1)	6.69	9.13
Russia-Canada	1	-2.30	-2.38	R=0((r>0)	13.23	19.99
Kussia-Callada	l	-2.50	-2.38	R=1(r>1)	5.12	9.13
Russia-USA	1	2 70**	2 91**	r=0((r>0)	15.47	19.99
KUSSIA-USA	1 -3.79**		-3.81**	R=1(r>1)	3.75	9.13

Asterisks denote levels of significance (\* for 10 percent, \*\* for 5 percent). The 5% and 10% critical values for tests with a drift are -3.42 and -3.10 respectively. Critical values were obtained from MacKinnon (1991).

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#### Cointegration tests' results – pairwise for Ukraine

 $\sum$ 

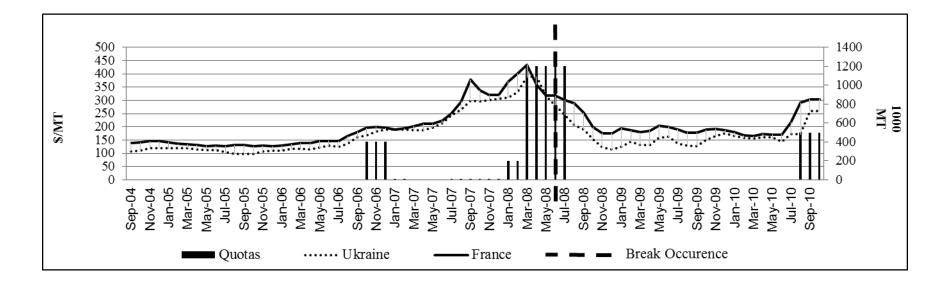
	Engel	and Granger	r procedure			
Pairs of series	# of lags			Johansen method		
			Ho(H1)	Trace	5%CV	
	2	0.22	2 (1*	R=0((r>0)	24.66**	19.99
Ukraine-France	3	-2.33	-3.64*	R=1(r>1)	5.28	9.13
Illeraina Canada	1	1.00	1.00	R=0((r>0)	12.48	19.99
Ukraine - Canada	1	-1.90	-1.99	R=1(r>1)	4.70	9.13
		2.01	2.01 2.24*		12.48	19.99
Ukraine-USA	2	-2.91	-3.24*	R=1(r>1)	4.56	9.13

Asterisks denote levels of significance (\* for 10 percent, \*\* for 5 percent). The 5% and 10% critical values for tests with a drift are -3.42 and -3.10 respectively. Critical values were obtained from MacKinnon (1991).

# Long-run dynamics

- The long-run elasticities are equal to
  - 1.04 (case of Russia-France)
  - 1.16 (case of Russia-USA)
  - □ 1.05 (case of Ukraine-France)

# The story of a structural break



	# of lags	ADF	PP	KPSS
Pre-break period	4	-2.61	-3.14*	0.17
Post-break period	2	-1.75	-2.74	0.35*

**Source:** IGC, 2011

# Short-run Dynamics

Dependent variable	Independent variable	# of lags	Speed of adjustment, α <sub>1</sub>	LM test	F <sub>test</sub>
Russia	France	2;2	-0.18 (-0.82)	1.87 [0.76]	12.78**
France	Russia	1;1	-0.59 (-2.82)**	1.73 [0.78]	10.08**
Russia	USA	1;1	-0.09 (-0.98)	3.28 [0.51]	14.78**
USA	Russia	1;1	-0.30 (-2.11)**	1.66 [0.79]	4.03**
Ukraine	France	2;2	-0.20 (-1.99)*	0.92 [0.63]	11.38**
France	Ukraine	1;1	-0.25(-2.63)**	2.76 [0.25]	9.13**

