Abstract — Agricultural production relies to a great extent on biological processes in natural environments. In addition to volatile prices, it is thus heavily exposed to risks caused by the variability of natural conditions such as rainfall, temperature and pests. With a view to the apparently lacking support of risky farm production program decisions through formal planning models, the objective of this paper is to examine whether, and eventually by how much, farmers’ “intuitive” program decisions can be improved through formal statistical analyses and stochastic optimization models. In this performance comparison, we use the results of the formal planning approach that are generated in a quasi ex-ante analysis as a normative benchmark for the empirically observed ones. To avoid benchmark solutions that would possibly exceed the respective farmer’s risk tolerance, we limit the formal search to a subset of solutions that are second-degree stochastically dominant compared to the farmer’s own decision. We furthermore compare the suitability of different statistical (time series) models to forecast the uncertainty of single gross margins.

Keywords — stochastic optimization, program planning, time series analysis

1. INTRODUCTION

Agricultural production relies to a great extent on biological processes in natural environments. It is heavily exposed to the variability of natural conditions such as rainfall, temperature and pests. Due to climate change these risks are likely to increase in the future, especially in developing countries. While neither being able to avoid production nor market risks, farmers always have adapted, and always will adapt, to changing conditions. One important strategy is the adaptation of the production program. Hence, the question arises whether the decisions associated with such adjustment processes can be supported by formal statistical analysis and stochastic program optimization.

Optimization procedures have been receiving a lot of attention in agro-economic research and teaching for several decades (see e.g. [1], [2], [3]). Nonetheless, even in developed countries, where technological innovations in general have been readily adopted by farmers, formal optimization has scarcely found its way into the farm planning practice. With regard to teaching the question has even been raised whether valuable time at universities is wasted with linear optimization [4].

Questioning the extra value of formal risk programming, a widespread opinion is that farmers - in particular with regard to program decisions in crop farming - make near-optimal choices based on experience, incremental learning and simple heuristics (see also [5]). This means that information such as the available acreage, the labour capacities, the potential cropping activities, the crop rotation requirements, the risk associated with various production activities and the individual risk aversion are considered implicitly in the mental model of the farmer. In contrast to that, the planning of the production program based on formal optimization procedures requires the explicit definition of the set of choices and restrictions and the objective function. Especially the difficulties arising from an explicit quantification of the uncertainty of single gross margins (see [6]) and the subjective risk attitude of the decision maker (see [7], [8]) are reasons why optimization has scarcely found its way into on-farm planning.

With a view to the apparently lacking support of risky farm decisions through formal planning models, the objective of this paper is to examine empirically whether, and eventually by how much, farmers’ “intuitive” program decisions can be improved through adequate and manageable stochastic optimization procedures. A final answer to this question requires the consideration of two issues (see e.g. [9]): (i) the overall performance of the formal planning tool if adequately used, and (ii) its practical usefulness and de facto po-
tential to increase profits given both the costs associated with using it and the cognitive constraints of potential users. This paper addresses only the former issue. It compares farmers' empirical performance with the performance that could have been realized if formally optimized programs had been implemented. We are the first, to our knowledge, to use the results of theoretical models generated in a quasi ex-ante analysis as a normative benchmark for empirically observed ones. From a methodical perspective, this requires the solution of two problems: firstly, the development of a manageable method to consider farmers' risk attitudes in the formal planning model and secondly, the examination of the suitability of different statistical models to quantify the uncertainty of single gross margins.

The rest of the paper is structured as follows: the database and the method are described in section two. This includes the statistical analysis of the single gross margin time series as well as the formal optimization model and the procedure of the benchmark comparison. Section three presents the results of the benchmark comparison. Finally, we provide conclusions in section four.

II. DATA AND METHOD

Our exemplary empirical analysis is a case study looking at the performance of four German cash crop farms over a period of six years. In the following we explain the empirical database in detail. Afterwards we describe the statistical models that are alternatively used to quantify the volatility of the single gross margins of each crop. The optimization model is described at the end of the section. Unlike the optimization approaches proposed until now, our approach generates a single alternative production program which has the following characteristics:

1. The uncertainty of the single gross margins (of wheat, barley etc.) is quantified through a statistical analysis of individual farm data. The resulting probabilistic information is used for the required one-year-ahead forecasts that are fed into the risk programming model.

2. The farmer's risk attitude is seen as being reflected by his choice of production program and his apparently accepted variance of the total gross margin. We consider this expression of his risk attitude and use the observed variance as an upper bound in the optimization model.

Finally, for each of the 24 planning occasions, the total gross margins which could have been realized if the formally optimized programs had been implemented are then ex-post compared to those that were actually realized by the farmers.

