Short and long-run returns to agricultural R&D in South Africa, or will the real rate of return please stand up?

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Abstract

This paper briefly presents the results of a total factor productivity (TFP) study of South African commercial agriculture, for 1947–1997, and illustrates some potential pitfalls in rate of return to research (ROR) calculations. The lag between R&D and TFP is analyzed and found to be only 9 years, with a pronounced negative skew, reflecting the adaptive focus of the South African system. The two-stage approach gives a massive ROR of 170%. The predetermined lag parameters are then used in modeling the knowledge stock, to refine the estimates of the ROR from short- and long-run dual profit functions. In the short run, with the capital inputs treated as fixed, the ROR is a more reasonable 44%. In the long run, with adjustment of the capital stocks, it rises to 113%, which would reflect the fact that new technology is embodied in the capital items. However, the long-run model raises a new problem since capital stock adjustment takes 11 years, 2 years longer than the lag between R&D and TFP. If this is assumed to be the correct lag, the ROR falls to 58%, a best estimate. The paper draws attention to the possible sensitivity of rate of return calculations to assumed lag structure, particularly when the lag between changes in R&D and TFP is skewed. © 2000 Elsevier Science B.V. All rights reserved.

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1. Introduction

Equivalent measures of technological change exist which are based on the dual relationships between the production, cost and profit functions. These measures are also equivalent to economic accounting measures, based on index number theory. Similarly, there are two econometric approaches to explain changes in agricultural productivity, which form the basis for calculating the returns to research (ROR). ¹ Evenson et al. (1987) call these the integrated approach (where the productivity-enhancing, or conditioning factors are included directly in a primal or dual representation of production) and the two-stage decomposition, in which changes in total factor productivity (TFP) are first calculated, and then explained, by the conditioning factors that are thought to account for growth.

¹ There is a huge literature on the returns to agricultural research. See, for example, Echeverria (1990) or Thirtle and Bottomley (1992).
In either case, changes in output, costs, profits, or TFP are usually explained by conditioning variables such as the stock of knowledge (accumulated research capital, generated by past research expenditures), extension services and farmer education, as well as changes in the weather. The basic argument is that R&D generates technology, extension diffuses it, and better-educated farmers are better at screening new technology. Consequently, they adopt technologies more quickly and also adapt technology, thereby adding an element of on-farm technology generation. For South Africa, spillovers through international technology transfers are also important, so international patents are included. Finally, the influence of the weather is considerable, so a weather index should reduce the unexplained errors.

Both integrated and two-stage approaches have advantages and disadvantages. The dual integrated approach has the advantage of minimizing restrictive separability assumptions, as well as avoiding the need for the assumptions of full static equilibrium, Hicks neutral technical change, and constant returns to scale, all of which are implicit in the two-stage approach. On the other hand, the two-stage method concentrates on the technology-related variables, so they can be modeled more satisfactorily. Thus, both routes are followed in this paper, which summarizes and develops past work in this area with respect to South Africa.

The primary focus of the paper is in illustrating some potential pitfalls in ROR calculations. Little attention is generally paid to correct lag length and shape parameters in the ROR literature, which are here found to be entirely different from the symmetric structures used previously and this is shown to change the ROR substantially. This result indicates the sensitivity of these ROR calculations to assumed lag structure, especially when the effects of agricultural R&D are skewed and have short lags as in South Africa's applied and adaptive research system.

Subsequent sections of the paper review the literature on the calculation of South African commercial TFP, and the necessary theory underlying short- and long-run generalized quadratic (GQ) residual profit functions. The rest of the required data, including international technology spillover variables, and their sources are then discussed, followed by the estimation, results, and calculation of ROR. In conclusion, we indicate some possible refinements to agricultural ROR calculations, without suggesting that the best methods are at hand.


Thirll et al. (1993) analyze TFP growth in South African agriculture for the period 1947–1992, using a Tornqvist–Theil index. The series has been updated to 1997 by Nick Vink at the University of Stellenbosch, in South Africa. Fig. 1 shows that the output index has grown by over 250% for the period, an annual rate of 2.8%. The index of inputs has more than doubled, growing at 1.5% a year, but considering the entire time period at once hides the fact that inputs grew at over 2.5% per annum until 1979 and since then have been falling at 0.5% per annum. This fall in inputs explains the recent growth in the TFP index. Over the full period, TFP grows rather slowly, at 1.2% per year, but there was no growth until 1965, then 2.13% per annum until 1983, a severe drought year, and a slower growth of 1.5% per annum since 1984.

To explain the causes of the aggregate TFP changes, it is necessary to look at the components of the index. Output can be decomposed into crops, horticulture and fruit, and animal products. Fruit and horticulture have grown more rapidly than crop output, while livestock outputs have grown the least. The growth rates were 3.9, 2.7 and 2.3%, respectively, and as a result, the share of horticulture and fruit has increased at the expense of the other enterprises. The indices behave in a reasonable manner; crop productivity later (Fig. 2).