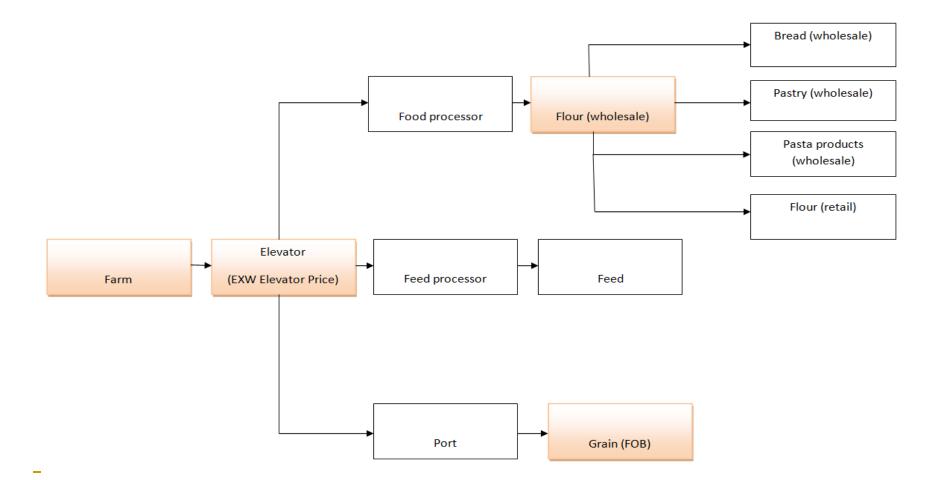
### Essay 2: Analysis of Asymmetric Price Transmission in the Ukrainian Wheat Supply Chain



### **Objective:**

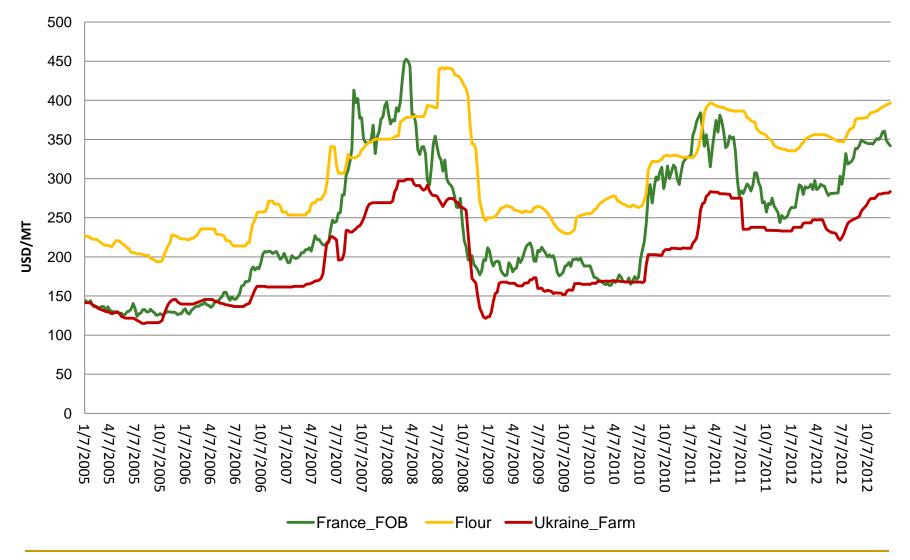
To investigate short- and long-run vertical price transmission behavior in the Ukrainian wheat supply chain

# Simplified Ukrainian wheat supply chain (without the retail level, except for flour)



Source: Canadian International Development Agency, 2007

#### Data



# Long-run Dynamics

$$P_t^{FL} = \beta_0 + \beta_1 P_t^{FW} + \varepsilon_t$$

**Or Dynamic OLS** 

$$P_{t}^{FL} = \theta_{0} + \theta_{1}P_{t}^{FW} + \sum_{i=-m}^{+m} \Delta P_{t-i}^{FW} + \epsilon_{t}$$

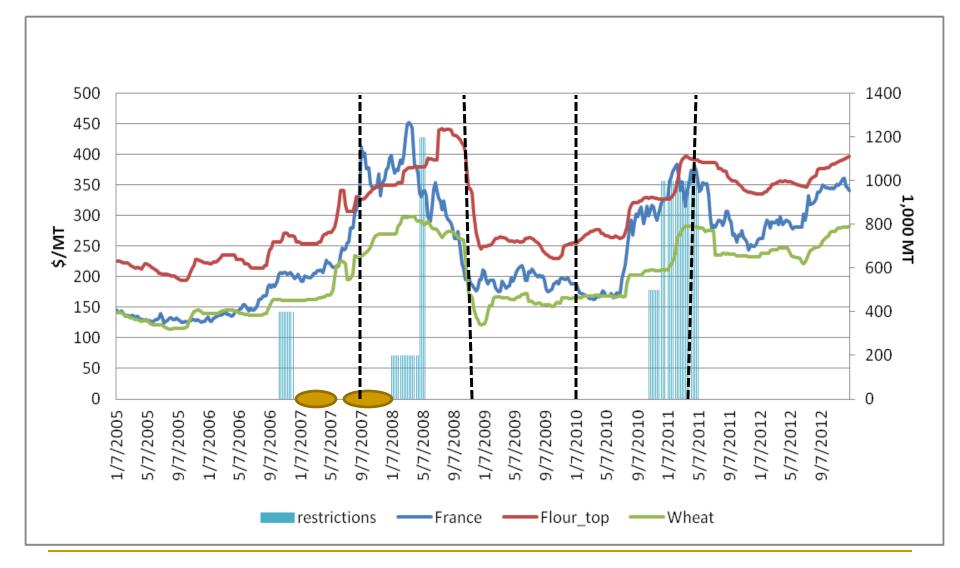
The long-run elasticities are equal to
 0.69 (case of Farm-France)
 0.74 (case of Flour-Farm)

