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**Consumer Willingness to Pay for Food Safety in Beijing:
A Case Study of Food Additives**

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Consumer Willingness to Pay for Food Safety in Beijing: A Case Study of Food Additives*

Abstract: Constructing a theoretical framework and using a survey data of 294 customers from 25 supermarkets in Beijing, this paper studies the willingness to pay (WTP) for additive-free Mooncakes in Beijing and finds that age and income are important for WTP for “food safety” in China. Income is positively correlated with the WTP and there is an inverted-U-shaped relationship between age and WTP. This study indicates that consumers in Beijing are willing to pay 5.80 Yuan more for an additive-free Mooncake, which provides a good policy benchmark for the government regulation on food additives. Furthermore, the theoretical framework also provides a good benchmark for understanding WTP in the future study of food safety.

Key words: Food Safety; Willingness to Pay; Double-Bounded Dichotomous Choice; Additive-Free Mooncakes

JEL: I12, Q18

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Introduction

The demand for food quality is increasing as income increases in China and food safety is a very important component in food quality (Yu and Abler 2009). Recent crises associated with food safety, such as the baby milk powder incident, Bovine Spongiform Encephalopathy (Mad Cow disease), Avian Influenza (Bird Flu) have raised public concerns about food safety not only in China but also in the world. Chinese government has adopted a series of policies to strengthen the regulation on food safety. The cost-benefit approach is prevailing for the evaluation of food safety policies; though in practice, it is a big problem to define costs and benefits accurately. Generally, we use the consumer willingness to pay (WTP) for food safety to measure the benefits for consumers (Golan and Kuckler 1999) .

The benefit of food safety is a non-market value, and it is difficult to get the information of the revealed preferences of consumers. We in general adopt the stated preference methods to assess the value¹, of which the contingent valuation method (CVM) is the most important and also the most popular one (Zhang et.al., 2003)². Golan and Kuckler (1999), and Antle (2001) are good reviews for both the theoretical and empirical studies of consumer WTP for food safety.

¹ Bockstael and Freeman (2005) is a good review of the development, status and controversies about the revealed and stated preference methods.

² Carson and Hanemann (2005) summarize the development and current situation of CVM.

As the concern about food safety is increasing in China, there is a large body of literature studying consumer willingness to pay for food safety (Wang, 2003; Zhou, 2004; Chen, 2006; Huang et al., 2006; Huang et al., 2007; Zeng et al., 2008).

However, the existing researches of the WTP for food safety in China have some shortcomings. (1) These papers are mainly focusing on the analysis of consumer attitudes, perceptions and willingness to buy; and they lack the analysis of consumers' willingness to pay, specifically short of calculation of WTP. For example, Dai et al. (2006) only analyzes consumers' purchase behavior and the determinants of WTP. There are only a few exceptions. For instance, the results of Zhou et al. (2006) indicate that the mean WTP for food safety to reduce pesticide residues of B. Chinensis is ¥ 5.36 per kilogram, which indicates a very high value of food safety as well as the benefits of government regulations. If we do not calculate a specific number of WTP, it can not give an appropriate evaluation of food safety policies. (2) Compared with a large volume of empirical studies, so far there are little theoretical frameworks to explain the empirical results. (3) There is little literature studying willingness to pay for "additive-free food".

There are many types of food additives, and the preservatives are one of the most prevailing but controversial varieties. Preservatives can help store food for a longer time, but they may harm consumer health. How to regulate food additives is strongly hinged on the consumer benefits from the regulation, which can be measured by the consumer willingness to pay.

In this study, we will take the Mooncakes as an example to study consumer willingness to pay for additive-free Mooncakes in China. The Mooncakes are Chinese traditional pastries consumed during the Mid-Autumn Festival (August 15 of the lunar calendar). The Mid-Autumn is very important for family unions and the Mooncakes are offered between friends and family members.

The structure of this paper is outlined as follows: first, this study constructs a theoretical framework to explain the willingness to pay for food safety; second, uses a survey data of willingness to pay for “additive-free food (Mooncakes)” of 294 customers from 25 supermarkets in Beijing to empirically study the determinants of willingness to pay for “additive-free food” in China, and to test the theoretical hypotheses as well; and finally, gives the conclusion.

