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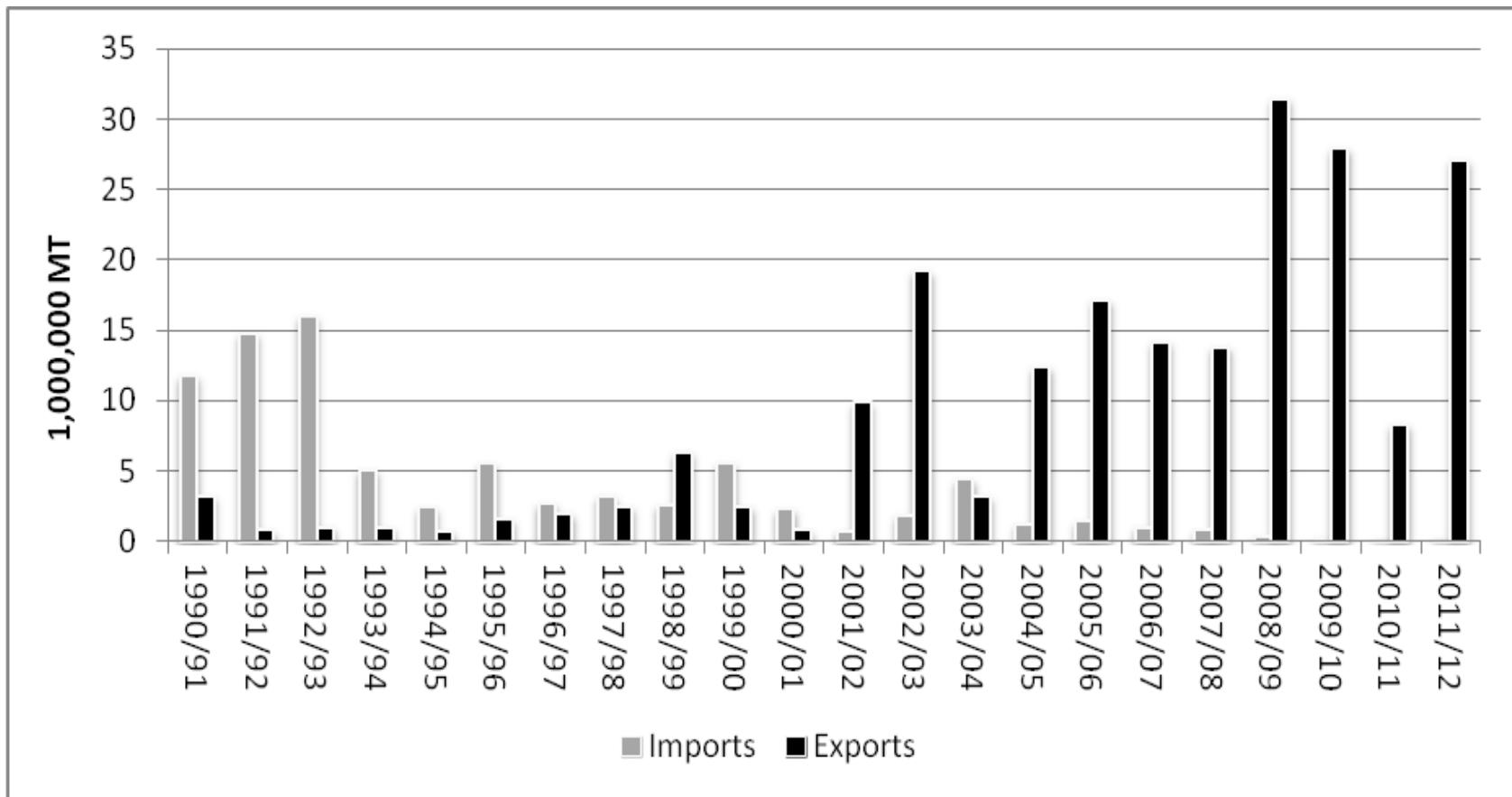
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Black Sea and World Wheat Market Price Integration Analysis



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Dynamics of the Russian and Ukrainian wheat exports and imports