A. Empirical database

We are investigating four cash crop farms in Brandenburg, North-East Germany. Farms 1, 2, and 3 are about 50 km west of Berlin, and farm 4 is about 100 km north of Berlin. The acreage of these farms has virtually not changed over the last six years. Their average size amounts to 729 ha (farm 1), 1111 ha (farm 2), 1210 ha (farm 3), and 175 ha (farm 4). With the farm owner/manager, three permanent workers are engaged on farm 1, five on farm 2, four on farm 3, and one on farm 4. The major production activities considered by farmers 1, 2, and 3 include winter and spring wheat, winter rye, winter and spring barley, winter canola, corn, and non-food canola or set-aside land. Having an otherwise similar crop mix, farmer 4 does not include spring crops in his repertoire.

For all farms, minor crops such as alfalfa, oil flax or peas are excluded from the performance comparison. These crops represent something close to hobbies, rather than serious production activities. Proportionately they are almost irrelevant on all farms. On the contrary, sugar beets are very relevant and profitable without question. They are left out of the analysis in that their acreage is a priori presumed to amount to the maximum level defined by the production quota allocated to each farm. Excluding both hobby activities and the most competitive crop from the model does not impede the insight to be gained from the analysis. It only leads to a shift in the level of both the empirical and the optimized total gross margins and is therefore irrelevant for the performance comparison.

The farmers were interviewed with respect to their factor endowment (human resources and farm land) over the last six years. We also inquired about the number of field working days, the maximum working hours per day, the time required for the various activities in the critical seasons (March/April, May/June, mid-July/mid-September, and mid-September/mid-November) and possible purchasing activities such as

1 See e.g. [10] for an investigation of the practical usefulness of decision tools in a virtual setting.
hiring seasonal labor. Furthermore, crop rotation constraints (minimum and maximum proportions of the particular crops) were taken into consideration.

Both production and market risks are embedded in the single gross margins of each crop. Thus, farm-specific time series of single gross margins are needed to specify the probabilistic information relevant to each farmer for his program decision. In principle, data sets should be as large as possible for time series analysis. However, due to the structural discontinuity at the beginning of the nineties (collapse of the centralized economy and transition to market economy), farm-specific gross margins before 1992/1993 in the new federal states of Germany often contain ambiguous information or are not available at all. Hence, while using individual farm data after 1992, we construct proxies for the years 1980 to 1992. These proxies are site-specific single gross margins which are based on yields obtained on comparable soils and natural conditions in the old federal states of Germany and on West German price data ([11],[12]).

B. Time series analysis

The time series available at the respective planning dates \( t' - 1 \) comprise the years \( t = 1980 \) to \( t' - 1 \). The considered planning dates \( t' - 1 \) are 1998, 1999, 2000, 2001, 2002 and 2003. Thus, a farm-specific time series of 19 data is available for each crop at the first planning date “fall 1998”. For each of the following planning dates, the available time series increases by one year. For each single gross margin and farm/year combination (1) static distributions, (2) linear time series models, and (3) unbiased (non-linear) time series models are tested as alternative statistical methods and thus forecasting models.

Forecasting model 1: static distributions

Forecasting model 1 determines a static parametric distribution for each of the single gross margins. According to the Chi-Square, Kolmogorov-Smirnov and Anderson-Darling tests, the normal distribution cannot be rejected for any of the considered single gross margin time series (at a significance level of 5%). However, when compared to the normal distribution, Beta, logistic and/or triangular distributions show a slightly better match with the empirical distributions in some cases. In agreement with the standard approach (cf. [13]), we nonetheless assume a normal distribution for all single gross margins.

Let \( GM_i^j \) be the gross margin per unit of production activity \( j \) observed at time \( t \). Then a static normal distribution for a single gross margin can be described as follows:

\[
GM_i^j = E(GM_i^j) + \chi_i^j = \frac{1}{N} \sum_{t'=1980}^{1980+N-1} GM_i^j + \chi_i^j, \quad (1)
\]

with \( N = t' - 1980 \)

The future gross margin \( GM_i^j \) results from its expected value \( E(GM_i^j) \) and a \( N[0,\sigma_i^j] \)-normally distributed random component \( \chi_i^j \) (error term, white noise). For a static distribution, the expected value corresponds to the mean. The standard deviation of the error term \( \sigma_i^j \) reflects the standard deviation of all observed values. We introduce notation (1) - even though it is rather unusual for simple distributions - on grounds of consistency with the notation for the stochastic processes.

Forecasting model 2: linear time series models

Allowing for stochastic processes implies that one examines the time-dependent pattern of random variables through time series analysis. Abstracting from discontinuities, a stochastic process represents the best estimation with regard to the variable’s distribution at future points in time. Auto-Regressive-Integrated-Moving-Average models of the order \( p, d \) and \( q \) (ARIMA\((p,d,q)\)-models) are linear time series models. Due to their flexibility, ARIMA\((p,d,q)\)-models are used to represent a multitude of stochastic economic processes (cf. e.g. [14]).