On the input side of the index, non-farm intermediate inputs and capital inputs have been substituted for the primary inputs, especially labor. Thus, the cost share of labor has more than halved, from 36 to 15%, and the growth rate of labor over the period was −0.7% per annum. Land has held its share but shows no growth. So all the input growth is accounted for by the rapid increase in intermediate inputs (4.2% per annum) and capital goods (1.67% per year). These aggregates over the period conceal the changes that are clear when analyzing the indices in more detail;
Fig. 1. Divisia output, input, and total factor productivity (TFP) indices, for 1948–1997.

Fig. 2. Divisia output indices for horticulture, crops, and livestock, for 1948–1997.
intermediate inputs grew at over 5% per annum before 1978 and declined at 0.19% a year since then; capital inputs increased at 2.53% per annum until 1983 and since then have decreased at 3.79% a year.

To better understand these changes, more disaggregated series are needed which allow the indices to be related to policy changes in the economy. The land index was growing during the period when cultivated area was expanding. Fig. 3 shows that it grew at 0.3% a year through 1960 and then began a slow decline of about 0.25% per annum for the rest of the period. The labor index grew rapidly until 1959, at 1.3% per annum, then wavered until the late 1960s, before beginning a decline of 1.18% a year, from 1968 to 1980. During the growth years up to 1970, the cultivated area under maize production increased in the summer rainfall areas, as oxen were replaced with tractors. Larger areas could be managed and labor use increased, partly due to the spread of fertilizer and high yield varieties. After 1970, the mechanization effect dominated, especially as the combine was introduced and it alleviated the heavy labor demand at harvest time (Sartorius von Bach and van Zyl, 1991). In this respect, the summer rainfall areas followed the pattern in the winter rain regions, where the expansion of cultivated area was largely complete by 1947 and the whole period would have been characterized by the substitution of machinery for labor.

Social policies contributed to the decline in employment, including measures restricting the mobility of labor (the Pass Laws), which became severe after 1965. The effects of economic policies were probably greater than social policies, as these included cheap credit, negative real interest rates and tax breaks, allowing capital equipment to be written off in the year of purchase. There can be little doubt that these price distortions and policies led to unwarranted substitution of capital for labor, imposing substantial social and economic costs on the rural poor (van Zyl et al., 1987).

Then, from 1980, there were 3 years of bad weather, followed by a recovery that by the mid-1980s took labor use back to the level of the early 1970s. Like the shedding of labor, the reversal of the downward trend in employment after 1981 can be explained by policies affecting the relative prices of labor and machinery.
The machinery input index grew at 7.57% a year until 1958; then, at 0.76% per annum until 1981, and has fallen 2.9% a year ever since, partly due to the negative effects of the drought, but in the longer run because of a dramatic change in relative prices. By 1980, economic sanctions protesting apartheid had begun to damage the economy and the gold price collapsed, forcing the government to devalue the Rand substantially and to end credit subsidies and tax concessions on machinery purchases. Thus, the price of imported machinery rose considerably relative to the price of a domestic input like labor and the price distortions which had maximized rural unemployment were ended. Farmers were forced to be more competitive and efficient, and at undistorted relative prices, substituted labor for machinery. This accounts for the partial recovery of the labor index in Fig. 3 and the decline in the machinery index in Fig. 4. Fig. 4 also shows the series for buildings and fixed improvements, which falls rapidly from the early 1970s. This is also policy-induced, as the Pass Laws became more severe from 1968 and subsidies for building housing for black workers were withdrawn.

Thus, in the South African case, the policy changes have been so extreme that policy variables explain a considerable proportion of the changes in productivity. We now turn to the effects of the economic variables that account for long-term TFP growth.

2.1. Explaining TFP growth

Changes in the TFP index should be explained by determining variables, such as R&D expenditures. Following the literature, and using the Cobb Douglas function, $Y_t$ is aggregate output, $X_{jt}$ are traditional inputs, $\Theta_{gt}$ are determining variables and $\beta_j$ and $\gamma_g$ are parameters. The production function can be written as

$$Y_t = \pi X_{jt}^{\beta_j} \pi \theta_{gt}^{\gamma_g}. \quad (1)$$

Rather than estimating the production function directly, a Tornqvist–Theil index was used to aggregate the outputs $Y_{it}$ and the conventional inputs $X_{jt}$. This allows (1) to be written in TFP form, as

$$\ln(\text{TFP})_t = \frac{\hat{Y}}{X} = \ln(\pi \theta_{gt}^{\gamma_g}) \quad (2)$$
where the aggregate index for the inputs is

\[ \hat{X} = \frac{1}{2} \sum_j (C_{jt} + C_{jt-1}) \ln \left( \frac{X_{jt}}{X_{jt-1}} \right) \]  

