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Confidence intervals for the variance component of random-effects linear models

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Abstract. We present the postestimation command **xtvc** to provide confidence intervals for the variance components of random-effects linear regression models. This command must be used after **xtreg** with option **mle**. Confidence intervals are based on the inversion of a score-based test (Bottai 2003).

Keywords: st0077, xtvc, variance components, confidence intervals, score test, random-effects linear models

1 Introduction

The random-effects linear model has been widely applied to different areas of data analysis (see, among many others, Breslow and Clayton 1993; Diggle, Liang, and Zeger 1994; Snijders and Bosker 1999; McCulloch and Searle 2001; Skrondal and Rabe-Hesketh 2004). The Stata **xtreg** command fits the random-effects linear regression model, which can be written as

$$y_{it} = \mathbf{x}_{it}\beta + u_i + e_{it}, \quad u_i \sim N(0, \sigma_u^2), e_{it} \sim N(0, \sigma_e^2)$$
(1)

where y_{it} is the *t*th observation taken on some random variable *Y* for the *i*th unit and $i = 1, ..., m, t = 1, ..., T_i$; \mathbf{x}_{it} is a covariate vector and β is a parameter vector of fixed effects; u_i is a unit-specific normal random effect with zero mean and variance σ_u^2 that is assumed to be non-negative; and e_{it} is the normal residual error with variance σ_e^2 that is assumed to be strictly positive. Also, u_i and e_{it} are assumed to be independent. Units can refer to individuals on whom repeated observations are taken, families whose members are sampled, or otherwise-defined groups within which observations may be correlated.

In such models, it is often of interest to make inference not only about the fixed and random effects but also about the variance components. In particular, testing homogeneity across units is equivalent to testing the null hypothesis

$$H_0: \sigma_u^2 = 0 \tag{2}$$

In general, testing whether a variance parameter is zero implies testing a parameter value on the boundary of the parameter space, the variance being non-negative. Several

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st0077

Confidence intervals

authors suggest using the large-sample likelihood-ratio test that adjusts for the boundary condition. In fact, under this irregular scenario, the asymptotic distribution of the usual likelihood-ratio test statistic follows a distribution that is a 50:50 mixture of a $\chi^2_{(1)}$ and the constant zero (Self and Liang 1987). The Stata command **xtreg** provides the upper-tail probability of the appropriate asymptotic distribution of the likelihood-ratio test statistic (Gutierrez, Carter, and Drukker 2001).

However, such a method cannot be used to construct confidence intervals for the variance of the random effect, σ_u^2 . Besides, the confidence intervals provided for the random-effect variance by **xtreg**, based on a Wald-type test, can be shown to be asymptotically wrong. To the best of our knowledge, no published work has provided methods for constructing likelihood-based confidence regions for the variance component that are asymptotically correct.

It can be shown that inference about the variance component σ_u^2 can be accommodated within the irregular problems of singular information. Such a connection had been noted several years ago (Chesher 1984; Lee and Chesher 1986), but only recently a general theory was developed for the singular-information case (Rotnitzky et al. 2000). Using the results derived for the singular-information problem (Bottai 2003), a method is implemented in the Stata command **xtvc** that is based on the inversion of a scoretype test, which provides asymptotically correct confidence intervals. Also, when testing the hypothesis of homogeneity across units (2), the proposed method is shown to have better small-sample properties than the one based on the likelihood-ratio test statistic.

The rest of the article is organized as follows: section 2 introduces the syntax of the command xtvc; section 3 provides an example in which the command xtvc is applied to real data; section 4 reports the observed rejection proportions of the confidence intervals generated by xtvc on simulated data; and some final remarks are presented in section 5.

2 The xtvc command

2.1 Syntax

The xtvc command is to be used after the xtreg command with the mle option for maximum likelihood estimation. The syntax of xtvc is as follows:

xtvc [, $\underline{l}evel(\#) \underline{h}O(\#)$]

2.2 Options

- level(#) specifies the confidence level, as a percentage, for the confidence interval of the variance component. The default is level(95) or as set by set level; see
 [U] 23.6 Specifying the width of confidence intervals.
- h0(#) performs the score-based test for the null hypothesis H_0 : sigma_u = #. The default null value is 0.

M. Bottai and N. Orsini

2.3 Saved Results

xtvc saves all the results of xtreg plus the following:

```
   Scalars
   e(score)
   score test statistic
   e(pval)
   p-value

   e(suuppb)
   upper bound of \sigma_u
   e(sulowb)
   lower bound of \sigma_u

   Macros
   e(pcmd)
   xtvc
```

3 Example: the NLSY data

xtvc is applied to the longitudinal data from a subsample of the NLSY data (Center for Human Resource Research 1989) described in many of the [XT] **xt** entries and available on the Stata Press web page (http://www.stata-press.com/data/r8/xt/). In this example, we fit a random-effects linear model for the variable ln_wage as a function of several variables as was done in the **xtreg** example; see [XT] **xtreg**.