The ARIMA\((p,d,q)\)-model that fits best to a particular time series can be determined with the Box-Jenkins test procedure [15]. According to this test, an \( AR(p) \)-process results for all single gross margin time series in our illustrative analysis:

\[
GM_i^j = E(GM_i^j) + \chi_i^j = \alpha_i^j + \sum_{u=1}^{p} \alpha_u^j \cdot GM_{t-u}^j + \chi_i^j, \quad (2)
\]

with \( \sum_{u=1}^{p} |\alpha_u^j| < 1 \)

\( \alpha_i^j \) denotes a constant, \( \alpha_u^j \) denote the weight factors that need to be estimated for the last \( p \) observation values \( GM_{t-u}^j \), and \( \chi_i^j \) describes a \( N[0,\sigma_i^j] \)-normally distributed error term. It should be noted that

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the expected value as well as the error term can, and mostly will, differ from the ones obtained with the static distribution model even if the same data are analyzed.

**Forecasting model 3: non-linear time series models**

Recent research emphasizes the importance of non-linear dependencies in time series (e.g. [16]). Non-linearity cannot be identified with standard statistical procedures such as the Box-Jenkins test procedure (cf. [15]). These “conventional” tests presume a priori linearity. Hence, more sophisticated statistical methods, which do not predispose linearity but facilitate the unbiased identification of both linear and non-linear processes, are needed. The method of “heuristic self-organizing time series models” originally described by [17] offers an alternative for the identification and specification of non-linear stochastic relationships (cf. [18]). A special class of self-organizing algorithms is the so-called Group Method of Data Handling (GMDH). GMDH-algorithms combine the connectionistic approach to artificial neural networks with the classical method of regression. They generate general polynomial process models (cf. [18]: 77):

\[
GM_j = \gamma_0 + \sum_{a=1}^{p} \sum_{l=1} \gamma_{a,l} \cdot GM_{t-a} + \chi^t
\]

\[
+ \sum_{a=1}^{p} \sum_{l=1} \sum_{v=1} \gamma_{a,v} \cdot GM_{t-a} \cdot GM_{t-v} + \sum_{a=1}^{p} \sum_{l=1} \sum_{v=1} \gamma_{a,v} \cdot GM_{t-a} \cdot GM_{t-v} \cdot GM_{t-v} + ... \tag{3}
\]

These polynomials take into account up to \( p \) preceding values with different weights as well as non-linear terms, the potential number of which exponentially increases with the number of the considered preceding values. Thus, the functional form of the polynomial may easily get quite large. Its general structure can be understood as an AR(\( p \))-process (upper line) which is extended by a non-linear component (lower line). While the GMDH-model has the principal capacity to account for any kind of distribution, we assume normally distributed error terms since the results of the Chi-Square, the Kolmogorov-Smirnov and the Anderson-Darlings tests show that the normal distribution cannot be rejected at a significance level of 5%.

**C. Optimization model and performance analysis**

Fig. 1 illustrates the methodical steps of data collection and processing that are needed to optimize the production program.

As described above, data are collected for each year and farm. This concerns the realized single gross margins up to the planning date, the constraints such as acreage, labor and crop rotation requirements, the empirically observed production program, and the single gross margins realized in the target year. In the next step the single gross margin time series up to the respective planning date are statistically analyzed. The resulting probabilistic information includes the correlations, variances, and expectation values of the single gross margins. These values as specified separately in each of the three variants of statistical analysis. Depending on the consequently differing model inputs, the optimization model, of which the general structure is described below, will thus provide three alternative planning variants for each target year and farm.

As described above, data are collected for each year and farm. This concerns the realized single gross margins up to the planning date, the constraints such as acreage, labor and crop rotation requirements, the empirically observed production program, and the achieved single gross margins realized in the target year. In the next step the single gross margin time series up to the respective planning date are statistically analyzed. The resulting probabilistic information includes the correlations, variances, and expectation values of the single gross margins. These values are specified separately in each of the three variants of statistical analysis. Depending on the consequently differing model inputs, the optimization model, of which the general structure is described below, will thus provide three alternative planning variants for each target year and farm.

**Optimization model**

The farm-specific constraints and the expectation values, variances and correlations of the single gross margins are fed into a quadratic optimization model. While considering the probabilistic information derived from the time series of single gross margins up to each planning date, we do not model the actual program decisions as dynamic or multi-period optimization problems.
Instead, the program for each year is optimized subject to the constraints that are effective in the respective year as indicated by the individual farmer. Interdependencies between the activities of subsequent years are considered through crop rotation constraints that may vary from year to year. For each farm and planning date $t^* (1998$ to $2003)$, formally optimized alternative programs are determined according to expectations regarding the respective target year $t^* (1999$ to $2004)$.