(3)

which is the logarithm of the ratio of two successive input quantities (the \(X_{jt}\)'s) weighted by a moving average of the share of the input in total cost (the \(C_{jt}\)'s). Since there are multiple outputs, as well as inputs, the same aggregation procedure is applied to give the Tornqvist–Theil output index:

\[ \hat{Y} = \frac{1}{2} \sum_i (S_{it} - S_{it-1}) \ln \left( \frac{Y_{it}}{Y_{it-1}} \right) \]  

(4)

which is the logarithm of the ratio of two successive output quantities weighted by a moving average of the share of the output in total revenue. All the indices reported are chained, so that each value is calculated relative to the previous observation, rather than a base year. The TFP index is simply the ratio of the chained output index and the chained input index.

The \(X_j\)'s include all the conventional inputs such as land, labor, capital, machinery, buildings, chemicals and other miscellaneous inputs. The \(\Theta_j\)'s are the stock of accumulated research capital \((K)\), extension expenditures \((X)\), the number of international patents pertaining to agricultural chemicals and machinery \((P)\), the US TFP index \((T)\), and farmer education \((E)\). US TFP is included to cover spillovers of technology from other research jurisdictions. The patent variable includes both private spillovers from abroad and the technology provided by multinational seed, chemical and machinery companies, regardless of whether they are undertaking research in South Africa or not.

Unfortunately, very little is known about R&D in the purely domestic South African private sector, but this need not prevent economic analysis of the returns to R&D, as this is already included in the market system. Indeed, as was noted by Griliches (1979), if agricultural inputs were supplied by a monopolist, and input statistics took proper account of quality adjustments, technical change emanating from the private sector input industries would be fully included in the input series. Such technological changes are in the farm inputs sector, not the farm sector itself (Kislev and Peterson, 1982), and would not present any difficulties. It is only to the extent that the input suppliers are not monopolists that the statistical sources fail to measure inputs in efficiency units, making it necessary to make some allowance for private R&D expenditures. However, due to these two factors, not all technical change in the input industries is correctly measured at source and there will be some spillover that is still captured in the measures of agricultural productivity. Thus, in estimating the returns to R&D, all the public expenditures should be included on the cost side as well as a proportion of private expenditures. For Europe and the US, the impact of the private sector may be more pronounced than we assume it is here (Schimmelpfennig and Thirtle, 1999).

Accumulated research capital \((KS)\), can be defined very simply as the sum of past R&D expenditures:

\[ KS = \sum R_{t-i} \]  

(5)

but if there is no research, there should be negative growth of \(K_t\). For example, in plant breeding, yield gains tend to be lost over time if research on a particular variety is not maintained, as pests and diseases evolve, making the variety susceptible to attack when it was previously immune. Hence, maintenance expenditures are required to prevent falling productivity (Blakeslee, 1987).

The alternative to including an arbitrary depreciation factor in the calculation of \(KS\) is to include a finite number of lagged \(R_{t-i}\)'s as explanatory variables. Initially, the effect of R&D on productivity is expected to be small and then to reach a peak, before diminishing to zero as the new technology becomes obsolete. Following this procedure, adding a constant \((A)\) and a stochastic error term gives the customary model

\[
\ln(TFP)_t = \ln A + \sum \alpha_i \ln R_{t-i} + \gamma_1 \ln X_{t-i} + \gamma_2 \ln E_{t-i} + \gamma_3 \ln P_{t-i} + \gamma_4 \ln T_{t-i} + \gamma_5 W_t + U_t
\]  

(6)

where TFP\(_t\) is the productivity index, \(R_{t-i}\) is R&D lagged \(i\) years and all the other variables except the weather have varying lags of \(i\) years. Khatri et al. (1998) have established that these variables are Granger prior to TFP. \(U_t\) is the remaining stochastic error.

This model was estimated by Thirtle and van Zyl (1994), imposing a 9-year, second degree polynomial lag on R&D and shorter lags on the other variables.
Only R&D, extension and education were significant, and the net internal ROR was found to be 64%.  

2.2. Modeling the distributed lags for R&D and patents

The primary focus of the paper is in determining, and using in an ROR calculation, the length and shape of the lag between R&D and TFP. The first stage of the analysis is to determine the length of the lag by estimating an unrestricted finite lag model, with lag length $k$. Using fewer lags than the true specification implies biased estimates and too many lags implies inefficient estimates. The lag length is found by searching over a range, using appropriate model selection criteria. Geweke and Meese (1981) investigate various model selection criteria for this purpose, including the Akaike information criterion (AIC), the final prediction error (FPE), the Bayesian estimation criterion (BEC) and the Schwartz criterion (SC). The tests are not always in agreement, but for R&D, all the tests confirmed a lag of 9 years.

