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MORE ON ESTIMATING ELASTICITIES USING LAGGED ENDOGENOUS VARIABLES

by

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MORE ON ESTIMATING ELASTICITIES USING  
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Introduction

In an article appearing in this journal, Bopp and Pendley<sup>1</sup>, hereafter abbreviated B-P, have considered the problem of constructing confidence intervals for long-term elasticities estimated from logarithmic equations containing lagged endogenous variables as explanatory variables. The estimated long-term elasticity for explanatory variable  $i$  is

$$\hat{\eta}_i = \frac{\hat{B}_i}{(1-\hat{\lambda})} \quad , \quad (1)$$

where  $\hat{B}_i$  is the estimated regression coefficient for explanatory variable  $i$ , or equivalently, the short-term elasticity for explanatory variable  $i$ , and  $\hat{\lambda}$  is the estimated coefficient for the lagged endogenous variable. Bopp and Pendley use an estimate of the asymptotic variance of  $\hat{B}_i/(1-\hat{\lambda})$  to construct approximate confidence intervals for  $\eta_i$ . The purpose of this note is to present a simple technique for constructing exact confidence intervals for  $\eta_i$ . This technique is also illustrated by application to the data used by B-P.

Exact Confidence Intervals for  
Long-Term Elasticities

The ratio  $\hat{B}_i/(1-\hat{\lambda})$  is, under the usual assumptions, a ratio of two normally distributed random variables with the numerator having an expected value of  $B_i$  and variance  $\sigma_{\hat{B}_i}^2$ , and the denominator having expected value of  $1-\lambda$  and variance  $\sigma_{\hat{\lambda}}^2$ . The covariance of the numerator and denominator is  $-\sigma_{\hat{B}_i, \hat{\lambda}}$ . Although Hinkley<sup>2</sup> has derived an exact

expression for the exact cumulative distribution function (c.d.f.) for such a ratio, evaluation of that function involves use of tabulations of the bivariate normal distribution function. Since use of these tabulations usually requires trivariate interpolation, construction of exact confidence intervals for  $\eta_i$  using this approach would be at best inconvenient.

However, a result due to Fieller<sup>3</sup> allows construction of exact confidence intervals for  $\eta_i$  and requires no more than solutions to quadratic equations. This result makes use of the fact that the ratio  $R = \hat{B}_i / (1 - \hat{\lambda})$  is equivalent to

$$\hat{B}_i - R(1 - \hat{\lambda}) = 0 . \quad (2)$$

A  $1 - \alpha$  percent confidence interval for  $R$  is given by those values of  $R$  for which

$$\text{Prob} \left\{ \frac{(\hat{B}_i - R(1 - \hat{\lambda}))^2}{s_{(1-\hat{\lambda})}^2 - 2R(\hat{B}_i(1-\hat{\lambda}) - t_{\alpha/2}^2 s_{\hat{B}_i, (1-\hat{\lambda})}) + \hat{B}_i^2 - t_{\alpha/2}^2 s_{\hat{B}_i}^2} \leq t_{\alpha/2}^2 \right\} = 1 - \alpha \quad (3)$$

where  $s_{\hat{B}_i}^2$  is the sample variance of  $\hat{B}_i$ ;  $s_{(1-\hat{\lambda})}^2$  is the sample variance of  $(1 - \hat{\lambda})$ ;  $s_{\hat{B}_i, (1-\hat{\lambda})}$  is the sample covariance of  $\hat{B}_i$  and  $(1 - \hat{\lambda})$ ; and  $t_{\alpha/2}^2$  is the squared value of the upper  $\alpha/2$  percentage point of the t-distribution with appropriate degrees of freedom, or equivalently the upper  $\alpha$  percentage point of the F-distribution with one degree of freedom in the numerator and the appropriate degrees of freedom in the denominator. For the problem at hand, the appropriate degrees of freedom for the t-distribution or the denominator of the F-distribution are the error degrees of freedom of the regression used to estimate  $\hat{B}_i$  and  $\hat{\lambda}$ . For the confidence interval to be closed,  $(1 - \hat{\lambda})$  must be significantly different