The optimization model can be described as follows:

$$\text{max } E(TGM^*) = \sum_{j,t} E(GM^*_{jt}) \cdot x^*_{jt}$$

s.t.

$$\sum_{j,t} a^*_{jt} \cdot x^*_{jt} \leq b^*_{jt}, \quad i = 1,2,\ldots,I$$

$$V \leq V_{emp}$$

$$x^*_{jt} \geq 0, \quad j = 1,2,\ldots,J$$

$E(TGM^*)$ denotes the expected total gross margin in the respective target year of the optimized production program. The objective function coefficients $E(GM^*_{jt})$ represent the expected gross margins per unit of production activity $j$. $x^*_{jt}$ describe the levels of the production activities. $b^*_{jt}$ denote the capacities (restrictions), and $a^*_{jt}$ represent the capacity requirements per unit of production activity. $V_{emp}$ is the total gross margin (TGM-) variance inherent to the empirically observed production program. $V^*$, in contrast, denotes the TGM-variance of the optimized program. Using $V_{emp}$ as an upper bound ensures that the reflection of the farmer’s risk attitude as observed in his own choice of production program is taken into account in the optimization. $^2$

The calculation of the variance is based on the results of the statistical analysis. Since the $J$ random variables (single gross margins) are normally distributed and additively combined, the TGM-variance of the empirical production program $V_{emp}$ can be calculated in a way analogous to a portfolio consisting of $J$ asset positions (cf. [19]: 150):

$$V_{emp} = \sum_{j=1}^{J} \left( x^*_{i},emp \cdot \sigma^*_{j} \right)^2 + 2 \sum_{j=1}^{J} \sum_{k<j} x^*_{i},emp \cdot \sigma^*_{j} \cdot x^*_{i},emp \cdot \sigma^*_{k} \cdot \rho^*_{j,k}$$

$^2$ We omit the subscript $t^*$ when referring to $E(TGM^*)$, $V_{emp}$ and $V$. They always correspond to the respective target year.

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\( \rho_{jk} \) denote the correlation coefficients between the single gross margins \( j \) and \( k \) (in the case of static distributions), or between their error terms (in the case of stochastic processes). \( \sigma_{j}^2 \) and \( \sigma_{k}^2 \) describe the respective standard deviations. The correlation coefficient and the standard deviation are determined on the basis of the data collected up to the particular planning date. \( x_{j, \text{emp}}^* \) and \( x_{k, \text{emp}}^* \) respectively, represent the weight (acreage) of the production activities in the farmer’s empirically observed program. The TGM-variance of the optimized program \( V \) is to be determined in a way analogous to (5). One merely needs to replace the observed production levels \( x_{j, \text{emp}}^* \) and \( x_{k, \text{emp}}^* \) by the optimized production levels \( x_{j, \text{emp}}^* \) and \( x_{k, \text{emp}}^* \).

Crop rotation requirements justifying the model’s linearity assumption (i.e. the assumption that resources used and revenues derived are linearly linked with the activity level) are represented in each year’s optimization model as upper and lower bounds on the crop acreages. In a first optimization run, we use the maximum and minimum crop proportions as indicated by the farmers in the interviews as constraints. We then cross-check with farmers whether they consider the resulting programs feasible and consistent both with their preferences and the specific farm situation. If necessary, constraints are changed or added to the model. This includes changes of the feasible minimum and maximum crop proportions which may vary depending on the production program(s) of the previous year(s). This feasibility check is repeated until no further modifications are needed. This step by step procedure ensures that the optimized programs, while differing from the farmers’ realized programs, are both feasible and acceptable for the real decision-makers.

**Benchmark comparison**

We finally calculate the normative benchmark, i.e. the total gross margin \( TGM \) which would have been realized in the target year if the optimized production program had been implemented:\n
\[
TGM = \sum_{j=1}^{l} GM_{j}^l \cdot x_{j}^* \tag{6}
\]

\( GM_{j}^l \) indicate the actually realized single gross margins in the target year. \( x_{j}^* \) denote the planned proportions of crops as derived from the formal planning model. An ex-post comparison between the benchmark \( TGM \) and the farmer’s realized total gross margin \( TGM_{\text{emp}} \) reveals whether an extra value could have been derived from the formal planning model. \( TGM_{\text{emp}} \) is to be determined in a way analogous to (6). One merely needs to replace the optimized production level \( x_{j}^* \) by the farmer’s observed production levels \( x_{j, \text{emp}}^* \). Regarding the validity of the performance comparison, it is to be noted that no informational advantage was accorded to the formal planner. The single gross margins realized in the target year are only used for the final benchmark comparison.

A brief comparison with conventional quadratic risk programming

While bearing resemblance to conventional expected-value variance (EV-) models the proposed optimization approach exhibits some particularities: EV-models handle the problem of unknown risk attitudes by carrying out variant calculations regarding the risk aversion coefficient, thus providing a set of efficient combinations of variance and expected total gross margin (cf. Fig. 2, left).