## Structural breaks

	Break date	Confidence interval	BIC	RSS
Break 1	8/17/2007	07/20/07 - 08/24/07	2.03	-947.44
Break 2	10/24/08	10/17/08 - 10/31/08		
Break 3	01/01/10	12/18/09 - 3/05/11		
Break 4	03/18/11	3/11/11 - 4/1/11		

Period 1 (January 2005 till August 2007) Period 2 (September 2007 – October 2008) Period 3 (November 2008 – January 2010) Period 4 (February 2010 – March 2011) Period 5 (April 2011 – December 2012)

#### Chronology of export quotas and structural breaks



#### Dynamic OLS results for different regime dummies

Variable	Coefficient <sup>e</sup>
Const	2.17 (0.35)***
Trend	0.0005(0.00)***
P <sup>FR</sup>	0.55 (0.06)***
Regime 1	0.01 (0.07)
Regime 2	-0.34 (0.08)***
Regime 3	-0.33 (0.17)*
Regime 4	-0.14 (0.07)**
Const_r1	-0.08 (0.41)
Const_r2	2.08 (0.48)***
Const_r3	1.6 (0.92)*
Const_r4	0.66 (0.38)*
$\Delta P_{t-1}^{FR}$	-0.43(0.09)***
$\Delta P_{t-2}^{FR}$	-0.31 (0.09)***
$\Delta P_{t+1}^{FR}$	-0.04 (0.09)
$\Delta P_{t+2}^{FR}$	-0.02 (0.08)
AIC	-1033.76

"The Newey-West heteroskedasticity and autocorrelation robust standard errors are reported in parenthesis. Asterisks denote levels of significance (\* for 10 percent, \*\* for 5 percent, \*\*\* for 1 percent).

#### Testing for asymmetric price transmission: M-TAR model

	France-Wheat	Wheat-Flour
Variable	Parameter estimate	Parameter estimate
γ <sub>1</sub>	-0.06 (-2.01)**	-0.03 (-0.99)
$\gamma_2$	-0.09(-3.44)**	-0.03 (-0.99) -0.10 (-4.33)**
$H_0: \gamma_1 = \gamma_2 = 0(\Phi)$	7.72**	9.75**
$H_0: \gamma_1 = \gamma_2(F)$	0.81[0.37]	4.61 [0.03]**
τ	0	0

# Short-run Dynamics

Dependent variable	Independent variable	# of lags	Speed of adjustment, $\rho_1$	Speed of adjustment, $\rho_1$	Speed of adjustment, $\rho_2$
FrenchFOB	Wheat	1;1	-0.0240 (-1.33)	-	-
Wheat	FrenchFOB	2;2	-0.0638 (-5.70)**	-	-
Flour	Wheat	1;1		-0.02 (-1.23)	-0.11(-6.08)**

**Essay 3:** Export Restrictions and Price Volatility in the Ukrainian Wheat Market: Evidence from a Dynamic Conditional Correlation GARCH Model



# Objective

- 1) to investigate the development of Ukrainian wheat farm price volatility during the export interventions compared to open trade
- to study interdependence of the world (French) and Ukrainian domestic wheat price volatility in the analyzed period

# GARCH-DCC

- DCC is a type of multivariate GARCH model.
   See Engle(2002).
- It approximates a dynamics conditional correlation matrix and permits to evaluate the level of interdependence between markets across time

# GARCH-DCC (cont.)

$$H_t = D_t R_t D_t$$

- $H_t$  is the conditional covariance matrix of the price returns' vector
- $D_t = diag\{\sqrt{h_{jj}}\}, j = 1,...,J,$  is a 3x3 matrix of the standardized disturbance variances from the univariate GARCH models
- $R_t = (diag(Q_t))^{-1/2}Q_t(diag(Q_t))^{-1/2} a$  symmetric dynamic correlations matrix of standardized residuals, where

$$Q_{t} = \{\rho_{ij,t}\} = (1 - \alpha - \beta)\overline{Q} + \beta Q_{t-1} + \alpha(u_{t-1}u_{t-1})$$