The second stage involves determining the shape of the lag, and the conventional inverted U-shaped second degree polynomial fitted the South African data poorly relative to asymmetric alternatives. Amongst the large number of alternative lag structures tested, generalizations of the exponential and Gamma distributions (Schmidt, 1974) gave the best results according to the model selection criteria. Fig. 5 shows the three structures that gave the best results and compares them with the second degree polynomial. The selection criteria clearly reject the second degree polynomial (PDL(2)). Of the other structures, the linear exponent exponential, with a 1-year lead time (EXP(1) LEAD 1) is preferred. The strong negative skew, with a peak in the second year, followed by a decline very much like geometric decay, is very different from the second degree polynomial. The early peak has a marked effect on the ROR calculation, which we will show is sensitive to the distribution of returns over time.

This result can be contrasted with similar tests for the UK system, where the lag was 16 years and the distribution was positively skewed, with the peak

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2 The gross ROR is based on the gross value of output, whereas net RORs are calculated using the value of output minus the value of inputs.

3 Since the technology variables are interdependent, the lag structures should be jointly estimated. Thus, the model selection criteria are applied to combinations of lags for the different variables in a grid search for the model that performs best. The procedure is arduous, both because of the interdependence and because several distributions need to be tested with different polynomial degrees. Thus, the grid search has to be performed over a large number of possible cases.
effects at 12–16 years (Khatri, 1994). This suggests that, whereas the UK system includes a high proportion of basic scientific research with a long gestation period, the South African research system is dominated by short-term developmental and applied work. This is despite the fact that government-funded research in the universities is included.

Applying the same method to the patent count series gave a lag of 12–13 years and again a negatively skewed exponential distribution. The longer lag on the patent series might be because mechanical innovations are embodied in capital items and the capital stocks adjust slowly. The lags on extension and education should be far shorter than those for R&D, so the same elaborate lag structures are not imposed. Indeed, the model selection criteria indicate that a one-period lag is the appropriate specification.

Incorporating the exponential lag structures in the estimation of Eq. (6) shows how much the lag structure matters. The R&D lag is so precise that none of the other variables except the weather have any explanatory power and even the net ROR is an implausible 170%. Thus, careful determination of the lag structures has the unfortunate effect of destroying the conventional two-stage model.

2.3. Using the lag structures in the profit function

The poor results from the refined version of the two-stage model suggest that an integrated approach may be preferable. Including the technology variables in the estimation of the residual profit function is less restrictive than the two-stage approach, but the model is too complex and there are insufficient degrees of freedom to allow much in the way of testing for the best formulation. Thus, the lags and depreciation rates in previous studies have often been chosen somewhat arbitrarily. In this case, prior estimation of the two-stage model provides information on the lag structures of the technology variables, that can be incorporated in the profit function. Thus, the R&D capital stock is constructed using the perpetual inventory model (PIM), which exhibits the same pattern of geometric decay as the exponential lag. This knowledge stock (KS), with R&D entering with a one-period lag, is then

$$ KS_t = RD_t + (1-\delta)KS_{t-1} $$

where $\delta$ is the rate of depreciation. This depreciation rate is set at 10%, so that 90% of the weight is on the first 9 years, the length selected for the exponential lag previously. There is no problem with choosing a starting value for KS because data is available back to 1929, so by 1947, the starting value of KS in 1929 is irrelevant.

The same approach is applied to the other variables, so the patent knowledge stock is approximated in the profit function by a PIM with a depreciation rate of 8.3%, which is appropriate for a 12-year lag. Lastly, extension expenditures and farmer education are included in the model with a simple 1-year lag, as before. Thus, all the information on lag structures gleaned from the two-stage approach are incorporated into the less restrictive integrated model, to which we now turn.

3. The residual profit function model

Assuming that farmers maximize expected profits, the normalized restricted profit function (Lau, 1976), with conditioning factors included as fixed inputs, is used to model farmer behavior. Consider a multiple output technology producing outputs $Y$, $(Y_1, \ldots, Y_m)$, with the respective expected output prices $P$, $(P_1, \ldots, P_m)$, using $n$ variable inputs $X$, $(x_1, \ldots, x_n)$ with prices $W$, $(w_1, \ldots, w_n)$. Variable expected profits are defined as

$$ \pi = \sum_{i=1}^{m} P_i y_i - \sum_{j=1}^{m} w_j x_j = P'Y - W'X $$

The expected prices are taken to be the actual prices from the previous year. This is in accordance with the findings of van Schalkwyk et al. (1994), who concluded that naive price expectations explain aggregate South African farmer behavior better than other expectation models.