Aiming at supporting practical decision-making, we maximize the expected total gross margin subject to the constraint of not exceeding the empirically observed willingness to accept risk. This means taking the variance of the total gross margin \( V_{\text{emp}} \), inherent to the production program chosen by the farmer, as an observable, albeit incomplete, reflection of his subjective risk attitude (cf. Fig. 2, right).

We do not argue that this reflection represents the farmer’s risk attitude completely. Nor do we presume that our procedure necessarily identifies the production program that maximizes the farmers’ utility. We rather focus on the manageability and applied usefulness of the approach which allows for a clear endogenous recommendation by reducing the efficient solution set to one single combination of variance and expected total gross margin. Technically speaking, we limit the set of feasible solutions to those yielding a higher or identical expected total gross margin at a lower or identical variance as the one previously accepted by the farmer. This is equivalent to limiting the formal search to an identifiable subset of solutions that are second-degree stochastically dominant compared to the farmer’s own decision. We are thus sure to increase (or at least meet) the farmer’s expected utility compared to the one resulting from his own program.
Fig. 2 Classical procedure in the EV-model (left) vs. practical decision support (right)

III. RESULTS

A. Profitability comparison

Until now, none of the decision-makers of the considered farms has been using formal optimization procedures to determine the production program. In fact, production planning is completely based on non-formal planning and intuition. Table 1 compares the average annual total gross margins realized by the farmers with those that could have been realized if the formally optimized programs had been implemented.

As an additional point of reference, the first row of Table 1 shows the hypothetical room for improvement (extra value) that would be generated if one had a “perfect forecasting model”, i.e. a model, which exactly predicts the single gross margins in the respective target years. If such perfect information was available, the total gross margin could be improved by 15.9% on average over all farms and all years. When interpreting this figure one should note that a perfect prognosis is never possible in reality because any time series contains unsystematic and unpredictable random errors. Hence, the figures depicted in the first row of Table 1 are only a first hint that it is worthwhile examining the potential for improvement that might be generated by using formal procedures of statistical analysis and stochastic optimization.

Rows 2 to 4 of Table 1 depict the extra value that is added by formal planning based on the three different forecasting models described above. Only farm 2 could have improved its average annual performance considerably through the formal approach 1. The increase would have been quite low in farm 4, and on farm 1 and 3 the average annual performance would even have been inferior to that of the actual programs. Although not shown in Table 1, a look at the crop mix of farm 2 reveals that the optimized program suggests a sharp decrease of the proportion of corn. Farmer 2 has in reality planted corn on 15% of his acreage on an average whereas it is virtually irrelevant in the programs of the other three considered farms. Its comparative competitiveness being low, the 15% proportion of corn on farm 2 must be interpreted as a serious planning mistake. It seems that the mistake is so serious that it could even have been alleviated through an optimization approach based on an ill-founded forecasting model. Despite its positive effect on farm 2, the overall change of performance caused by approach 1 over all farms and all years is nearly zero. We must therefore conclude that a “standard” risk programming approach which incorporates risk through static distributions may even be inferior to informal decision-making of reasonably good farm managers.
Table 1 Average annual total gross margins (in €) of realized and optimized production programs

<table>
<thead>
<tr>
<th></th>
<th>Farm 1</th>
<th></th>
<th>Change</th>
<th>Farm 2</th>
<th></th>
<th>Change</th>
<th>Farm 3</th>
<th></th>
<th>Change</th>
<th>Farm 4</th>
<th></th>
<th>Change</th>
<th>Sum over all farms</th>
<th></th>
<th>Change</th>
</tr>
</thead>
<tbody>
<tr>
<td>&quot;Perfect forecasting model&quot;</td>
<td>323 002</td>
<td>360 063</td>
<td>11.5%</td>
<td>325 636</td>
<td>413 944</td>
<td>27.1%</td>
<td>432 118</td>
<td>482 390</td>
<td>11.6%</td>
<td>78 894</td>
<td>87 826</td>
<td>11.3%</td>
<td>1 59 650 1 344 224</td>
<td>15.9%</td>
<td></td>
</tr>
<tr>
<td>Approach 1: assuming static distributions</td>
<td>323 002</td>
<td>307 874</td>
<td>-4.7%</td>
<td>325 636</td>
<td>360 656</td>
<td>10.8%</td>
<td>432 118</td>
<td>420 384</td>
<td>-2.7%</td>
<td>78 894</td>
<td>80 286</td>
<td>1.8%</td>
<td>1 59 650 1 169 200</td>
<td>0.8%</td>
<td></td>
</tr>
<tr>
<td>Approach 2: assuming linear time series</td>
<td>323 002</td>
<td>344 679</td>
<td>6.7%</td>
<td>325 636</td>
<td>373 466</td>
<td>14.7%</td>
<td>432 118</td>
<td>450 237</td>
<td>4.2%</td>
<td>78 894</td>
<td>82 878</td>
<td>5.1%</td>
<td>1 59 650 1 251 259</td>
<td>7.9%</td>
<td></td>
</tr>
<tr>
<td>Approach 3: allowing for non-linear time series</td>
<td>323 002</td>
<td>349 028</td>
<td>8.1%</td>
<td>325 636</td>
<td>374 385</td>
<td>15.0%</td>
<td>432 118</td>
<td>446 061</td>
<td>3.2%</td>
<td>78 894</td>
<td>82 619</td>
<td>4.7%</td>
<td>1 59 650 1 252 093</td>
<td>8.0%</td>
<td></td>
</tr>
</tbody>
</table>