from zero. The exact  $1-\alpha$  confidence limits for  $\eta_i$  are those values of  $R$  for which

$$R^2 \left( (1-\hat{\lambda})^2 - t_{\alpha/2}^2 s_{(1-\hat{\lambda})}^2 \right) - 2R \left( \hat{B}_i (1-\hat{\lambda}) - t_{\alpha/2}^2 s_{\hat{B}_i (1-\hat{\lambda})} \right) + B_i^2 - t_{\alpha/2}^2 s_{\hat{B}_i}^2 = 0. \quad (4)$$

Note that the solution to this quadratic equation in  $R$  need not, and generally will not, be symmetric about  $\hat{\eta}_i$ . Thus, the use of the asymptotic variance of  $\hat{\eta}_i$  to construct approximate confidence intervals for  $\eta_i$  centered upon  $\hat{\eta}_i$  can be misleading.

#### An Application

In order to illustrate the construction of exact confidence intervals for long-term elasticities, we make use of the data analyzed by B-P. They estimated a logarithmic equation expressing residential electrical consumption in Connecticut ( $Q_t$ ) as a function of real disposable income ( $Y_t$ ), own marginal price ( $P_t$ ), and lagged consumption ( $Q_{t-1}$ ) using annual observations from 1960 to 1975. The estimated short- and long-term elasticities and their respective variances are summarized in Table 1. The estimated coefficient of adjustment ( $1-\hat{\lambda}$ ) was 0.152, and the relevant covariance terms were estimated as follows:  $s_{\hat{Y}, (1-\hat{\lambda})} = -s_{\hat{Y}, \hat{\lambda}} = 0.005$ ; and  $s_{\hat{P}, (1-\hat{\lambda})} = -s_{\hat{P}, \hat{\lambda}} = 0.0008$ . The error degree of freedom is 12. With this information, exact confidence intervals can be constructed for both the short- and long-term elasticities. These are also displayed in Table 1.

Note that the exact confidence intervals for the long-term elasticities are much wider than their respective short-term counterparts. Also, the exact confidence intervals for the long-term elasticities are not symmetric about their respective point estimates, a result not

Table 1. Exact Confidence Intervals for Short- and Long-Term Elasticities

Variable	Estimated Elasticity <sup>a</sup>		Exact 95% Confidence Intervals	
	Short-term	Long-term	Short-term	Long-term
Income	0.17 (0.1221)	1.13 (0.474)	-0.10 to 0.44	-1.51 to 1.78
Price	-0.21 (0.0480)	-1.4 (0.587)	-0.31 to -0.11	-4.08 to -0.52

a. Standard errors of estimated elasticities are shown in parentheses. The standard errors for long-term elasticities are asymptotic.

detected by the use of the asymptotic variance of  $\hat{\eta}_i$  in constructing appropriate confidence intervals for  $\eta_i$ . Although the exact confidence interval for the short-term income elasticity is wider than the interval for short-term own price elasticity, the relative widths of the long-term elasticity intervals are reversed, a result agreeing with the findings of B-P.

#### Summary and Conclusions

The purpose of this note is to demonstrate a technique for constructing exact confidence intervals for long-term elasticities from logarithmic equations containing lagged endogenous variables as regressors. This technique involves only simple computations based upon results generated in the course of parameter estimation.

Notes

<sup>1</sup>Bopp, Anthony E., and Robert E. Pendley. "A Note on Estimating Economic Elasticities Using Lagged Endogenous Variables," Intermountain Economic Review (Fall 1978), pp. 111-7.

<sup>2</sup>Hinkley, D. V. "On the Ratio of Two Correlated Normal Random Variables," Biometrika 56(1969), pp. 635-9.

<sup>3</sup>Fieller, E. C. "The Distribution of the Index in a Normal Bivariate Population," Biometrika 24(1932), pp. 428-40.

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