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Difference-based semiparametric estimation of partial linear regression models

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Abstract. This article describes the `plreg` command, which implements the difference-based algorithm for fitting partial linear regression models.

Keywords: `st0109`, `plreg`, nonparametric regression, difference-based estimator, partial linear regression

1 Introduction

Only in rare cases, economic theory implies a particular functional form for an empirical model specification. Incorrect parameterization of the regression equation might result in inconsistent estimates. Sometimes, the researcher might feel more confident about the functional form of some parts of the regression equation but be less confident about the form of the other parts. Combining the parametric and nonparametric techniques to yield the semiparametric regression model could then help obtain consistent estimates of the parameters of interest.

In this article, I describe implementing the difference-based algorithm to fit the partial linear regression model. The econometric problem of fitting a partial linear model arises in a variety of settings. For example:

- Yatchew (1997) estimates the relationship between variable costs of distributing electricity per customer as a nonlinear function of the scale of operation as measured by the number of customers. The other control variables in the model include measures of customer density, remaining life of distribution assets, and a proxy for local wage rates.
- Yatchew (1998) applies the partial linear regression technique to estimating the hedonic price of housing attributes. Parametric variables include lot size, area of living space, and presence of various amenities. The location effect, which has no natural parametric specification, is incorporated nonparametrically.
- Mesnard and Ravallion (2001) estimate the effect of wealth on business startups among migrants returning to their home country, Tunisia. The paper tests for nonlinear wealth effects on the transition to self-employment, consistent with the argument that the extent of aggregate business activity in the economy depends on the distribution of wealth.

2 Methods

Consider a semiparametric regression,

$$y_i = f(z_i) + x_i\beta + \epsilon_i \quad (1)$$

where z is a random variable, x is a p -dimensional random variable, $E[y|x, z] = f(z) + x\beta$, and ϵ_i is an independently and identically distributed mean-zero error term, such that $\text{Var}[y|x, z] = \sigma_\epsilon^2$. The function f is a smooth, single-valued function with a bounded first derivative. In this model, the parametric ($x\beta$) and nonparametric [$f(z)$] parts are additively separable.

Following the approach suggested by Yatchew (1997), to fit the partial linear model (1), we first rearrange (sort) the data such that $z_1 < z_2 < \dots < z_T$, where T is the number of observations in the sample. Then the first difference of (1) results in the following:

$$\begin{aligned} \{y_{i(n)} - y_{i(n-1)}\} &= [f\{z_{i(n)}\} - f\{z_{i(n-1)}\}] \beta \{x_{i(n)} - x_{i(n-1)}\} + \\ &\quad \epsilon_{i(n)} - \epsilon_{i(n-1)} \quad n = 2, \dots, T \end{aligned} \quad (2)$$

When the sample size increases, $f\{z_{i(n)}\} - f\{z_{i(n-1)}\} \rightarrow 0$ because the derivative of f is bounded. Under standard assumptions, (2) could be estimated by ordinary least squares (OLS). The vector of estimated parameters $\hat{\beta}_{\text{diff}}$ has the approximate sampling distribution

$$\hat{\beta}_{\text{diff}} \rightarrow N\left(\beta, \frac{1}{T} \frac{1.5\sigma_\epsilon^2}{\sigma_u^2}\right)$$

where σ_u^2 is the conditional variance of x given z . The error term in (2) has a first-order moving average structure, thus reducing efficiency of the OLS estimator. The efficiency could be improved by using higher-order differences (Yatchew 1997). The generalization of (2) for the m th-order differencing can be expressed as

$$\sum_{j=1}^m d_j y_{i-j} = \beta \left(\sum_{j=1}^m d_j x_{i-j} \right) + \sum_{j=1}^m d_j f(z_{i-j}) + \sum_{j=1}^m d_j v_{i-j} \quad (3)$$

where d_0, \dots, d_m are differencing weights satisfying the conditions

$$\sum_{j=1}^m d_j = 0 \quad \text{and} \quad \sum_{j=1}^m d_j^2 = 1 \quad (4)$$

The first condition in (4) ensures that the differencing removes the nonparametric component in (3) as the sample size increases. The second normalization condition implies that the residuals in (3) have variance of σ_u^2 . With the optimal choice of weights, (3) could be estimated by OLS. If m is large enough, the estimator approaches asymptotic efficiency.¹

1. The Monte Carlo simulations using the `plreg` command demonstrate that the noticeable gains in efficiency from the higher-order differencing occur on samples with 30,000 or more observations. On the smaller samples (1,000–3,000 observations), using the higher-order differencing might result in biased estimates.