Theoretical Framework

In order to simplify the study, we assume that there are only two types of Mooncakes with different safety levels in a market: one contains preservatives, and the other does not. If a consumer only consumes additive-free Mooncakes, the indirect utility function is $V(p_F, h, m, Z)$, given a price of the additive-free Mooncake p_F , consumers' health stock h , current income m , and a vector of some other exogenous variables Z .

The additive-free Mooncakes usually have a shorter stock period, and this will increase the cost of producers, so that the price might be higher. If the Mooncakes contain preservatives, the price of Mooncake will decrease to $p_F - t$.

Assuming the market is competitive and in equilibrium and the information about additives is symmetric, t is a mark-up in cost after adding the preservatives from the perspective of producers; on the other hand, t is the consumer willingness to pay for additive-free Mooncakes from the perspective of consumers. It is also known that the preservatives can damage consumers' health, and then the health stock of the consumer may decrease from h to $h - d$ if he chooses to consume the Mooncakes with preservatives. So, if a consumer only consumes the Mooncakes containing the preservatives, the indirect utility function will be $V(p_F - t, h - d, m, Z)$. Market equilibrium shows that

$$V(p_F - t, h - d, m, Z) = V(p_F, h, m, Z) \quad (1)$$

Taking the first-order approximation of $V(p_F - t, h - d, m, Z)$, we have

$$V(p_F - t, h - d, m, Z) \approx V(p_F, h, m, Z) - \frac{\partial V}{\partial p_F} t - \frac{\partial V}{\partial h} d \quad (2)$$

Combining (1) and (2), we have

$$t = -\frac{\partial V / \partial h}{\partial V / \partial p_F} d \quad (3)$$

By Roy's identity,

$$t = \frac{d}{x_n} \frac{\partial V / \partial h}{\partial V / \partial m} \quad (4)$$

Where x_n is the Marshallian demand for the additive-free Mooncakes. Assuming the demand elasticity of Mooncakes is very small. That is, regardless of whether the Mooncake contains preservatives or not, the total consumption of Mooncakes is constant for the consumer. Define $k = \frac{d}{x_n}$; denoting the health damage from consuming per unit Mooncake, and for a certain consumer, k is a constant. Rewriting equation (4),

$$t = k \frac{\partial V / \partial h}{\partial V / \partial m} \quad (5)$$

$\partial V / \partial h$ is defined as the marginal utility of health; $\partial V / \partial m$ is defined as the marginal utility of money. By Equation (5), we can give two hypotheses as follows:

1. Consumers' willingness to pay food safety is positively correlated with the marginal utility of health. In our real world, the youth have longer life expectations than the elder, so that it is expected that the youth's marginal utility of health is bigger than the elder due to time. But, considering that young people are easy to recover from injury, very young person's marginal utility of health may be very low. Adding up the two effects, $\partial V / \partial h$ would first increase and then decrease as the ages of consumers increase. According to this, we can have the first testable hypothesis: consumers' willingness to pay for additive-free food would first increase and then decrease as the age increases.

2. Consumer willingness to pay for food safety is negatively correlated with the marginal utility of money. In our real world, the rich usually have smaller $\partial V / \partial m$, so

that we can have the second testable hypothesis: consumer willingness to pay would increase as income increases.

In the next part, we will use the survey data of willingness to pay for additive-free Mooncake of 294 consumers from 25 supermarkets in Beijing of China to test the above hypotheses and give the policy implications as well.