• The primary focus of the GARCH-DCC model is on obtaining conditional correlations in R<sub>t</sub>:

$$r_{ij,t} = \rho_{ij,t} / \sqrt{\rho_{ii,t}} \sqrt{\rho_{jj,t}}$$

### Data

• Weekly prices for:

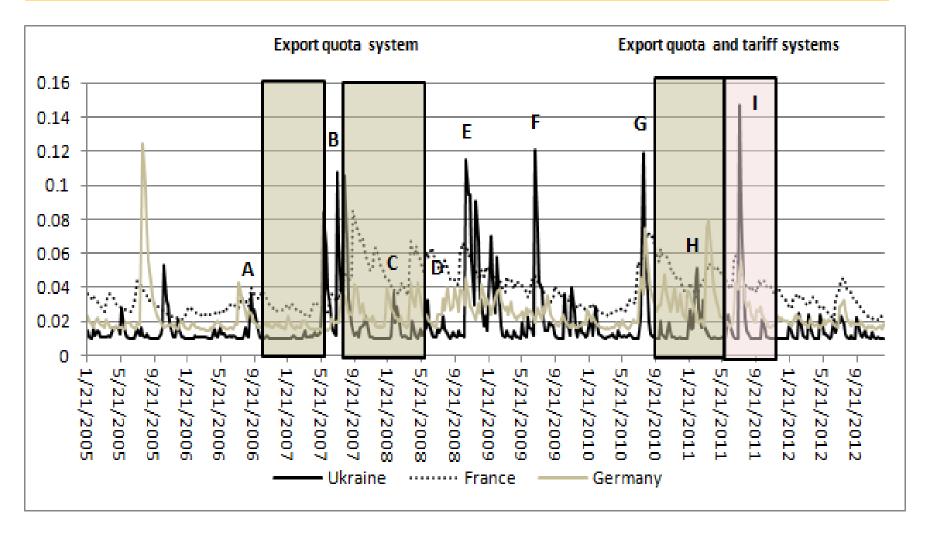
OUkrainian ex warehouse prices of milling wheat of Class III
OGerman warehouse delivery price of bread wheat
OFrench FOB soft wheat price (Rouen)

- Time span: from January 2005 till December 2012
- Sources: **APK-Inform, AMI, HGCA**

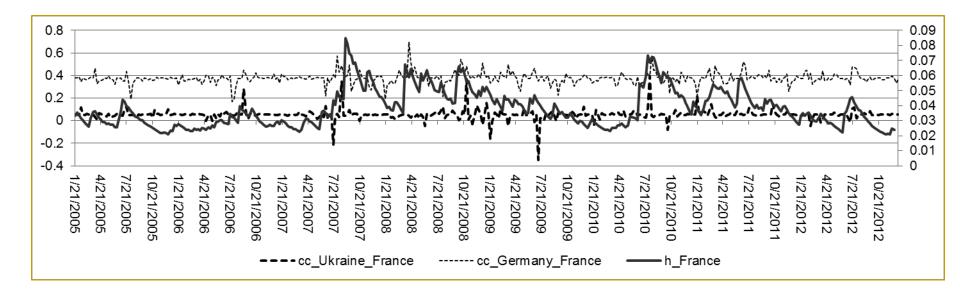
# Univariate GARCH models' results

	France	Ukraine	Germany			
Volatility equations						
$\alpha_0$	-0.00 (0.95)	-0.00 (0.75)	0.00 (0.91)			
$\alpha_1$	0.14 (0.02)**	0.74 (0.00)***	0.37 (0.01)**			
$\beta_1$	0.83 (0.00)***	0.25 (0.00)***	0.56 (0.00)***			
$\alpha_1 + \beta_1$	0.97	0.99	0.93			
Log	809	1270	1041			
Likelihood						

### Conditional variances



### Dynamic correlations



	Ukraine-France	Germany-France
$r_{ij,t}min$	-0.34	0.17
r <sub>ij,t</sub> max	0.39	0.69
r <sub>ij,t</sub> mean	0.05	0.38
χ²-test: R <sub>t</sub> = R	305.5***	442.9***

# Volatility Transmission

	Nakatani & T	eräsvirta test	Hafner & Herwartz test		
	π	p-value	τ	p-value	
France- Ukraine	6.04	0.19	8.67	0.07*	
France- Germany	9.15	0.06*	9.13	0.06*	

### THANK YOU!

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