Define $\Delta \mathbf{y}$ to be the $(T - m) \times 1$ vector with elements $(\Delta \mathbf{y})_i = \sum_{j=1}^m d_j y_{i-j}$ and $\Delta \mathbf{x}$ to be the $(T - m) \times p$ matrix with elements $(\Delta \mathbf{x})_i = \sum_{j=1}^m d_j x_{i-j}$. Then,

$$\widehat{\boldsymbol{\beta}}_{\text{diff}} = (\Delta \mathbf{x}' \Delta \mathbf{x})^{-1} \Delta \mathbf{x}' \Delta \mathbf{y} \rightarrow N \left\{ \boldsymbol{\beta}, \frac{1}{T} \left(1 + \frac{1}{2m} \right) \sigma_\epsilon^2 \Sigma_{x|z}^{-1} \right\} \quad (5)$$

$$s_{\text{diff}}^2 = \frac{1}{T} \left(\Delta \mathbf{y} - \Delta \mathbf{x} \widehat{\boldsymbol{\beta}}_{\text{diff}} \right)' \left(\Delta \mathbf{y} - \Delta \mathbf{x} \widehat{\boldsymbol{\beta}}_{\text{diff}} \right) \rightarrow \sigma_\epsilon^2 \quad (6)$$

$$\widehat{\Sigma}_{x|z} = \frac{1}{T} (\Delta \mathbf{x})' \Delta \mathbf{x} \rightarrow \Sigma_{x|z} \quad (7)$$

This method allows performing inferences on $\boldsymbol{\beta}$ as if there were no nonparametric component f in the model. Once $\widehat{\boldsymbol{\beta}}_{\text{diff}}$ is estimated, a variety of nonparametric techniques could be applied to estimate f as if $\boldsymbol{\beta}$ were known. Formally, subtracting the estimated parametric part from both sides of (1), we get

$$y_i - x_i \widehat{\boldsymbol{\beta}}_{\text{diff}} = x_i (\boldsymbol{\beta} - \widehat{\boldsymbol{\beta}}_{\text{diff}}) + f(z_i) + \epsilon_i \cong f(x_i) + \epsilon_i$$

Because $\widehat{\boldsymbol{\beta}}_{\text{diff}}$ converges quickly enough to true $\boldsymbol{\beta}$, the consistency, optimal rate of convergence, and construction of confidence intervals for f remain valid and could be estimated by standard smoothing methods.

Using estimates (5), one can perform the differencing test for the parametric specification of f . Suppose that $g(z, \pi)$ is the known function of z and some unknown parameter π . We want to test the null hypothesis that $y_i = g(z_i, \pi) + x_i \beta_p$ against the alternative hypothesis that $y_i = f(z_i) + x_i \boldsymbol{\beta}$. Parameters π and β_p and mean square residual s_{res}^2 could be obtained by estimating the parametric regression of y on x and z . Then,

$$\mathbf{V} = \sqrt{mT} (s_{\text{res}}^2 - s_{\text{diff}}^2) / s_{\text{diff}}^2 \rightarrow N(0, 1) \quad (8)$$

3 The plreg command

The `plreg` command uses two alternative sets of differencing weights, d_1, \dots, d_m . Optimal weights do not have analytical expressions but have been tabulated (up to $m = 10$) by Hall, Kay, and Titterton (1990) and by Yatchew (1998). In contrast to the former's optimal weight sequences, Yatchew's decline in absolute values toward zero. The nonlinear function f is estimated by Stata's `lowess` procedure. `plreg` also outputs the result of a significance test on z , which is a special case of (8) where $g(z, \pi)$ is a constant function, so that the restricted model is a linear regression function, $y_i = \pi + x_i b$.