Econometric Model

The contingent valuation method (CVM) has many different elicitation formats, and different formats would influence the results greatly. Researchers have gradually developed two types of methods to elicit consumer WTP: (1) the Continuous Method, in practice, including the payment card (PC) approach and the open-ended (OE) approach; (2) the Discrete Method, in practice, mainly including dichotomous choice (DC) approach. Ready, Buzby and Hu (1996) point out that a continuous format generates a lower estimated WTP than a dichotomous choice format due to more yes-saying among DC respondents. In practice, the discrete method is more popular. In 1993, the National Oceanic and Atmospheric Administration (NOAA) panel gave some important guiding principles about the application of CVM, and NOAA recommended the dichotomous choice approach for eliciting WTP for non-market goods (Arrow et al., 1993)³. The DC approach also has different elicitation formats. Single-bounded dichotomous choice (SBDC) and double-bounded dichotomous

³ This format had many advantages. Such as, it can have a better simulation of the market price, and reduce the strategic bias to get more reliable and accurate valuation of the WTP.

choice (DBDC) are the most important ways. Technically, the estimation methods of SBDC and DBDC are completed by Hanemann et al., in 1984 and 1991, respectively.

Based on the principles of utility maximization, consumers would choose different levels of food safety. According to McFadden (1974)'s random utility model (RUM), the economic principle of CVM can be described as follows: other things being equal, when the level of food safety rises from a relatively low level Q_0 (additive food) to a higher level Q_1 (additive-free food), consumers can gain more utility, as mentioned above, that is,

$$V_1(Q_1, p_F, m, Z, \varepsilon_1) > V_0(Q_0, p_F, m, Z, \varepsilon_0)$$

ε_0 and ε_1 are the random error terms. CVM uses the survey method to reveal consumer preferences, and we can derive the equilibrium utility at different levels of food safety from the above theoretical framework, so that

$$V_1(Q_1, p_F + t, m, Z, \varepsilon_1) = V_0(Q_0, p_F, m, Z, \varepsilon_0)$$

Then, we can use statistical methods to derive t , which represents the consumer willingness to pay (Zhou et al., 2006).

This paper uses the DBDC approach as the specific elicitation format, and the following part will introduce its principles and the mathematical derivation. The DBDC approach was first proposed by Hanemann (1985) and then developed by Hanemann et al. (1991). It asks the respondents to engage in two rounds of bidding: participants respond to a first dollar amount and then face a second question involving

another dollar amount, higher or lower depending on the response to the first question (Hanemann et al., 1991).

In this paper, respondents are presented with the following questions: “If the price of the Mooncake without preservatives is B_i Yuan per unit higher than the conventional Mooncake, are you willing to pay?” then followed up with: “What about B_i^u (or B_i^d)?” B_i is the initial bid, B_i^u is the second bid if the response to the first bid was “yes”; B_i^d is the second bid if the response was “no”. In this way, the respondent’s answers will be four possible combinations :(yes, yes), (no, no), (yes, no), (no, yes). Hanemann et al. (1991) first constructed the log-likelihood function of the DBDC approach, and verified that the DBDC approach was shown to be asymptotically more efficient than the conventional SBDC approach, although the analysis of data is more complex.

Following Watson and Ryan (2007), Let t_1 be the base bid at the initial dichotomous choice question (DC1) and t_2 be the follow up bid at the second dichotomous choice question (DC2). The above possible responses are:

- 1) When respondent’s answer is “yes-yes”, $WTP \geq t_2$
- 2) When respondent’s answer is “no-no”, $WTP < t_2$
- 3) When respondent’s answer is “yes-no”, $t_1 \leq WTP < t_2$
- 4) When respondent’s answer is “no-yes”, $t_1 > WTP \geq t_2$

Following this:

$$WTP_{ij} = \beta' x_{ij} + \varepsilon_{ij} \quad (6)$$

where WTP_{ij} is the WTP of individual j , and $i = 1, 2$ represents DC1 and DC2, respectively; x_{ij} ($i = 1, 2$) is a vector of explanatory variables, including the bids (B), consumers' demographic characteristics (such as income, age, gender, education and so on), supermarket's characteristics (S, for example, the size of supermarket); β is a corresponded vector of coefficients. The error term, ε_{ij} , incorporates both individual and question specific error.