Normalizing the profit function with respect to an output or input price ($w_0$ in Eq. (9), which is the price of livestock inputs) has the practical advantages of ensuring that the homogeneity requirement is met, reducing the number of parameters to be estimated, formulating the problem in a manner consistent with economic theory, and negating the need to deflate prices. The optimal solutions to maximizing (8) would be equivalent to those obtained from the maximization
of normalized restricted profits, and thus, the normalized expected profit function can be represented as
\[ \Pi^* = \Pi^* \left( \frac{P}{w_0}, \frac{W}{w_0}; Z, \theta \right) = \frac{\pi^*(P; W; Z, \theta)}{w_0} \]  
(9)
where \( \Pi \) represents the normalized restricted profit function, \( Z \) is the vector of fixed and quasi-fixed inputs, \( \Theta \) is the vector of technology variables, or conditioning factors (also treated as quasi-fixed in Eq. (9)) and (*) indicates optimized levels. The theoretical restrictions on (9) are that the normalized restricted profit function (hereafter called the profit function) is non-decreasing in \( P \), non-increasing in \( W \), linearly homogeneous in prices, twice continuously differentiable and convex in prices, and concave in the quasi-fixed factors (Lau, 1976).

Many studies using time series data employ a time trend as an index of technology. Clark and Youngblood (1992) have demonstrated that the use of a deterministic time trend with difference stationary production and price data is inconsistent. By including the (productivity shifting) conditioning factors directly in the objective function, this criticism is addressed in a manner consistent with their recommendations. Quasi-fixed factors (capital stocks) are those that are fixed in the short-run (one production period), but can be varied in the longer run. Fixed inputs and conditioning variables, including public and private research expenditures, the international stock of knowledge, extension, farmer education and the weather, are factors of production that cannot be varied by the farmer even in the long run. Thus, profit maximization is assumed to be subject to the levels of these factors.

The functional form employed is the GQ. The GQ profit function is defined as
\[ \Pi = \alpha_0 + \alpha' \hat{P} + \delta' \Theta + \frac{1}{2} \hat{P}' \beta \hat{P} + \frac{1}{2} \Theta' \phi \Theta + \hat{P}' \gamma \Theta \]  
(10)
where \( \hat{P} \) hat is the stacked vector of normalized output and non-numeraire input prices, \( (P, W)' \) and \( \Theta \) is the stacked vector of \( k \) quasi-fixed, one fixed (for the short-run function) and conditioning factors. The vector \( \alpha \) (\( \alpha_1, \ldots, \alpha_{m+n-1} \)) and matrices \( \beta \) (\( \beta_{ij}, i, j = 1, \ldots, m + n - 1 \), \( \phi_{gh}, g, h = 1, \ldots, K + L \)) and \( \gamma \) (\( \gamma_{ig}, i = 1, \ldots, m + n - 1 \), \( g = 1, \ldots, k + l \)) contain the parameter coefficients to be estimated. The vector of parameters \( \gamma \) is of particular interest as these are the shadow prices of the fixed factors and technology variables. Applying Hotelling’s lemma, we derive the (short-run) optimal levels of output supply and input demands:
\[ -x_i^* = \alpha_i + \sum_{j=1}^{m} \beta_{ij} p_j + \sum_{j=m+1}^{m+n-1} \beta_{ij} w_j + \sum_{g=1}^{k+1} \gamma_{ig} \Theta' g', \ i = m+1, \ldots, m+n-1 \]  
(11)
\[ y_i^* = \alpha_i + \sum_{j=1}^{m} \beta_{ij} p_j + \sum_{j=m+1}^{m+n-1} \beta_{ij} w_j + \sum_{g=1}^{k+1} \gamma_{ig} \Theta' g', \ i = 1, \ldots, m \]  
(12)
The long-run profit function differs only in treating all the factors of production as variable inputs. Thus, machinery, land, and the animal capital stocks are assumed to be at their long-run equilibrium levels and there are no quasi-fixed factors. The vector \( \Theta \) only consists of conditioning factors, such as the weather and technology-related variables, that are always exogenous to the farmer’s decision process. The prices of inputs that were quasi-fixed factors in the short-run model are included in the vector \( w \). Variable input demands for these factors are obtainable using Hotelling’s lemma, and are estimated jointly with other input demands and other supply equations.

The price elasticities are derived as logarithmic derivatives of the supply and derived demand equations with respect to prices. If the elements of \( \Theta \) are viewed as short-run constraints on production, it is possible to derive the effects of relaxing the \( \Theta \) variable constraints on the output and variable input levels. These fixed factor elasticities are derived as logarithmic derivatives of the supply and derived demand equations with respect to the elements of \( \Theta \) (Lass, 1985; Khatri et al., 1997).