3.1 Syntax

```
plreg depvar indepvars [if] [in], nlf(varname) [generate(newvar)
      order(#) wh level(#) collinear lowess_options ]
```

where *depvar* is a dependent variable in (1) and *indepvars* is a vector of variables in a linear (parametric) portion of regression (1).

3.2 Options

nlf(*varname*) is required and specifies the argument of an unknown function, *f*.

generate(*newvar*) creates a new variable *newvar* containing the smoothed values of the argument of *f*. These values are estimated by the locally weighted regression using **lowess**. A corresponding graph of the estimated function *f* could also be output; see [R] **lowess**.

order(#) specifies the differencing order. Tenth-order differencing is the maximum allowed. If order() is not specified, the model is fitted by first-order differencing.

wh specifies a form of the vector of differencing weights d_1, \dots, d_m , as in (3). By default, Yatchew (1998) weights are used. If wh is specified, Hall, Kay, and Titterton (1990) weights are used for differencing.

level(#) sets the confidence level; the default is level(95).

collinear specifies that collinear variables be kept.

lowess_options control the way **lowess** generates the smoothed values for the argument of the nonlinear function.

3.3 Saved results

In addition to the standard results saved by **regress**, **plreg** saves the following in **e()**:

Scalars

e (s2diff)	residual variance (6)
e (s2lin)	variance of the residual in specification that assumes that <i>f</i> is a constant function
e (order)	order of differencing
e (stest)	value of the test on the significance of the variable that enters (1) nonlinearly

Matrices

e (b)	matrix of coefficients of differencing (5)
e (V)	variance-covariance matrix of differencing (5)

3.4 Postestimation commands

Most postestimation commands available after **regress** are also available after **plreg**. The postestimation commands are based on the estimation of the difference (3).

4 Example

I illustrate using the `plreg` command by replicating the example from Yatchew (2003). Data for that example come from the survey of 81 municipal electricity distributors in Ontario, Canada, during 1932.²

The cost of distributing electricity is modeled in a simple Cobb–Douglas framework,

$$tc_i = f(\text{cust}_i) + \beta_1 \text{wage}_i + \beta_2 \text{pcap}_i + \beta_3 \text{puc}_i + \beta_4 \text{kwh}_i + \beta_5 \text{life}_i + \beta_6 \text{lf}_i + \beta_7 \text{kmwire}_i + \epsilon_i$$

where `tc` is the log of total cost per customer, `cust` is the log of number of customers, `wage` is the log of wage rate, `pcap` is the log price of capital, `puc` is the dummy variable for the public utility commissions that deliver additional services and may benefit from economy of scope, `kwh` is the log of kilowatt hours per customer, `life` is the remaining life of distribution assets, `lf` is the load factor, and `kmwire` is the log of kilometers of distribution wire per customer. The objective of the analysis is to assess scale economies in electricity distribution.

The parametric effect, β , is estimated by first-order differencing, as in (2). We also estimate by OLS a pure parametric specification, where the scale effect, f , is modeled with a quadratic polynomial. Stata output of the estimations using `plreg_example.dta` is shown below.

```
. use plreg_example
. plreg tc wage pcap puc kwh life lf kmwire, nlf(cust) generate(func) bwidth(1)
> nodraw
```

Partial linear regression model with Yatchew's weighting matrix

Source	SS	df	MS	Number of obs =	80
Model	1.765078594	7	.252154085	F(7, 73)	= 12.663
Residual	1.453568962	73	.019911904	Prob > f	= 0.0000
				R-squared	= 0.5484
				Adj R-squared	= 0.5051
Total	3.219	80	.040233094	Root MSE	= 0.1411

tc	Coef.	Std. Err.	t	P> t	[95% Conf. Interval]
wage	.4484555	.3695674	1.213	0.229	-.2880912 1.185002
pcap	.458975	.0760358	6.036	0.000	.3074359 .6105141
puc	-.0856378	.042962	-1.993	0.050	-.1712609 -.0000148
kwh	-.0105118	.0879159	-.1196	0.905	-.185728 .1647045
life	-.506133	.1318116	-3.84	0.000	-.7688332 -.2434328
lf	1.25216	.4595468	2.725	0.008	.3362849 2.168036
kmwire	.3516307	.0943774	3.726	0.000	.1635368 .5397245

Significance test on cust: V = 5.757 P>|V| = 0.000

2. I use the data for this example with the permission of Dr. Yatchew. You can download the data from <http://www.chass.utoronto.ca/~yatchew/>.