Russian and Ukrainian Wheat Exports (1000 MT)

| | Country | 2008/09 | 2009/10 | 2010/11 | 2011/12 |
|---|----------------|---------|---------|---------|---------|
| 1 | United States | 27,101 | 24,172 | 35,977 | 28,071 |
| 2 | Australia | 13,450 | 13,764 | 18,477 | 23,041 |
| 3 | Russia | 18,393 | 18,556 | 3,983 | 21,627 |
| 4 | Canada | 18,674 | 18,992 | 16,768 | 17,603 |
| 5 | EU-27 | 25,351 | 22,115 | 22,906 | 16,439 |
| 6 | Argentina | 8,651 | 5,255 | 7,742 | 11,949 |
| 7 | Kazakhstan | 5,701 | 7,871 | 5,519 | 10,619 |
| 8 | Ukraine | 13,037 | 9,337 | 4,302 | 5,436 |
| 9 | Turkey | 2,342 | 4,363 | 2,945 | 3,680 |

> 50% of total wheat is exported to **North Africa** and **Near East Asia**

Objective of the study

- to investigate short- and long-run wheat price dynamics between Ukraine and Russia and other major wheat exporters - United States, European Union (EU), and Canada.

More specifically the goals are...

- To check whether Black sea wheat markets are integrated with the world grain markets
- To analyze if the price transmission is symmetric (for the pairs of series that are cointegrated)
- To investigate the short run dynamics between cointegrated series