Shadow values are given by the partial derivatives of the profit function with respect to the \( \Theta \) variables (Diewert, 1974). The derived shadow values can be interpreted equivalently as the marginal change in profits for an increment in a particular element of \( \Theta \) or as the imputed rental value for an additional unit of that factor. From the short-run formulation,
the shadow values of capital, livestock and land can be derived as well as the values for public research, extension, patents and education, which are also calculated from the long-run version. The difference between the rental value and the shadow value indicates whether the factor is over, under or optimally utilized. Finally, the shadow value of research can be used to derive the rate of ROR investment (Huffman, 1987).

Dual measures of technological biases can also be obtained from the profit function. Huffman (1987) suggests a summary measure which provides input and output biases with respect to the conditioning factors and Khatri (1994) generalizes the conditioning factor biases for a multiple output technology.

4. Data

The national farm-level production data for the period 1947–1992 were obtained from several sources, largely RSA (1994), and are described in some detail in Thirtle et al. (1993). For both the short- and long-run profit function specifications, the three output aggregates are crops, horticulture, and livestock and livestock products.

For the short-run profit function, the variable inputs are Divisia aggregated into four groups: (1) hired labor; (2) machinery running costs (fuel, machinery repairs and other); (3) intermediate inputs (fertilizer, other chemicals, and packing material) and (4) livestock feed and dips. Vehicles and fixed capital in the form of buildings and other fixed improvements (especially irrigation equipment) are assumed to be quasi-fixed in the short run, as are the stocks of animals. The total area of land in the commercial sector is included as a fixed input.

For the long-run specification, all the conventional inputs are variable. These were Divisia aggregated into the following groups: (1) hired labor; (2) machinery running costs (fuel, machinery repairs and other); (3) intermediate inputs (fertilizer, other chemicals, packing material, feed, and dips); (4) capital, particularly vehicles and other capital in the form of buildings and fixed improvements; (5) livestock and (6) land. The capital stocks are calculated using US depreciation rates (Jorgenson and Yun, 1991, Table 13B, p. 82) in a PIM that assumes geometric decay, as in Ball (1985). The rental prices of the capital stocks are calculated using Jorgenson's formula to derive long-run capital service prices from the assumed depreciation rate and the real rate of interest. 4

The conditioning factors, that are treated as fixed inputs in both the short- and long-run specifications, are public research expenditures, public extension expenditures, a rainfall index, world patents 5 and a farmer education index (ED). The farmer education index is the average number of years of secondary education of farmers, which was kindly provided by the South African Agricultural Union (SAAU).

5. Estimation and results

There are too many parameters in the short-run profit function (10) to estimate the full model in one stage, so the residual profit function approach (Bouchet et al., 1989; Khatri et al., 1997) is used. The system of supply and demand equations, Eqs. (11) and (12), is estimated in the first stage and then the remaining variables are used to explain the residual. The estimated shadow prices and the input biases involve both the parameters from the supply and demand system, and the residual profit function. However, as the majority of the parameters for the shadow price and input bias equations are in the supply and demand system, the parameters used in the residual profit function (most of which are significant) can be treated as constants. This allows the derivation of indicative significance bounds for the shadow price and input bias estimates. 6

The system of output supply and variable input demand equations are estimated for both the short-run and the long-run using the iterative Zellner or seemingly unrelated procedure. Each system, with symmetry imposed, produces a set of parameter estimates (not reported here), most of which are significant at the 5% confidence level. The coefficients of determination ($R^2$'s) of the estimated individual supply and demand equations (for both the short-run and long-run

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4 We thank Eldon Ball for constructing these series.

5 The patent data comes from the US patent database compiled at the University of Reading by John Cantwell. The series are patent counts, for all agriculture-related chemical and mechanical patents registered in the US.

6 The US TFP index was not included in the profit function estimation, both to preserve degrees of freedom and because variable deletion tests applied to the two-stage approach indicated that it had no explanatory power.
specifications) vary between 0.87 and 0.99, which is high even for time series models. The Durbin–Watson statistics indicate that there are no problems of serial correlation in the individual equations. Further, although homogeneity remains a maintained assumption (implicitly imposed when normalizing), symmetry and monotonicity, which are necessary conditions for global convexity, are both satisfied by the estimated systems. The estimated profit functions are thus found to be acceptable both with respect to their statistical performance and theoretical consistency.

The results obtained with the short-run and long-run profit functions (that can be estimated in one stage), specifically the elasticity estimates, are in accordance with expectations. The elasticities of the outputs and variable inputs in the long- and short-run estimations are remarkably similar. They are not repeated here as they are almost identical to the short-run profit function results reported by Khatri et al. (1995). As expected, the long-run elasticities for land, capital and livestock are consistently higher than the short-run elasticities. It can therefore be concluded that the results from the short- and long-run models are consistent.