By Equation (6), for instance, the probability of respondent j answering “yes” to DC1 and “no” to DC2 is expressed as:

$$\Pr(\text{yes-no}) = \Pr(WTP \geq t_1, WTP < t_2).$$

That is,

$$\Pr(\text{yes-no}) = \Pr(\beta' x_{1j} + \varepsilon_{1j} \geq t_1, \beta' x_{2j} + \varepsilon_{2j} < t_2).$$

Then, incorporating all response combinations in the likelihood function, gives

$$\begin{aligned} L_j(\beta' x_{ij} | t) &= \Pr(\beta' x_{1j} + \varepsilon_{1j} \geq t_1, \beta' x_{2j} + \varepsilon_{2j} < t_2)^{YN} \\ &\times \Pr(\beta' x_{1j} + \varepsilon_{1j} > t_1, \beta' x_{2j} + \varepsilon_{2j} \geq t_2)^{YY} \\ &\times \Pr(\beta' x_{1j} + \varepsilon_{1j} < t_1, \beta' x_{2j} + \varepsilon_{2j} < t_2)^{NN} \\ &\times \Pr(\beta' x_{1j} + \varepsilon_{1j} < t_1, \beta' x_{2j} + \varepsilon_{2j} \geq t_2)^{NY} \end{aligned} \quad (7)$$

Assuming the error terms ε_{1j} and ε_{2j} are normally distributed with mean zero and variances σ_1^2 and σ_2^2 , respectively, and the correlation coefficient between DC1 and DC2 is expressed by ρ . The Equation (7) can be estimated using the bivariate probit model (Cameron and Quiggan, 1994)⁴. Thus, we can get the estimators for the

⁴ A restricted version of the bivariate probit model is the interval data model (Hanemann et al., 1991).

constant α^* and the coefficients β_M^* 、 β_S^* 、 β_B^* , so that we can calculate the mean WTP:

$$E(WTP) = -\frac{\alpha^* + \beta_M^* E(M) + \beta_S^* E(S) - \frac{n_1}{n}}{\beta_B^*} \quad (8)$$

where β_M^* 、 β_S^* 和 β_B^* are the estimated coefficients for consumers' demographic characteristics, supermarket's characteristics and the bids respectively. $E(\bullet)$ represents the mean of the corresponded variables. n is the whole sample, n_1 is the amounts of respondents whose answer is “yes”.

Data Description

Data used in this paper is from a survey of willingness to pay for “additive-free Mooncakes” in Beijing, conducted by the School of Agricultural Economics and Rural Development at Renmin University of China in October 2006. Using the face-to-face interview, this survey covered the main areas of Beijing. Based on the results of pre-survey and the study of Cooper (1993), we finally adopted three sets of bids (1、1.6、2.5);(1.6、2.5、1);(2.5、3.5、1.6). This survey includes 294 effective samples. Table 1 shows the descriptive statistics of the sample.

[Insert Table 1 & Figure 1]

From Table 1, we can find that the sample size is distributed evenly in each set of the bids, and each set has about 100 questionnaires. However, as shown in Figure 1, we find that the number of answering “yes-yes” is a little bit high, and the proportion

is 65.31%, perhaps resulting from (1) yea-saying bias and (2) starting point bias as Ready Buzby and Hu stated (1996). Also, the number of the respondents with a bachelor degree shares about half of the sample, and there may exist some bias in the sample, even though China's overall educational level is increasing. However, Beijing is a cultural center in China which might be another reason to explaining a higher proportion of high-educated people. In our sample, about 58.8% are women; it might be explained by the fact that women usually play a more important role in family food shopping. We also find that the average age of respondents is 34.87 years old, and the median and mode are 30-year-old and 23-year-old respectively; and it might indicate that the analysis in this study might bias to the youth.

We choose monthly income as the indicator of family welfare status. In our sample, monthly income below 3,000 Yuan accounts for 53.1%, and above 8,000 Yuan only 7.1 percent. There are 127 or 43.6% unmarried persons and the rest 164 or 56.4% are married. This proportion is close to some existing studies (Zhou et al 2006). We have two questions to survey the “sensitive” groups (Zeng et al., 2007); and in the survey, 32.3% of the families have children under 12, while 62.9% have old people above 65, because the two groups within the family may affect the respondents’ preferences.