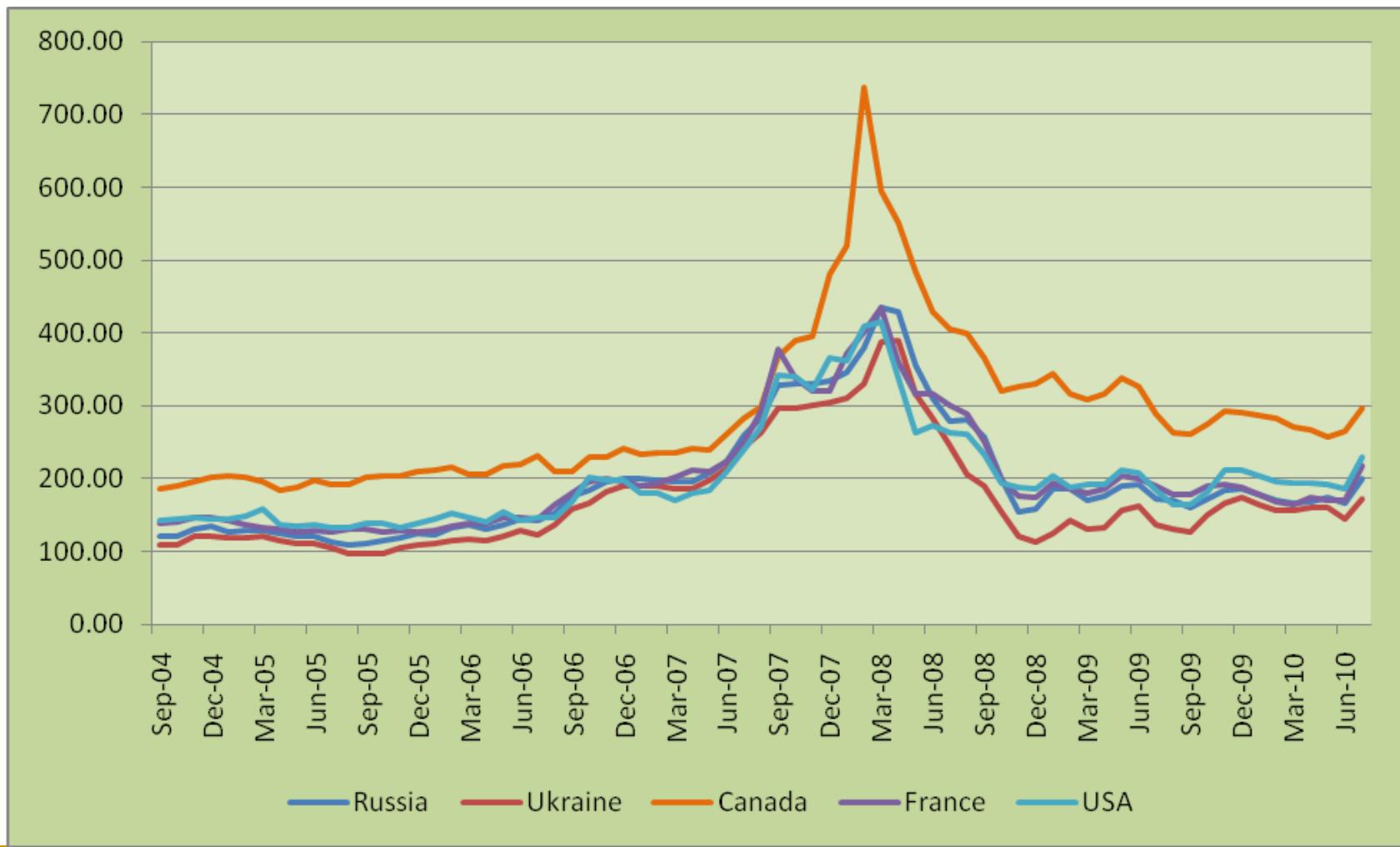
Methods used

- Testing for Unit Roots (ADF, PP, and KPSS)
 - Testing for Cointegration (long-run relationship)
 - Residual based test [primary one]
 - Johansen's Maximum Likelihood test
 - TAR, M-TAR (asymmetric price adjustment)
 - Error-Correction Model (short-run dynamics)

Data

- Monthly wheat FOB prices for:
 - **Russian Soft Wheat (Black Sea ports)**
 - **Ukrainian Feed Wheat (Black Sea ports)**
 - **Canadian Western Red Spring Wheat (St. Lawrence)**
 - **US Soft Red Winter Wheat (Gulf ports)**
 - **French Soft Wheat (Rouen)**
- Time span: **from July 2004 till October 2010**
- Prices were obtained from **the International Grain Council**

Comparison of the analyzed wheat price series, \$ per ton



Step 1: Testing data stationary

- Augmented Dickey-Fuller (ADF)
- Philips-Perron (PP) tests
- Kwiatkowski, Phillips, Schmidt, and Shin (KPSS) test

 **The results suggested that all series are I(1) stationary**

Step 2: Cointegration tests

- ❑ **Cointegration** presupposes that observable variables exhibiting non-stationary behavior will nonetheless be linked in the long-run
- ❑ Two methods:
 - ❑ Johansen Maximum Likelihood Method (both multiple and pairwise comparisons)
 - ❑ Engel and Granger residual based test (only pairwise comparisons)

Cointegration test results – Johansen ML on multiple series (trace test)

| Ho(Rank=r) | H1(Rank>r) | Trace | 5% CV |
|------------|------------|----------|-------|
| 0 | 0 | 112.08** | 75.74 |
| 1 | 1 | 63.47** | 53.42 |
| 2 | 2 | 27.87 | 34.8 |
| 3 | 3 | 12.26 | 19.99 |
| 4 | 4 | 5.89 | 9.13 |

Cointegration tests' results – pairwise for Russia

| Pairs of series | Engel and Granger procedure | | | | Johansen method | | |
|-----------------|-----------------------------|---------|---------|-----------|-----------------|-------|--|
| | # of lags | ADF | PP | | | | |
| | | | Ho(H1) | Trace | 5%CV | | |
| Russia-France | 2 | -5.32** | -5.24** | R=0((r>0) | 25.98** | 19.99 | |
| | | | | R=1(r>1) | 6.69 | 9.13 | |
| Russia-Canada | 1 | -2.30 | -2.38 | R=0((r>0) | 13.23 | 19.99 | |
| | | | | R=1(r>1) | 5.12 | 9.13 | |
| Russia-USA | 1 | -3.79** | -3.81** | r=0((r>0) | 15.47 | 19.99 | |
| | | | | R=1(r>1) | 3.75 | 9.13 | |

Asterisks denote levels of significance (* for 10 percent, ** for 5 percent). The 5% and 10% critical values for tests with a drift are -3.42 and -3.10 respectively. Critical values were obtained from MacKinnon (1991).