The syntax of this command does not specify an option for the order of differencing using the default first-order differencing; `bwidth(1)` is an option of the Stata `lowess` command, which determines the bandwidth for nonparametric smoothing.

The significance test of the variable (`cust`) that enters the specification nonlinearly (7) indicates that the log of number of customers is highly significant (p -value of 0.000). The estimation of the fully parametric model with a quadratic polynomial of the log of number of customers shows that although the effect of exogenous variables is qualitatively similar between these two specifications, the magnitudes of some coefficients are different. For example, the effect of log wage on log of total cost per customer declines from 0.83 in the fully parametric model to 0.45 in the partial linear model estimation.

```
. regress tc cust custsq wage pcap puc kwh life lf kmwire
```

Source	SS	df	MS			
Model	2.76114864	9	.306794293	Number of obs =	81	
Residual	1.70734029	71	.024047046	F(9, 71) =	12.76	
Total	4.46848893	80	.055856112	Prob > F =	0.0000	
				R-squared =	0.6179	
				Adj R-squared =	0.5695	
				Root MSE =	.15507	

tc	Coef.	Std. Err.	t	P> t	[95% Conf. Interval]	
cust	-.832789	.1749502	-4.76	0.000	-1.18163	-.4839481
custsq	.0402137	.0091974	4.37	0.000	.0218746	.0585529
wage	.8325415	.32466	2.56	0.012	.1851878	1.479895
pcap	.5620181	.0741125	7.58	0.000	.414242	.7097941
puc	-.0705723	.0388506	-1.82	0.074	-.1480383	.0068937
kwh	-.0174608	.0889375	-0.20	0.845	-.1947972	.1598756
life	-.602922	.1192685	-5.06	0.000	-.8407367	-.3651073
lf	1.243992	.4343841	2.86	0.005	.3778543	2.110129
kmwire	.4452568	.085974	5.18	0.000	.2738295	.616684
_cons	2.750979	2.138662	1.29	0.203	-1.513392	7.01535

Using these two estimations, we can conduct a test of quadratic versus nonparametric scale effect. Substituting corresponding values into (8), we get $V = \sqrt{81}(0.240 - 0.199)/0.199 = 1.854$, which corresponds to a p -value of 0.032.

`plreg` uses the Stata `lowess` routine to generate the nonparametric smoothing of nonlinear function f . Figure 1 illustrates the nonparametric and fully parametric estimates of the return to scale in electricity distribution plotted against the log of the total number of customers. Quadratic specification fits the data closely to the nonparametric specification.

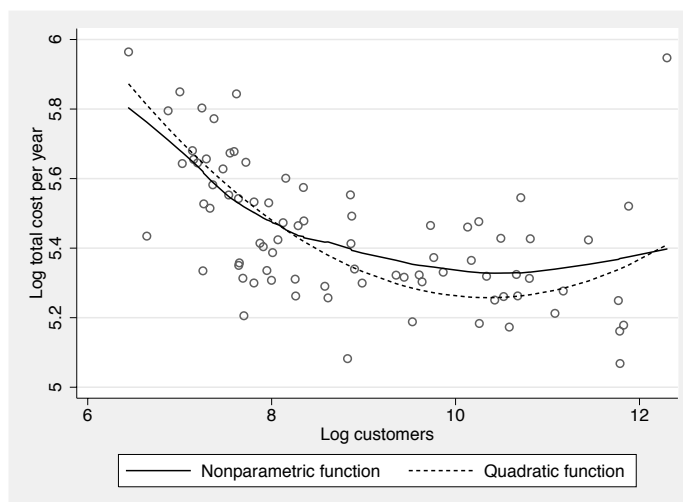


Figure 1: Electricity distribution plotted against the log of the total number of customers

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