In our survey, consumers’ concern about food safety is measured by five-point Scale. Statistic results show that, the level of concern is very high. We use the question “have you heard of the incidents of unqualified Mooncakes?” to specifically

measure the level of consumers' cognition of food safety, and the majority of consumers have heard some negative incidents. Mooncakes are very traditional pastries in China, and our survey finds that 82% of the respondents consume Mooncakes during the Mid-Autumn Festival.

In addition, consumer habits may differ in different supermarkets due to the hierarchical effect. For instance, the size of the supermarket may matter (Zeng et. Al 2008). We use small/medium/large to measure the size of supermarkets. In our survey, we request that every district must have at least one of the three types (small, medium and large). Usually, there are fewer customers in the small-scale supermarket, so that we get relatively a small sample of respondents in the small supermarkets, only accounting for 16.3% of the whole sample.

Empirical results

As we have pointed out, there are many factors that influence consumers' willingness to pay for the food safety, including food's prices, the level of consumers' awareness, consumers' purchasing habits, consumers' socio-economic variables as well as the characteristics of supermarkets. At the same time, it is worth noting that pre-set value of the bids (Bid) will also have an impact on consumers' WTP. We also expect that consumers' response to additive-free Mooncakes is that the higher "price" (Bid) is, the lower the probability consumers are willing to pay for it.

[Insert Table 2]

We collect 294 valid observations. Table 2 gives the explanations to the variables included in the regression. Based on the theoretical framework above, we mainly concern about the influence of consumers' age and income on WTP, controlling other variables. The specific bivariate probit model ⁵ can be set as follows:

$$\begin{aligned}
 Wtp1 &= \alpha_0 + \alpha_1 Bid1 + \alpha_2 Income + \alpha_3 Age + \alpha_4 Age2 + \alpha_5 Elder + \alpha_6 Edu \\
 &\quad + \alpha_7 Favor + \alpha_8 Cog + \alpha_9 Concern + \alpha_{10} Size2 + \alpha_{11} Size2 + \varepsilon_1 \\
 Wtp2 &= \beta_0 + \beta_1 Bid2 + \beta_2 Income + \beta_3 Age + \beta_4 Age2 + \beta_5 Elder + \beta_6 Edu \quad (9) \\
 &\quad + \beta_7 Favor + \beta_8 Cog + \beta_9 Concern + \beta_{10} Size2 + \beta_{11} Size2 + \varepsilon_2
 \end{aligned}$$

Equation (9) is estimated by Stata 9.2 software. Table 3 reports the estimation results. The chi-square test for the model is statistically significant at 1%, which indicates that the model fits the data very well.

[Insert Table 3]

As shown in Table 3, the coefficients for Bid, Income, Age, Age2, Cog and Size2 are statistically significant. The most important results are: (1) income is positively correlated with the WTP; (2) there is an inverted-U-shaped relationship between age and WTP, and the turning point is at 32-year-old. Both the signs of the interested coefficients are highly consistent with the theoretic framework.

The level of consumers' cognition of food safety (Cog) has a positive impact on WTP. Compared with the consumers who have not heard of the incidents of unqualified Mooncakes, those who have heard are willing to pay a higher price for the additive-free Mooncake. Consumers who pay attention to the information of food

⁵ The dummy variables of "Baby" and "Marital Status" are dropped in the regression due to multicollinearity with age and "Elder".

safety also have higher level of awareness, and their demand for food will transfer from “eating fully”, “eating well” to “eating safely”. So, they will pay more for additive-free Mooncake.

Compared to those in the large-scale supermarkets, consumers shopping in the medium-sized supermarkets (Size2) have a lower probability of willing to pay. The greater the size of the supermarket is, the more consumers’ WTP for the additive-free Mooncake; and this is consistent with the reality. The reason may be that consumers who concern more about food safety usually have a higher possibility to choose the large-scale supermarket for shopping (Zeng et al., 2008).