The elasticities show the low supply response of South African agriculture, even over the long run. This result corresponds very closely with the findings of van Zyl (1986) and Sartorius von Bach and van Zyl (1991). It also confirms earlier comments on the abnormal development path of South African agriculture (World Bank, 1994; Kirsten and van Zyl, 1996).

The estimates of the factor-saving biases of technical change show that the Hicks neutrality assumption implicit in the two-stage model is rejected, which is contrary to assumptions implicit in the two-stage approach and may have biased the results. Public R&D has been capital and intermediate input using, and labor and animal input saving. This pattern is typical of developed countries and is hardly appropriate for a country with abundant cheap labor and high rural unemployment. Thus, the public R&D system has exacerbated the damage done by apartheid policies discussed above.

### 5.1. Shadow prices

Khatri et al. (1998a) established that the shadow price of land is positive, but for livestock, it was not significantly different from zero and for capital it was negative, indicating that policies to reduce capital use could have increased profit. More importantly, they found that the capital stock took 11 years to adjust to changes in input prices. This paper concentrates on the shadow values of the technology variables, which can be interpreted as the marginal change in profits from a unit increment in a technology-related variable. The shadow value for R&D can be used to derive the rate of ROR investment (see Khatri et al., 1996). Note, however, that the shadow values of the short- and long-run conditioning variables are not directly comparable due to differences in the units of measurement of the capital items. These shadow values are reported in Table 1 for both the short- and long-run specifications, evaluated at the variable means.

<table>
<thead>
<tr>
<th>Factor</th>
<th>Short-run specification</th>
<th>Long-run specification</th>
</tr>
</thead>
<tbody>
<tr>
<td>Public research</td>
<td>4.04</td>
<td>323.5</td>
</tr>
<tr>
<td>International patents</td>
<td>0.23</td>
<td>342.8</td>
</tr>
<tr>
<td>Public extension</td>
<td>-0.012</td>
<td>-290.6</td>
</tr>
<tr>
<td>Farmer education</td>
<td>-1378.5</td>
<td>-6867.2</td>
</tr>
</tbody>
</table>

*Note: All the shadow prices are significant at the 0.05 level; shadow values of the short- and long-run conditioning variables are not directly comparable due to differences in the units of measurement of the capital items.*
tures were too high. Similarly, the education index appears to have considerable explanatory power in both the short- and long-run formulations, but the shadow price is negative. Since education is a proxy for managerial ability, this is contrary to expectations, but this negative result and the weak contribution of extension expenditures are probably related.

The shadow value of education was positive until the early 1960s, but has become increasingly negative since then, and the shadow price of extension has been falling over the period, which suggests that South African commercial farmers have become less dependent on public extension advice. This corresponds with the findings of Koch et al. (1991), who show that government extension officers spend increasingly more time on administrative duties and do very little actual extension work. Thus, the decline in extension effort could explain the low payoff, but so could reduced need for extension. The education level of South African commercial farmers is relatively high, so it is entirely possible that the minimum level required to assimilate mass produced research and extension messages has been reached.

But there is a more radical explanation that is far more specific to South Africa. The unreported fixed factor elasticities show that all the input and output elasticities with respect to education are positive, and all but one are highly significant. Thus, education augments output but it also augments input use, more than proportionately in the case of non-labor inputs. As crop production expanded into climatically marginal and more risky areas, intermediate input use and mechanization increased considerably in the period from 1965 to the early 1980s (van Zyl et al., 1995). There was evidence of over-mechanization (van Zyl et al., 1987) and fertilizer was often applied (on extension service advice) up to levels where it actually decreased output (Korentajer et al., 1989). This was especially disastrous in the bad climatic conditions of the early and late 1980s (van Rensburg and Groenewald, 1987). Sartorius von Bach et al. (1992) clearly show that it was the better educated farmers who adopted these practices to a greater extent. Thus, educated farmers did respond more strongly to extension service advice, but because maximum physical production (as opposed to maximizing profit) was the major goal and focus of the agricultural research and extension system, the effects were negative.

5.2. Internal rate of return

The shadow prices reported above are hard to interpret because the technology variables have no obvious prices, and thus, there is no way to compare whether there has been over or under investment. Therefore, the effectiveness of public R&D expenditures can be better interpreted by calculating rates of return (ROR). The shadow values represent the imputed marginal value of a unit increase in knowledge stocks. Thus, to estimate the marginal internal rate of return (MIRR) to research, the additional flow of research investment required to change current knowledge stocks by one unit must be calculated and this will depend on the length and shape of the lag. The MIRR will vary with the choice of the number of periods over which the incremental research is distributed. Research is found to affect productivity for 9 years, so the MIRRs reported in Table 2 are for this period of incremental research investments (resulting in a unit change in the knowledge stock). The rate of interest that equates this incremental research expenditure to the shadow price is the MIRR (Ito, 1991).