. webuse nlswo (National Long . iis idcode . xtreg ln_w g Fitting consta	ork, clear gitudinal Surv grade age ttl ant-only mode	vey. Young _exp tenure l:	Women 14- e not_smsa	26 years a south,	s of age in 19 mle	968)
(output omit	ted)					
(output omit	ted)					
Random-effects Group variable	s ML regressio e (i): idcode	on		Number Number	of obs = of groups =	28091 4697
Random effects	s u_i ~ Gauss:	ian		Obs per	group: min = avg = max =	1 1 6.0 15
Log likelihood	d = −9218.977	73		LR chi2 Prob >	2(6) = chi2 =	6861.27 0.0000
ln_wage	Coef.	Std. Err.	Z	P> z	[95% Conf.	Interval]
grade	.0691186	.0017232	40.11	0.000	.0657412	.072496
age	003809	.0000491	-3.90	0.000	0031412	0020907
toruro	.030151	.0011135	16 08	0.000	.02/900/	.0323334
not smga	- 1200780	0071709	-18 13	0.000	- 1//0337	- 11502470
south	- 0941264	0071354	-13 19	0.000	- 1081115	- 0801413
_cons	.7566548	.0267764	28.26	0.000	.7041741	.8091355
/sigma_u /sigma_e	.2503043 .2959207	.003531 .0013704	70.89 215.94	0.000	.2433837 .2932348	.2572249 .2986065
rho	.4170663	.0074739			.4024786	.4317692

Likelihood-ratio test of sigma_u=0: chibar2(01)= 7277.75 Prob>=chibar2 = 0.000

Confidence intervals

We then use the xtvc command:

•	xtvc

	ln_wage	ML Estimate	[95% Conf. Interval]	
/sigma_u .2503043 .2488335 .2630834	/sigma_u	. 2503043	.2488335 .2630834	

Score test of sigma_u=0: chi2(1)= 39399.39 Prob>=chi2 = 0.000

The point estimate for the random-effects standard deviation σ_u is exactly the same as the one given by **xtreg**, but the confidence interval provided by **xtvc** is slightly shifted to include greater values. Both the score-type test provided by **xtvc** and the likelihood-ratio test provided by **xtreg** reject the null hypothesis that the standard deviation σ_u is equal to zero. With the h0 option of the **xtvc** command, it is also possible to test any value for the standard deviation σ_u , not only zero. For example, we can test the value $\sigma_u = 0.25$, which is included in the 95% confidence interval.

. xtvc, h0(0.25)

ln_wage	ML Estimate	[95% Conf. Interval]	
/sigma_u	.2503043	.2488335 .2630834	

Score test of sigma_u=0.25: chi2(1)= 2.63 Prob>=chi2 = 0.105

4 Simulated data

The **xtvc** command was applied to simulated data. Three thousand samples were pseudo-randomly generated for model (1) under a grid of values for the random-effect standard deviation $\sigma_u = 0, 0.01, \ldots, 0.09, 0.10, 10$, and for different numbers of units or groups m = 10, 100, 1000. The residual-error standard deviation σ_e was set constant to the value one for all the simulations. Two covariates were pseudo-randomly generated from a uniform(-1, 1) and a uniform(0, 2) distribution, respectively, with $\beta = (1, 2)^T$. The observed rejection proportions over the simulated samples of the 95% confidence intervals provided by **xtvc** are shown in table 1. For the samples generated under the value $\sigma_u = 0$, the observed rejection proportion of the adjusted likelihood-ratio test at the 5% level provided by **xtreg** is also reported.

(Continued on next page)

M. Bottai and N. Orsini

Table 1: Observed rejection proportions of xtvc and xtreg (using chibar2(01)) among 3,000 simulated samples generated under different values of σ_u and number of units or groups for the random-effects linear model (1) (simulation error $\pm 0.78\%$).

σ_u	m = 10	m = 100	m = 1000
xtvc			
0.00	5.20	5.23	4.63
0.01	5.17	5.43	5.37
0.02	5.03	5.23	4.93
0.03	5.33	5.60	4.57
0.04	5.30	5.07	5.63
0.05	4.73	5.63	5.00
0.06	5.77	5.17	4.93
0.07	5.30	5.63	5.30
0.08	5.27	5.40	4.53
0.09	5.47	5.43	5.30
0.10	4.80	5.20	4.07
10.0	4.57	5.03	4.90
xtreg			
0.00	2.43	4.13	4.27

Regardless of the number of units or groups, m, the observed rejection proportion is uniformly close to its nominal level of 5% across the values of the standard deviation σ_u . Although based on a large-sample test, **xtvc** shows acceptable behavior in small samples as well.

The adjusted likelihood-ratio test provided by **xtreg** was applied only to the samples simulated under the value $\sigma_u = 0$. In the present simulation, when the number of units or groups m = 10, its observed rejection proportion is 2.43%, well below its nominal level of 5%. In other extensive simulation experiments not reported here, we observed that the rejection proportion becomes satisfactorily close to the nominal level only when the number of units or groups is no smaller than a thousand.

The observed rejection proportion of the confidence regions obtained by inverting the Wald-type test, as provided by **xtreg**, is wrong in small samples as well as large samples. Depending on the values of σ_u and m, its rejection probability can be as high as 15% or as low as 0.5%. Besides, its confidence intervals may happen to include negative values, which are out of the feasible space of the variance parameter.

5 Final remarks

The xtvc command is the only solution for those seeking to construct confidence intervals for the variance component of a random-effects linear regression model. The method can be extended to more general models, such as generalized linear mixed mod-

Confidence intervals

els, whose estimation is based on the likelihood function. In the present version, the command xtvc only provides interval estimates when the number of units or groups is greater than eight. For balanced data, explicit solutions for the upper and lower bounds of the confidence intervals are available but are not implemented in the command xtvc. Instead, in the unbalanced case, the bounds of the confidence intervals are obtained by iterative algorithms. Equations are solved by bisection methods, which usually take little time to converge. In later versions, the Newton–Raphson optimization could be used instead, should the command take too long.

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M. Bottai and N. Orsini

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