The results also show that the coefficients for Edu, Elder, Freq, Concern, Size3 and other variables don’t have significant influence on WTP statistically.

One of the main purposes of food safety research is to calculate the value of WTP, and it can provide a benchmark for assessing the policy. According to the method of calculating the mean WTP as shown in equation (8), we can put the estimated coefficients and corresponding mean of all the variables into equation (8), and we calculated the mean WTP for the additive-free Mooncake is 5.80 Yuan per unit. Using Krinsky and Robb Monte Carlo simulation, we obtain a 95% confidence interval for WTP which is 5.34~6.43 Yuan (Jeanty 2007). It is a relatively high value for the willingness to pay for an additive-free Mooncake in China. The results also indicate that the consumers’ concern on food additives is very high in China and

therefore the benefits of the government regulation on food additives are also very high.

By the way, we added a question of willingness to pay for additive-free Mooncakes with an open-ended format in the same questionnaire as the previous studies did. We calculated the mean of the WTP of the open-ended format, and find that the WTP for the additive-free Mooncakes is about 5.5 Yuan, which is highly consistent with the above result from the DBDC format .

Conclusion

This study finds that income could be positively correlated with the WTP and there is an inverted-U-shaped relationship between age and WTP from a theoretical perspective. These hypotheses are tested by the survey data of WTP for additive-free Mooncakes from 294 customers in 25 supermarkets in Beijing. Such a theoretical framework provides a good benchmark of understanding WTP in the future study of food safety.

This study also indicates that consumers in Beijing are willing to pay 5.80 Yuan more for an additive-free Mooncake, which shows consumers' high concern about the problem of food preservatives is very high in Beijing. The result provides a good policy benchmark of the benefits from the government's regulation on food additives. On the other hand, it may stimulate producers to develop substitutional technologies for food additives to increase consumers' welfare.

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Table 1 Descriptive Statistics

Variables	Description	Num.	Percentage	Variables	Description	Num.	Percentage
B_i	1	87	29.6	B_i^u (or B_i^d)	0.5	14	4.8
	1.6	105	35.7		1	23	7.8
	2.5	102	34.7		1.6	94	32
Wtp1	no	58	19.7		2.5	83	28.2
	yes	236	80.3		3.5	80	27.2
Cognition	no	146	49.7		Wtp2	no	94
	yes	148	50.3	yes		200	68
Preference for Mooncake	Eat every year	176	59.9	Size of supermarket	large	127	43.2
	Eat frequently	65	22.1		middle	119	40.5
	Almost do not eat	53	18		small	48	16.3
Education	Illiterate	4	1.4	Gender	female	172	58.5
	Primary school	2	0.7		male	122	41.5
	Junior high school	59	20.1	Baby	no	199	67.7
	Senior high school	86	29.3		yes	95	32.3
	College	127	43.2	Elder	no	109	37.1
	Master and above	16	5.4		yes	185	62.9
The average monthly family income (Yuan)	<500	2	0.7	Age	<25	95	32.3
	500—1000	21	7.1		25-30	59	20.1
	1000—2000	54	18.4		31-40	49	16.6
	2000—3000	79	26.9		41-50	42	14.3
	3000—4000	49	16.7		51-60	26	8.9
	4000—5000	35	11.9		>60	23	7.8
	5000—8000	33	11.2	Concern for food safety	Very much	123	41.8
	>8000	21	7.1		Relatively	124	42.2
Marital status	Married	164	56.4	Average	31	10.5	
	Unmarried	127	43.6	Indifferent	8	2.7	
	others	0	0	Not care	8	2.7	