Table 2 Average crop proportions (in %) realized by farmers compared to those derived from the superior planning approach 3

<table>
<thead>
<tr>
<th></th>
<th>Farm 1</th>
<th></th>
<th>Increase*</th>
<th>Decrease*</th>
<th>Farm 2</th>
<th></th>
<th>Increase*</th>
<th>Decrease*</th>
<th>Farm 3</th>
<th></th>
<th>Increase*</th>
<th>Decrease*</th>
<th>Farm 4</th>
<th></th>
<th>Increase*</th>
<th>Decrease*</th>
<th>Sum over all farms</th>
<th></th>
<th>Increase*</th>
<th>Decrease*</th>
</tr>
</thead>
<tbody>
<tr>
<td>Winter wheat</td>
<td>-1.8</td>
<td>3</td>
<td>3</td>
<td>3.8</td>
<td>3</td>
<td>3</td>
<td>8.7</td>
<td>3</td>
<td>2</td>
<td>-1.6</td>
<td>2</td>
<td>4</td>
<td>0.8</td>
<td>11</td>
<td>12</td>
<td>12</td>
<td></td>
<td></td>
<td></td>
<td></td>
</tr>
<tr>
<td>Spring wheat</td>
<td>0.9</td>
<td>1</td>
<td>0</td>
<td>0.0</td>
<td>0</td>
<td>0</td>
<td>-</td>
<td>0</td>
<td>0</td>
<td>0.6</td>
<td>1</td>
<td>0</td>
<td>0.4</td>
<td>2</td>
<td>0</td>
<td>0</td>
<td></td>
<td></td>
<td></td>
<td></td>
</tr>
<tr>
<td>Winter rye</td>
<td>16.6</td>
<td>5</td>
<td>0</td>
<td>22.1</td>
<td>6</td>
<td>0</td>
<td>7.3</td>
<td>5</td>
<td>1</td>
<td>11.4</td>
<td>6</td>
<td>0</td>
<td>16.1</td>
<td>22</td>
<td>1</td>
<td>1</td>
<td></td>
<td></td>
<td></td>
<td></td>
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<tr>
<td>Winter barley</td>
<td>-13.6</td>
<td>1</td>
<td>5</td>
<td>-5.6</td>
<td>0</td>
<td>5</td>
<td>-17.5</td>
<td>0</td>
<td>6</td>
<td>-7.2</td>
<td>0</td>
<td>6</td>
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<td>1</td>
<td>22</td>
<td>22</td>
<td></td>
<td></td>
<td></td>
<td></td>
</tr>
<tr>
<td>Spring barley</td>
<td>-1.2</td>
<td>0</td>
<td>1</td>
<td>2.5</td>
<td>3</td>
<td>0</td>
<td>-</td>
<td>1</td>
<td>0</td>
<td>1.5</td>
<td>1</td>
<td>0</td>
<td>1.2</td>
<td>5</td>
<td>1</td>
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* Number of years - out of the six planning years considered for each farm - in which changes should have been made according to formal modeling. Cases in which the crop proportion remains unchanged can be calculated as residuals.
On the contrary, very encouraging results were found for the other two planning approaches (see the third and fourth row of Table 1). The average total gross margin could have been improved noticeably on all four farms if farmers had used formal optimization based on probabilistic information derived from systematic time series analysis: farmer 1, for instance, achieved in reality an average total gross margin of €323,002. Optimized production programs based on forecasting model 3 would have increased that amount to €349,028. In other words: the average annual total gross margin for farm 1 could have been increased by 8.1% (or €26,026 per annum). For farm 2 the respective figures amount to 15.0% (or €48,749 per annum), in the case of farm 3 to 3.2% (or €13,943 per annum), and for farm 4 to 4.7% (or €3,725 per annum).

Averaged over all farms, the potential for improvement compared to the farmers’ actual programs amounts to nearly 8% (7.9% for approach 2, and 8.0% for approach 3). That is, even though much more effort went into the statistical analysis of approach 3, the results are only slightly superior to those derived from linear time series models. This might be interpreted as preliminary evidence that, with GMDH models which allow for non-linear time series, we have arrived at a stage of model sophistication where the marginal returns of increasing planning efforts sharply decrease.