Cointegration tests' results – pairwise for Ukraine

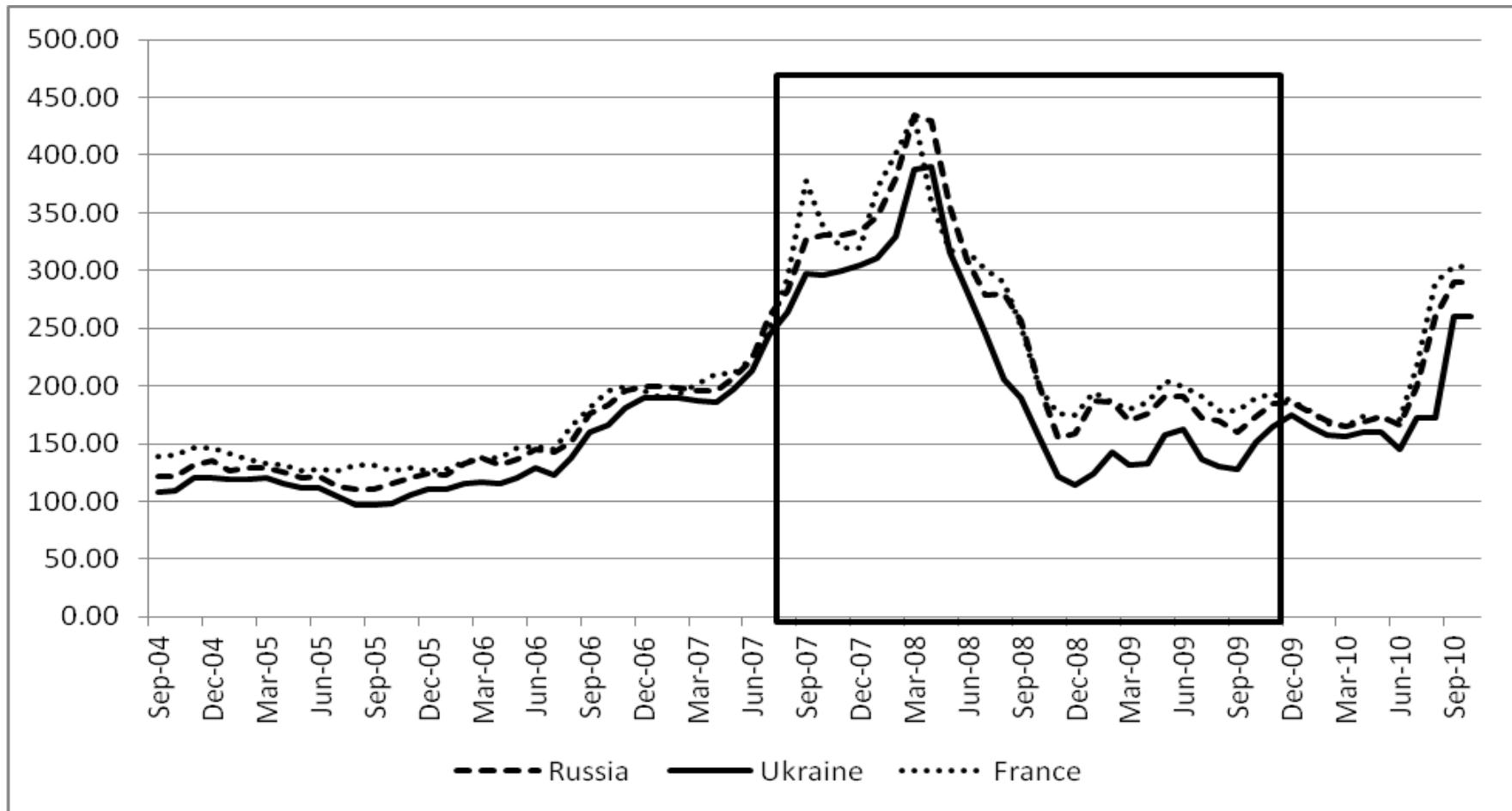
| Pairs of series | Engel and Granger procedure | | | | Johansen method | | |
|------------------|-----------------------------|-------|--------|-----------|-----------------|-------|--|
| | # of lags | ADF | PP | Ho(H1) | | | |
| | | | | | | | |
| Ukraine-France | 3 | -2.33 | -3.64* | R=0((r>0) | 24.66** | 19.99 | |
| | | | | R=1(r>1) | 5.28 | 9.13 | |
| Ukraine - Canada | 1 | -1.90 | -1.99 | R=0((r>0) | 12.48 | 19.99 | |
| | | | | R=1(r>1) | 4.70 | 9.13 | |
| Ukraine-USA | 2 | -2.91 | -3.24* | r=0((r>0) | 12.48 | 19.99 | |
| | | | | R=1(r>1) | 4.56 | 9.13 | |

Asterisks denote levels of significance (* for 10 percent, ** for 5 percent). The 5% and 10% critical values for tests with a drift are -3.42 and -3.10 respectively. Critical values were obtained from MacKinnon (1991).