Table 2 Variables included in the Estimation

Variables	Description	Mean	Std. Dev.
Bid1	The base bid at the initial dichotomous choice question	1.7347	0.6088
Wtp1	Respondent's answer to the initial question, 1=yes, 0=no	0.8027	0.3986
Bid2	The follow up bid at the second dichotomous choice question	2.2718	0.9211
Wtp2	Respondent's answer to the second question, 1=yes, 0=no	0.6803	0.4672
Income	The average monthly family income (Yuan) 1~8 represents eight levels from low to high, respectively	4.6803	1.6952
Age	Respondent's age	34.8742	15.0651
Age2	The square of age	1442.3910	1281.2420
Elder	Whether respondent's family have old people above 65, 1=yes, 0=no	0.6293	0.4838
Edu	Respondent's education, 1~6 represents six levels from low to high, respectively	4.2857	0.9599
Freq	The frequency of eating Mooncakes in the Mid-Autumn Festival, representing the level of consumers' preference 1= Eat every year, 2= Eat frequently, 3= Almost do not eat	1.5816	0.7784
Cog	Have you heard of the incidents of unqualified Mooncakes? representing the level of cognition, 1=yes, 0=no	0.5034	0.5008
Concern	Respondent's concern for food safety, 1=Very much, 2=Relatively, 3=Average, 4=Indifferent, 5=Not care	1.8231	0.9215
Size1	Large-scale supermarket, 1=yes, 0=no	0.4320	0.4962
Size2	Medium-sized supermarket, 1=yes, 0=no	0.4048	0.4917
Size3	Small-scale supermarket, 1=yes, 0=no	0.1633	0.3702

Table 3 Estimation of the Bivariate Probit Model

Variables	Coef.	Stand Error
Bid (1&2)	-0.3830	0.0448***
Income	0.1462	0.0473***
Age	0.0504	0.0270*
Age2	-0.0008	0.0003**
Elder	-0.0973	0.1588
Edu	0.0001	0.0835
Freq	-0.0244	0.1002
Cog	0.3552	0.1534**
Concern	-0.0581	0.0794
Size2	-0.2976	0.1659*
Size3	-0.2696	0.2221
Constant	0.3275	0.6460
ρ	0.9989	0.6690
Log Likelihood	-270.9069	
Mean / Median WTP	5.80	
95% C.I.	5.34—6.43	
Number of Obs.	294	

Note: (***), (**), (*) denote significance at the 1%, 5% and 10% levels, respectively.

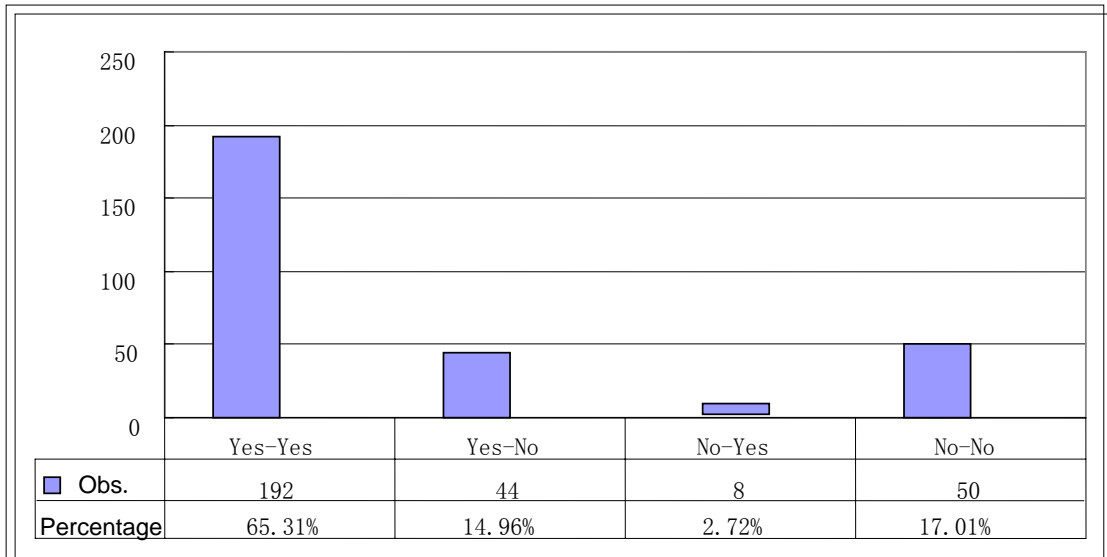


Figure 1 Distribution of WTP