Going beyond the consideration of averages, the most essential results can be summarized as follows: firstly, the optimized programs derived from approach 3 surpasses in performance the empirical ones in 23 out of the 24 cases. It can be added that the same applies to approach 2. Approach 1, in contrast, outperforms the realized programs in merely 11 out of the 24 cases. Second, being the odd exception, the total gross margin realized in the year 2001 by farmer 3 is higher than the one that would have been achieved with approach 3. At the planning date in the year 2000, however, the expected total gross margin of the optimized production program was 1.5% higher than the one of the farmer’s program. This underlines the well-known fact that in an uncertain environment a faulty decision may by chance result in higher profits, but uninformed choices will not be superior in the long run.

B. Comparison of production programs

Table 2 provides a rough characterization of production programs by comparing the farmers’ crop mix with the optimized and more profitable mix that would have been derived from approach 3. While only commenting on the benchmark comparison with approach 3, identical conclusions are to be drawn from the comparison with approach 2. Comparing the programs and identifying the main divergences provides first evidence for systematic planning mistakes made by farmers.

The most noticeable result of the comparison is that, according to the superior approach 3, the share of (winter) rye should have been increased considerably in all farms and in 22 out of the 24 cases modeled. Rye is very drought-resistant. From a crop science perspective it is thus especially well suited for the natural conditions of Brandenburg which is characterized by low and uncertain rainfalls and poor and quickly draining soils. Searching for the most common and apparent change suggested by the formal approach 3, we may conclude that, besides farmer’s 2 particular planning mistake regarding corn, all farmers should reduce their proportion of winter barley in favor of rye.

This result is an indication for systematic planning errors made by the farmers in the past. Evolutionary economics (cf. e.g. [20]) could be called upon to explain this finding: even important economic decisions are often not supported by formal decision models. Rather, they are based on the decision-maker’s experience gathered in the course of the past decades and on simple heuristics such as “never make any decision that differs very much from past ones” (cf. e.g. [5]). In an environment where the relative competitiveness of different crops, for instance, changes quickly, decision-makers may thus not be quick enough to adapt to changed conditions. In other words: we might ask the question whether (boundedly rational) farmers making routine production program decisions learn too slowly.

In the considered context it seems reasonable to speculate that farmers have not yet adapted their routines to account fully for two major changes of their relevant environment: on the one hand, the enormous progress in rye breeding over the last years which brought rye production up to competitive levels, and, on the other hand, the increasingly precarious rainfalls in Brandenburg, possibly caused by climatic change. Slow learning and adaptation, in turn, justifies the use of formal decision aids by farmers and management consultants.

C. Exploratory research regarding model robustness

Using a very cautious wording, the above-presented results provide evidence that cases exist where an improvement of farm production program planning can
be provided through formal stochastic optimization. However, the variability of circumstances as well as the small number (24) of analyzed cases does not allow for statistical generalizations. Aiming to identify directions for relevant future research we carried out a preliminary exploration regarding the influence of differing soil qualities on the model’s robustness to provide superior results.

To do so, in addition to the four farms analyzed above we looked at five more crop farms in North-East Germany. Since farm-specific single gross margin data had only been recorded on these five farms since 1998, we only considered the target years 2002, 2003 and 2004. Forming equal-sized groups, we attributed three farms to category I (low-quality soil; below 30 points according to relative the German soil quality classification scheme from zero to 100 points), three to category II (medium-quality soil; between 31 and 44 points), and three to category III (high-quality soil; between 45 and 52 points).

As shown in Table 3 in the three analysed years the largest performance improvement of 12% on average could have been realized in the three farms of the medium-quality soil category II. This figure is down to 3.5% in the high-quality soil category, and to 7.1% in the low-quality soil category. While not being able to fully explain this, some educated guesses may serve as hypotheses or research questions in further studies: Let us first look for a possible explanation for the lower improvement potential in the farms of high-quality soil category compared to those on medium-quality soils: the number of crops to be considered by farmers on high-quality soils may be very small and include only very few high-yield cropping activities. In other words: if the relative competitiveness of different crops is clearly differing and if the farmers’ respective knowledge is adequate to allow for a significant reduction of the planning complexity through an a priori exclusion of those crops that are “out of question”, little support is needed, and can be given, by a sophisticated planning approach.