Therefore,

- Based on the results we confirm cointegration of Russian-French, Russian-US and Ukrainian-French pairs of prices
- The long-run elasticities are equal to
 - 1.04 (case of Russia-France)
 - 1.16 (case of Russia-USA)
 - 1.05 (case of Ukraine-France)

Possibility of a structural break?



Source: IGC, 2011

Chronology of government decisions on grain export restrictions and export quotas, starting from 2006 (1000 MT)

| Period | Wheat | Barley | Corn |
|-----------------------|---|------------------|------------------|
| 10/17/2006-12/31/2006 | 400 | 600 | 600 |
| 12/14/2006-06/30/2007 | 3 | 600 | 500 |
| 02/15/2007-06/30/2007 | 228 | 606 | 30 |
| 02/26/2007-06/07/2007 | | Quotas cancelled | Quotas cancelled |
| 06/08/2007-06/30/2007 | Quotas cancelled | | |
| 07/01/2007-10/31/2007 | 3 | 3 | 3 |
| 01/01/2008-03/31/2008 | 200 | 400 | 600 |
| 01/01/2008-07/01/2008 | 1,200 | 900 | |
| 05/21/2008 | Quotas are cancelled | | |
| 08/2010 | 500 | 500 | 2,000 |
| 10/2010-12/2010 | 500 | 200 | 2,000 |
| 12/2010 | 1,000 | 200 | 3,000 |
| 03/2011 | 1,000 | 200 | 5,000 |
| 05/2011 | Quotas are cancelled | | |
| 05/2011-01/2012 | Tariffs are introduced | | |
| 10/2011 | Tariffs cancelled, except for barley (01/01/2012) | | |

TAR model snapshot

$$\Delta \bar{\varepsilon}_t = \gamma_1 \bar{\varepsilon}_{t-1} + \sum_{i=1}^p \gamma_{i+1} \Delta \bar{\varepsilon}_{t-i} + \omega_t$$

→ $\Delta \bar{\varepsilon}_t = I_t \gamma_1 \bar{\varepsilon}_{t-1} + (1 - I_t) \gamma_2 \bar{\varepsilon}_{t-1} + \varphi_t$, where

$$I_t = \begin{cases} 1 & \text{if } \bar{\varepsilon}_{t-1} \geq \tau \\ 0 & \text{if } \bar{\varepsilon}_{t-1} < \tau \end{cases}$$

Step 3: Testing for asymmetric price transmission – TAR model*

| | Russia - France | Russia-USA | Ukraine –France |
|--------------------------------------|--------------------|--------------------|--------------------|
| Variable | Parameter estimate | Parameter estimate | Parameter estimate |
| γ_1 | -0.80 (-5.31)** | -0.26 (-2.05)* | -0.22 (-1.18) |
| γ_2 | -0.74 (-4.76)** | -0.36 (-2.30)* | -0.38 (-2.75)** |
| $H_0: \gamma_1 = \gamma_2 = 0(\Phi)$ | 22.08** | 4.38** | 6.82** |
| $H_0: \gamma_1 = \gamma_2 (F)$ | 0.11[0.74] | 0.24 [0.62] | 0.33 [0.57] |
| τ | -0.019 | -0.04 | 0.016 |

*M-TAR model provided similar results

Step 3: Testing for asymmetric price transmission – cont.

- The results show that for all three cointegrated pairs of wheat prices the price transmission is **symmetric**
- This implies that the adjustment towards the equilibrium is of the same magnitude regardless of the direction of the change.

Step 4: Error-Correction Model

- Error Correction Models (ECMs) estimate the speed at which a dependent variable returns to equilibrium after a change in an independent variable
- Before the ECM can be formed, there first has to be evidence of cointegration

ECM results

| | Speed of adjustment, α_1 | Test F-value | Time of adjustment |
|------------------|------------------------------------|--------------|-----------------------|
| Russia-France | -0.48** | 12.32** | 3.5 months |
| Russia - USA | -0.20** | 10.69** | 10 months |
| Ukraine - France | -0.20** | 11.38** | 10 months |

Policy Implications

- Trade liberalization issues
 - Long-run transmission results indicate that Ukraine and Russia are integrated with the world market
 - Transmission is symmetric
- Estimation of elasticities
 - Modeling global wheat market behavior
- Future research



THANK YOU!