The net (output value minus input costs) ROR for the short-run profit function is 44%, while the long-run result is 113%. The lower value is perhaps reasonable for a research system that is to some extent free-riding on the investments made by others. The long-run result is not as high as for the two-stage model, but it is 2.6 times the short-run result. The return will be higher because of the effects of new technology embodied in new capital equipment. Thus, the difference between the two results should depend on the fact that the capital stocks are allowed to adjust in the long-run case. This raises a new problem, since Khatri et al. (1995) found that the capital stock adjusted to changes in the real rate of interest with a long lag of 11 years. The

<table>
<thead>
<tr>
<th>Table 2</th>
<th>Estimated returns to R&amp;D</th>
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<tr>
<td>Lag length</td>
<td>Short-run specification (%)</td>
</tr>
<tr>
<td>5 years</td>
<td>44</td>
</tr>
<tr>
<td>11 years</td>
<td>–</td>
</tr>
</tbody>
</table>

*aLag length appropriate for variable long-run inputs (see discussion in Section 5.2).  
bLag length appropriate for an average of fixed and quasi-fixed long-run factors.
12–13-year lag for the patent series, reported above, corroborates this period of adjustment. This suggests that, although the negatively skewed 9-year lag, with a peak effect after 2 years, may be appropriate for the short-run profit function, it is hard to reconcile with long-run adjustments in fixed or quasi-fixed factors. Therefore, the short lag may be appropriate for R&D on variable inputs like seed varieties and agronomic improvements, but R&D on capital items like irrigation equipment, cultivation implements and other specialized machinery must take longer than 2 years to have a peak effect. The difference between the short-run and long-run RORs should result from the technology that is embodied in the capital items. If this is so, then the effects will only occur when the capital stock has adjusted.

To take the adjustment period of 11 years into account, without knowing the lag distribution, the MIRR for the long run is calculated with a lag of 11 years between the unit increment in the knowledge stock and the shadow value. This gives the last figure in Table 2, of 58%, which may be a more realistic estimate of the long-run net ROR. These are certainly respectable rates of return on public expenditure, with the usual qualifications that the figures may be somewhat diminished if we adjust for the dead-weight losses associated with tax collection and the possibility that public funding may be crowding out private sector research. Of greater interest is the sensitivity of ROR calculations to assumed lag structure, particularly when the lag in the effect of R&D on TFP is skewed, negatively in this case, so benefits are exposed to less discounting than if the lag structure were to be symmetric.

6. Conclusion

This paper reports TFP calculations for South African commercial agriculture that show the damage caused by the policies of the apartheid era, both in terms of low productivity growth and the unwarranted substitution of scarce capital for abundant labor. The shadow price of capital is found to be negative, which leads to restructuring problems when it takes the capital stock 11 years to adjust to changes in input prices. This long lag, from the fixed nature of capital investments, is no doubt also related to apartheid policies which are now gone, but still might imply that over a decade is required to turn around an over-capitalization problem. In fact, since the collapse of the Rand with the gold price that occurred in the early 1980s, we would not expect the overcapitalization to be corrected by the end of the data series used here.

TFP growth is explained by public R&D and extension expenditures, farmer education and spill-ins of private R&D. This two-stage approach allows the lags to be investigated and the lag between public R&D and TFP is found to be only 9 years long, with a strong negative skew, giving a peak effect after only 2 years. This is because of the adaptive focus of the South African research system, and can be compared with an 18-year positively skewed lag for the UK system, which undertakes far more basic research. This lag shape results in a rate of return to R&D of 170% and has the effect of making all the other variables insignificant.

To improve the results, the lag structures established in the two-stage approach are used in constructing the technology-related variables included in the estimation of short- and long-run residual profit functions. These profit functions conform to theoretical requirements and produce reasonable estimates of the shadow values of the technology-related variables. The short-run net ROR is found to be 44%, but the long-run ROR is not clear. If the short lag is accepted, the ROR is 113%, but if the capital stock adjustment lag is taken into account, the ROR falls to 58%. More generally, the results show that the ROR is critically dependent on the length and especially the shape of the lag. There is also a problem in determining the long-run ROR in any case, where the estimated lag on R&D is less than the adjustment period of the capital stock. This issue has not arisen in studies of the developed countries because the R&D lag for systems that perform substantial basic research is far longer than for South Africa, which concentrates on adaptive research. The question remains how sensitive ROR calculations might be, even in developed settings, to a significantly skewed distribution of research benefits.

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References