Let us now search for an explanation for the lower improvement potential in the farms of the low-quality soil category compared to those on medium-quality soils: according to their empirically observed programs the latter also accept a higher total gross margin variance, which, in turn, can be attributed partly to a lower risk aversion and partly to a steeper risk efficient production frontier. The higher performance potential can thus be seen as an indication that the use of the production factor risk generates decreasing marginal returns, not only due to the decreasing slope of the production function, but also to the decreasing ability of boundedly rational decision-makers to realize the potential returns. This can again be attributed to the higher complexity of the planning problem.

| Table 3 Potential for improvement (in % of the total gross margin) in differing soil categories |
|---------------------------------|------------------|-------------------|
| category | Average change | Standard deviation |
| I (low-quality soil) | 7.1 | 6.3 |
| II (medium-quality soil) | 12.0 | 19.8 |
| III (high-quality soil) | 3.5 | 4.1 |

* Three farms and years; planning approach 3.

We may thus formulate as a plausible hypothesis that the room for improvement opened up by formal optimization increases with an increasing complexity of the planning problem.

IV. CONCLUSION

At first view, this paper seems to revive the discussion about the benefits of optimization models in applied agricultural program planning - a discussion that has been virtually closed in the past for the so-called good reason that agricultural practitioners have found neither need nor want to use formal approaches in on-farm planning. Our illustrative analysis of 24 farm/year combinations, however, suggests that a different attitude may be needed.

The formal approach used in this study provides practical assistance for dealing with the problem of individual risk attitudes. It includes the variance of the total gross margin inherent to the production program chosen by the farmer as an additional restriction in the stochastic optimization model. Any practical decision support procedure which relies on this approach thus requires that farmers first specify their “own” production programs without the formal planning aid, thus providing an observable reflection of their risk attitude. Afterwards, one can search for alternative programs - with the same or even less variance - lead to superior or at least equal expected total gross margin.
margins. With regard to decision theory it must be recognized that the recommended alternative does not necessarily represent the solution that maximizes the farmer’s utility. Instead the search is limited to second-degree stochastically dominant alternatives, thus providing decision support without needing to elicit the farmer’s risk aversion.

Our methodical comparison of different variants of statistical analysis (static distributions vs. stochastic processes) indicates that the extra value to be derived from formal optimization methods depends on the available data being adequately processed and used. Data may be time-dependent and exhibit a trend. This simple fact suffices to show that approaches which prima facie resort to the mean and variance of past values considered equal in weight may not represent good forecasting models in many cases. Inserting too simplistic assumptions into formal planning models may well cause their performance to be inferior to that of planning on a rule of thumb basis. This could be an explanation why some of the formal optimization approaches proposed in the past have not been accepted by farmers, and rightly so.

While the analyzed sample is small, the identified capacity to outperform farmers’ informal decisions provides first evidence that the efficiency of on-farm decision-making might be improved through formal optimization. However, this needs to be investigated through further research. The identified dimension and the continuity of the efficiency gains found in the case study warrant the effort to do so. In other words, the model’s robustness to provide superior results should be tested by applying it to a larger number of farms in different regions and with different sizes, production structures and operating figures. In this context the following extensions of the model may prove valuable:

- As long as one looks at program planning on large farms, basically any crop rotation requirement can be translated into a respective proportion of crops in any one year. This may be different on small farms where one needs to consider that fields, being of smaller and differing size, are not to be further subdivided for different crops. Furthermore, soil quality may differ from one field to the other. Extending the model to include such field-specific information requires additional effort but could be implemented in principle.
- We included the total gross margin variance inherent to the farmer’s observed program as a fixed restriction in the formal optimization model. Thus, we did not consider that farmers might be prepared, for instance, to accept some additional volatility if the expected increase of the total gross margin covers their risk premium. That is, we cannot be sure to have found the utility maximizing production program. This does not impede the insights and clear-cut results of the analysis. In fact, the consequence is just that there may be even more room for improvement. One could investigate this by letting farmers choose from alternatives derived from a stepwise relaxation of the variance restriction. However, this means abandoning the model-endogenous recommendation of a single superior solution.

- In our case study, extending the time series model to include non-linear structures did not add much extra value compared to simple linear models. This can be seen as evidence that increasing planning efforts and further model sophistication are not feasible due to decreasing marginal returns. It might nonetheless be worthwhile to search for models that perform better still. Such a search could include models which allow for process parameters that are variable over time. Explicit GARCH-models could, for instance, be used in the case of a time-variable variance (cf. [21]).

Decisions regarding the resources to be spent for planning are, like all economic choices, subject to efficiency considerations. Thus, additional efforts such as the introduction of formal and more sophisticated planning models need to be justified by additional benefits. Before plunging into any of the above-mentioned activities, it should be checked whether the gain in information justifies the additional costs. We can assume that the critical farm size allowing for sufficient economies of scale depends on the costs (including learning costs) associated with the introduction of optimization models. These costs, in turn, depend on the knowledge and the skills of farm managers and thus, amongst other things, on the quality of their training. Better trained agricultural managers and consultants will need less time and effort to adopt more sophisticated approaches because they have less learning costs.

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